

Intravenous magnesium sulphate and sotalol for prevention of atrial fibrillation after coronary artery bypass surgery: a systematic review and economic evaluation

J Shepherd, J Jones, GK Frampton, Ł Tanajewski,
D Turner and A Price



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Abstract

Intravenous magnesium sulphate and sotalol for prevention of atrial fibrillation after coronary artery bypass surgery: a systematic review and economic evaluation

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Objectives: To assess the clinical and cost-effectiveness of magnesium sulphate compared with sotalol, and to assess the clinical effectiveness of magnesium sulphate compared with placebo in the prevention of atrial fibrillation (AF) in patients who have had a coronary artery bypass graft (CABG).

Data sources: Major electronic databases were searched from December 2003 to May 2007.

Review methods: Selected studies were assessed, subjected to data extraction using a standard template and quality assessment using published criteria. A simple short-term economic model was developed, informed by a systematic review of economic evaluations and populated with data from a review of costing/resource-use studies and other published studies. The cost-effectiveness of magnesium sulphate as prophylaxis was estimated for a set of base-case assumptions and the robustness of these results was assessed using deterministic and probabilistic sensitivity analysis.

Results: Twenty-two papers met the inclusion criteria reporting 15 trials which all compared magnesium sulphate with placebo or control. They ranged in size from 15 to 176 patients randomised, and were conducted in Europe, the USA and Canada. The standard of reporting was generally poor, with details of key methodological attributes difficult to elucidate. No trials were identified that specifically aimed to compare magnesium sulphate with sotalol. Of 1070 patients in the pooled magnesium group, 230 (21%) developed postoperative AF, compared with 307 of 1031 (30%) patients in the placebo or (control) group. Meta-analysis using a fixed-effects model generated a pooled odds ratio (OR) that was significantly less than 1.0 [OR = 0.65, 95% confidence interval (CI) 0.53 to 0.79, test for overall effect $p < 0.0001$], but with statistically significant heterogeneity ($I^2 = 63.4\%$,

$p = 0.0005$). Two randomised controlled trials (RCTs) were notable as they had relatively lower ORs in favour of magnesium sulphate. When these were removed from the analyses the pooled OR remained statistically significant, but heterogeneity no longer remained significant. These two studies tended to impart a highly significant reduction in the odds of AF to whichever subgroup they were analysed in. When studies were ordered by total duration of prophylaxis, an apparent relationship between duration and odds of AF was evident, with decreasing odds of AF as duration of prophylaxis increased. This was confirmed by linear regression analysis ($R^2 = 0.743$, $p < 0.001$). When the data were grouped into three classes according to duration, a statistically significant intervention effect was only present for the longest duration (OR = 0.12, 95% CI 0.06 to 0.23, $p = 0.00001$). Statistically significant intervention effects were associated with the initiation of prophylaxis 12 hours or more before surgery (OR 0.26; 95% CI 0.16 to 0.44, test for overall effect $p = 0.00001$, fixed-effects model) and less than 12 hours before surgery or during the surgery itself (OR = 0.73, 95% CI 0.56 to 0.97, test for overall effect $p = 0.03$, fixed-effects model), but not when prophylaxis was initiated at the end of surgery or postsurgery (OR = 0.85, 95% CI 0.59, 1.22, $p = 0.37$, fixed-effects model). When studies were ordered by total dose of intravenous magnesium sulphate (<25 g), the odds of AF were independent of the dose. A notable exception was that for a total dose of 9 g magnesium sulphate; here the odds of AF were significantly reduced relative to the control group, although this may be explained by the fact that these studies had excluded patients who were on antiarrhythmic drugs and so may have been at higher risk of AF. Sixty-three potentially relevant references about cost-effectiveness were identified, but no

economic evaluations of intravenous magnesium alone as prophylaxis against AF following CABG, compared with sotalol as prophylaxis or no prophylaxis, were identified. Studies reporting resource use by patients with AF following CABG suggest that while AF significantly increased inpatient stays, by up to 2.3 days in the intensive care unit (ICU) and 3.4 days on the ward, differences in length of stay and costs between patients receiving prophylaxis and those not receiving prophylaxis were not statistically significant. In the base-case analysis, magnesium sulphate prophylaxis resulted in 0.081 fewer cases of AF at an incremental cost of £2.55. The incremental cost-effectiveness ratio (ICER) was £32 per AF case avoided. The estimated difference in average length of stay between the prophylaxis and no-prophylaxis strategies was only 0.24 days, despite a large assumed difference of 3 days for patients experiencing AF in each group (1 extra day in the ICU and 2 extra days on the ward). In a deterministic sensitivity analysis the greatest variation in ICERs was observed for input parameters relating to the baseline risk of AF following CABG and the effectiveness of prophylaxis, cost of prophylaxis and the resource consequences of postoperative AF. The largest ICER (£2092) in the sensitivity analysis was associated with increasing the length of patients' preoperative stay. In the base case it was assumed that admission routines would be identical under both strategies. However, patients receiving prophylaxis by intravenous infusion may have longer preoperative stays. In a probabilistic analysis the majority of the simulations were associated with improved outcomes (in this case fewer cases of AF), but also higher costs. Prophylaxis was the dominant strategy (better outcome at lower cost) in about 41% of the simulations using the base-case assumptions. Under an alternative scenario where patients receiving prophylaxis are admitted for longer before their operation, to receive their initial infusion, the proportion of simulations where prophylaxis dominates fell to around 5%. The

probability of being cost-effective was 99% at a willingness to pay (WTP) threshold of £2000 per AF case avoided and 100% at a WTP threshold of £5000 per AF case avoided under the base-case assumptions. Under the alternative scenario of longer preoperative stays the probability of being cost-effective at these two threshold values fell to 48% and 93%, respectively. It is unclear what the appropriate decision threshold should be, given that this model used intermediate rather than final outcomes.

Conclusions: No RCTs were identified that specifically aimed to compare intravenous magnesium with sotalol as prophylaxis for AF in patients undergoing CABG. Intravenous magnesium, compared with placebo or control, is effective in preventing postoperative AF, as confirmed by a statistically significant intervention effect based on pooled analysis of 15 RCTs. It was also found that AF was less likely to occur when a longer duration of prophylaxis was used, and the earlier that prophylaxis is started; however, this finding was associated with two RCTs that had more favourable results than the other trials. No clear relationship between dose and AF was observed, although a lower constant dose rate was associated with the lowest odds of AF. Further research should investigate the relationship between dose, dose rate, duration of prophylaxis, timing of initiation of therapy and patient characteristics, such as degree of risk for AF. This will provide stronger evidence for the optimum delivery of intravenous magnesium in patients undergoing CABG. In the base-case analysis in the economic model, magnesium sulphate prophylaxis reduced the number of postoperative AF cases at a modest increase in cost. The results of the economic analysis are highly sensitive to variation in certain key parameters. Prophylaxis is less likely to be a cost-effective option if it requires changes in admission routines that result in longer preoperative stays than would be the case without prophylaxis.



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Glossary and list of abbreviations

Technical terms and abbreviations are used throughout this report. The meaning is usually clear from the context, but a glossary is provided for the non-specialist reader. In some cases, usage differs in the literature, but the term has a constant meaning throughout this review.

Glossary

Atrial arrhythmia Altered atrial rhythm (includes atrial fibrillation, atrial flutter and atrial tachycardia).

Atrial fibrillation Uncoordinated atrial pulsation.

Atrial flutter Increased but coordinated atrial pulsation.

Atrial tachycardia Increased atrial beat rate.

F waves (ECG) Regular, rapid atrial waves indicative of atrial flutter.

Left ventricular ejection fraction The fraction of blood pumped out of a ventricle with each heart beat; one of the most important predictors of prognosis.

P wave (ECG) The wave of depolarisation that spreads from the sinoatrial node throughout the atria.

PQ interval (ECG) The time between the beginning of atrial depolarisation and the beginning of ventricular depolarisation.

QRS complex (ECG) Deflections in the tracing comprising the Q, R and S waves indicating currents generated when the ventricles depolarise before their contraction.

QT interval The time between the start of the Q wave and the end of the T wave in the heart's electrical cycle.

Supraventricular arrhythmia A rhythmic abnormality of the heart caused by impulses originating above the ventricles, e.g. in the atrioventricular (sinoatrial) node; may be synonymous with atrial arrhythmia.

List of abbreviations

ACE	angiotensin-converting enzyme	i.v.	intravenous
AF	atrial fibrillation	LVEF	left ventricular ejection fraction
BMI	body mass index	MI	myocardial infarction
BNF	British National Formulary	NA	not applicable
BP	blood pressure	NICE	National Institute for Health and Clinical Excellence
bpm	beats per minute	NR	not reported
BSA	body surface area	ns	not significant
CABG	coronary artery bypass graft	NYHA	New York Heart Association
CDSR	Cochrane Database of Systematic Reviews	OR	odds ratio
CEAC	cost-effectiveness acceptable curve	PCV	pharmacological cardioversion
CI	confidence interval	POAT	postoperative atrial tachycardia
COPD	chronic obstructive pulmonary disease	PSA	probabilistic sensitivity analysis
CPB	cardiopulmonary bypass	PVD	peripheral vascular disease
CRD	Centre for Reviews and Dissemination	QALY	quality-adjusted life-year
D ₅ W	5% dextrose water solution	RCT	randomised controlled trial
df	degrees of freedom	SD	standard deviation
ECV	electrical cardioversion	SDU	step-down unit
FEV ₁	forced expiratory volume in 1 second	SHTAC	Southampton Health Technology Assessments Centre
HRG	Healthcare Resource Group	SR	systematic review
ICER	incremental cost-effectiveness ratio	SVA	supraventricular arrhythmia
ICU	intensive care unit	SVT	supraventricular tachycardia
ITT	intention-to-treat	WTP	willingness to pay

All abbreviations that have been used in this report are listed here unless the abbreviation is well known (e.g. NHS), or it has been used only once, or it is a non-standard abbreviation used only in figures/tables/appendices in which case the abbreviation is defined in the figure legend or at the end of the table.



Executive summary

Background

Atrial fibrillation (AF) is a supraventricular arrhythmia characterised by abnormal heart rhythm, with symptoms such as palpitations and nausea. It is one of the most common complications after coronary artery bypass graft (CABG) and 20–40% of patients experience AF following cardiac or thoracic surgery. AF increases the risk of mortality and morbidity from stroke, heart failure, myocardial infarction and thromboembolism. This can result in prolonged hospitalisation, hospital readmission, excess utilisation of hospital resources and increased health service costs. Risk factors include advanced age (particularly over the age of 50), previous history of AF, male gender, hypertension, diabetes, smoking, myocardial infarction and valvular heart disease.

Clinical guidelines recommend that β -blockers are used routinely as first choice for the prophylaxis of AF in all patients undergoing cardiac surgery. It is also recommended that sotalol hydrochloride, a β -blocker with class III antiarrhythmic activity, is used. Magnesium may also be given to patients undergoing cardiothoracic surgery to reduce hypomagnesaemia, a common occurrence following this kind of surgery. However, it is not a first line choice for prophylaxis and it is not known to what extent it is used in current practice.

Objective

The aim of this research is to conduct a systematic review and economic evaluation of the clinical and cost-effectiveness of magnesium sulphate compared with sotalol, and to assess the clinical effectiveness of magnesium sulphate compared with placebo in the prevention of atrial fibrillation in patients who have had a CABG.

Methods

Methods for assessing clinical effectiveness

A systematic review was conducted to compare intravenous magnesium sulphate with placebo (or

control), and intravenous magnesium sulphate with sotalol given as prophylaxis before the onset of AF, in patients over 18 years, undergoing elective isolated CABG. Studies of other magnesium compounds (e.g. chloride, hydroxide or unspecified) were excluded. The primary outcome was incidence of postoperative AF. Supraventricular arrhythmias other than AF (e.g. tachycardias and atrial flutter) and all other non-atrial arrhythmias were excluded. Patient length of postoperative stay and the total length of hospital stay were additional outcomes.

A comprehensive search strategy was developed to identify relevant randomised controlled trials (RCTs) and systematic reviews. As this systematic review updates a previous published systematic review the searches were limited to studies published after the cut-off date for literature searching in that review (December 2003). The strategy was applied to ten general and specialist health and biomedical databases. Titles and abstracts were screened systematically against the inclusion criteria and full papers were ordered for further investigation. All included trials were subjected to data extraction using a standard template and quality assessment using published criteria. Data were analysed by narrative synthesis and quantitative meta-analysis, with sensitivity analyses where necessary. *A priori* defined subgroup analyses were performed to assess the effects of different delivery strategies for intravenous magnesium, including different total doses, timing of the initiation of prophylaxis and total duration of prophylaxis.

Methods for assessing cost-effectiveness

A systematic literature search was undertaken to identify economic evaluations of intravenous magnesium sulphate alone as prophylaxis against AF following CABG compared with sotalol as prophylaxis or no prophylaxis. A secondary aim of this review was to identify economic evaluations of other agents used for prophylaxis against postoperative AF or studies that reported cost/resource-use differences for patients undergoing CABG who developed AF. The purpose of reviewing these studies was to identify the scope and methods adopted in previous economic evaluations of prophylaxis against postoperative

AF and to identify the impact of postoperative AF on patients' resource use, which would inform the development of an economic model.

A comprehensive search strategy was developed to identify relevant economic evaluations and costing studies. The strategy was applied to a number of general and specialist health and biomedical databases. Titles and abstracts were screened against the inclusion criteria and full papers were ordered for further investigation. Included studies were discussed in a narrative review.

A simple short-term economic model was developed, informed by the systematic review of economic evaluations and populated with data from the review of costing/resource-use studies and other published studies. The cost-effectiveness of magnesium sulphate as prophylaxis was estimated for a set of base-case assumptions and the robustness of these results was assessed using deterministic and probabilistic sensitivity analysis.

Results of the assessment of clinical effectiveness

The review identified 206 potentially relevant references. Of these, 22 papers met the inclusion criteria, comprising 17 papers that reported parallel-group RCTs (15 RCTs altogether) and five systematic reviews.

Of the 15 trials included, all compared magnesium sulphate with placebo or control. No trials were identified that specifically aimed to compare magnesium sulphate with sotalol. The 15 trials ranged in size from 15 to 176 patients randomised, and were conducted in Europe, the USA and Canada. The standard of reporting was generally poor, with details of key methodological attributes (e.g. method of randomisation and concealment of allocation) difficult to elucidate.

Of 1070 patients in the pooled magnesium group, 230 (21%) developed postoperative AF, compared with 307 of 1031 (30%) patients in the placebo or (control) group. Meta-analysis using a fixed-effects model generated a pooled odds ratio (OR) that was significantly less than 1.0 [OR = 0.65, 95% confidence interval (CI) 0.53 to 0.79, test for overall effect $p < 0.0001$], but with statistically significant heterogeneity ($I^2 = 63.4%$, $p = 0.0005$). Two RCTs were notable as they had relatively lower ORs in favour of magnesium sulphate. When these were removed from the analyses the pooled OR remained statistically significant, but

heterogeneity no longer remained significant. These two studies tended to impart a highly significant reduction in the odds of AF to whichever subgroup they were analysed in.

When studies were ordered by total duration of prophylaxis, an apparent relationship between duration and odds of AF was evident, with decreasing odds of AF as duration of prophylaxis increased. This was confirmed by linear regression analysis ($R^2 = 0.743$, $p < 0.001$). When the data were grouped into three classes according to whether duration of prophylaxis was 1 day or less, 2–4 days, or 5 days or greater, a statistically significant intervention effect was only present for the longest duration group (OR = 0.12, 95% CI 0.06 to 0.23, $p = 0.00001$).

Statistically significant intervention effects were associated with the initiation of prophylaxis 12 hours or more before surgery (OR 0.26; 95% CI 0.16 to 0.44, test for overall effect $p = 0.00001$, fixed-effects model) and less than 12 hours before surgery or during the surgery itself (OR = 0.73, 95% CI 0.56 to 0.97, test for overall effect $p = 0.03$, fixed-effects model), but not when prophylaxis was initiated at the end of surgery or postsurgery (OR = 0.85, 95% CI 0.59, 1.22, $p = 0.37$, fixed-effects model).

When studies were ordered by total dose of intravenous magnesium sulphate (<25 g), the odds of AF were independent of the dose. A notable exception was that for a total dose of 9 g magnesium sulphate, the odds of AF were significantly reduced relative to the control group, based on three studies that used this dose, including the two RCTs mentioned above that appeared to contribute to heterogeneity. This may be explained by the fact that each had excluded patients who were on antiarrhythmic drugs. They may have been at higher risk of AF compared with patients in other studies and, if so, might have benefited more from prophylactic magnesium. Within the subgroup of eight studies that maintained a constant dose rate there appears to be a relationship between the dose rate of magnesium sulphate and the odds of AF, with the largest prophylactic effects being seen at the lowest dose rates.

Results of the assessment of cost-effectiveness

Sixty-three potentially relevant references were found. No economic evaluations of intravenous

magnesium alone as prophylaxis against AF following CABG, compared with sotalol as prophylaxis or no prophylaxis, were identified. Four studies were included in the secondary review. One of the included studies was a report of an economic evaluation of oral amiodarone for prophylaxis against AF following CABG. The evaluation suggested that the principal determinant of the cost-effectiveness of prophylaxis against AF is likely to be the length of stay in the intensive care unit (ICU) and on hospital wards. A simple economic model, using a decision tree, was constructed. A flow diagram developed from this decision tree was assessed for its relevance to UK clinical practice and applicability to modelling the cost-effectiveness of magnesium sulphate prophylaxis. The diagram was taken to be a reasonable representation of current practice for patients developing AF following CABG, subject to modifications that would make it more consistent with current UK and European clinical guidelines.

Studies reporting resource use by patients with AF following CABG suggest that, while AF significantly increased inpatient stays, by up to 2.3 days in the ICU and 3.4 days on the ward, differences in length of stay and costs between patients receiving prophylaxis and those not receiving prophylaxis were not statistically significant. The lack of significant findings, with respect to differences in length of stay or cost, may reflect clinical trials being powered to detect differences in clinical outcome and not differences in resource use. However, the lack of significant differences may also reflect the fact that, since postoperative AF affects a minority of patients (albeit a large minority), the difference in resource use between patients with and without AF may be diluted when looking at mean values across a cohort of patients.

A simple economic model was developed to estimate the cost-effectiveness of magnesium sulphate prophylaxis against AF following CABG. This was populated with data on the baseline risk of AF following CABG and the relative risk of AF with magnesium sulphate prophylaxis from the meta-analysis, along with cost and resource-use data from published sources. In the base-case analysis, magnesium sulphate prophylaxis resulted in 0.081 fewer cases of AF at an incremental cost of £2.55. That is, the cost of prophylaxis was slightly higher than the expected savings due to reduced ICU and ward stays resulting from the reduction in AF cases. The incremental cost-effectiveness ratio (ICER) was £32 per AF case

avoided. The estimated difference in average length of stay between the prophylaxis and no-prophylaxis strategies was only 0.24 days, despite a large assumed difference of 3 days for patients experiencing AF in each group (1 extra day in the ICU and 2 extra days on the ward). In the deterministic sensitivity analysis the greatest variation in ICERs was observed for input parameters relating to the baseline risk of AF following CABG and the effectiveness of prophylaxis, the cost of prophylaxis and the resource consequences of postoperative AF. The largest ICER (£2092) in the sensitivity analysis was associated with increasing the length of patients' preoperative stay. In the base case it was assumed that admission routines would be identical under both strategies. However, patients receiving prophylaxis by intravenous infusion may have longer preoperative stays.

In the probabilistic analysis the majority of the simulations were associated with improved outcomes (in this case fewer cases of AF), but also higher costs. Prophylaxis was the dominant strategy (better outcome at lower cost) in about 41% of the simulations using the base-case assumptions. Under an alternative scenario where patients receiving prophylaxis are admitted for longer before their operation, to receive their initial infusion, the proportion of simulations where prophylaxis dominates falls to around 5%. Analysis using an acceptability curve showed that the probability of magnesium sulphate prophylaxis being cost-effective, compared with surgery with no prophylaxis, increases with willingness to pay (WTP) for a unit of outcome. The probability of being cost-effective was 99% at a WTP threshold of £2000 per AF case avoided and 100% at a WTP threshold of £5000 per AF case avoided under the base-case assumptions. Under the alternative scenario of longer preoperative stays the probability of being cost-effective at these two threshold values fell to 48% and 93%, respectively. It is unclear what the appropriate decision threshold should be, given that this model used intermediate rather than final outcomes.

Conclusions

No RCTs were identified that specifically aimed to compare intravenous magnesium with sotalol as prophylaxis for AF in patients undergoing CABG. Such a comparison does not appear to be clinically meaningful. Intravenous magnesium, compared with placebo or control, is effective in preventing postoperative AF, as confirmed by a statistically

significant intervention effect based on pooled analysis of 15 RCTs. It was also found that AF was less likely to occur when a longer duration of prophylaxis was used, and the earlier that prophylaxis is started; however, this finding was associated with two RCTs that had more favourable results than the other trials, but with no clear explanation as to why. No clear relationship between dose and AF was observed, although a lower constant dose rate was associated with the lowest odds of AF.

In the base-case analysis in the economic model, magnesium sulphate prophylaxis reduced the number of postoperative AF cases at a modest increase in cost. The results of the economic analysis are highly sensitive to variation in certain

key parameters. Prophylaxis is less likely to be a cost-effective option if it requires changes in admission routines that result in longer preoperative stays than would be the case without prophylaxis.

Recommendations for further research

Further research should investigate the relationship between dose, dose rate, duration of prophylaxis, timing of initiation of therapy and patient characteristics, such as degree of risk for AF. This will provide stronger evidence for the optimum delivery of intravenous magnesium in patients undergoing CABG.

Chapter I

Aim of the review

The aim of this research is to conduct a systematic review and economic evaluation of the clinical effectiveness and cost-effectiveness of magnesium sulphate compared with sotalol, and

to assess the clinical effectiveness of magnesium sulphate compared with placebo in the prevention of atrial fibrillation (AF) in patients who have had a coronary artery bypass graft (CABG).

Chapter 2

Background

Description of the underlying health problem

AF is a supraventricular arrhythmia characterised by uncoordinated atrial activation with consequent deterioration of atrial mechanical function, and is one of the most common complications after CABG. Between 20 and 40% of patients experience AF following cardiac or thoracic surgery.¹ AF usually occurs early in the postoperative period, with 70% of events developing within 4 days. However, AF sometimes occurs after discharge.²

During an episode of AF patients can experience symptoms such as palpitations, nausea and malaise. The management of AF aims to restore sinus rhythm through pharmacological or electrical cardioversion, and to reduce the risk of thromboembolism with the use of antithrombotic drugs (e.g. warfarin or aspirin). Rate control is another management goal, with the use of drugs such as digoxin (for non-acute episodes), β -blockers and rate-limiting calcium antagonists. Rate control has been shown to be as effective as rhythm control.³

AF increases the risk of mortality and morbidity from stroke, heart failure, myocardial infarction, thromboembolism and bleeding from anticoagulation.¹ This results in prolonged hospitalisation, hospital readmission, excess utilisation of hospital resources and increased health service costs.² Consequently, primary prevention of AF after CABG is of great importance.

The exact cause of AF after CABG is thought to be multifactorial. Risk factors include advanced age (particularly over the age of 50), previous history of AF, male gender, hypertension, diabetes, smoking, myocardial infarction and valvular heart disease.³ Magnesium is essential to the functioning of the cardiovascular system and patients with cardiac problems often exhibit abnormal magnesium metabolism. Cardiac surgical procedures may also cause rapid and acute changes in magnesium status.

Approximately 23,000 CABG operations are performed annually in England.⁴ For adults undergoing elective CABG the procedure may be on- or off-pump (i.e. the patient's circulation is, or is not, diverted through a pump oxygenator machine). During surgery the heart is beating when a patient is off-pump and can be either beating or artificially stopped when the patient is on-pump.

AF is detected in an ECG by the presence of rapid oscillations or fibrillatory waves that vary in size, shape and timing. A distinction is made between AF and atrial flutter, the latter being a more organised arrhythmia characterised by a sawtooth pattern of regular atrial flutter waves in the ECG.

Current service provision

European clinical guidelines recommend that β -blockers are used routinely as first choice for the prophylaxis of AF in all patients undergoing cardiac surgery.¹ β -blockers act as antiarrhythmic drugs principally by attenuating the effects of the sympathetic system on automaticity and conductivity within the heart.⁵ Commonly used β -blockers include metoprolol, bisoprolol and atenolol. Amiodarone, a class III antiarrhythmic drug, is recommended for all patients undergoing cardiac surgery in whom β -blocker therapy is not possible.

In 2006 the National Institute for Health and Clinical Excellence (NICE) issued a clinical guideline on the management of AF.⁶ In terms of prophylaxis of postoperative AF it recommends that one of the following drugs be used: amiodarone, a β -blocker, sotalol (class III β -blocker), or a rate-limiting calcium antagonist (e.g. verapamil or diltiazem). In addition, patients undergoing cardiac surgery on pre-existing β -blocker therapy should continue with this unless contraindications develop (e.g. bradycardia or hypotension). However, the role of magnesium was not assessed and therefore it is not covered by their recommendations.

Description of the intervention

Magnesium may be given to patients undergoing cardiothoracic surgery to reduce hypomagnesaemia, a common occurrence following this kind of surgery. Magnesium may exhibit antiarrhythmic activity through a number of mechanisms including inhibiting L-type calcium channels, which reduces sinus node rate firing.⁷ However, the mechanisms are not fully understood.⁸ Magnesium is not a first line choice for prophylaxis of AF and there appear to be few data on the extent to which it is used in current practice.

Magnesium may be used in combination with other drugs such as β -blockers (e.g. bisoprolol or sotalol).⁹ The drug is usually administered intravenously at a dose of 1–4 g in 10–20% solution at a rate not exceeding 1.5 ml of 10% solution, or equivalent per minute, or intravenous infusion of 4 g in 250 ml of 5% dextrose at a rate not exceeding 3 ml per minute. It may also be given as an intravenous bolus, intramuscularly or orally in the form of magnesium glycerophosphate. Prophylaxis with magnesium can begin before, during or after cardiothoracic surgery. The duration of prophylaxis varies from a matter of hours to several days.

Magnesium is well tolerated by patients and is unlikely to cause drug reactions such as plaques

associated with antiarrhythmic agent use and the side-effects of drowsiness and lethargy from using β -blockers. However, it has wide effects on basic biological mechanisms and is unlikely to be particularly targeted in action.

Sotalol hydrochloride (Beta-cardone[®], Celltech; Sotacor[®]; Bristol-Myers Squibb; non-proprietary) is a non-selective β -blocker with class III antiarrhythmic activity, used in the prophylaxis of paroxysmal supraventricular arrhythmias. By blocking the potassium channels, sotalol prolongs repolarisation, therefore lengthening the QT interval and decreasing automaticity.

Sotalol has an appreciable class III action only at high doses (240–480 mg per day). At low doses commonly prescribed in the UK (80–160 mg per day) the main antiarrhythmic effect is its class II (i.e. β -blocker) action. Side-effects may include ventricular proarrhythmias.² It is available orally for prophylaxis at a dose of 80 mg daily in one or two divided doses, increased gradually at intervals of 2–3 days to a usual dose of 160–320 mg daily in two divided doses.⁵ Sotalol via injection is not recommended routinely for prophylaxis. No estimates of the extent to which it is currently used were available, but expert opinion suggests that most clinicians use β -blockers other than sotalol.

Chapter 3

Methods of assessing clinical effectiveness

A systematic review was conducted according to the scope and methods outlined in the protocol issued in March 2007 (based on the HTA commissioning brief).

Inclusion and exclusion criteria

Populations

For studies to be included, patients had to be aged over 18 years, undergoing elective isolated CABG (either on-pump or off-pump, with any number of grafts or any conduit type). Studies were excluded if patients received other concomitant surgical procedures (e.g. valvular operations), unless the proportion of patients undergoing isolated CABG was also clearly documented. Studies in which methods of detecting AF were not specified, in which patients had a history of AF or history of AF was not reported, and/or studies in which length of follow-up was not reported were included in this review if they met the other inclusion criteria.

Interventions

Studies using intravenously administered magnesium sulphate alone, as either a bolus or a continuous infusion of clearly specified dosage and duration, given as prophylaxis before the onset of AF, were included. Studies of other magnesium compounds (e.g. chloride, hydroxide or unspecified) were excluded. Studies in which patients received other drugs (e.g. β -blockers) were included in this review only if the drugs were administered as usual patient care and did not differ between the randomised study groups.

Comparators

The study compared:

- intravenous magnesium sulphate versus different administration strategies of intravenous magnesium sulphate
- intravenous magnesium sulphate versus placebo/control
- intravenous magnesium sulphate versus sotalol.

Outcomes

The primary outcome was incidence of AF after CABG. Supraventricular arrhythmias other than

AF (e.g. tachycardias and atrial flutter) and all other non-atrial arrhythmias were excluded.

Study types

Randomised controlled trials (RCTs) were included. Systematic reviews were identified for context, but their results were not extracted and their methodological quality was not appraised.

Search strategy

A comprehensive search strategy was developed to identify RCTs and systematic reviews that have investigated the effects of prophylactic intravenous magnesium sulphate on AF after CABG. The search strategy (Appendix 1) aimed to identify systematically all relevant studies that met the inclusion criteria reported above. As the current review provides an update to a previous systematic review by Alghamdi and colleagues¹⁰ (as requested in the HTA commissioning brief), the searches for intravenous magnesium versus placebo/control were limited to studies published after 2003 (the cut-off date for literature searching in Alghamdi's systematic review being December 2003). No date restriction was applied to searches for sotalol studies.

The search strategy was applied to the following general and specialist health and biomedical databases (Appendix 1): Cochrane Library (2007 Issue 2), Ovid MEDLINE® (1950 to May 2007 week 1), Ovid MEDLINE® In-process and Other Non-indexed Citations, EMBASE (1980 to 2007 week 19), Database of Abstracts of Reviews of Effectiveness (DARE), Health Technology Assessment (HTA) Database, Centre for Reviews and Dissemination (CRD) database, National Research Register (NRR), Current Controlled Trials, including the Medical Research Council (MRC) Trials Database (controlled-trials.com) and Clinical Trials (clinicaltrials.gov). Each database was searched once, during 14–17 May 2007.

Study inclusion

All references identified by the literature searches were imported into a Reference Manager

bibliographic database. After duplicate references had been deleted from the database, the title and (where available) abstract of each reference were screened systematically against the inclusion criteria reported above, to assess the relevance of the study for inclusion in the review. This initial screening step was carried out by a reviewer using a standard decision tree (Appendix 2). Cases of uncertainty were resolved by discussion with a second reviewer. For those references that did not fulfil the inclusion criteria (owing to inappropriate population, study design, intervention or outcome) the reason for exclusion was recorded. Full papers for those references that appeared relevant on the basis of title and (where available) abstract were retrieved. Full papers were checked for their relevance using the same decision tree independently by two independent reviewers. Any disagreements between the reviewers were resolved by discussion and, if necessary, consultation with a third reviewer. Reasons for excluding papers at this stage were also recorded. Reference lists of relevant systematic reviews that were identified using the search strategy were checked for additional relevant literature not identified in the systematic searches.

Data extraction

Data were extracted from the included studies using a predesigned and piloted data extraction template to report the study design, patient populations, interventions, outcomes, analyses and any study limitations (Appendix 3). Data were extracted for incidence of AF and other relevant AF outcomes (e.g. time to onset of AF). If available, data on the patients' length of stay in the intensive care unit (ICU), postoperative stay and total hospital stay were also extracted. Data from each study were extracted by one reviewer and checked by a second reviewer. Any disagreements were resolved by consensus, if necessary involving a third reviewer.

Quality assessment

The methodological quality of each of the included studies was assessed systematically according to guidelines provided by the NHS CRD.¹¹ Study quality was assessed independently by two reviewers and reported in *Table 2* for each of seven criteria [randomisation, treatment allocation concealment, homogeneity of patient populations, blinding of outcome assessors, presentation of outcome data, intention-to-treat

analysis (ITT), and description of withdrawals and dropouts]. Any disagreements in quality classification were resolved by consensus, if necessary involving a third reviewer.

Data synthesis

Two forms of synthesis were conducted, a narrative synthesis and quantitative meta-analysis. Both forms of analysis were conducted according to standard principles, and using accepted methods.^{11,12} The key characteristics of the included studies (e.g. details of the study populations, intervention characteristics and outcomes measured) and the results (e.g. incidence of AF, length of hospital stay and adverse events) were summarised narratively and tabulated.

The feasibility and appropriateness of meta-analysis were considered once narrative synthesis had been completed. Meta-analysis to quantify the effects of prophylaxis on incidence of AF was performed using Cochrane Review Manager Software (RevMan, version 4.2). A similar approach to meta-analysis used by Alghamdi and colleagues was followed here.¹⁰ The proportion of patients experiencing AF in the intervention and comparator groups and the total number of patients in each study group were entered into RevMan. Data for the ITT population were entered where available; however, in the majority of studies it was not clear whether a true ITT analysis had been performed owing to poor reporting. Study authors were not routinely contacted to supply missing data or to clarify their analysis (with the exception of Nurözler and colleagues,²⁵ see section 'Subgroup analyses: dose', p. 19).

A fixed-effects analysis was performed, with random-effects analysis reserved for cases where statistical heterogeneity could not be explained. Heterogeneity was defined by a statistically significant χ^2 test ($p > 0.10$) and quantified by the I^2 statistic, where a figure greater than 50% indicates substantial heterogeneity.¹²

Sensitivity analyses were performed in cases of statistical heterogeneity (e.g. based on study methodological quality). Subgroup analyses were performed to assess the effects of different delivery strategies for intravenous magnesium, including different total doses, timing of the initiation of prophylaxis and total duration of prophylaxis. These subgroups were defined *a priori*

in the research protocol, based on the HTA commissioning brief for this assessment which requested analysis of optimal delivery strategies. For each subgroup, trials were assigned to mutually exclusive categories for analysis. (The categories themselves were not defined in the protocol, but were devised by the authors during the course of the systematic review.) For the total duration of infusion the categories were: 1 day or less, 2–4 days, and 5 days or more. The studies were also ordered by duration, from shortest to longest. For timing of the initiation of prophylaxis the categories were: 12 hours or more presurgery, less than 12 hours presurgery or during surgery, and termination of surgery or immediately postsurgery. The total dose of magnesium was not

split into separate subgroups as it was considered that defining subgroups on the basis of dose thresholds would be an arbitrary process. Expert clinical opinion suggested that there is no consensus over high or low dose thresholds. Alternatively, all doses were displayed graphically in a forest plot from lowest to highest to permit examination of the relationship between dose and odds of AF. A fourth delivery strategy specified in the research protocol, bolus versus intravenous infusion, was not assessed as only two trials were included that reported administering magnesium via bolus.

Effect measures were expressed as the odds ratio (OR) with 95% confidence intervals (CIs).

Chapter 4

Clinical effectiveness results

Quantity and quality of research available

The search strategy (Appendix 1), together with the previous systematic review by Alghamdi and colleagues,¹⁰ yielded 204 potentially relevant references. Two additional relevant references^{13,14} were identified during checking of the reference lists of the retrieved systematic reviews. Of these initial 206 papers, 158 were excluded on screening of titles and abstracts and a further 26 were subsequently excluded on screening of the full papers (Appendix 4). The remaining 22 papers that met the inclusion criteria for this review comprised 17 that reported parallel-group RCTs (15 trials altogether)^{13–29} and five systematic reviews.^{7–10,30}

Systematic reviews

As mentioned earlier, the current systematic review updates a systematic review published by Alghamdi and colleagues (2005).¹⁰ The inclusion criteria for the current review were based on those used by Alghamdi and colleagues.¹⁰ Seven of the eight studies included in that systematic review are included here. (NB. The additional study³¹ appears to have been erroneously included by Alghamdi and colleagues; it evaluated magnesium chloride, yet their inclusion criteria specified only magnesium sulphate, hence it is not included in the current report.) Note that of the additional eight studies included here, three were new studies^{16,22,29} published since Alghamdi and colleagues' review, and five were studies published before 2003, of which two were not found by Alghamdi and colleagues but identified by the present authors' search of reference lists of other systematic reviews^{13,14} and three were studies that Alghamdi and colleagues excluded.^{15,24,26} Of these three exclusions, two were rejected because the primary outcome measure was a broader measure of arrhythmia including other forms of arrhythmia as well as AF (e.g. atrial flutter or atrial tachycardia).^{15,24} However, these studies are included in the current report as it was discernible from the published papers that the proportion of arrhythmic events that were AF was relatively high

(e.g. greater than 90%). The third study²⁶ was excluded by Alghamdi and colleagues because their inclusion criteria prohibited studies in which patients had a prior history of AF. However, it is included in the current review as that particular study excluded patients with chronic AF, and the proportion of patients in the trial with previous non-chronic AF was relatively low (less than 10%). It should be pointed out that some of the studies included here and in the review by Alghamdi and colleagues¹⁰ did not report whether or not patients had experienced previous AF, and consequently it cannot be guaranteed that all patients in these trials were being treated for first occurrence of AF.

Four other relevant published systematic reviews were identified from the literature search.^{7–9,30} The systematic reviews varied slightly in terms of inclusion criteria and when they were conducted, meaning that each contained a different set of studies. The systematic reviews with the highest degree of overlap with the current systematic review were by Burgess and colleagues (2006),³⁰ Miller and colleagues (2005)⁷ and Shiga and colleagues (2004),⁸ which included up to 11 of the 15 studies included here. These reviews included other studies that are excluded from the current review, reflecting differences in inclusion criteria. For example, some of the reviews included studies of magnesium chloride.

One of the key differences between the current review and the five published systematic reviews is that none of those reviews specifically aimed to assess the role of delivery characteristics in the effectiveness of magnesium prophylaxis, with the exceptions of Henyan and colleagues⁹ and Miller and colleagues,⁷ who performed subgroup analyses based on dose and timing of the initiation of therapy. Another key difference is that none of the five published systematic reviews aimed to assess the effectiveness of magnesium sulphate compared with sotalol. The systematic review by Burgess and colleagues³⁰ also included a set of studies comparing sotalol to placebo, and a Cochrane systematic review has also compared sotalol with placebo;³² however, such a comparison is outside the scope of the current review.

In summary, this systematic review updates and expands a previous review by Alghamdi and colleagues,¹⁰ adding in three newly published RCTs and a further five older RCTs that were not previously included. There is some degree of overlap with other published systematic reviews, but none specifically aimed to assess the role of intravenous magnesium delivery or to compare intravenous magnesium with sotalol.

Characteristics of the included RCTs

Most of the trials included in the current report were conducted in Turkey (five trials^{14,17,18,23,25,27}) and the USA (four trials^{15,16,19,22}), with single trials also in Canada,²⁴ Finland,²⁶ Israel¹³ and Switzerland²⁸. The number of patients randomised to each intervention ranged from 15 to 176 (Table 1). In nine trials the comparator was explicitly reported as a placebo group, whereas in six trials it was referred to as a control group. One of the trials, by Forlani and colleagues,^{20,21} compared magnesium and sotalol in combination with a control group, and against magnesium alone and sotalol alone (Table 1). According to the authors the study was not designed to compare

magnesium against sotalol directly, and no statistical tests were reported for the comparison of the two drugs.

The standard of reporting in these trials was generally poor. In the majority of studies the adequacy of randomisation, concealment of treatment allocation, blinding of assessors and analysis of missing data could not be elucidated (Table 2). Although most of the studies reported baseline characteristics of their patient populations, the quantity of information given and whether differences between groups were tested statistically were highly variable among the studies. None of the studies reported the ethnic or socio-demographic status of their recruited patients. Reporting of the methods of detecting AF and the definitions of AF used in the studies were variable (Appendix 5).

Assessment of effectiveness

Incidence of AF

As mentioned earlier, only one RCT included both a sotalol and an intravenous magnesium arm.^{20,21} However, the trial was designed to compare the clinical efficacy of combined sotalol and

TABLE 1 Characteristics of 15 RCTs that met the criteria for inclusion in this review

Study	Population	Intervention	N	Comparator	N	Outcomes			
						AF	Length of stay		
							ICU	Postop.	Hospital
Bert <i>et al.</i> , 2001 ¹⁵	CABG + V	Mg	63	Control	60	(✓)			✓
Bhudia <i>et al.</i> , 2006 ¹⁶	CABG + V	Mg	174	Placebo	176	✓	✓	✓	✓
Caspi <i>et al.</i> , 1995 ¹³	CABG ^a	Mg	50	Placebo	48	✓			
Dagdelen <i>et al.</i> , 2002, ¹⁷ 2003 ¹⁸	CABG	Mg	93	Control	55	✓			
Fanning <i>et al.</i> , 1991 ¹⁹	CABG	Mg	49	Placebo	50	(✓)			
Forlani <i>et al.</i> , 2002, ²⁰ 2003 ²¹	CABG	Mg	54	Control	50	✓		✓	
		Sotalol	51						
		Mg + sotalol	52						
Hazelrigg <i>et al.</i> , 2004 ²²	CABG	Mg	105	Control	97	✓	✓		✓
Kaplan <i>et al.</i> , 2003 ²³	CABG	Mg	100	Placebo	100	✓	✓		✓
Karmy-Jones <i>et al.</i> , 1995 ²⁴	CABG + V	Mg	46	Placebo	54	(✓)	✓		✓
Nurözler <i>et al.</i> , 1996 ²⁵	CABG	Mg ^b	25	Placebo	25	✓			
Parikka <i>et al.</i> , 1993 ²⁶	CABG	Mg	69	Placebo	71	✓			
Toraman <i>et al.</i> , 2001 ²⁷	CABG	Mg	100	Control	100	✓	✓	✓	
Treggiari-Venzi <i>et al.</i> , 2000 ²⁸	CABG	Mg	49	Placebo	53	✓	✓		
Yilmaz <i>et al.</i> , 2000 ¹⁴	CABG ^a	Mg ^c	15	Control	15	✓			
Zangrillo <i>et al.</i> , 2005 ²⁹	CABG	Mg	80	Placebo	80	✓	✓		✓

Unless stated otherwise, CABG was isolated (^anot reported whether CABG was isolated).

Hospital, total hospital time; ICU, time in intensive care unit; Mg, magnesium sulphate given by intravenous delivery (^bunclear whether all magnesium administered was given as sulphate; ^cnot reported whether delivery was intravenous); Postop, total postoperative time; V, valvular surgery; ✓, directly reported; (✓), inferred indirectly.

TABLE 2 Quality of the RCTs assessed systematically according to the criteria of the NHS CRD

Study	Randomisation	Concealment of allocation	Baseline characteristics	Blinding of assessors	Primary outcome data	ITT analysis	Missing values
Bert <i>et al.</i> , 2001 ¹⁵	Adequate	Unknown	Reported	Adequate	Adequate	Inadequate	Partial
Bhudia <i>et al.</i> , 2006 ¹⁶	Unknown	Adequate	Reported	Unknown	Adequate	Inadequate	Unknown
Caspi <i>et al.</i> , 1995 ¹³	Unknown	Unknown	Reported	Adequate	Partial	Inadequate	Unknown
Dagdelen <i>et al.</i> , 2002, ¹⁷ 2003 ¹⁸	Unknown	Unknown	Reported	Adequate	Adequate	Inadequate	Unknown
Fanning <i>et al.</i> , 1991 ¹⁹	Unknown	Unknown	Reported	Unknown	Adequate	Inadequate	Unknown
Forlani <i>et al.</i> , 2002, ²⁰ 2003 ²¹	Adequate	Unknown	Reported	Unknown	Adequate	Inadequate	Inadequate
Hazelrigg <i>et al.</i> , 2004 ²²	Inadequate	Inadequate	Reported	Inadequate	Partial	Inadequate	Inadequate
Kaplan <i>et al.</i> , 2003 ²³	Inadequate	Inadequate	Reported	Inadequate	Adequate	Inadequate	Inadequate
Karmy-Jones <i>et al.</i> , 1995 ²⁴	Unknown	Unknown	Reported	Adequate	Partial	Inadequate	Unknown
Nurözler <i>et al.</i> , 1996 ²⁵	Unknown	Unknown	Reported	Unknown	Adequate	Inadequate	Unknown
Parikka <i>et al.</i> , 1993 ²⁶	Unknown	Unknown	Reported	Unknown	Adequate	Inadequate	Adequate
Toraman <i>et al.</i> , 2001 ²⁷	Inadequate	Inadequate	Reported	Adequate	Adequate	Inadequate	Inadequate
Treggiari-Venzi <i>et al.</i> , 2000 ²⁸	Partial	Adequate	Reported	Unknown	Adequate	Inadequate	Inadequate
Yilmaz <i>et al.</i> , 2000 ¹⁴	Unknown	Unknown	Unknown	Unknown	Adequate	Inadequate	Unknown
Zangrillo <i>et al.</i> , 2005 ²⁹	Adequate	Unknown	Reported	Inadequate	Adequate	Adequate	Adequate
Overall (modal class)	Unknown	Unknown	Reported	Unknown	Adequate	Inadequate	Unknown

intravenous magnesium against the two agents separately, and against a control. No statistical tests were performed for the comparison between sotalol and intravenous magnesium. In the study postoperative AF occurred in 12% of patients receiving sotalol and in 15% of patients receiving magnesium.

No RCTs were identified that compared different delivery strategies of intravenous magnesium. The remainder of this section therefore focuses on the results of the RCTs that compared intravenous magnesium with placebo or control.

Of 1070 patients in the pooled magnesium group, 230 (21%) developed postoperative AF, compared with 307 of 1031 (30%) patients in the control (or placebo) group. *Figure 1* shows the forest plot of the meta-analysis of all 15 RCTs on the incidence of AF. Initial analysis using a fixed-effect model gave a pooled OR significantly less than 1.0 (OR = 0.65, 95% CI 0.53 to 0.79, test for overall effect $p < 0.001$), but with statistically significant heterogeneity ($I^2 = 63.4%$, $p = 0.0005$).

All individual studies with an OR that was significantly different from 1.0, and those with an OR bordering on significance (whose CI only just included 1.0) were in favour of magnesium sulphate. The overall effect favouring magnesium was driven by a relatively small proportion of the studies, with the majority (nine out of 15) exhibiting ORs very close to 1.0. Two RCTs, by Dagdelen and colleagues (2002)¹⁷ and Toraman and colleagues (2001),²⁷ are notable as they had

lower ORs in favour of magnesium sulphate than all the other studies.

Sensitivity analyses

Sensitivity analyses were performed to explore the likely reasons for heterogeneity. When the analysis was restricted to the two RCTs that were judged to have adequately concealed random allocation^{16,28} (*Table 2*), the intervention effect was no longer statistically significant (OR = 0.84, 95% CI 0.57 to 1.26, test for overall effect $p = 0.40$), although heterogeneity was absent ($I^2 = 0%$) and no longer significant ($p = 0.92$). A similar pattern was evident when analyses were restricted to the three RCTs whose randomisation procedures were judged to be adequate^{15,20,21,29} (*Table 2*) or the five RCTs whose blinding was considered adequate^{13,15,17,18,24,27} (forest plots not shown). Given that the number of studies meeting the criteria of methodological adequacy for these attributes was comparatively low, these results should be interpreted with caution.

In a meta-analysis of eight RCTs, which included seven of the RCTs included in the current report, Alghamdi and colleagues¹⁰ also identified a statistically significant intervention effect and statistically significant heterogeneity. When their analysis was restricted to the six highest quality studies scored using the Van Tulder methodological quality assessment scale, their intervention effect remained statistically significant, but heterogeneity was reduced and no longer statistically significant. The two studies that were excluded were Dagdelen (2002)¹⁷ and Toraman (2001).²⁷ These two trials

TABLE 3 Dosage and duration of magnesium sulphate administration in the intervention group (indicated only where this differed from the control or placebo comparator group)

Study	Method of delivery	Time	MgSO ₄ dose or concentration	Carrier	Duration
Bert et al., 2001 ¹⁵	i.v. infusion	1. After termination of CPB	2 g	50 ml normal saline	30 minutes
		2. On arrival in ICU	2 g	50 ml normal saline	30 minutes
		3. Each morning for first 4 days postsurgery	2 g × 4	50 ml normal saline × 4	30 minutes × 4
		<i>Overall total</i>	<i>12 g</i>	<i>300 ml normal saline</i>	<i>3 h in 5 days</i>
Bhudia et al., 2006 ¹⁶	CPB circuit i.v.	1. During CPB	3.6 to 4.8 mg dl ⁻¹	–	–
		2. During anaesthesia induction	0.78 g (32 mmol)	100 ml normal saline	15 minutes
		3. After anaesthesia induction	3.16 g (130 mmol)	100 ml normal saline	24 h
		<i>Overall total</i>	<i>3.94 g (162 mmol) + CPB conc.</i>	<i>200 ml normal saline + CPB prime</i>	<i>24.25 h</i>
Caspi et al., 1995 ¹³	i.v. via syringe pump	1. During interval between induction of anaesthesia and aortic cross-clamp	16 mmol	20 ml saline	–
		i.v.	2. After release of cross-clamp	32 mmol	20 ml saline
	<i>Overall total</i>	<i>48 mmol</i>	<i>40 ml saline</i>	<i>~24 h</i>	
Dagdelen et al., 2002, ¹⁷ 2003 ¹⁸	i.v. infusion	1. 1 day presurgery	1.5 g	100 ml 0.9 NaCl solution (25 ml h ⁻¹)	1 day
		2. Just after surgery	1.5 g	100 ml 0.9 NaCl solution (25 ml h ⁻¹)	1 day
		3. Once daily for 4 days postsurgery	1.5 g × 4	100 ml 0.9 NaCl solution (25 ml h ⁻¹) × 4	4 days
		<i>Overall total</i>	<i>9 g</i>	<i>600 ml 0.9 NaCl solution</i>	<i>6 days</i>
Fanning et al., 1991 ¹⁹	i.v. infusion	1. First 24 h postsurgery	40 mEq l ⁻¹ at 100 ml h ⁻¹ (total 96 mEq at 4 mEq h ⁻¹)	5% dextrose in water + 20 mEq potassium chloride l ⁻¹ at 100 ml h ⁻¹	1 day
		2. Next 72 h postsurgery (25–96 h)	40 mEq l ⁻¹ at 25 ml h ⁻¹ (total 72 mEq at 1 mEq h ⁻¹)	5% dextrose in water + 20 mEq potassium chloride l ⁻¹ at 25 mL h ⁻¹	3 days
		<i>Overall total</i>	<i>168 mEq</i>	<i>5% dextrose in water + 80 mEq potassium chloride</i>	<i>4 days</i>

continued

TABLE 3 Dosage and duration of magnesium sulphate administration in the intervention group (indicated only where this differed from the control or placebo comparator group) (cont'd)

Study	Method of delivery	Time	MgSO ₄ dose or concentration	Carrier	Duration
Forlani et al., 2002, ²⁰ 2003 ²¹ Magnesium group	i.v.	From just before CPB until 5 days postsurgery	1.5 g (12 mEq) day ⁻¹ × 6	NR	6 days
		<i>Overall total</i>	9 g (72 mEq)	–	6 days
Forlani et al., 2002, ²⁰ 2003 ²¹ Magnesium + sotalol group	i.v. magnesium + oral sotalol	From just before CPB until 5 days postsurgery	As above plus sotalol 80 mg orally twice daily	NR	6 days
		<i>Overall total</i>	9 g (72 mEq) MgSO ₄ + 960 mg sotalol	–	6 days ^a
Hazelrigg et al., 2004 ²²	Bolus (no other details) i.v. drip infusion	1. 30 minutes before CPB	80 mg kg ⁻¹ (ideal body weight)	100 ml of D ₅ W	30 minutes
		2. Subsequent 48 h	8 mg kg ⁻¹ (ideal body weight) h ⁻¹	100 ml of D ₅ W	48 h
		<i>Overall total</i>	464 mg kg ⁻¹ (ideal body weight)	200 mL of D ₅ W	48.5 h
Kaplan et al., 2003 ²³	i.v. infusion	1. 12 h presurgery	3 g (24.34 mEq) at 50 ml h ⁻¹	100 ml saline at 50 ml h ⁻¹	2 h
		2. After termination of CPB or anastomosis of last graft	3 g (24.34 mEq) at 50 ml h ⁻¹	100 ml saline at 50 ml h ⁻¹	2 h
		3. First postsurgery dose in ICU	3 g (24.34 mEq) at 50 ml h ⁻¹	100 ml saline at 50 ml h ⁻¹	2 h
		4. 1–3 days postsurgery	3 g (24.34 mEq) at 50 ml h ⁻¹ daily × 3	100 ml saline at 50 ml h ⁻¹	2 h × 3
		<i>Overall total</i>	18 g (146.04 mEq)	600 ml saline	12 h in 4.5 days
Karmy-Jones et al., 1995 ²⁴	i.v.	1. At termination of CPB	2.4 g (19.2 mEq)	50 mL of D ₅ W	20 minutes
		2. Every 4 h for a further five doses	2.4 g (19.2 mEq) × 5	50 ml of D ₅ W	100 minutes
		<i>Overall total</i>	14.4 g (115 mEq)	100 mL of D ₅ W	2 h in 20.3 h
Nurözler et al., 1996 ²⁵	i.v. infusion	1. In cardioplegia ^b	16 mmol l ^{-1b}	NR	NR
		2. First postoperative day	100 mEq	1000 ml of D ₅ W day	1 day
		3. Postoperative days 2–5	25 mEq day ⁻¹ × 4	1000 ml of D ₅ W day ⁻¹ × 4	4 days
		<i>Overall total</i>	200 mEq	5 litres of D ₅ W	5 days
Parikka et al., 1993 ²⁶	i.v. infusion	1. First 24 h after surgery (started within 2 h of operation)	40 mmol	1 litre of 5% glucose solution	1 day
		2. Next 24 h after surgery	30 mmol	500 ml of 5% glucose solution	1 day
		<i>Overall total</i>	70 mmol	1.5 l of 5% glucose solution	2 days

continued

TABLE 3 Dosage and duration of magnesium sulphate administration in the intervention group (indicated only where this differed from the control or placebo comparator group) (cont'd)

Study	Method of delivery	Time	MgSO ₄ dose or concentration	Carrier	Duration
Toraman <i>et al.</i> , 2001 ²⁷	i.v. infusion	1. One day presurgery	6 mmol	100 ml 0.9% NaCl (25 ml h ⁻¹)	1 day
		2. Just after CPB	6 mmol	100 ml 0.9% NaCl (25 ml h ⁻¹)	1 day
		3. Once daily for 4 days postsurgery	6 mmol × 4	100 ml 0.9% NaCl (25 ml h ⁻¹) × 4	4 days
		<i>Overall total</i>	<i>36 mmol</i>	<i>600 ml 0.9% NaCl</i>	<i>6 days</i>
Treggiari-Venzi <i>et al.</i> , 2000 ²⁸	i.v. infusion	For 72 h, starting within 1 h of arrival in ICU	4 g (16 mmol) (32 mEq) day ⁻¹ × 3	NR ^c	3 days
		<i>Overall total</i>	<i>12 g (48 mmol) (96 mEq)</i>	–	3 days
Yilmaz <i>et al.</i> , 2000 ¹⁴	Bolus (no other details)	At initiation of CPB ^d	0.4 mmol kg ⁻¹ (15%) ^e	NR	NR
		<i>Overall total</i>	<i>0.4 mmol kg⁻¹ (15%)^e</i>	–	–
Zangrillo <i>et al.</i> , 2005 ²⁹	Infusion	For 30 minutes immediately after central venous cannulation	2.5 g (20 mEq) ^f	100 ml of normal saline	30 minutes
		<i>Overall total</i>	<i>2.5 g (20 mEq)</i>	<i>100 ml of normal saline</i>	<i>30 minutes</i>

CPB, cardiopulmonary bypass; D₅W, 5 % dextrose water solution; NR, not reported.
^a Sotalol was subsequently continued at 40 mg orally twice daily for a further 4 weeks.
^b Magnesium compound and volume (litres) administered in cardioplegia not reported.
^c Noted that placebo was 0.9% NaCl.
^d Magnesium sulphate was also administered to both groups in the cardioplegia solution.
^e Parameter not stated (may refer to body weight or ideal body weight?).
^f 2.5 g magnesium sulphate was also administered to both groups in the ICU within the first 24 h postsurgery.

were conducted by the same team of investigators and appear to be similar in many intervention and study design characteristics.

In the current meta-analysis quality scores were not assigned to studies as this is not recommended practice.³³ Critical appraisal of the various aspects of the methodological quality of these two RCTs (Table 2) concluded that, in terms of key methodological attributes such as method of randomisation or concealment of allocation, either they were inadequate or details were not reported clearly enough to allow an informed judgement to be made (although it should be noted that these two studies are not atypical, as standards of reporting were generally poor in most trials). From visual inspection of Figure 1 these two studies appeared to contribute to the statistical heterogeneity, with relatively small ORs. Removing

them from the analysis meant that heterogeneity was no longer statistically significant ($p = 0.35$, $I^2 = 9.4\%$), but increased the pooled OR to 0.78 and widened the CIs (95% CI 0.63 to 0.97), although a statistically significant intervention effect remained ($p = 0.02$). The analyses presented in the next section illustrate that these two trials appear to be different from the other trials in all subgroup analyses.

Subgroup analyses: total duration of prophylaxis

The total duration of magnesium sulphate prophylaxis ranged from 30 minutes to 6 days (0.5 to 144 hours) (Table 3). When studies were ordered by duration of prophylaxis, an apparent relationship between duration of prophylaxis and odds of AF was evident, with decreasing odds of AF as duration of prophylaxis is increased

(Figure 2). Note that the forest plot displaying this relationship includes all studies except for one by Yilmaz and colleagues¹⁴ (for which the duration of prophylaxis was not reported), and consequently the heterogeneity and overall effect are similar to those reported above in Figure 1. As noted above, two studies, by Dagdelen^{17,18} and Toraman²⁷ warrant careful consideration because they appear to influence strongly the overall pattern in the forest plot. Both these studies excluded patients on β -blockers and might therefore have analysed populations who were at relatively high risk of AF, who might have been more likely to respond to prophylaxis for AF. A third study that warrants consideration is by Nurözler and colleagues.²⁵ In their study, the full dose and duration of magnesium sulphate administration are uncertain; as well as giving magnesium intravenously it appears to have been applied in the cardioplegia solution used during surgery, with unspecified details of dose (Table 3). Omitting these three studies from the meta-analysis subgroup would result in a forest plot that still appears to support the relationship between duration of prophylaxis and odds of AF (Figure 3). Excluding these studies

also considerably reduces the heterogeneity of the overall effect ($I^2 = 5.8\%$, $p = 0.39$, fixed-effects model).

The relationship between the duration of prophylaxis and the odds of AF appears approximately linear, irrespective of whether the OR is plotted on an arithmetic (Figure 4) or a logarithmic scale. A linear metaregression with \ln OR as the dependent variable (random-effects least-squares model in which studies are weighted according to the inter-trial and intra-trial variances) based on 14 studies (omitting only the study by Yilmaz,¹⁴ which did not report duration) would confirm the linear relationship ($R^2 = 0.743$, $p < 0.001$). The linear relationship would still be supported by this model if a further three studies by Dagdelen,^{17,18} Nurözler²⁵ and Toraman²⁷ (discussed above) were excluded ($R^2 = 0.704$, $p = 0.001$).

It is difficult to determine how (or indeed whether) the studies should be grouped for analysing effects of prophylaxis duration on odds of AF. If the data are grouped into three classes

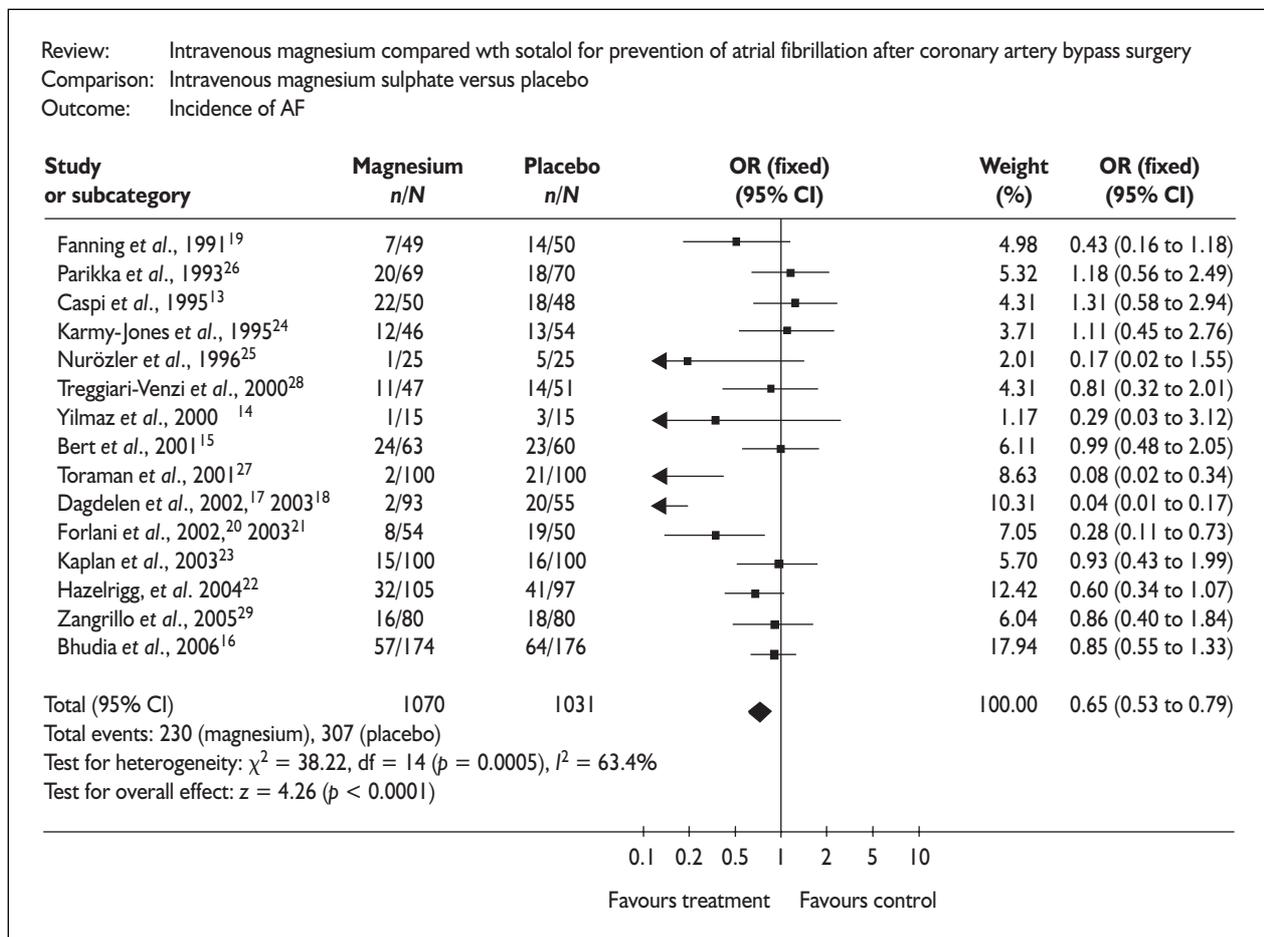


FIGURE 1 Meta-analysis of incidence of AF (fixed-effects model)

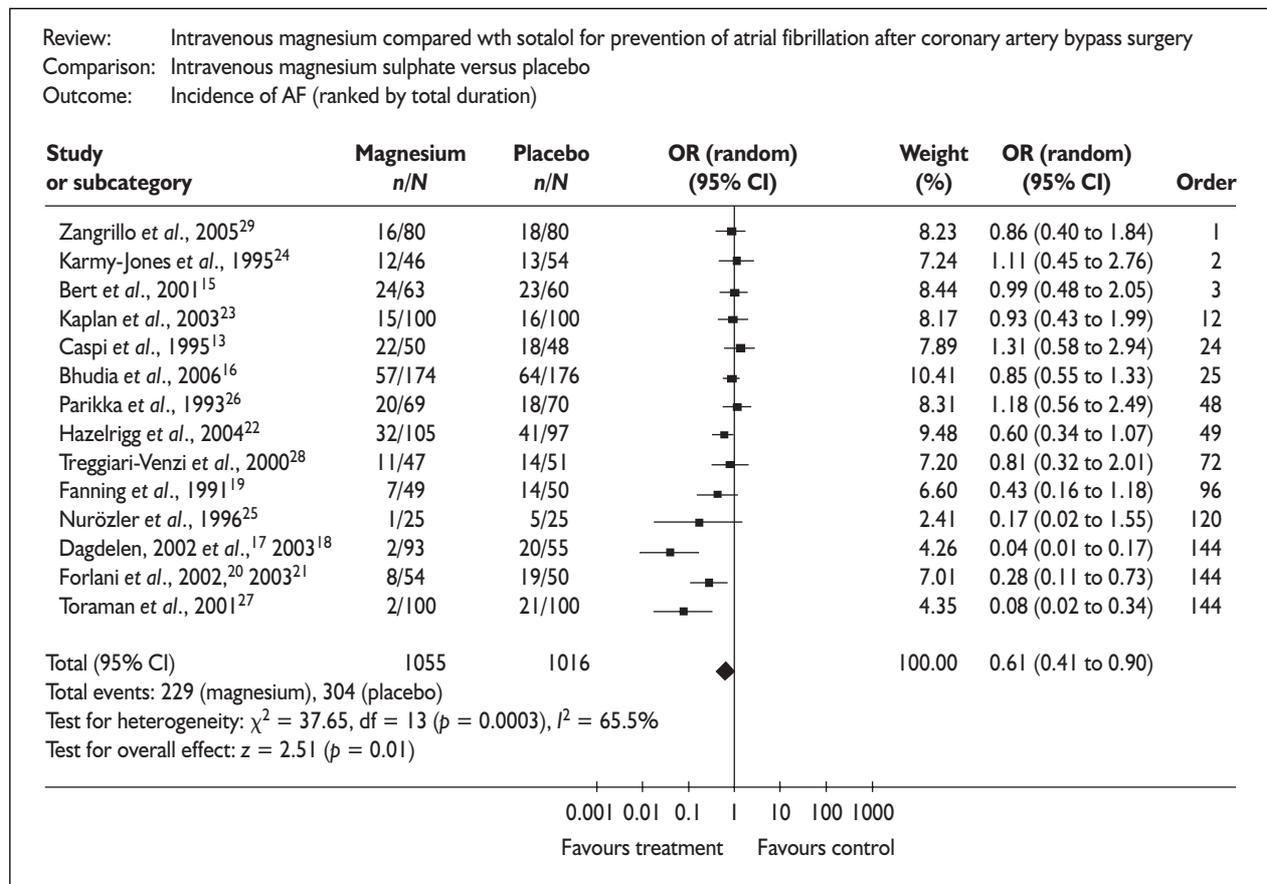


FIGURE 2 Studies ordered by duration of magnesium sulphate intervention (random-effects model). The data column 'Order' shows the duration of intervention (hours, rounded to the nearest 1 hour).

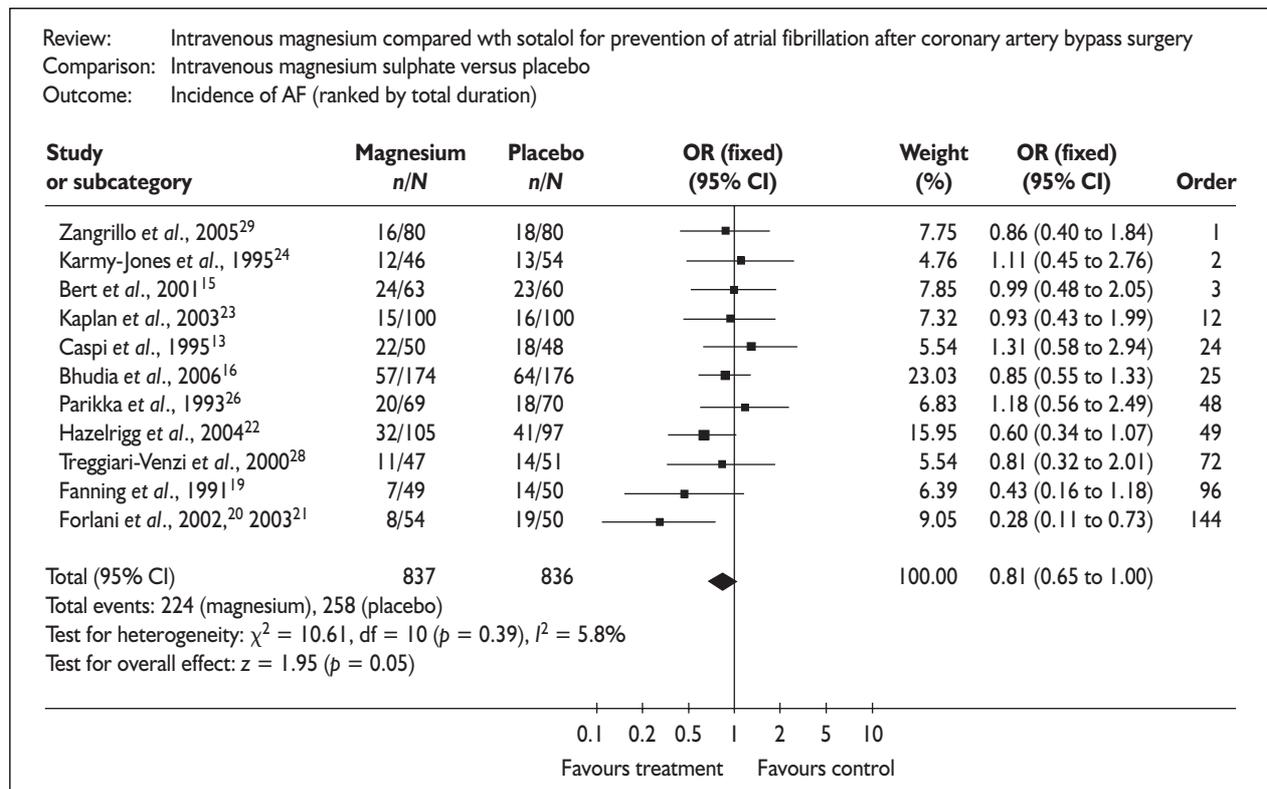


FIGURE 3 Studies ordered by duration of magnesium sulphate intervention, excluding studies by Dagdelen et al.,^{17,18} Nurözler et al.²⁵ and Toraman et al.²⁷ (fixed-effects model; other details as in Figure 2)

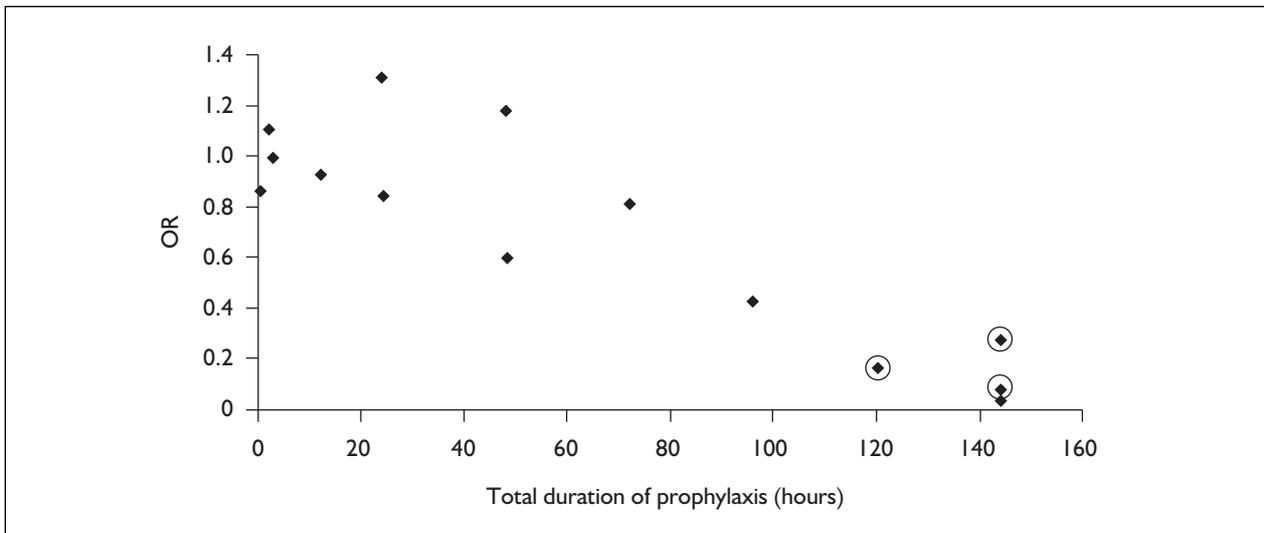


FIGURE 4 Relationship between the duration of magnesium sulphate prophylaxis and the odds of AF. Studies by Dagdelen,^{17,18} Nurözler²⁵ and Toraman²⁷ are encircled; a study by Yilmaz¹⁴ is omitted (see text).

according to whether duration of prophylaxis was 1 day or less, 2–4 days, or 5 days or greater, a statistically significant intervention effect is only present for the longest duration group (OR = 0.12, 95% CI 0.06 to 0.23, test for overall effect $p = 0.00001$) (forest plot not shown). In this subgroup heterogeneity is not statistically significant ($p = 0.12$), although the I^2 value (48%), is just below the suggested threshold of substantial heterogeneity (50%).¹² The effect for a duration of 5 days or greater would remain statistically significant after excluding the trials by Toraman²⁷ and Dagdelen^{17,18} (data not shown).

Subgroup analyses: timing of initiation of prophylaxis

Most of the studies did not report the exact timing of the start of magnesium prophylaxis, but the studies can be grouped according to whether prophylaxis commenced 12 hours or more presurgery (Figure 5), within 12 hours of surgery (including during surgery itself) (Figure 6) or at the termination of surgery or immediately postsurgery (Figure 7). For the first two subgroups there was a statistically significant intervention effect, while in the third there was no significant effect.

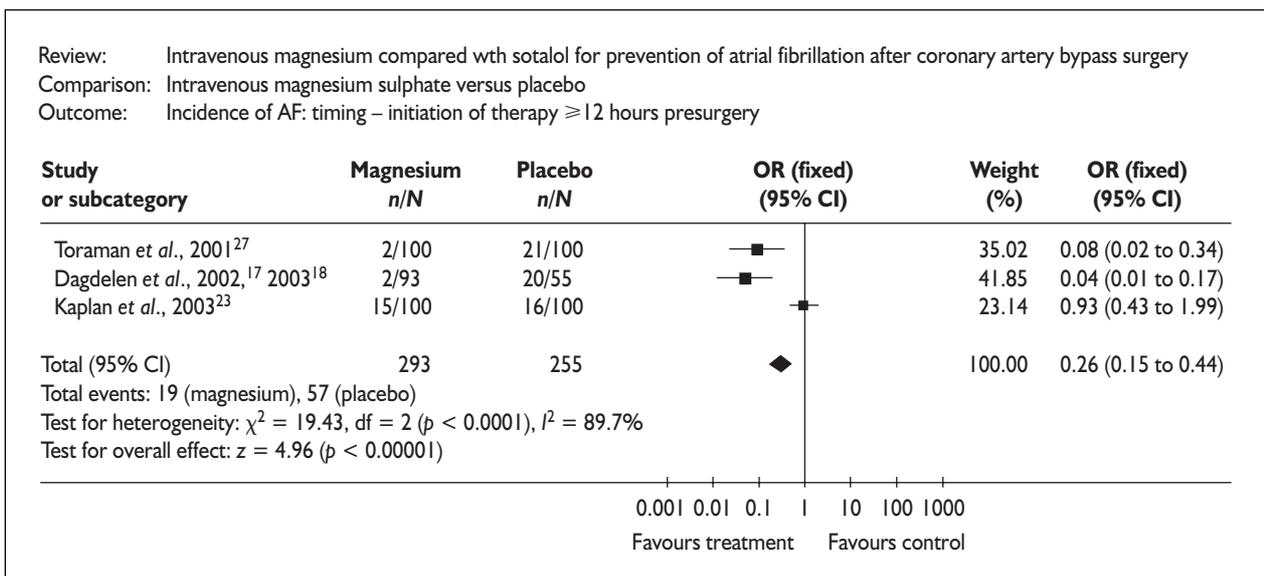


FIGURE 5 Subgroup analysis: timing of initiation of prophylaxis (greater than or equal to 12 hours presurgery) (fixed-effects model)

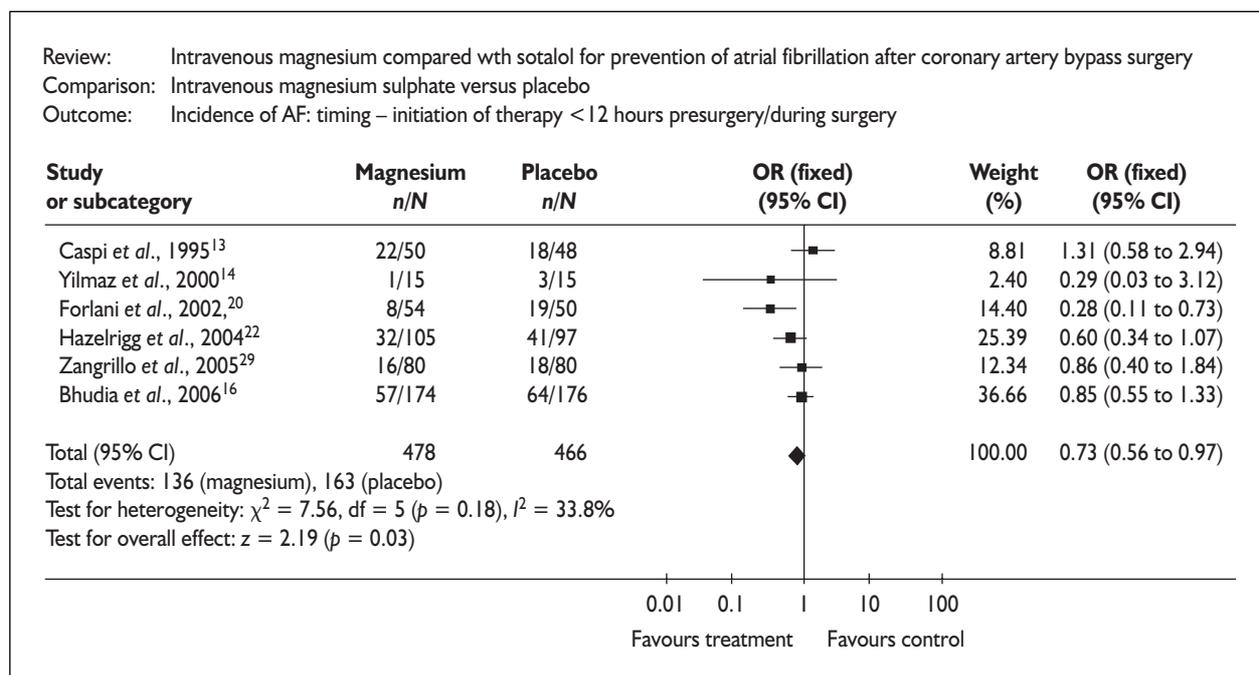


FIGURE 6 Subgroup analysis: timing of initiation of prophylaxis (less than 12 hours presurgery or during surgery) (fixed-effects model)

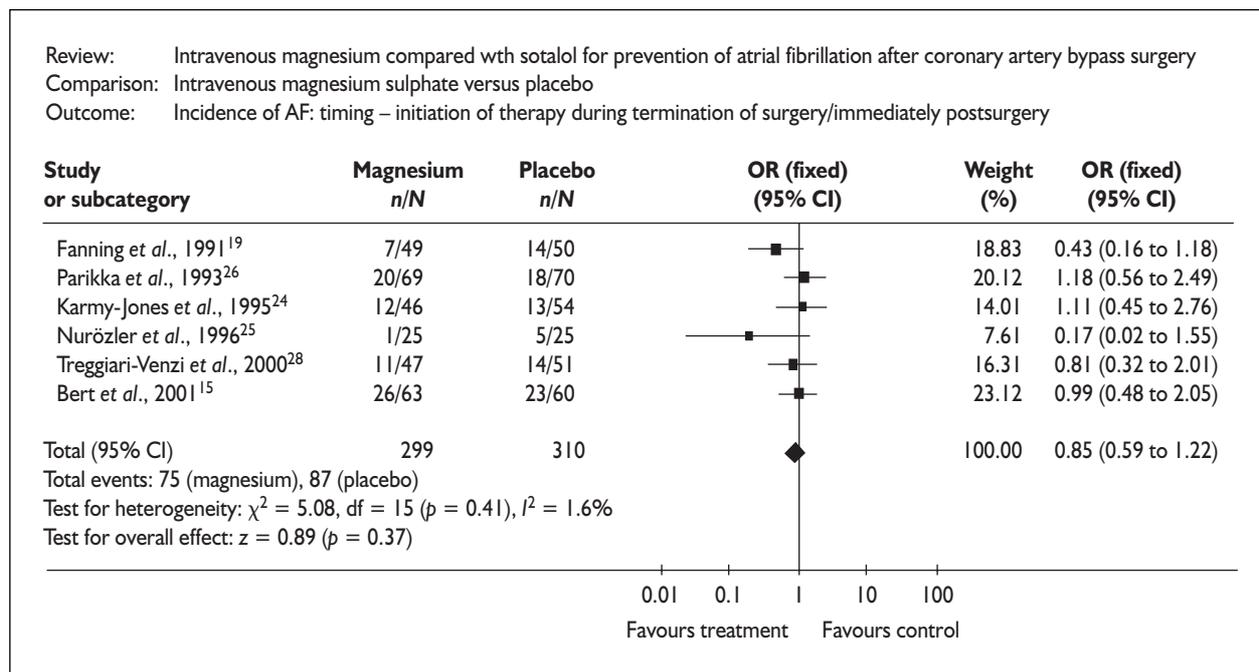


FIGURE 7 Subgroup analysis: timing of initiation of prophylaxis (at termination of surgery or immediately postsurgery) (fixed-effects model)

For the subgroup of studies that initiated prophylaxis 12 hours or more before surgery the OR was 0.26 (95% CI 0.15 to 0.44, test for overall effect $p = 0.00001$, fixed-effects model) (Figure 5). However, there was statistically significant heterogeneity ($I^2 = 89.7\%$, $p < 0.0001$). This is likely to be due to the different effects in the

Toraman²⁷ and Dagdelen^{17,18} trials, compared with the trial by Kaplan and colleagues.²³ When a random-effects model was used the OR was reduced to 0.15, but the CIs included 1.0 (95% CI 0.02 to 1.38, test for overall effect = 0.09) and the effect was therefore no longer statistically significant.

For studies that initiated prophylaxis less than 12 hours before surgery or during the surgery itself, there was a significant intervention effect (OR = 0.73, 95% CI 0.56 to 0.97, $p = 0.03$, fixed-effects model) without statistically significant heterogeneity ($I^2 = 33.8\%$, $p = 0.18$) (Figure 6).

The subgroup of studies that initiated prophylaxis at the end of surgery or postsurgery was not associated with a significant intervention effect (OR = 0.85, 95% CI 0.59 to 1.22, $p = 0.37$, fixed-effects model); heterogeneity was also not statistically significant in this subgroup ($I^2 = 1.6\%$, $p = 0.41$) (Figure 7).

Subgroup analyses: dose

In eight of the 15 trials the dose of magnesium sulphate was reported in grams. In five trials the dose in grams was converted by the reviewers from millimoles or milliequivalents, by assuming 1 g \equiv 4 mmol \equiv 8 mEq of magnesium sulphate (based on a conversion formula provided by Shiga and colleagues⁸). The total dose of magnesium sulphate administered in 13 of the trials ranged from 2.5 to 25 g (Table 3). In the remaining two trials, total doses of magnesium sulphate were given per kilogram of unspecified ideal body weight²² or per an unspecified parameter,¹⁴ precluding them from meta-analysis.

The dose–response relationship has not previously been characterised for effects of magnesium sulphate on AF. Meta-analysis (Figure 8) shows that for total doses of magnesium sulphate less than 21 g, the odds of AF were independent of the dose. A notable exception is that for a total dose of 9 g magnesium sulphate the odds of AF were significantly reduced relative to the control group (OR = 0.12, test for overall effect $p < 0.00001$, heterogeneity $p = 0.06$). Previous studies have identified serum magnesium concentration as a predictor of AF, but it is unclear why a profound prophylactic effect of magnesium sulphate would be found only at a total dose of 9 g. It is notable that the three studies that administered magnesium sulphate at the total dose of 9 g (Dagdelen,^{17,18} Forlani^{20,21} and Toraman²⁷) had each excluded patients who were on antiarrhythmic drugs (Appendix 6). As mentioned previously, the patients included in these studies might have been at higher risk of AF compared with patients in other studies and, if so, might have benefited more from prophylactic magnesium. The studies by Dagdelen^{17,18} and Toraman²⁷ in particular tend to impart a highly significant reduction in the odds of AF to whichever subgroup these studies are analysed in (see above).

If the study by Nurözler and colleagues²⁵ is included in the meta-analysis, the forest plot (Figure 8) would suggest a possible dose–response relationship, as the highest doses tested (21 and 25 g) are associated with a reduction in the odds of AF, although for these individual doses the CIs of the OR include 1 ($p > 0.05$). Repeating the meta-analysis with the two highest doses grouped together ($n = 2$) would yield a significant OR for this high-dose (≥ 21 g) group (OR = 0.35, 95% CI 0.14 to 0.88, test for overall effect $p = 0.02$, heterogeneity $p = 0.45$; forest plot not shown). The Nurözler and colleagues study is problematic, however, because in addition to the total of 200 mEq magnesium sulphate that they administered intravenously, they gave an unspecified dose of an unspecified magnesium compound only to the intervention group, during cardioplegia. This effectively confounds their intravenous magnesium sulphate dose with an unknown variable. Excluding the Nurözler and colleagues data from the forest plot is therefore advisable unless the details of how magnesium was administered in cardioplegia can be clarified. (The authors were contacted during preparation of the current review, but no response was received at the time of submission of this report.) Excluding the Nurözler and colleagues study would weaken the support for a dose–response relationship.

A potential limitation of comparing studies in terms of the total magnesium sulphate dose administered is that total dosage does not take into account the dose rate, which may have clinical relevance. An analysis of whether the dose rate of magnesium sulphate affects the incidence of AF was possible for eight trials in which the dose rate was kept constant (Table 4). In the remaining seven trials either the dose rates were varied during the course of a study, or they were unclear^{13,14,16,19,22,25,26} (Table 3).

Within the subgroup of eight studies that maintained a constant dose rate there appears to be a relationship between the dose rate of magnesium sulphate and the odds of AF, with the largest prophylactic effects being at the lowest dose rates (Figure 9). In the forest plot the dose rates (mg hour⁻¹) are ordered from lowest (62.5 mg hour⁻¹) to highest (7200 mg hour⁻¹). The forest plot illustrates a large and statistically significant degree of heterogeneity ($I^2 = 76.6\%$, $p = 0.0001$, random-effects model). It is notable that two studies that appear influential in this relationship (by Dagdelen^{17,18} and Toraman²⁷) have been highlighted previously in terms of their low ORs and possibly atypical patient populations

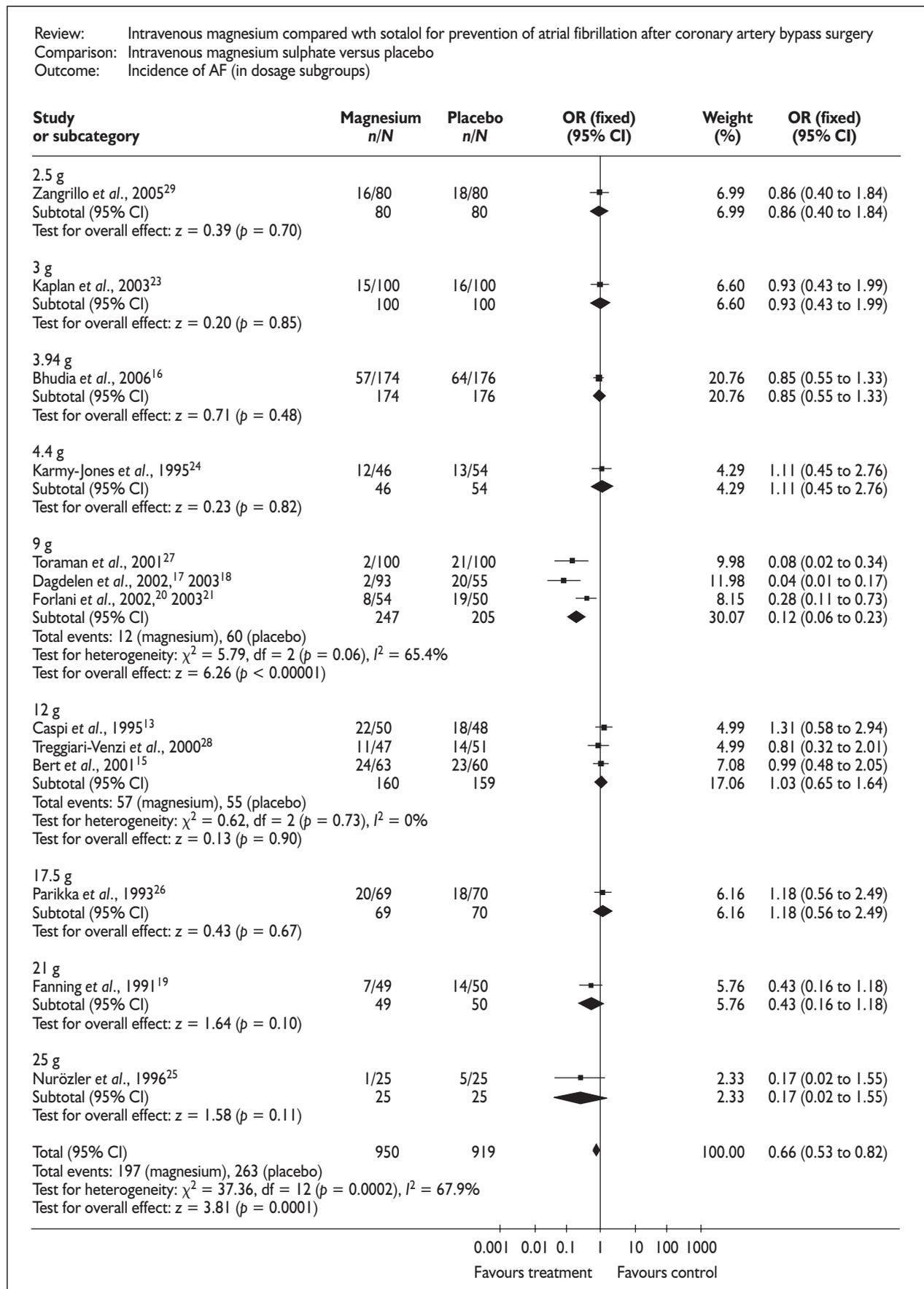
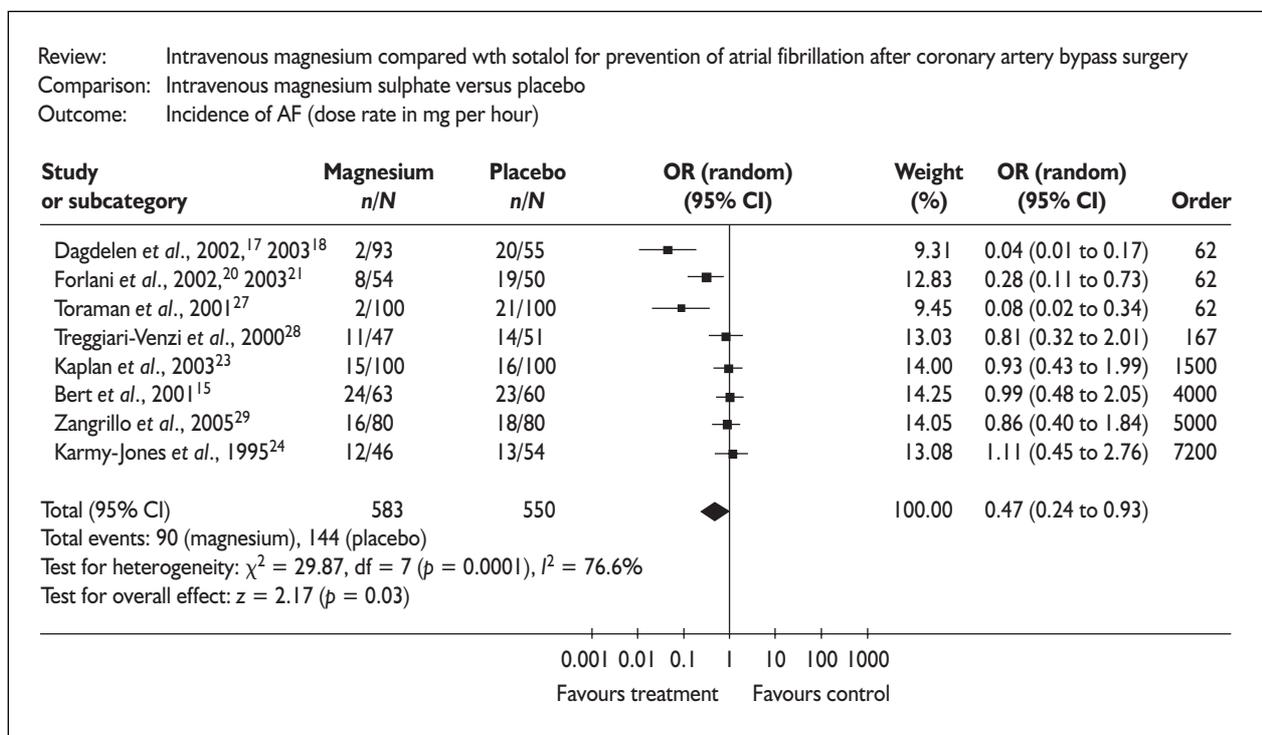


FIGURE 8 Subgroup analysis: total magnesium dose (fixed-effects model). NA, not applicable.

TABLE 4 Studies in which the dose rate of magnesium sulphate was kept constant

	Dose rate reported	Dose rate (mg h ⁻¹)	Total time (h)	Total dose (g)	i.v. carrier
Bert <i>et al.</i> , 2001 ¹⁵	2 g in 30 minutes	4000	3	12	50 ml saline
Dagdelen <i>et al.</i> , 2002, ¹⁷ 2003 ¹⁸	1.5 g day ⁻¹	62.5	144 (= 6 days)	9	100 ml saline
Forlani <i>et al.</i> , 2002, ²⁰ 2003 ²¹	1.5 g day ⁻¹	62.5	144 (= 6 days)	9	NR
Kaplan <i>et al.</i> , 2003 ²³	3 g in 2 h	1500	12	18	100 ml saline
Karmy-Jones <i>et al.</i> , 1995 ²⁴	2.4 g in 20 minutes	7200	2	14.4	50 ml dextrose water
Toraman <i>et al.</i> , 2001 ²⁷	6 mmol (≡ 1.5 g) day ⁻¹	62.5	144 (= 6 days)	9	100 ml saline
Treggiari-Venzi <i>et al.</i> , 2000 ²⁸	4 g day ⁻¹	167	72 (= 3 days)	12	NR
Zangrillo <i>et al.</i> , 2005 ²⁹	2.5 g in 30 minutes	5000	0.5	2.5	100 ml saline

**FIGURE 9** Subgroup analysis: all studies that maintained a constant dose rate (random-effects model). Studies are ordered from lowest dose rate (top) to highest (bottom). Order: dose rate in mg h⁻¹.

(see above). Excluding these two trials from the sub-group reduces the heterogeneity to a non-significant level ($I^2 = 11.6\%$, $p = 0.34$, fixed-effects model), but also weakens the evidence for an effect of dose rate, with only one out of the remaining six studies exhibiting an OR that is clearly and significantly different from 1.0 (Figure 10).

Adverse events

Studies varied in whether they reported adverse events, and how such events were reported; for example, in some cases, AF was reported as an

adverse event. Ten of the 15 studies mentioned whether mortality differed between the magnesium sulphate and control (or placebo) groups (Table 5). Overall, numbers of deaths were too low for differences to attain statistical significance.

Potentially adverse events that were reported included a wide range of electrocardiological conditions (e.g. ventricular tachycardia, ischaemia, other arrhythmias), haemodynamic changes (e.g. altered blood pressure) and biochemical events (e.g. changes in serum electrolytes and creatinine

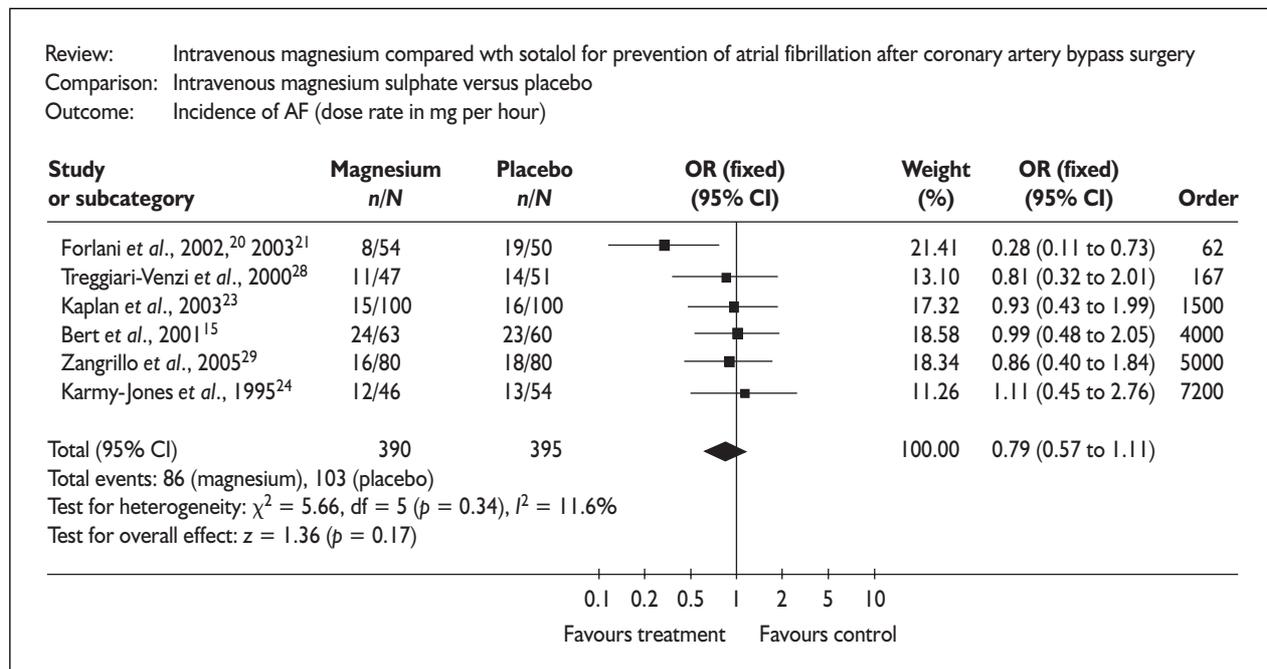


FIGURE 10 Subgroup analysis: studies that maintained a constant dose rate, excluding those by Dagdelen et al.^{17,18} and Toraman et al.²⁷ (fixed-effects model). Studies are ordered from lowest dose rate (top) to highest (bottom). Order: dose rate in mg h⁻¹.

TABLE 5 Summary of mortality reported in patients receiving magnesium sulphate compared with control (or placebo) patients

Study	Postoperative in-study mortality (n)	Mortality reported after discharge or main follow-up (n)
Bert et al., 2001 ¹⁵	0 Mg, 0 control	NR
Bhudia et al., 2006 ¹⁶	1 Mg, 1 placebo	1 Mg, 4 placebo ($p = 0.2$)
Caspi et al., 1995 ¹³	1 Mg, 0 placebo	NR
Dagdelen et al., 2002, ¹⁷ 2003 ¹⁸	NR	NR
Fanning et al., 1991 ¹⁹	0 Mg, 1 placebo	NR
Forlani et al., 2002, ²⁰ 2003 ²¹	1 Mg, 0 control	NR
Hazelrigg et al., 2004 ²²	1 Mg, 2 control	NR
Kaplan et al., 2003 ²³	1 Mg, 1 placebo	NR
Karmy-Jones et al., 1995 ²⁴	0 Mg, 2 placebo	NR
Nurözler et al., 1996 ²⁵	NR	NR
Parikka et al., 1993 ²⁶	NR	1 (group not specified)
Toraman et al., 2001 ²⁷	NR	NR
Treggiari-Venzi et al., 2000 ²⁸	0 Mg, 0 placebo (inferred)	0 Mg, 1 placebo
Yilmaz et al., 2000 ¹⁴	NR	NR
Zangrillo et al., 2005 ²⁹	0 Mg, 0 placebo	NR

kinase concentrations). One study, by Hazelrigg and colleagues,²² reported a statistically significant, but not clinically significant, lower arterial pressure in magnesium-treated patients immediately after surgery (no p -value reported). They also reported a significantly greater excretion ($p < 0.001$) of magnesium and calcium

over a 24-hour period in the magnesium sulphate-treated patients, but the timing relative to surgery was unclear. In all other studies where adverse events differed significantly between the groups, it was the magnesium sulphate group that had the more favourable incidence (or concentration).

Chapter 5

Economic analysis

The aim of this section is to assess the cost-effectiveness of intravenous magnesium sulphate alone as prophylaxis against AF following CABG compared with no prophylaxis. The economic analysis comprises:

- a systematic review of the literature on the cost-effectiveness of intravenous magnesium sulphate alone as prophylaxis against AF following CABG
- identification of appropriate methods to model the cost-effectiveness of prophylaxis against AF following CABG
- identification of key resource-use differences between post-CABG patients with and without AF
- presentation of the economic model and cost-effectiveness evaluation.

Methods for economic analysis

A systematic literature search was undertaken to identify economic evaluations comparing intravenous magnesium sulphate as prophylaxis against AF following CABG compared with sotalol or no prophylaxis. The details of the search strategy are documented in Appendix 7.

Titles and abstracts of studies identified by the search strategy were assessed for potential eligibility by three health economists independently. The inclusion criteria for the review are documented in Appendix 8. Economic evaluations were eligible for inclusion if they reported the cost-effectiveness of intravenous magnesium sulphate alone as prophylaxis against AF following CABG compared with sotalol as prophylaxis or compared with no prophylaxis. Studies reporting the economic evaluation of other agents for prophylaxis against AF following CABG, and studies reporting costs or resource use associated with postoperative AF and its management were included in a secondary review. The purpose of the secondary review was to identify appropriate methods for modelling the cost-effectiveness of prophylaxis against AF following CABG, and key resource-use differences between post-CABG patients with and without AF.

On the basis of the secondary review a short-term economic model was developed to assess the cost-effectiveness of magnesium sulphate prophylaxis against AF following CABG. The model needed to take account of any additional resource use by patients with AF, including the costs of managing postoperative AF during the patient's admission. The outcome used in the economic analysis was cases of AF avoided. Long-term modelling to estimate gains in life expectancy or quality-adjusted life expectancy was not attempted.

Results of the systematic review of economic evaluations

Tailored searches, using the terms reported in Appendix 7, were applied to specialist health, biomedical and health economic databases: Ovid MEDLINE® (1950 to May week 1 2007), EMBASE (1980 to 2007 week 19), Ovid MEDLINE In-Process and Other Non-Indexed Citations (14 May, 2007), and NHS EED and Econlit. A total of 63 references was downloaded into a Reference Manager bibliographic database. References were coded according to whether they were identified by drug-specific filters (eight magnesium studies and 55 sotalol studies were coded). None of these studies met the prespecified criteria for inclusion in the primary review of economic evaluations of intravenous magnesium sulphate alone as prophylaxis against AF following CABG compared with sotalol as prophylaxis or no prophylaxis.

On initial assessment of title and abstract, 49 references were excluded from the review, while two references were included in the secondary review. The remaining 12 references were classified as unclear on screening of title and abstract.

The 12 references classified as unclear on screening of title and abstract were reassessed once the full papers were retrieved. Of these, two were included in the secondary review (one was an economic evaluation of another agent for prophylaxis against AF following CABG, and the other was a meta-analysis reporting resource-use differences) and all other references were excluded.

Summary

- Sixty-three references were identified by the literature search. No studies met the inclusion criteria for the primary review of economic evaluations of intravenous magnesium sulphate as prophylaxis against AF following CABG compared with sotalol as prophylaxis or no prophylaxis.
- Four of the 63 references were included in a secondary review of economic evaluation of other agents for prophylaxis against AF following CABG, and studies reporting costs or resource use associated with postoperative AF and its management.

Secondary review of economic evaluations

Since no studies met the inclusion criteria for the primary review, a review was conducted of papers that reported economic evaluations of other agents used for prophylaxis against postoperative AF, or papers that reported cost/resource-use differences for patients undergoing CABG who developed AF. Of the four included studies, one is an economic evaluation of oral amiodarone for prophylaxis,³⁴ two are reviews of clinical trials^{35,36} that also reported economic outcomes (including trials of amiodarone and sotalol) and one is a clinical trial report²⁸ (including amiodarone, magnesium sulphate and placebo) that reported resource use for patients receiving prophylaxis and for patients with or without AF. The review is broken down into a discussion of:

- the scope of analysis and methods adopted in previous economic evaluation of prophylaxis against postoperative AF
- resource-use differences for patients with and without AF reported in the included studies.

Scope and methodology of economic evaluations

Reddy and colleagues³⁴ developed a simple decision-tree model to assess the cost-effectiveness of amiodarone as prophylaxis against AF for patients undergoing CABG. The proportions of patients developing AF with or without prophylaxis (31.2% and 25.5%, respectively) were taken from a controlled trial.³⁷ Other data used to populate the model were taken from published literature and a clinical database at the Hartford Hospital Cardiology Department. The time-frame for the model was the initial hospitalisation for cardiac surgery, and outcomes were expressed as costs to discharge and AF events occurring during hospital stay following CABG.

Patients who did not develop AF simply accrued costs associated with oral amiodarone prophylaxis (if in the intervention group) and hospital stay, in the ICU and the cardiac step-down unit (SDU). Costs of hospitalisation were based on length of stay in the ICU and SDU, estimated from the literature (the paper contains no information on how these length of stay estimates were derived, other than giving reference to source papers), and unit costs (cost per day) estimated from the local hospital cost database. Cost per day was 32% higher for ICU than for SDU. The same unit costs were applied for patients with and without AF. However, ICU costs were 43% higher and SDU costs 54% higher for patients developing AF, owing to the longer lengths of stay (4.32 versus 3.02 days in the ICU and 5.7 versus 3.7 days in SDU).

It was assumed that AF developed on the second day after surgery; therefore, this was the date on which treatment for AF would start. All patients with AF were assumed to be treated with digoxin, with an initial loading dose and subsequent daily maintenance doses. Patients with AF of greater than 48 hours' duration were also treated with anticoagulants, receiving 2 days of heparin and 8 days of warfarin. Patients received two activated partial thromboplastin (aPTT) tests per day while on heparin and one international normalised ratio (INR) per day while on warfarin. The proportion of patients with AF lasting for more than 48 hours was taken from the hospital cardiology department's clinical database and was estimated at 16.7%.

In addition to costs of hospitalisation and drug treatments for AF, patients may require rhythm control (using electrical or pharmacological cardioversion) or rate control (using drugs such as digoxin). The proportion of patients with AF following CABG who spontaneously convert to normal sinus rhythm (33%) and the proportion requiring electrical cardioversion (13%) were taken from the literature.³⁸ The proportion receiving rate control, which was assumed to continue until discharge from hospital, was the complement of these two proportions (i.e. 54%). The proportion of patients having successful electrical cardioversion or successful pharmacological cardioversion, which was assumed to follow unsuccessful electrical cardioversion, was also taken from the literature.

The flow of patients implied by the model adopted by Reddy and colleagues³⁴ appears to be a reasonable representation of current practice for

patients developing AF following CABG, subject to modifications that would make it more consistent with current clinical guidelines related to England and Wales⁶ and Europe.¹

The model appears to have been evaluated using Monte Carlo simulation. It is difficult to assess the validity of this approach as only the central estimates (with minima and maxima) for parameter values adopted in the model are reported, with no information on how values were selected in any given simulation or whether distributions were assigned to parameters. The base-case results are presented as average cost per AF case avoided for each strategy rather than an incremental analysis for the prophylaxis strategy compared with no prophylaxis. However, the analysis shows that the most influential variables were the cost of hospitalisations and the frequency of AF. The dominant effect of hospital costs is not surprising given the unit costs of US\$1080 and \$1420 for step-down care and ICU bed days, compared with \$64 for prophylaxis with oral amiodarone or \$152 for electrical cardioversion.

Resource-use differences, with and without AF

Treggiari-Venzi and colleagues²⁸ suggest that prophylaxis with either amiodarone or magnesium sulphate will not be cost-effective since the proportion of patients developing AF with prophylaxis is not significantly different, at the 5% level, from placebo [placebo 27% versus magnesium sulphate 23% ($p = 0.82$) and amiodarone 14% ($p = 0.14$)]. They support their argument with a brief analysis which suggests that length of stay in ICU will be longer with amiodarone prophylaxis than without. This is based on the observation that, while patients in the trial who developed postoperative AF had a median length of ICU stay 2 days longer than those without AF (median 5 versus 3 days), patients receiving amiodarone had a median length of ICU stay 1 day longer than placebo patients (median 4 versus 3 days). With an

avoidable risk of AF with amiodarone prophylaxis calculated as 13% (27–14%) they estimate that routine amiodarone prophylaxis for all patients undergoing CABG would be associated with an extra 740 days in the ICU for every 1000 patients compared with no prophylaxis (Table 6).

In effect, Treggiari-Venzi and colleagues²⁸ argue, on the basis of the comparison of oral amiodarone prophylaxis against placebo, that prophylaxis against postoperative AF is dominated by the no-prophylaxis option (with symptomatic management of AF) and therefore no consideration of the cost of prophylaxis was required. They reinforced this conclusion by stating that the oral amiodarone group experienced a higher adverse event rate than did placebo patients, which “may be attributable to the study drug”. The same does not apply for magnesium sulphate prophylaxis which, while not having a statistically significant effect on the occurrence of postoperative AF, could save 80 days of ICU stay using a similar analysis as reported for oral amiodarone (Table 6), which was not reported by Treggiari-Venzi and colleagues.²⁸ Treggiari-Venzi and colleagues²⁸ did not report length of stay on ward following discharge from ICU or total hospital stay, whereas other studies have suggested that AF may increase ward stay as well as ICU stay.

Reddy and colleagues³⁵ reviewed six trials of amiodarone (either oral or intravenous) and one trial of sotalol as prophylaxis against AF following CABG. The main economic outcomes reported were length of stay (reported for all trials and given as total hospital stay or broken down by ICU and ward stay) and hospital costs (reported in five of the amiodarone trials). Trials included in the review typically reported length of stay for each arm of the trial (prophylaxis versus placebo) and do not indicate the difference in length of stay for patients with and without AF. Reddy and colleagues³⁵ reviews additional studies that reported length of stay for patients with AF,

TABLE 6 Days in ICU with and without prophylaxis (from Treggiari-Venzi and colleagues²⁸)

Prophylaxis type	Days in ICU for cohort of 1000 patients		
	With AF ^a	Without AF ^b	Total
Oral amiodarone	840	3440	4280
None	1350	2190	3540
Magnesium sulphate	1150	2310	3460

^a Length of stay × probability of AF × 1000.
^b Length of stay × (1 – probability of AF) × 1000.

suggesting that AF may increase total stay by between 3.2³⁹ and 6⁴⁰ days. Breaking this down by ICU and ward stay, one study showed that AF increased ICU stay by 2.3 days (5.7 versus 3.4 days) and ward stay by 3.4 days (10.9 versus 7.5 days).⁴¹ In the clinical trials of prophylaxis against postoperative AF, while all studies reported a lower average length of stay for patients receiving prophylaxis (and five of the studies showing a significant decrease in the proportion of patients with AF), only one study reported a statistically significant reduction in length of stay.³⁸ This may be a reflection of the trials being underpowered to show significant differences in resource use, since they will typically be powered only to detect differences in clinical outcome. The lack of significant differences may also reflect the fact that, while it may substantially increase length of stay in the ICU or on the ward for some patients, postoperative AF only affects a minority of patients, so the effect of this difference is diluted when looking at mean length of stay across a cohort of patients. Among those studies that reported costs, only one study reported a statistically significant difference³⁸ (with lower costs for patients receiving prophylaxis). The results with respect to cost were less consistent, with two studies reporting lower costs for placebo patients.

Zimmer and colleagues³⁶ included similar studies to those reviewed by Reddy and colleagues³⁵ in their meta-analysis. This included meta-analysis of the effect of amiodarone or sotalol prophylaxis on length of hospital stay and on hospital costs.

Summary

A simple decision tree can be used to estimate the short-term cost-effectiveness of magnesium sulphate prophylaxis against AF following CABG. The form of the analysis in this report will be cost-effectiveness analysis with AF events avoided as the outcome measure.

The principal determinant of the cost-effectiveness of prophylaxis against AF following CABG is likely to be length of stay in the ICU and on hospital wards.

The structure presented by Reddy and colleagues³⁴ is reasonable, but requires some modification to make it consistent with current clinical guidelines relevant to the NHS.

The economic analysis will need to examine potential resource differences that may be associated with statistically non-significant differences in outcome, particularly in terms of length of stay.

A comparison of the cost-effectiveness of prophylaxis with intravenous magnesium sulphate compared with sotalol will not be performed. No RCTs or economic evaluations of this comparison were identified in the present systematic review, and expert clinical opinion suggests that it is not a clinically meaningful comparison.

SHTAC economic model methods

Estimation of net benefits (taking account of disbenefits)

The outcome of interest in the economic model in this report is the proportion of patients experiencing AF following CABG, with or without prophylaxis with magnesium sulphate; or, alternatively, the number of patients experiencing AF in a given patient cohort, for example 1000 patients. The pooled risk ratio derived in the meta-analysis (estimated as a relative risk, rather than OR as presented in 'Incidence of AF', p. 10) will be applied to the estimated baseline proportion of patients experiencing AF following CABG with prophylaxis. This will provide an estimate of the proportion of AF cases averted by prophylaxis with magnesium sulphate.

As reported in 'Description of the underlying health problem' (p. 3), the typical range reported for AF following CABG, without prophylaxis, is between 20 and 40% of patients. For the economic model a point estimate of 30%, which is the average across the placebo arms from trials included in the meta-analysis, will be used. *Table 7*

TABLE 7 Baseline risk of AF following CABG and relative risk with magnesium sulphate prophylaxis: inputs to economic model

	Point estimate Limit/maximum	Lower confidence Limit/minimum	Upper confidence
Relative risk of AF with magnesium sulphate prophylaxis	0.73	0.63 ^a	0.84 ^a
Baseline risk of AF following CABG	0.30	0.20	0.40
^a 95% confidence interval.			

summarises the point estimates and confidence limits/ranges used in the analysis.

Applying 27% risk reduction to the baseline estimate of risk of AF following CABG produces an estimate of the proportion of patients receiving prophylaxis who experience AF of 21.9%. In terms of the estimated 23,000 operations performed annually in England (see 'Description of underlying health problem', p. 3) this would translate to 1863 fewer patients experiencing AF following CABG, if all patients were assumed to be eligible for prophylaxis. In resource terms, assuming that patients with AF stay in the ICU 1 day longer than those without AF, this could result in the avoidance of 1836 ICU days attributable to AF following CABG.

There would be additional resource savings in terms of avoided drug treatments for AF, reduced requirements for anticoagulation and associated tests, and averted cardioversions as the number of patients with AF is reduced. Treatments for AF are not without complications, so a reduced risk of AF following CABG may not only reduce the potential morbidity burden in terms of AF and consequent risk of stroke and other conditions (discussed in 'Description of underlying health problem', p. 3), but also avoid morbidity of adverse events associated with AF treatment (not quantified in the model).

As discussed earlier, a short-term model was adopted that only estimates the reduction in the proportion of patients experiencing AF and resource use up to hospital discharge.

Estimation of net costs

The net costs of consumables for providing magnesium sulphate prophylaxis are calculated as £9.84 per infusion. These are based on unit costs of £2.91 for a 2-ml ampoule of magnesium

sulphate (which corresponds to 4 mmol of magnesium sulphate) and £2.01 for a 50-ml ampoule of sodium chloride intravenous infusion solution (0.9% NaCl). Unit costs are taken from the British National Formulary (BNF), No. 53 (March 2007).⁵ A range of concentrations has been adopted in clinical trials: the European Association of Cardiothoracic Surgeons guidelines recommend a 6-mmol infusion in 100 ml 0.9% NaCl solution at a rate of 25 ml per hour.¹ The recommended regimen is to provide one infusion on the day before surgery, one just after bypass and an infusion once daily for 4 days postoperatively. If the typical length of stay in the ICU following surgery is 1–2 days, it can be assumed that three of the infusions occur alongside usual care and would not require any additional set-up time or monitoring. However, the preoperative and two of the postoperative infusions would take place on the wards and would have additional staff costs for set-up and monitoring.

To estimate the staff costs associated with providing magnesium sulphate prophylaxis by intravenous infusion, it was assumed that 10 minutes of doctor time would be required to initiate and terminate the infusion. In addition, 10 minutes of nurse time would be required each hour of the infusion for patient monitoring. If the infusion occurs at a rate of 25 ml per hour, this implies that each infusion takes 4 hours. Using hourly costs for a specialist registrar and for nurse staff on a 24-hour ward (Unit Costs of Health and Social Care, 2006⁴²) gives a cost of £11.33 for medical staff and £14.67 for nursing staff per infusion (Table 8). Following the recommended regimen of six infusions and assuming that three of these occur in settings where additional costs for initiating and monitoring infusions would be minimal, the total net cost of providing magnesium sulphate prophylaxis by intravenous infusion was estimated at £137.04.

TABLE 8 Unit costs applied in the economic model (£)

	Point estimate	Lower limit	Upper limit
Magnesium sulphate (per infusion)	9.84		
Medical staff cost (per infusion)	11.33	5.67	17.00
Nurse staff cost (per infusion)	14.67	7.33	22.00
CABG (excluding critical care costs) ^a	8172		
Coronary intensive care (per day) ^b	925	693	1196
CABG – excess bed days ^a	334	231	369

Lower and upper limit are based on variation of $\pm 50\%$ unless otherwise stated.
^a HRG – E04: coronary bypass; elective inpatient. (NHS Reference Costs 2005/06⁴³).
^b HRG – CC6L2: cardiac intensive care unit – level 2 care (NHS Reference Costs 2005/06⁴³).
 For NHS reference costs the lower and upper limits are the lower and upper quartiles.

Unit costs for the coronary artery bypass admission, cost per bed day in coronary intensive care and costs for additional bed days associated with postoperative AF were taken from NHS Reference Costs (2005/06).⁴³ The elective inpatient cost for Healthcare Resonance Group (HRG) E04 (coronary bypass) was taken as the unit cost for the coronary artery bypass admission; this excludes costs associated with critical care. The cost for postoperative intensive care was taken as level 2 care in the cardiac intensive care unit (HRG CC6L2). The Intensive Care Society defines level 2 care as being for “patients requiring more detailed observation or intervention, single failing organ system or postoperative care, and higher levels of care.”⁴⁴ The excess bed-day cost for HRG E04 was taken as the unit cost for additional bed days associated with postoperative AF. Reference costs, along with the lower and upper quartiles, are reported in *Table 8*.

To estimate the cost of a typical CABG admission, using NHS Reference Costs, an estimate of ICU use must be added. The Bristol Royal Infirmary *Cardiac services, adult cardiac surgery audit report*⁴⁵ reported a mean cardiac ICU stay between 1.7 and 2.3 days, with a median of 1 day. This analysis assumes a normal postoperative ICU stay of 1 day. Since this estimate is common to all patients in the model (as is the cost of CABG), it has no impact on the cost-effectiveness estimates, only on total cost. Using these assumptions an uncomplicated

CABG admission may be expected to cost £9097 using the unit costs adopted for this analysis.

The greatest anticipated savings from the use of prophylaxis against AF following CABG are likely to be a reduction in the proportion of patients having longer stays in the ICU and a reduction in the proportion of patients having longer overall stays. Using the 27% risk reduction discussed in the previous section would suggest eight fewer patients experiencing AF per 100 operations. To estimate the additional inpatient resource use associated with postoperative AF, studies reporting additional ICU and ward stays associated with postoperative AF were briefly reviewed. These studies were selected from the list of references identified by the search for economic evaluations and costing studies of magnesium sulphate and sotalol as prophylaxis against AF following CABG reported in the section ‘Results of the systematic review of economic evaluations’ (p. 23). Additional studies were identified from the reference lists of included papers.

Where studies have reported length of stay in ICU following CABG and postoperative length of stay for patients with and without postoperative AF, they have consistently shown longer stays for patients with postoperative AF (*Table 9*), although the difference is not always shown to be statistically significant at the 5% level. The range for extra time spent in the ICU is between 8 hours and 2.3 days

TABLE 9 Estimated increase in length of stay, by location of care, associated with patients with AF after CABG

Study	Length of stay		
	ICU (days)	Ward (days)	Overall (days)
Creswell <i>et al.</i> , 1993 ⁴¹	2.3	3.4	
Mendes <i>et al.</i> , 1995 ⁴⁶			1.7
Mathew <i>et al.</i> , 1996 ⁴⁷	0.54	2	
Aranki <i>et al.</i> , 1996 ⁴⁰			4.9 ^a
Kowey <i>et al.</i> , 1997 ⁴⁸			3 ^b
Almassi <i>et al.</i> , 1997 ⁴⁹	1.6 ^b		3 ^b
Zaman <i>et al.</i> , 1997 ⁵⁰			1.1
Borzak <i>et al.</i> , 1998 ⁵¹	1		2.9 ^a
Tamis and Steinberg, 2000 ³⁹			3.2
Treggiari-Venzi <i>et al.</i> , 2000 ²⁸	2 ^b		
Zaman <i>et al.</i> , 2000 ⁵²			1.9
Kim <i>et al.</i> , 2001 ⁵³			1.2 ^a
Hrvanek <i>et al.</i> , 2002 ⁵⁴			1.4
Thompson <i>et al.</i> , 2002 ⁵⁵			5.5
Tamis-Holland <i>et al.</i> , 2006 ⁵⁶		1.1 ^a	
Hosokawa <i>et al.</i> , 2007 ⁵⁷	0.33		3

^a Adjusted for confounding factors.
^b Median (mean not reported).

and for extra days spent on the ward range the range is 1.1–3.4 days. The range for overall postoperative stay is 1.1–5.5 days. These values are similar to those adopted by Reddy and colleagues³⁴ in their economic evaluation of amiodarone prophylaxis, where the base-case estimates were 1.3 additional days in the ICU (range 1.3–1.85) and 2 additional days on the ward (range 0–5.95) associated with postoperative AF.

Postoperative AF is associated with patient factors (such as age) that may also be associated with longer inpatient stays. For example, the Society for Cardiothoracic Surgery’s National Adult Surgical Database Report (2000/01)⁵⁸ reported average postoperative stays between 6 and 10 days, with older patients having the longer average length of stay. Therefore, the unadjusted lengths of stay reported in *Table 9* may overestimate the effect of AF on postoperative length of stay. Those studies that have sought to control for confounding factors^{40,51,53,56} have typically shown that the effect of AF alone is reduced, but that differences between patients with and without AF remain significant. Variables that were found to be significant confounders, and included in the adjustments, are patient sex,^{40,51,53} age,^{40,51,53,56} postoperative complications,^{40,56} preoperative unstable angina⁵³ and digoxin use before surgery.⁵⁶ On the basis of these studies it was assumed that postoperative AF results in 1 additional day in the ICU (range 0–2) and 2 additional days on the ward (range 0–4) (*Table 10*).

The additional cost per 100 patients undergoing CABG, with a 30% baseline risk of postoperative AF and using the unit costs in *Table 8*, would be £47,785 if overall stay were increased by 3 days for patients experiencing AF following CABG (1 extra day in the ICU and 2 extra days on the ward). If prophylaxis with magnesium sulphate is associated with a relative risk of postoperative AF of 0.73, the reduction in costs associated with extra bed days due to AF could be £12,902 per 100 patients.

Costs of drugs used for pharmacological cardioversion (PCV) are based on an assumed

body weight of 84 kg for a male patient undergoing CABG. This was derived from the distribution of body surface area (BSA) reported in the Society for Cardiothoracic Surgery’s National Adult Surgical Database Report (2003),⁵⁹ which has a mean of approximately 2.00 m². Using the equation reported to derive BSA [see equation (1)] and an assumed average height of 175 cm (5’10”), the patient weight associated with a BSA of 2.00 can be estimated. A range for body weight was calculated using heights of 170 and 180 cm, giving a weight range of 80–89 kg.

$$BSA = 7.184 \times 10^{-3} \times m^{0.425} \times h^{0.725} \quad (1)$$

where *m* is weight (kg) and *h* is height (cm).

Using a dosing protocol for amiodarone of a loading dose of 5 mg/kg⁻¹ for 30 minutes as intravenous infusion, followed by 25 mg per hour for 24 hours and continuing treatment for 24 hours after cardioversion, gives a consumables cost of £40.88 for the loading dose and 48 hours of treatment (assuming that cardioversion occurs with 24 hours) (*Table 11*). This excludes staff costs for initiation of the intravenous infusions and for patient monitoring, as it was assumed that the majority of PCV will be initiated in the ICU. The reported effectiveness of amiodarone, for cardioversion, ranges from 77 to 93% of patients with postoperative AF converting to sinus rhythm at 24 hours.¹

An alternative strategy for PCV was costed, using sotalol with a loading dose of 1 mg kg⁻¹, followed by 0.2 mg kg⁻¹ for 12 hours as an intravenous infusion. If cardioversion occurs patients receive a maintenance dose of 160 mg oral sotalol, taken twice daily, for 2 weeks. The estimated consumables cost of this strategy was £20.11 for intravenous infusions and £3.55 for oral drugs, giving a total cost of £23.66 (*Table 11*). The reported effectiveness of sotalol, for cardioversion, ranges from 85 to 100% of patients with postoperative AF converting to sinus rhythm at 24 hours,¹ although these values were derived from small studies.

TABLE 10 Lengths of stay with and without postoperative AF used in the model

	Point estimate	Lower limit	Upper limit
CABG, excluding ICU, without AF	7		
ICU stay, without AF	1		
Additional ICU days with AF	1	0	2
Additional ward days with AF	2	0	4

TABLE 11 Costs for treating AF: cardioversion or rate control (£)

	Point estimate	Lower limit	Upper limit
Amiodarone – PCV	40.88	20.44	61.32
Sotalol –PCV	23.66	11.83	35.49
Direct current cardioversion	117	58.50	175.50
Digoxin – rate control	14.41	7.21	21.62
Anticoagulation	22.84	21.62	34.26

Lower and upper limit are based on variation of $\pm 50\%$ unless otherwise stated.

Costs for electrical cardioversion (ECV) were taken from Buxton and colleagues.⁶⁰ This gave a cost of £100 at 2001 prices, which was updated to 2005/06 prices using the Hospital and Community Health Services Pay and Prices Index.⁴² Similarly, costs of anticoagulation were taken from Reddy and colleagues.³⁴ Costs for warfarin and heparin reported in 1998 US dollars were converted to UK pounds using 1998 purchasing power parities⁶¹ and then updated to 2005/06 prices using the Hospital and Community Health Services Pay and Prices Index.⁴² Patients who did not convert to sinus rhythm within 48 hours received 2 days of intravenous heparin and 8 days of warfarin.

Rate control was costed using digoxin with a loading dose of 0.5 mg followed by 0.25 mg per 2 hours up to a maximum of 1.5 mg. This was followed by an oral maintenance dose of 0.25 mg daily. The estimated consumables cost of the intravenous loading dose was £13.74 and oral maintenance dosing was costed at £0.08 per day (based on a unit cost of £2.35 per packet of 28 × 250- μ g tablets).⁵

Estimation of cost-effectiveness

Model structure

Figure 11 illustrates the model developed to estimate the cost-effectiveness of intravenous magnesium as prophylaxis against AF following CABG. The rounded rectangles (CABG and Discharge) indicate the “start” and “end” states. The non-rounded rectangles are “outcomes”:

- presence or absence of postoperative AF
- success or failure of rhythm or rate control.

The diamonds indicate “decisions”:

- whether patients spontaneously revert to sinus rhythm (effectively a decision not to initiate treatment for AF, and observe)
- choice of rhythm or rate control (and choice of electrical or pharmacological rhythm control).

The flowchart allows for patients with AF, who fail to convert spontaneously to sinus rhythm, to receive pharmacological rate control rather than electrical or pharmacological rhythm control. It also allows for patients who fail PCV to move directly to rate control rather than ECV. The flow chart was converted to a decision tree (Figure 12) to perform the cost-effectiveness analysis.

Model inputs

Table 7 lists the point estimates (with lower and upper limits) for key parameters related to the baseline risk of AF following CABG and the relative risk of AF with magnesium sulphate prophylaxis used in the base-case analysis and in the deterministic sensitivity analyses. Tables 8 and 11 report the point estimates (with lower and upper limits) for unit costs used in the model, while Table 10 reports the point estimates (with lower and upper limits) for length of stay, with and without postoperative AF, used in the base-case analysis and in the deterministic sensitivity analyses.

Table 12 reports all parameter inputs to the model, with their point estimates along with the distribution types (and distribution parameters) applied in the probabilistic sensitivity analysis (PSA).

Methods

The decision tree can be used to estimate the cost-effectiveness, in terms of incremental cost per AF case avoided with magnesium sulphate prophylaxis compared with no prophylaxis. In addition to reporting the incremental costs and incremental effectiveness, the average length of stay for the prophylaxis and no-prophylaxis options is reported, as well as the average additional ICU and ward days under each option. Key areas of uncertainty in the model are examined using deterministic and probabilistic sensitivity analyses.

Deterministic sensitivity analysis is used to address particular areas of uncertainty in the model related to:

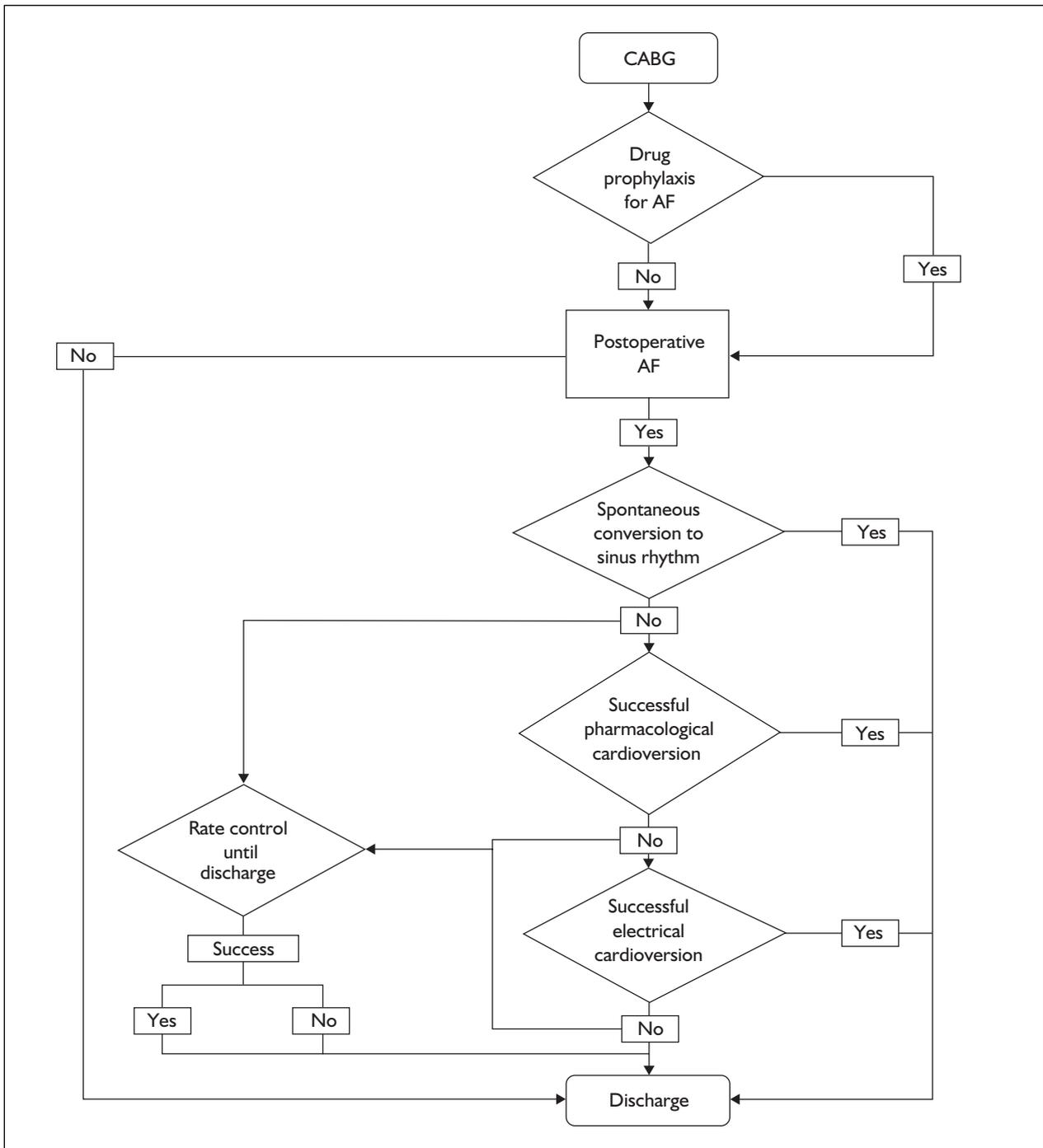


FIGURE 11 Flowchart for drug prophylaxis for, and management of, AF following CABG

- model structure
- methodological assumptions
- parameters around which there is considerable uncertainty or that may be expected, *a priori*, to have disproportionate impact on study results.

The purpose of this analysis is to identify clearly the impact of this uncertainty and to test the robustness of the cost-effectiveness results to

variation in structural assumptions and parameter inputs.

Parameter uncertainty is addressed using PSA. Probability distributions are assigned to the point estimates used in the base-case analysis, using the point estimates and ranges listed in Tables 7, 8, 10 and 11. Table 12 reports the variables included in the PSA, the form of distribution

TABLE 12 Parameter input values in the economic model

Parameter	Point estimate	Distribution (parameters of distribution)
<i>Probabilities</i>		
Baseline risk of AF following CABG	0.30	Beta $\alpha = 59.8621; \beta = 139.6782$
Relative risk of AF with magnesium sulphate prophylaxis	0.73	Normal $\mu = \ln(0.73); \sigma = 0.0734$
Probability of spontaneous conversion to sinus rhythm ^{34,38}	0.35	Beta $n = 40; r = 14$
Probability of cardioversion ^{34,38}	0.60	Beta $n = 40; r = 24$
Probability of successful PCV ^{62,63}	0.736	Beta $n = 72; r = 53$
Probability of successful ECV following unsuccessful PCV ^{62,63}	0.842	Beta $n = 19; r = 16$
<i>Additional resource use associated with AF</i>		
Additional days in ICU due to postoperative AF	1	Gamma $\alpha = 6.1117; \beta = 0.1636$
Additional days on ward due to postoperative AF	2	Gamma $\alpha = 9.5495; \beta = 0.2094$
<i>Costs</i>		
Magnesium sulphate prophylaxis	137.04	
Rate control – digoxin	14.41	
PCV with amiodarone	40.88	
PCV with sotalol	23.66	
Anticoagulation – heparin and warfarin	22.84	
Cost of ECV	117	
Cost of CABG admission	8172	
Coronary intensive care (per bed day)	925	Gamma $\alpha = 88.7514; \beta = 3.7613$
Hospital ward (cost per excess bed day)	344	Gamma $\alpha = 52.1459; \beta = 17.7421$

used for sampling and the parameters of the distribution.

Results of the economic model

Base-case results

The base-case results of the analysis are reported in *Table 13*. Costs under the strategy of providing magnesium sulphate as prophylaxis against postoperative AF are slightly higher than with no prophylaxis (£2.55). Prophylaxis is associated with a reduction in the proportion of patients experiencing AF following CABG (0.081), giving an incremental cost-effectiveness ratio (ICER) of £32 per AF case avoided.

To identify which factors have contributed most to these results, *Table 14* reports the average total inpatient stay estimated for each strategy and the average number of additional ICU and ward days associated with AF for each strategy.

As with the studies investigating the effect of prophylaxis on hospital stay (reviewed in the section ‘Secondary review of economic evaluations’, p. 24), there was only a small difference in length of stay (in the ICU or on the ward) between the prophylaxis and no-prophylaxis strategies (8.66 and 8.90, respectively), despite large assumed differences in length of stay for patients with and without postoperative AF (average of 11 versus 8 days, respectively). This

TABLE 13 Base-case results of economic analysis

Strategy	Cost (£)	Incremental cost (£)	AF cases	AF cases avoided	Incremental cost per AF case avoided (£)
No prophylaxis	9595		0.300		
Prophylaxis	9598	2.55	0.219	0.081	31.51

TABLE 14 Estimated total stays for CABG patients and additional days due to AF

Strategy	Total stay (days)	Additional ICU days	Additional ward days	Costs of inpatient days (£)
No prophylaxis	8.90	0.30	0.60	9581
Prophylaxis	8.66	0.22	0.44	9450

analysis assumes that admission routines would be identical for patients under both the no-prophylaxis and prophylaxis strategies.

Deterministic sensitivity analysis

A sensitivity analysis was conducted to consider the effect of uncertainty around model structure and variation in model parameters. The method adopted was univariate sensitivity analysis; that is, varying one parameter at a time, leaving all other variables unchanged. This is to highlight the impact, if any, of each selected parameter alone on the cost-effectiveness results.

Table 15 reports the results of the sensitivity analysis. The table is divided to distinguish between analyses undertaken due to uncertainties over parameters related to:

- the risk of postoperative AF and effectiveness of magnesium sulphate prophylaxis
- the cost of prophylaxis
- resource consequences of postoperative AF
- resources used in the management of postoperative AF.

Table 15 shows large variations in cost-effectiveness estimates relating to uncertainty over both the baseline risk and relative risk of postoperative AF with magnesium sulphate prophylaxis. At the upper limit of baseline risk of AF (40%) the prophylaxis option dominates, providing better outcome at lower cost. At the lower limit incremental cost is positive, the prophylaxis option costs more than no prophylaxis and the incremental effectiveness is reduced, leading to a substantial increase in ICER. This is further illustrated in Figure 13.

Variation in the relative risk of AF with prophylaxis shows a similar pattern, with the

prophylaxis dominating at the lower limit (larger treatment effect). At the upper limit the prophylaxis option has positive incremental cost and reduced incremental effect, leading to a large increase in ICER. This is further illustrated in Figure 14.

The cost-effectiveness estimates also seem to be highly sensitive to the overall cost of prophylaxis, and to variation in components of the cost of prophylaxis. Varying the total cost of prophylaxis by plus or minus 50% shows wide variation, from a situation where prophylaxis dominates (at lower cost of prophylaxis) to an increase in ICER equivalent to that for the lower limit in baseline risk of AF. Varying assumptions over staff inputs to initiate intravenous infusions and patient monitoring show large variations in ICER. A strategy where nurses initiate and terminate intravenous infusions (labelled “no medical input” in Table 15) reduces the cost for the prophylaxis option below the cost for no prophylaxis, using all other base-case assumptions, causing prophylaxis to be dominant.

It is possible that the prophylaxis option may increase preoperative length of stay, since one recommended regimen for magnesium sulphate includes an infusion on the day before surgery. If patients having prophylaxis stay on average half a day longer than those without prophylaxis the incremental cost increases by almost £167 and the ICER increases to £2092 per AF case avoided. This is the largest ICER for any option included in the one-way sensitivity analyses.

Considering the resource consequences of AF, the ICERs are highly sensitive to assumptions over the number of additional ICU days due to AF, and to a lesser extent the number of additional ward days. The range of ICERs was

TABLE 15 Univariate sensitivity analysis

Variable (base case)	Lower limit Upper limit	Incremental cost (£)	Incremental effectiveness	ICER (£/AF case averted)	Range
<i>Risk of AF and effectiveness of prophylaxis</i>					
Probability of AF (0.30)	0.20	47.38	0.054	877.44	1268.89
	0.40	-42.28	0.108	Dominant	
Relative risk of AF (0.73)	0.63	-42.28	0.108	Dominant	1620.40
	0.84	57.34	0.048	1194.66	
<i>Cost of prophylaxis</i>					
Prophylaxis cost (137.04)	68.52	-65.97	0.081	Dominant	1691.85
	205.56	71.07	0.081	877.44	
Minutes of doctor time for prophylaxis (20)	10	-14.45	0.081	Dominant	419.75
	30	19.55	0.081	241.39	
Minutes of nurse time for prophylaxis (40)	20	-19.45	0.081	Dominant	543.21
	60	24.55	0.081	303.12	
Nurse initiates infusion and monitors patient: no medical input	30	-42.45	0.081	Dominant	814.81
	60	-9.45	0.081	Dominant	
	90	23.55	0.081	290.77	
Number of prophylaxis infusions with additional staff costs (3)	6	80.55	0.081	994.47	NA
Additional preoperative days (0)	0.50	169.46	0.081	2092.13	NA
<i>Resource consequences of postoperative AF</i>					
Additional ICU days (1)	0.00	77.48	0.081	956.59	1850.40
	2.00	-72.40	0.081	Dominant	
Additional ward days (2)	0.00	56.63	0.081	699.08	1335.37
	4.00	-51.54	0.081	Dominant	
Additional bed days for spontaneous converters (3)	0	47.70	0.081	588.92	NA
Cost of ICU bed day (925)	693.00	21.36	0.081	263.66	502.94
	1196.00	-19.38	0.081	Dominant	
Cost of excess bed day (344)	231.00	19.21	0.081	237.12	275.95
	369.00	-3.15	0.081	Dominant	
<i>Resources used in managing postoperative AF</i>					
PCV/ECV attempted (0.6)	0.30	4.76	0.081	58.79	54.57
	0.90	0.34	0.081	4.23	
Spontaneous conversion (0.35)	0.175	1.16	0.081	14.30	34.43
	0.525	3.95	0.081	48.72	
Cost of ECV (117)	58.50	4.40	0.081	54.33	45.63
	175.50	0.70	0.081	8.70	
Cost of amiodarone PCV (40.88)	20.44	2.72	0.081	33.61	4.21
	61.32	2.38	0.081	29.41	
Use sotalol for PCV	23.66	2.70	0.081	33.28	NA
Cost of sotalol PCV (23.66)	11.83	2.79	0.081	34.50	2.44
	35.49	2.60	0.081	32.07	

lower for variation in the cost of ICU bed days or of excess ward days.

Tamis-Holland and colleagues⁵⁶ suggest that patients with postoperative AF who spontaneously

convert to sinus rhythm do not experience longer stays than do patients without AF. Setting the additional bed days to zero for patients with postoperative AF who spontaneously convert increases the incremental cost of the prophylaxis

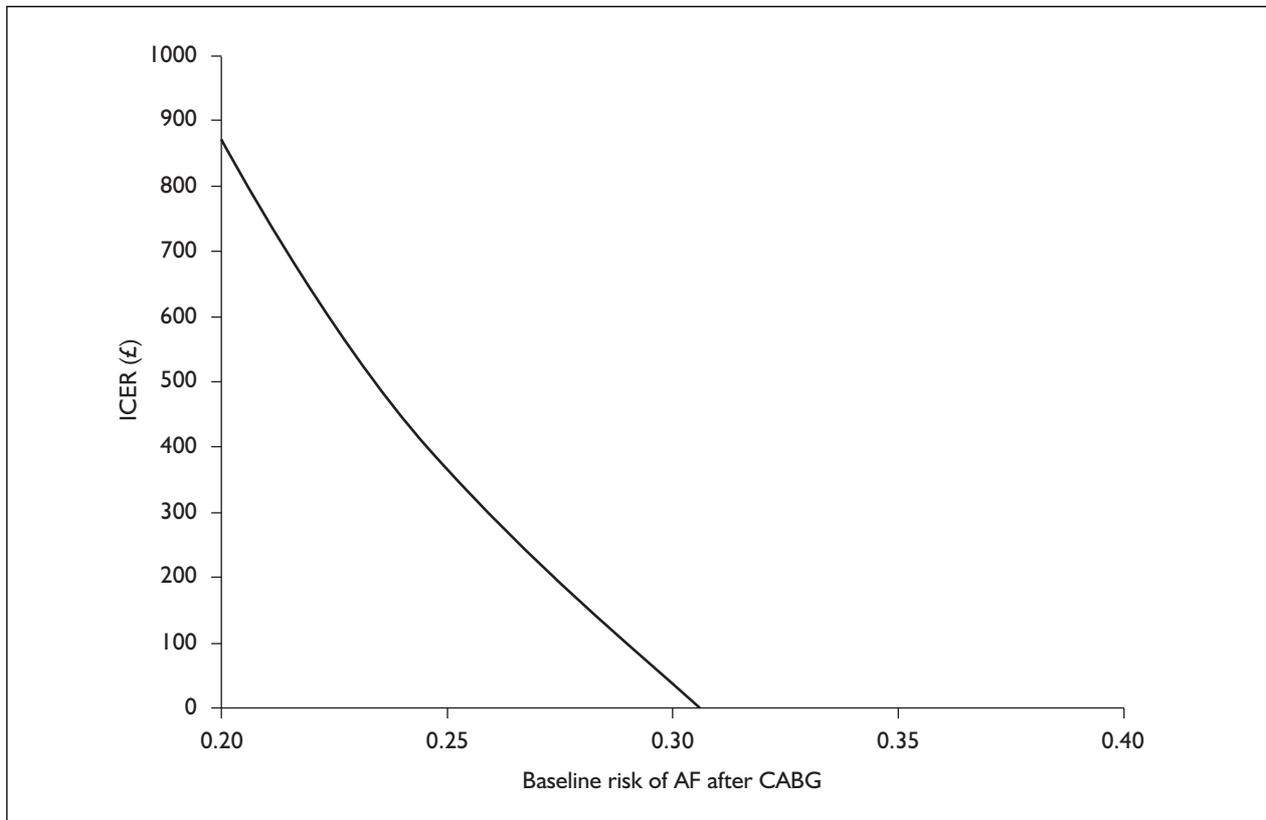


FIGURE 13 Sensitivity analysis on baseline risk of AF following CABG

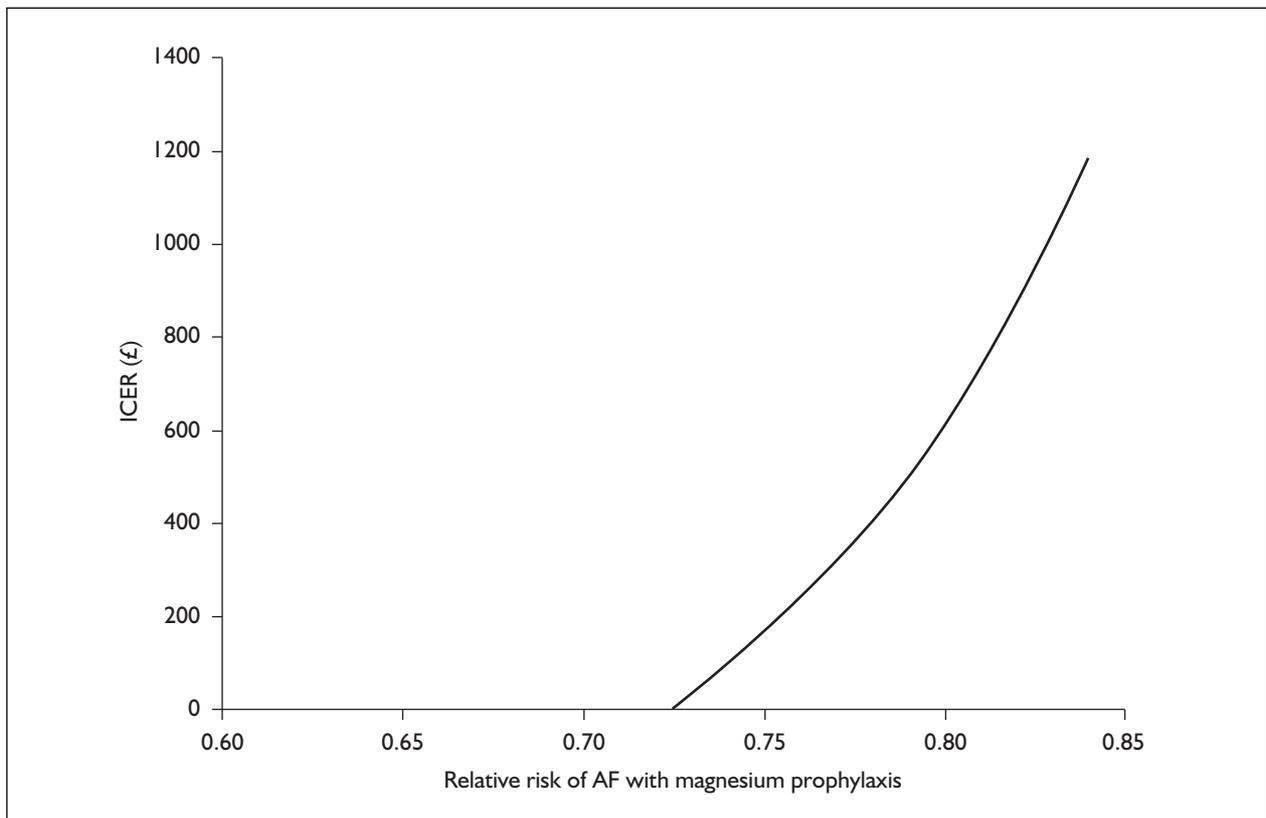


FIGURE 14 Sensitivity analysis on relative risk of AF following CABG with magnesium sulphate prophylaxis

strategy and increases the ICER to almost £600 per AF case avoided.

Table 15 shows that the ICERs are largely insensitive to variation in values of parameters related to management of postoperative AF.

Probabilistic sensitivity analysis

In a PSA of magnesium sulphate prophylaxis, where probabilities of postoperative AF, relative risk of AF with magnesium sulphate prophylaxis and additional bed days (ward or ICU) associated with AF were sampled across the distributions described in Table 12, just over half of the simulations are associated with better outcomes (fewer cases of AF), but also increased costs (Figure 15). Many of the simulations (41%) have negative incremental costs. In this area magnesium sulphate prophylaxis is the dominant strategy, achieving better outcome at lower costs. Simulations where costs for prophylaxis are lower are most likely to be associated with high estimates of additional stay on the ward or in the ICU associated with AF.

In this analysis magnesium sulphate prophylaxis had a probability of being cost-effective (compared with surgery without prophylaxis) of 99% at a willingness-to-pay (WTP) threshold of £2000 per AF case avoided and 100% at a WTP threshold of £5000 per AF case avoided (Figure 16, (CEAC: base-case assumptions).

The base-case analysis assumes identical admission routines for patients under the no-prophylaxis and prophylaxis strategies. An alternative scenario, where patients receiving magnesium sulphate prophylaxis by intravenous infusion are admitted early to receive their initial infusion, was included in the PSA (mean additional stay = 0.5 days, using a uniform distribution with minimum = 0 and maximum = 1). The effect of including this assumption in the PSA is to shift the scatter of points on the cost-effectiveness plane (Figure 15) upwards to a mean incremental cost of approximately £170, so that only around 5% of simulations have negative incremental costs. This increase in mean incremental costs also changes the location of the cost-effectiveness acceptability

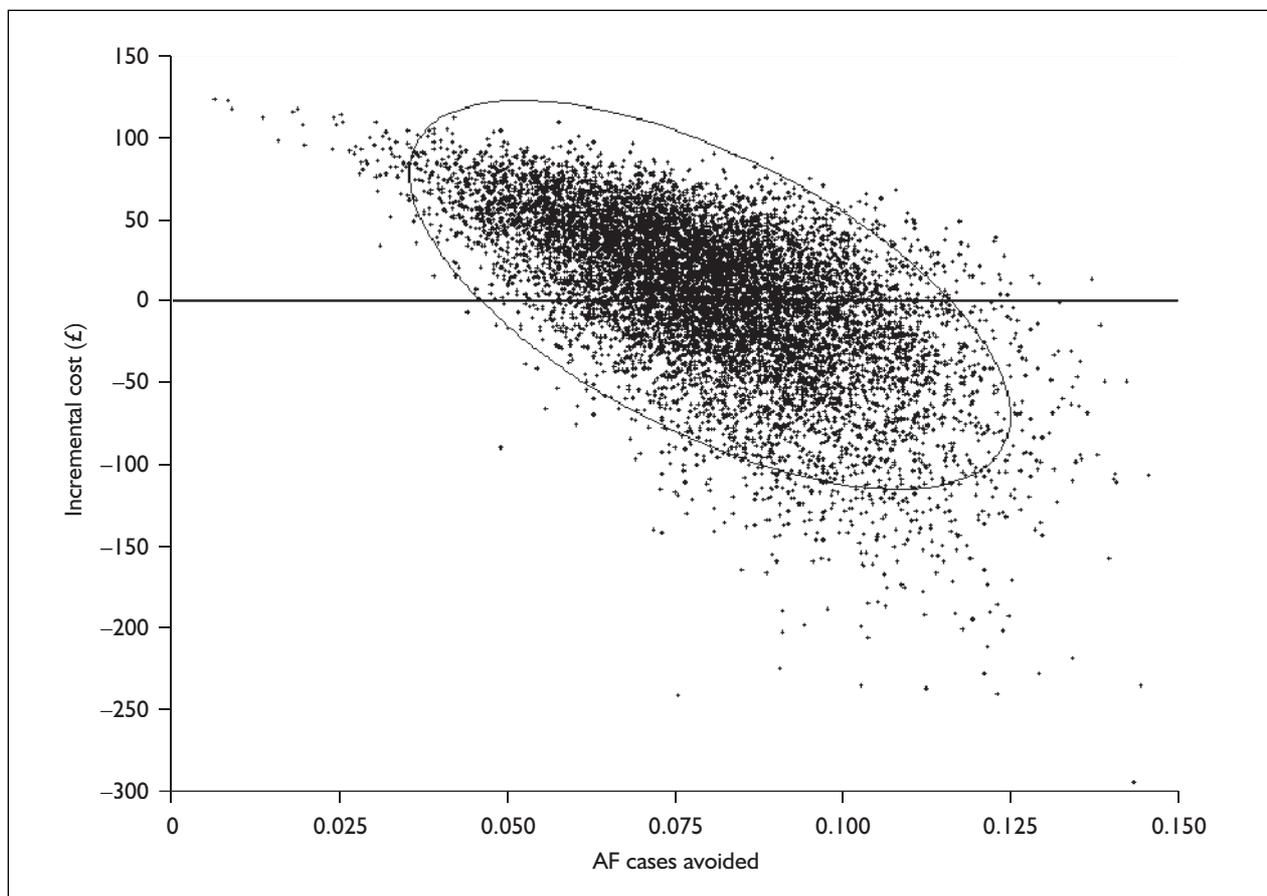


FIGURE 15 Cost-effectiveness plane: incremental cost and AF cases avoided for magnesium sulphate prophylaxis against postoperative AF

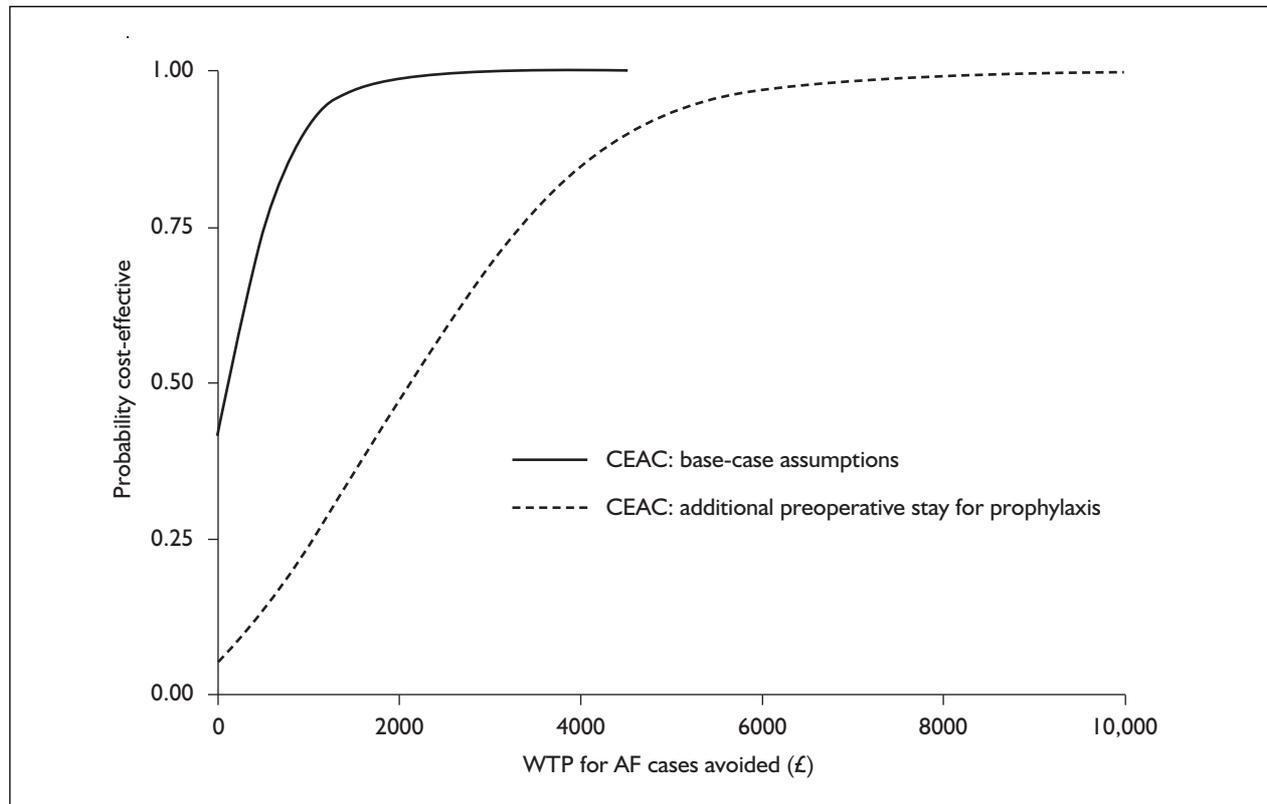


FIGURE 16 CEAC: magnesium sulphate prophylaxis PSA result

curve (CEAC) (*Figure 16*) and the probability of prophylaxis being cost-effective at each threshold of WTP. Now the probability of magnesium sulphate prophylaxis being cost-effective (compared with surgery without prophylaxis) is 48% at a WTP threshold of £2000 per AF case avoided and 93% at a WTP threshold of £5000 per AF case avoided.

In reporting this probabilistic analysis, much lower thresholds were used than would commonly be applied in models reporting outcomes in terms of quality-adjusted life-years (QALYs). It is not clear

in this model how the avoidance of postoperative AF would translate to final outcomes, such as gains in life expectancy or quality-adjusted life expectancy. The European Association of Cardiothoracic Surgeons guideline¹ suggests that AF following cardiac surgery doubles the risk of stroke, with a baseline risk of between 1.4%⁴¹ and 2.5%⁴⁹ for patients in sinus rhythm in the studies reviewed. At these levels of risk it is unlikely that the reduction in risk, modelled in this analysis, would translate to substantial gains in life expectancy or quality-adjusted life expectancy.

Chapter 6

Discussion

Clinical effectiveness

This systematic review did not identify any RCTs that aimed to compare intravenous magnesium with sotalol. Expert clinical advice suggests that in practice this would not be a meaningful comparison, as the decision would be more likely to be whether or not to use these (and other) drugs in combination, as opposed to choosing between them. The clinical effectiveness of magnesium sulphate and sotalol combination therapy has been evaluated in RCTs, although it is outside the scope of the current review.^{20,21,64} This would be a worthwhile topic for future evidence synthesis.

This systematic review has updated and expanded the review published by Alghamdi and colleagues in 2005,¹⁰ adding in a further eight RCTs. In that systematic review it was found that intravenous magnesium administered perioperatively to patients undergoing CABG was effective, with a statistically significant intervention effect. The current review comes to a similar conclusion, that intravenous magnesium is associated with a statistically significant reduction in the odds of a postoperative AF.

Effects of total dose, dose rate and duration of magnesium sulphate intervention

If the studies by Dagdelen and colleagues^{17,18} and Toraman and colleagues²⁷ are excluded, the current review would suggest that the total magnesium sulphate dose up to 21 g is probably not a useful predictor of effects on AF. It should be noted, however, that a robust analysis of total dose should, ideally, include co-variables, but this was not possible owing to the heterogeneity of reporting in the primary studies. The possibility of effects at higher total doses (>21 g) is unclear owing to ambiguity within the study by Nurözler and colleagues²⁵ (25 g total dose), in which confounding of variables might have occurred. Analyses of the duration of prophylaxis and the dose rate suggest that these may be better predictors of effects on AF, but the dose rate and duration of prophylaxis were not generally independent and their relative importance cannot

be ascertained from the available trials [only one total dose (12 g) was administered at more than one dose rate]. An hypothesis arising from these findings that may warrant further investigation is that continuous low-dose administration of magnesium sulphate would be more effective for prophylaxis of AF than a similar total dose administered over a shorter period.

Two of the five systematic reviews identified in the searches assessed the influence of total dose of magnesium sulphate on AF.^{7,9} Henyan and colleagues⁹ divided the studies according to whether they administered magnesium sulphate in low doses (<10 g) or moderate to high doses (≥ 10 g), whereas Miller and colleagues⁷ divided the studies according to whether they administered low doses (<35 mmol), medium doses (35–50 mmol) or high doses (>50 mmol). These dose subgroups appear to have been defined on an arbitrary basis, and it is unclear whether they have clinical or biochemical relevance. Only one of the subgroups in each review exhibited a significant intervention effect (the low-dose group of Henyan and colleagues⁹ and the medium-dose group of Miller and colleagues⁷) and it is notable that the studies by Dagdelen and colleagues^{17,18} and Toraman and colleagues²⁷ were in these sub-groups. Given that the studies by Dagdelen^{17,18} and Toraman²⁷ differed in other respects besides their dose of magnesium sulphate, the findings that low-dose⁹ or medium-dose⁷ magnesium sulphate significantly reduced the incidence of AF appear questionable and are not supported in the current review.

Possible influence of β -blocker therapy and other factors on effects of magnesium sulphate

As noted above, the studies by Dagdelen^{17,18} and Toraman²⁷ impart a significant intervention effect to almost any subgroup in which the studies are analysed. It is difficult to determine which aspects of these studies drive the significant intervention effect, as the studies differed in several respects from the other studies included in this review. For example, they were the only studies that explicitly excluded patients who were on β -blockers, and they started prophylactic therapy earlier than most

other studies, continued prophylaxis over a longer duration and used a lower dose rate than other studies. It is possible that an interaction between these (and possibly other) factors was responsible for the uniquely low ORs reported in these two studies. β -Blockers prevent the loss of intracellular magnesium and it has been suggested that prophylactic magnesium sulphate may be more effective in patients who are not offered preoperative β -blockers.²² The relative importance of β -blockers, the time of initiation and the duration of prophylaxis appear to warrant further investigation, given the overall positive results of these studies for prophylaxis of AF.

Heterogeneity of exclusion criteria in the primary studies

The criteria used to exclude patients from the trials were reported in variable detail, with up to 15 main criteria listed in one study, whereas three of the studies did not report any exclusion criteria (Appendix 6). The selection of exclusion criteria could have an important bearing on the risk classification of the recruited populations. As discussed above, excluding patients who are on antiarrhythmic drugs might have led to the included patients being at higher risk of AF. Conversely, excluding high-risk patients (e.g. those with ventricular dysfunction or history of AF) might have led to the included patients being at lower risk of AF. Ideally, meta-analysis should be stratified by patient risk to account for such interpopulation differences. This was not possible, however, as in most cases the exclusion criteria were weakly defined or undefined (e.g. unclear whether supraventricular arrhythmia included AF), and because it is difficult to weigh up objectively the relative importance of individual exclusion criteria in the risk of AF (Appendix 6).

Cost-effectiveness

The systematic review of economic evaluations did not identify any studies comparing magnesium sulphate prophylaxis with sotalol or with no prophylaxis. It is unlikely that economic evaluations of intravenous magnesium sulphate compared with sotalol have been performed, given the lack of RCT evidence for this comparison. The secondary review of economic evaluations of other agents as prophylaxis for AF found one study of oral amiodarone conducted using a short-term model. Short-term models such as this only follow up patients to discharge from hospital and use intermediate outcomes. It is unclear from the current evidence base how a reduction in

postoperative AF cases relates to changes in final outcomes.

The secondary review suggests that postoperative AF leads to increases in resource use, particularly increased stay in the ICU and on hospital wards. Estimates of the additional resource use associated with postoperative AF were developed using the results of the literature search. However, the data were not suitable for a quantitative synthesis. The applicability of all the included studies to current UK clinical practice may be questioned as the publication dates of the studies ranged over 14 years. Moreover, the included studies were conducted in a number of countries, which may have significantly different criteria for admission or discharge from the ICU and patterns of postoperative management. Existing UK cardiac surgical databases, such as the Society of Cardiothoracic Surgeons' National Adult Cardiac Surgical Database, could be investigated for suitable data to refine the inputs to the economic model. Such databases could also be used to define current patterns of practice in management of surgical patients and of postsurgical patients with AF.

The analysis presented here suggests that prophylaxis may achieve a reduction in AF cases at a modest additional cost and that this result is not sensitive to assumptions made regarding the management of postoperative AF, but is sensitive to assumptions over the length of stay for patients with AF. The results were also sensitive to the baseline risk of AF and to the effectiveness of prophylaxis, as would be expected. The PSA suggests that prophylaxis is most likely to be associated with increased treatment costs, although these may be modest. If intensive care facilities are in short supply it may be reasonable to accept a slight increase in treatment costs to release scarce resources. However, if prophylaxis requires earlier admission of patients than would be the case under the no-prophylaxis option, a more careful evaluation of the relative benefits of increasing ward stays compared with freeing up some intensive care facilities is required.

A key question that is beyond the scope of this report is the extent to which avoiding postoperative AF cases translates to changes in life expectancy or quality-adjusted life expectancy for patients undergoing coronary artery bypass surgery. Evidence reviewed in the European Association of Cardiothoracic Surgeons guideline¹ suggests that postoperative AF is associated with an approximate doubling in the risk of stroke

following CABG (from 1.4–2.5% for patients in sinus rhythm to 3.3–5.2% for patients who experience postoperative AF). Patients experiencing postoperative AF tend to have higher preoperative risks of stroke, such as older age and hypertension,⁴⁹ although a retrospective study by Stamou and colleagues⁶⁵ demonstrated an independent effect for postoperative AF on risk of stroke in multivariate analysis (OR = 1.7, 95% CI 1.4 to 2.2). While the epidemiological evidence is currently insufficient to support the construction of a robust, long-term model of outcomes for patients experiencing postoperative AF, these data suggest that for every 1000 patients undergoing

CABG between one and two strokes could be avoided through the use of magnesium sulphate prophylaxis (by averting 81 cases of postoperative AF). In addition to quality of life impacts resulting from the mortality and morbidity associated with strokes, costs of strokes have been estimated to be between £16,000 and £91,000 (2004 prices) depending on severity of stroke and time-horizon; the lower figure is for a mild stroke over 5 years, the higher figure is a lifetime cost for a major ischaemic stroke.⁶⁶ (Costs were reported in US dollars. These have been converted to UK sterling by the authors of this report, using purchasing power parities in the original article⁶⁶.)

Chapter 7

Conclusions

No RCTs were identified that specifically aimed to compare intravenous magnesium with sotalol as prophylaxis for AF in patients undergoing CABG. Intravenous magnesium, when compared with placebo or control, is effective in preventing postoperative AF, as confirmed by a statistically significant intervention effect based on pooled analysis of 15 RCTs. It was also found that AF was less likely to occur with a longer duration of prophylaxis and when prophylaxis was started earlier, although there were uncertainties associated with the trials in these subgroups. No clear relationship between dose and AF was observed, although a lower constant dose rate was associated with the lowest odds of AF. Further primary research should investigate the relationship between dose, dose rate, duration of prophylaxis, timing of initiation of therapy and patient characteristics, such as degree of risk for AF. This will provide stronger evidence for the optimum delivery of intravenous magnesium in patients undergoing CABG.

In terms of cost-effectiveness, magnesium sulphate prophylaxis can lead to reductions in the number of postoperative AF cases at low or no additional costs, owing to offsetting the additional costs of prophylaxis with reductions in ICU and ward stays as well as avoided costs of managing postoperative AF.

The results of the cost-effectiveness analysis are sensitive to plausible changes in certain

assumptions. In particular, the results were sensitive to variation in the baseline risk of AF, the relative risk of postoperative AF with magnesium sulphate prophylaxis, the cost of prophylaxis (in particular the possibility of a longer preoperative stay to allow for the first infusion to occur on the day before surgery) and the number of additional ICU or ward days associated with AF.

The probabilistic analysis showed that the proportion of simulations where magnesium sulphate prophylaxis was dominant depended on assumptions over the admission routines for patients receiving prophylaxis. If these were the same for both strategies then the proportion of simulations where prophylaxis was dominant was 41%. This reduced to 5% if it was assumed that patients receiving prophylaxis are admitted, on average, half a day early. The majority of simulations in both scenarios were associated with increased costs. All simulations had positive incremental effectiveness (i.e. magnesium prophylaxis was always effective in reducing the number of AF cases).

The CEAC shows that the probability of prophylaxis being cost-effective increases with increasing WTP for a unit of outcome. It is unclear what the appropriate decision threshold (threshold WTP) should be, with outcome in the model being expressed as AF cases averted.



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Contribution of authors

Jonathan Shepherd (Principal Research Fellow) was responsible for project management and report editing; analysis and interpretation of

clinical effectiveness studies; screening clinical effectiveness studies for inclusion; and report writing. Alison Price (Information Scientist) carried out the literature search. Geoff Frampton (Research Fellow) did the screening of clinical effectiveness studies for inclusion; data extraction and critical appraisal of clinical effectiveness studies; analysis and interpretation of clinical effectiveness studies; and report writing. Lukasz Tanajewski (Specialist) contributed to the data extraction and critical appraisal of clinical effectiveness studies; analysis and interpretation of clinical effectiveness studies; screening clinical effectiveness studies for inclusion; conception and design of economic model; and analysis and interpretation of results of economic model. Jeremy Jones (Principal Research Fellow) and David Turner (Senior Research Fellow) carried out the screening clinical effectiveness studies for inclusion; conception and design of economic model; analysis and interpretation of results of economic model; and report writing.



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65. Stamou SC, Hill PC, Dargas G, Pfister AJ, Boyce SW, Dullum MKC, *et al.* Stroke after coronary artery bypass: incidence, predictors, and clinical outcome editorial comment: Incidence, predictors, and clinical outcome. *Stroke* 2001;**32**:1508–13.
66. Miller PSJ, Andersson FL, Kalra L. Are cost benefits of anticoagulation for stroke prevention in atrial fibrillation underestimated? *Stroke* 2005;**36**:360–6.

Appendix I

Search strategy for RCTs and systematic reviews

Cochrane Library

2007 (Issue 2)

Searched 14 May 2007

- #1 MeSH descriptor Coronary Artery Bypass
explode all trees
- #2 (CABG or coronary artery bypass graft*):ti,ab
- #3 (coronary artery bypass NEAR/3
surgery):ti,ab
- #4 coronary revascular*:ti,ab
- #5 (#1 OR #2 OR #3 OR #4)
- #6 MeSH descriptor Magnesium Sulfate
explode all trees
- #7 magnesium (sulphate* or sulfate*):ti,ab
- #8 (#6 OR #7)
- #9 MeSH descriptor Atrial Fibrillation explode
all trees
- #10 (heart NEAR/3 fibrillat*):ti,ab
- #11 (atrial and fibrillat*):ti,ab
- #12 atrial fibrillat*
- #13 (#9 OR #10 OR #11 OR #12)
- #14 (#5 AND #8), from 2004 to 2007
- #15 (#13 AND #8), from 2004 to 2007
- #16 (#14 OR #15)
- #17 MeSH descriptor Sotalol explode all trees
- #18 sotalol
- #19 (#17 OR #18)
- #20 (#5 OR #13)
- #21 (#20 AND #19)
- #22 (#8 AND #19)
- #23 (#19 AND #5)

Number of hits (download file)

Cochrane Database of Systematic Reviews (CDSR)
magnesium: 1
CDSR sotalol: 8
CENTRAL magnesium: 8
CENTRAL magnesium + sotalol: 1
CENTRAL sotalol: 16

Ovid MEDLINE

1950 to May week 1 2007

Searched 15 May 2007

- 1 "Atrial Fibrillation"/ (19465)
- 2 Atrial Flutter/ (3991)
- 3 (atrial adj3 fibrillat\$.mp. (25464)
- 4 (heart adj3 fibrillat\$.mp. (950)
- 5 atrial fibrillat\$.mp. (25294)

- 6 (atrium adj3 flutter\$.mp. (23)
- 7 (atrium adj3 fibrillat\$.mp. (0)
- 8 (auricular\$ adj3 fibrillat\$.mp. (736)
- 9 (auricular\$ adj3 flutter\$.mp. (274)
- 10 Tachycardia Supraventricular/ (3781)
- 11 or/1-10 (30644)
- 12 "Coronary Artery Bypass"/ (31512)
- 13 (CABG or coronary artery bypass graft\$.ti,ab.
(16229)
- 14 (coronary artery bypass adj3 surgery).ti,ab.
(7332)
- 15 (coronary adj3 revasculari\$.ti,ab. (4439)
- 16 (myocardial adj3 revascular\$.ti,ab. (4063)
- 17 (coronary adj6 graft).ti,ab. (6832)
- 18 (coronary adj6 bypass).ti,ab. (26085)
- 19 (coronary adj6 surgery).ti,ab. (15257)
- 20 (aortocoronary adj6 bypass).ti,ab. (2453)
- 21 (valve\$ adj6 surgery).ti,ab. (4554)
- 22 Cardiac Surgical Procedures/ (23335)
- 23 or/12-22 (71271)
- 24 MgSO4.mp. (1029)
- 25 magnesium sulfate/ (3419)
- 26 magnesium sulfate\$.ti,ab. (1735)
- 27 exp magnesium compounds/ (10471)
- 28 magnesium/ (56302)
- 29 (magnesium and (sulphate\$ or sulfate\$)).ti,ab.
(3192)
- 30 magnesium sulphate\$.ti,ab. (729)
- 31 magnesium sulfate\$.ti,ab. (1735)
- 32 7487-88-9.rn. (3419)
- 33 Magnesium Deficiency/ (3387)
- 34 or/24-33 (67330)
- 35 3930-20-9.rn. (1695)
- 36 Sotalol/ (1695)
- 37 sotalol.ti,ab. (1997)
- 38 or/35-37 (2418)
- 39 11 and 23 and 34 (48)
- 40 23 and 34 (242)
- 41 11 and 34 (151)
- 42 39 or 40 or 41 (345)
- 43 limit 42 to (humans and english language and
yr="2004 - 2007") (45)
- 44 RANDOMISED CONTROLLED TRIAL.pt.
(234701)
- 45 CONTROLLED CLINICAL TRIAL.pt. (74869)
- 46 RANDOMISED CONTROLLED TRIALS.sh.
(48464)
- 47 RANDOM ALLOCATION.sh. (57810)
- 48 DOUBLE BLIND METHOD.sh. (91140)
- 49 SINGLE BLIND METHOD.sh. (10900)

- 50 or/44-49 (397972)
 51 CLINICAL TRIAL.pt. (435563)
 52 exp CLINICAL TRIALS/ (190865)
 53 (clin\$ adj25 trial\$).ti,ab. (129659)
 54 ((singl\$ or doubl\$ or trebl\$ or tripl\$) adj25
 (blind\$ or mask\$)).ti,ab. (90466)
 55 PLACEBOS.sh. (26144)
 56 placebo\$.ti,ab. (102017)
 57 random\$.ti,ab. (370373)
 58 RESEARCH DESIGN.sh. (47354)
 59 or/51-58 (843224)
 60 COMPARATIVE STUDY.sh. (0)
 61 exp Evaluation Studies/ (596080)
 62 FOLLOW UP STUDIES.sh. (337594)
 63 PROSPECTIVE STUDIES.sh. (220463)
 64 (control\$ or prospectiv\$ or volunteer\$).ti,ab.
 (1775327)
 65 or/60-64 (2564190)
 66 or/50,59,65 (2918050)
 67 43 and 66 (32)
 68 from 67 keep 1-32 (32)
 69 from 68 keep 1-32 (32)
 70 (review or review-tutorial or review-
 academic).pt. (1275632)
 71 (Medline or medlars or embase).ti,ab,sh. (23800)
 72 (scisearch or psychinfo or psycinfo).ti,ab,sh.
 (1277)
 73 (Psychlit or psyclit).ti,ab,sh. (709)
 74 cinahl.ti,ab,sh. (2178)
 75 ((hand adj59 search\$) or (manual\$ adj9
 search\$)).mp. [mp=title, original title,
 abstract, name of substance word, subject
 heading word] (3888)
 76 (electronic database\$ or bibliographic
 database\$ or computeri#ed database\$ or
 online database\$).mp. [mp=title, original title,
 abstract, name of substance word, subject
 heading word] (4129)
 77 (pooling or pooled or mantel haenszel).mp.
 [mp=title, original title, abstract, name of
 substance word, subject heading word] (25915)
 78 (peto or dersimonian or der simonian or fixed
 effect).mp. [mp=title, original title, abstract,
 name of substance word, subject heading word]
 (1092)
 79 71 or 72 or 73 or 74 or 75 or 76 or 77 or 78
 (52839)
 80 70 and 79 (19955)
 81 meta-analysis.pt. (15271)
 82 meta-analysis.sh. (7414)
 83 (meta-analys\$ or meta analys\$ or
 metaanalys\$).mp. [mp=title, original title,
 abstract, name of substance word, subject
 heading word] (27299)
 84 (systematic\$ adj9 review\$).mp. [mp=title,
 original title, abstract, name of substance
 word, subject heading word] (13485)
 85 (systematic\$ adj9 overview\$).mp. [mp=title,
 original title, abstract, name of substance
 word, subject heading word] (416)
 86 (quantitativ\$ adj9 review\$).mp. (1754)
 87 (quantitativ\$ adj9 overview\$).mp. [mp=title,
 original title, abstract, name of substance
 word, subject heading word] (142)
 88 (quantitativ\$ adj9 synthesis\$).mp. (1376)
 89 (methodologic\$ adj9 review\$).mp. (2493)
 90 (methodologic\$ adj9 overview\$).mp. (153)
 91 (integrative research review\$ or research
 integration).mp. (51)
 92 or/81-91 (41150)
 93 80 or 92 (53669)
 94 43 and 93 (6)
 95 from 94 keep 1-6 (6)
 96 11 or 23 (99880)
 97 34 and 38 and 96 (10)
 98 from 97 keep 1-10 (10)
 99 38 and 93 and 96 (27)
 100 from 99 keep 1-27 (27)
 101 100 (27)
 102 38 and 66 and 96 (313)
 103 limit 102 to (humans and english language)
 (268)
 104 103 not 99 (246)
 105 23 and 104 (45)
 106 104 not 105 (201)
 107 23 and 38 and 66 (52)
 108 limit 107 to (humans and english language)
 (51)
 109 23 and 38 and 92 (6)
 110 108 not 109 (47)

Number of hits (download file)

Systematic reviews (SRs) magnesium: 5

SRs magnesium + sotalol: 2

SRs sotalol: 6

RCTs magnesium: 24

RCTs magnesium + sotalol: 10

RCTs sotalol: 47

EMBASE

1980 to 2007 week 19

Searched 16 May 2007

- 1 heart atrium fibrillation/ (21875)
- 2 heart atrium flutter/ (3362)
- 3 exp Heart Atrium Arrhythmia/ (35266)
- 4 (atrial adj3 fibrillat\$).mp. (17803)
- 5 (heart adj3 fibrillat\$).mp. (31983)
- 6 atrial fibrillat\$.mp. (17577)
- 7 (atrium adj3 flutter\$).mp. (3380)
- 8 (atrium adj3 fibrillat\$r).mp. (0)
- 9 (auricular\$ adj3 fibrillat\$).mp. (108)
- 10 (auricular\$ adj3 flutter\$).mp. (29)

- 11 Supraventricular Tachycardia/ (5812)
 12 or/1-11 (47152)
 13 coronary artery bypass graft/ (22523)
 14 Coronary Artery Bypass Surgery/ (6944)
 15 exp coronary artery surgery/ (39530)
 16 (CABG or coronary artery bypass graft\$.ti,ab.
 (14959)
 17 (coronary artery bypass adj3 surgery).ti,ab.
 (6804)
 18 (coronary adj3 revasculari\$.ti,ab. (4087)
 19 (myocardial adj3 revascular\$.ti,ab. (3015)
 20 (coronary adj6 graft).ti,ab. (6026)
 21 (coronary adj6 bypass).ti,ab. (23026)
 22 (coronary adj6 surgery).ti,ab. (13234)
 23 (aortocoronary adj6 bypass).ti,ab. (1469)
 24 (valve\$ adj6 surgery).ti,ab. (3618)
 25 or/13-24 (50320)
 26 MgSO4.mp. (1011)
 27 magnesium sulfate/ (6009)
 28 Magnesium Derivative/ (989)
 29 magnesium/ (23019)
 30 (magnesium and (sulphate\$ or sulfate\$)).ti,ab.
 (2910)
 31 magnesium sulphate\$.ti,ab. (642)
 32 magnesium sulfate\$.ti,ab. (1548)
 33 7487-88-9.rn. (6022)
 34 Magnesium Deficiency/ (1203)
 35 or/26-34 (30923)
 36 3930-20-9.rn. (7170)
 37 Sotalol/ (7144)
 38 sotalol.ti,ab. (2048)
 39 or/36-38 (7279)
 40 12 and 25 and 35 (88)
 41 25 and 35 (209)
 42 12 and 35 (667)
 43 41 or 42 (788)
 44 limit 43 to (humans and english language and
 yr="2004 - 2007") (229)
 45 12 or 25 (95066)
 46 35 and 39 and 45 (149)
 47 exp Postoperative Complication/ (188542)
 48 exp meta analysis/ (30150)
 49 meta#analy\$.ab,sh,ti. (30151)
 50 methodologic\$ review\$.ab,sh,ti. (128)
 51 methodologic\$ overview\$.ab,sh,ti. (31)
 52 (integrative research adj5 review\$.mp. or
 research integration.ab,ti. [mp=title, abstract,
 subject headings, heading word, drug trade
 name, original title, device manufacturer, drug
 manufacturer name] (28)
 53 quantitat\$ synthesis.ab,sh,ti. (91)
 54 quantitat\$ review\$.ab,sh,ti. (260)
 55 quantitat\$ overview\$.ab,sh,ti. (60)
 56 systematic\$ review\$.ab,sh,ti. (23698)
 57 systematic\$ overview\$.ab,sh,ti. (287)
 58 48 or 49 or 50 or 51 or 52 or 53 or 54 or 55
 or 56 or 57 (43716)
 59 44 and 58 (28)
 60 from 59 keep 1-2,4-5,8-9,12-13,15,17-20,22-
 25,28 (18)
 61 randomisation/ (22296)
 62 controlled study/ (2402316)
 63 single blind procedure/ (6576)
 64 placebo/ (98219)
 65 double blind procedure/ (63899)
 66 clinical trial/ (423526)
 67 crossover procedure/ (18621)
 68 placebo\$.tw. (98196)
 69 blind\$ fashion.tw. (3585)
 70 random\$.tw. (335411)
 71 clinical trial?.tw. (99670)
 72 or/61-71 (2809080)
 73 limit 72 to human (1793268)
 74 44 and 73 (120)
 75 74 not 59 (92)
 76 from 75 keep 5,7,10,16,18,20,22-25,27,30-
 31,33,39-40,43,45,48,50-51,58,60,65,68-
 77,80,84,86,88,90-92 (41)
 77 46 and 58 (18)
 78 77 not (59 or 75) (4)
 79 from 78 keep 1-4 (4)
 80 46 and 73 (69)
 81 limit 80 to english language (66)
 82 81 not 74 (34)
 83 from 82 keep 2-4,7,10-12,14-15,17,19-24,26-
 27,31,34 (20)
 84 12 and 25 and 39 and 58 (18)
 85 from 84 keep 1,3 (2)
 86 84 not (78 or 59 or 75 or 81) (12)
 87 from 86 keep 1,3-5,7-8,10 (7)
 88 12 and 25 and 39 and 73 (102)
 89 limit 88 to (human and english language) (91)
 90 89 not 84 (74)
 91 from 90 keep 1-6,9,15,17-20,25-
 27,30,36,38,41,43-46,49,52-53,55-63,65-74
 (45)

Number of hits (download file)

SRs magnesium: 18
 SRs magnesium + sotalol: 4
 SRs sotalol: 20
 RCTs magnesium: 41
 RCTs magnesium + sotalol: 20
 RCTs sotalol: 45

Ovid MEDLINE In-Process and Other Non-Indexed Citations

14 May 2007
 Searched 15 May 2007
 1 (atrial adj3 fibrillat\$.mp. (728)
 2 (heart adj3 fibrillat\$.mp. (32)
 3 atrial fibrillat\$.mp. (727)

- 4 (atrium adj3 flutter\$.mp. (1)
 5 (atrium adj3 fibrillat\$.mp. (0)
 6 (auricular\$ adj3 fibrillat\$.mp. (16)
 7 (auricular\$ adj3 flutter\$.mp. (0)
 8 or/1-7 (747)
 9 (CABG or coronary artery bypass graft\$.ti,ab.
 (482)
 10 (coronary artery bypass adj3 surgery).ti,ab.
 (220)
 11 (coronary adj3 revasculari\$.ti,ab. (134)
 12 (myocardial adj3 revascular\$.ti,ab. (83)
 13 (coronary adj6 graft).ti,ab. (173)
 14 (coronary adj6 bypass).ti,ab. (638)
 15 (coronary adj6 surgery).ti,ab. (353)
 16 (aortocoronary adj6 bypass).ti,ab. (22)
 17 (valve\$ adj6 surgery).ti,ab. (117)
 18 or/9-17 (968)
 19 MgSO4.mp. (38)
 20 (magnesium and (sulphate\$ or sulfate\$)).ti,ab.
 (98)
 21 magnesium sulphate\$.ti,ab. (23)
 22 magnesium sulfate\$.ti,ab. (55)
 23 19 or 20 or 21 or 22 (129)
 24 sotalol.ti,ab. (30)
 25 8 or 18 (1659)
 26 23 and 25 (3)
 27 from 26 keep 1-3 (3)
 28 24 and 25 (14)
 29 from 28 keep 5,8,10-13 (6)

Number of hits (download file)

RCTs magnesium: 3
 RCTs sotalol: 6

DARE (on CRD databases)

magnesium and (atrial fibrillat* or coronary)
 Sotalol and (atrial fibrillat* or coronary)

Number of hits (download file)

Magnesium: 7
 Sotalol: 10

HTA Database (on CRD databases)

Magnesium and (atrial fibrillat* or coronary)
 Sotalol and (atrial fibrillat* or coronary)

Number of hits

0

National Research Register

- #1. magnesium 197
 #2. (atrial next fibrillat*) 442
 #3. ((coronary or heart or atri*) and surgery) 1991
 #4. (#1 and (#2 or #3)) 10

Number of hits (download file)

Single-centre projects (ongoing): 2
 Single-centre (completed): 7
 Lead centre for multicentre projects (completed):
 1

Current Controlled Trials including MRC Trials Database

<http://controlled-trials.com/>
 Magnesium and (atrial fibrillat* or coronary)

Number of hits (download file)

Magnesium: 1 (UK)

Clinical Trials.gov

<http://clinicaltrials.gov/>
 Magnesium and (atrial fibrillation or coronary)
 Sotalol and (atrial fibrillation or coronary)

Number of hits (download file)

Magnesium: 1
 Sotalol: 5

TOTAL KEYWORDED

SRs magnesium: 15
 SRs magnesium + sotalol: 5
 SRs sotalol: 26
 RCTs magnesium: 45
 RCTs magnesium + sotalol: 28
 RCTs sotalol: 71

Appendix 2

Decision tree for screening abstracts and full papers

Trial name or number: Adult patients with AF after elective CABG	Yes ↓ next question	Unclear ↓ next question	No → Exclude	Type: Exclude 1 (not the correct patient group)
Either on-pump or off-pump CABG techniques; any number of grafts; any conduit type (CABG must be specified; with or without valvular surgery)				
Design: RCT or systematic review	Yes ↓ next question	Unclear ↓ next question	No → Exclude	Exclude 2 (not the right study design)
Intervention^a IV magnesium sulphate as bolus or continuous infusion, or sotalol (oral or IV) or IV magnesium sulphate combined with sotalol, each of a specified dose and duration, as a prophylactic measure before onset of AF. (NB. Oral Mg, Mg chloride and Mg hydroxide are outside the scope of the review)	Yes ↓ next question	Unclear ↓ next question	No → Exclude	Exclude 3 (not the right intervention)
Report one or more of primary outcomes: incidence of AF after CABG (include broader arrhythmia definitions only if proportion with AF reported), measured using a continuous electrocardiogram (ECG) and confirmed by standard 12-lead ECG	Yes ↓ next question	Unclear ↓ next question	No → Exclude	Exclude 4 (not the right outcome measures)
Final decision	Include	Unclear (discuss)	Exclude	Results of discussion:
IV, intravenous. ^a Compared with placebo and/or sotalol.				

Appendix 3

Data extraction templates for the 15 RCTs included in this review

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Bert et al.¹⁵</p> <p>Year: 2001</p> <p>Country: USA</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: NR</p>	<p><i>Interventions for which data extracted:</i></p> <ol style="list-style-type: none"> 1. IV magnesium sulphate 2. Control <p><i>Other groups (data not extracted):</i></p> <ol style="list-style-type: none"> 3. Digoxin 4. MgSO₄ + digoxin 5. Propranolol 6. MgSO₄ + propranolol <p><i>Intervention details:</i></p> <ol style="list-style-type: none"> 1. 2 g magnesium sulphate diluted in 50 ml of normal saline infused i.v. during 30 minutes after termination of CPB (after protamine neutralisation) and again on arrival in the ICU. A daily dose of 2 g magnesium sulphate then infused each morning for the first 4 days postoperation for a total of 12 g in 96 h. No patients received magnesium pre-CPB 2. Received no antiarrhythmic agent (not stated whether normal saline was administered as in group 1) 	<p><i>Number of participants:</i></p> <p>Intervention: 63 Control: 60</p> <p><i>Sample attrition/dropout:</i></p> <p>Three patients were excluded after randomisation, but it was not reported whether these were from the magnesium and/or control groups, nor whether any data from these patients were analysed</p> <p><i>Inclusion criteria for study entry:</i></p> <p>Patients scheduled for primary CABG surgery were eligible for enrolment, including patients with additional aortic valve replacement surgery</p> <p><i>Exclusion criteria for study entry:</i></p> <p>Patients with: a history of atrial or ventricular arrhythmias, or any rhythm other than sinus rhythm on the ECG obtained the evening before surgery; LVEF ≤20% on angiography; airway disease requiring bronchodilator therapy; renal failure requiring haemodialysis; other surgical procedures (mitral valve surgery aortic arch replacement, or ventricular aneurysmectomy)</p> <p><i>Baseline characteristics of subjects:</i></p> <p><i>Mean ± SD age (years):</i></p> <p>Intervention: 62.7 ± 9.7^b Control: 63.6 ± 9.6^b (difference was tested only across six study groups)</p> <p><i>Gender (M/F):</i></p> <p>Intervention: 56/7 Control: 50/10 (difference was tested only across six study groups)</p> <p>Magnesium/control group values appeared similar for: LVEF, % β-blockers, Ca channel blockers, digoxin, diuretics, number of bypass grafts, % incomplete revascularisation, % inotropic use, number of patients with CABG + valve surgery (3/1) and mean cross-clamp time (59.6/55.1 minutes) (differences were tested only across six study groups)</p>	<p><i>AF outcome(s):</i></p> <p>Incidence of POAT^a (primary outcome)</p> <p><i>Other outcomes extracted:</i></p> <p>Duration of hospital stay</p> <p><i>Other outcomes (data not extracted):</i></p> <p>Time to extubation, MI, ventricular ectopic activity, initial ventricular rate</p> <p><i>Adverse symptoms:</i></p> <p>No deaths occurred in either the magnesium or control group during the study. The proportions of patients with MI or ventricular ectopic activity appeared similar in these groups</p> <p><i>Length of follow-up:</i></p> <p>4 days after CABG</p> <p><i>Recruitment dates:</i></p> <p>3-year period (dates not specified)</p>
<p>^a POAT (post-operative atrial tachyarrhythmia) includes AF, atrial flutter or an SVT that was sustained for >5 minutes and warranted pharmacological therapy.</p> <p>^b Variance measure not stated.</p>			
			<i>continued</i>

Results			
AF outcome	Intervention (n = 63)	Control (n = 60)	Difference
Incidence of POAT, n (%) ^c	24 ^d (38%)	23 ^d (38%)	p = 0.98
<p>^c Overall, across all six study groups, 90.4% of the POAT were AF (the proportion of AF in the magnesium and control groups was not reported).</p> <p>^d Calculated by reviewers.</p>			
Other outcomes	Intervention (n = 63)	Control (n = 60)	Difference
Duration of hospital stay (days)	8.2 ± 3.1 ^e	8.0 ± 2.9 ^e	Not reported for this comparison; for comparison across six groups p = 0.69 (ns)
<p>^e Assumed by reviewers (not stated) to be the mean value; variance measure not stated.</p>			
Methodological comments			
<p><i>Allocation to treatment groups:</i> Stated that patients were randomised to the study groups based on a table of random numbers (no other details provided).</p> <p><i>Blinding:</i> Stated that all arrhythmia tracings were reviewed by a cardiologist unaware of the patient's treatment group. No other details of blinding were given.</p> <p><i>Comparability of treatment groups:</i> Baseline characteristics of the groups appear similar; differences were not significant if all six study groups were tested together (no pairwise tests of baseline characteristics were reported).</p> <p><i>Method of data analysis:</i></p> <ul style="list-style-type: none"> • One-way analysis of variance and 2 × 6 contingency tables with χ^2 test were used to compare baseline characteristics, intraoperative measures and POAT among the six treatment groups. • 2 × 2 contingency tables with χ^2 test were used to compare the incidence of POAT between each pair of treatment groups. • An independent samples t-test was used to compare age between POAT and non-POAT groups (data not extracted). • Statistical tests were not supported by hypotheses; normality of data was not reported. <p><i>Sample size/power calculation:</i> NR.</p> <p><i>Attrition/dropout:</i> Three patients were excluded after randomisation, but it was not reported whether these were from the magnesium and/or control groups, nor whether any data from these patients were analysed. Patients experiencing AF were regarded as having completed the study protocol (the end-point and duration of study therefore varied among patients).</p>			
General comments			
<p><i>Generalisability:</i> No details of ethnicity or social background; single-location study in USA.</p> <p><i>Outcome measures:</i> POAT mainly comprised AF (90.4% overall), but AF was not reported separately for individual treatment comparisons.</p> <p><i>Intercentre variability:</i> NA.</p> <p><i>Conflict of interests:</i> None reported.</p>			
<p>LVEF, left ventricular ejection fraction; MI, myocardial infarction; ns, not significant; POAT, postoperative atrial tachycardia; SVT, supraventricular tachycardia.</p>			

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Bhudia et al.¹⁶</p> <p>Year: 2006</p> <p>Country: USA</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: NR</p>	<p>Interventions:</p> <p>1. IV magnesium sulphate</p> <p>2. Placebo control</p> <p>Intervention details:</p> <p>1. 780 mg (32 mmol) magnesium sulphate in 100 ml of normal saline given i.v. over 15 minutes during anaesthesia induction, followed by 3160 mg (130 mmol) in 100 ml of normal saline over 24 h; the CPB circuit was primed with magnesium sulphate to a concentration of 3.6 to 4.8 mg dl⁻¹</p> <p>2. Patients received normal saline given i.v. in the CPB circuit and over 24 h in bags and syringes indistinguishable from those used for group 1</p>	<p>Number of participants:</p> <p>Intervention: 174</p> <p>Control: 176</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry: Patients undergoing elective on-pump CABG [magnesium 71 (41%), placebo 57 (32%)], valve surgery [magnesium 59 (34%), placebo 65 (37%)] or combined CABG + valve surgery [magnesium 44 (25%), placebo 54 (31%)]</p> <p>Exclusion criteria for study entry: Patients with preoperative AF or renal impairment (this may be an incomplete list^a)</p> <p>Baseline characteristics of subjects:</p> <p>Mean ± SD age (years):</p> <p>Intervention: 64 ± 12</p> <p>Control: 64 ± 13</p> <p>Difference: <i>p</i> = 0.7 (ns)</p> <p>Gender (M/F):</p> <p>Intervention: 133/41 (24% F)</p> <p>Control: 137/39 (22% F)</p> <p>Difference: <i>p</i> = 0.8 (ns)</p> <p>Differences between groups were not statistically significant (<i>p</i> > 0.5) for: BMI, diabetes, hypertension, COPD, carotid disease, smoking history, MI history, preoperative AF, previous cardiac surgery, NYHA class, moderate or greater aortic regurgitation, mitral regurgitation and/or mitral stenosis, number undergoing isolated CABG or CABG + valve surgery, perfusion time, ischaemic time, education duration and standards</p> <p>Differences between groups were statistically significant for peripheral vascular disease (intervention 77, control 59, <i>p</i> = 0.04) and moderate or greater aortic stenosis (intervention 36, control 57, <i>p</i> = 0.01)</p>	<p>AF outcome(s):</p> <p>Incidence of AF (safety outcome)</p> <p>Other extracted outcomes:</p> <p>Length of stay in ICU and hospital</p> <p>Other outcomes (data not extracted): MI, stroke, return to operating room for treatment, blood transfusion units, neurological and psychological assessments (primary outcome), plasma magnesium concentration, renal insufficiency, respiratory insufficiency, septicemia</p> <p>Adverse symptoms:</p> <p>One death occurred in each group in hospital. Within 3 months of enrolment there was one death in the magnesium group and there were four deaths in the placebo group (<i>p</i> = 0.2, ns); these were not related to stroke or neurological injury</p> <p>Length of follow-up:</p> <p>Not stated for the outcomes extracted here (median length of postoperative stay was 6 days)</p> <p>Recruitment dates:</p> <p>February 2003 to September 2003</p>
<p>^a The authors refer to Figure E1 and Appendix E1 for exclusion criteria, but these sources were not included in the primary article and no links for locating them elsewhere were provided</p>			
Results			
AF outcome	Intervention (n = 174)	Control (n = 176)	Difference
No. (%) of patients with AF	57 (33%)	64 (36%)	<i>p</i> = 0.5 (ns)
<i>continued</i>			

Other outcomes	Intervention (n = 174)	Control (n = 176)	Difference
Median length of stay (days) in ICU (15th, 85th percentiles)	2 (1, 3)	2 (1, 3)	$p > 0.9$ (ns)
Median length of postoperative stay (days) (15th, 85th percentiles)	6 (5, 9)	6 (5, 9)	$p > 0.9$ (ns)
Median length of hospital stay (days) (15th, 85th percentiles)	7 (5, 11)	7 (5, 10)	$p = 0.3$ (ns)

Methodological comments

Allocation to treatment groups: Patients were randomised 1:1 at operation, with a block size of 2 and 4, independently for each of 11 operating rooms (no other details).

Blinding: The pharmacy department who prepared study medications and the anaesthetic team who administered the 24-h infusions were blinded to the study. The perfusion teams who dosed the CPB circuits were not blinded; they were required to monitor and maintain magnesium levels during CPB.

Comparability of treatment groups: Significantly more of the magnesium group had PVD, while significantly more of the control group had aortic stenosis; other variables did not differ between the groups (p -values were provided for all variables).

Method of data analysis:

- Continuous data with a skewed (unspecified) distribution were analysed with a Wilcoxon rank-sum test.
- Categorical data were analysed using a χ^2 test, or Fisher's exact test when the frequency was <5 .
- Presentation of results did not indicate which tests were actually used for each outcome; normality of outcomes was not reported.
- Statistical tests were not supported by hypotheses.

Sample size/power calculation: Stated that the study was powered for 300 patients, based on neurological assessments as the primary outcome, but no details given.

Attrition/dropout: NR.

General comments

Generalisability: No details of ethnicity or social background; single-location study in USA.

Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: None reported.

BMI, body mass index; COPD, chronic obstructive pulmonary diseases; NYHA, New York Heart Association; PVD, peripheral vascular disease.

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Caspi <i>et al.</i>¹³</p> <p>Year: 1995</p> <p>Country: Israel</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: NR</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Placebo control <p>Intervention details:</p> <ol style="list-style-type: none"> Two doses of magnesium sulphate continuously with a syringe pump (Graseby 3100; Watford, Herts, UK): <ul style="list-style-type: none"> 16 mmol from the time of anaesthetic induction to aortic cross-clamping 32 mmol after the release of aortic cross-clamping until 24 h later 20 ml of saline solution 	<p>Number of participants:</p> <p>Intervention: 50</p> <p>Control: 48</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry: Patients with unstable angina (grade IV) undergoing CABG, the criteria for unstable angina: the presence of chest pain at rest lasting for more than 15 minutes, associated with transient ST segment changes and with or without haemodynamic instability</p> <p>Exclusion criteria for study entry: NR</p> <p>Baseline characteristics of subjects:</p> <p>Mean age (range) (years):</p> <p>Intervention: 60 (41–78)</p> <p>Control: 62 (43–76)</p> <p>Gender (M/F) (n):</p> <p>Intervention: 34/16</p> <p>Control: 38/10</p> <p>Authors state that other baseline values (e.g. previous MI, use of β-blockers and Ca^{2+} blockers, ejection fraction, left main coronary artery disease) do not differ significantly between groups. No statistical tests reported. Two patients in intervention group and one patient in control group required intra-aortic balloon support before operation because of low cardiac output</p>	<p>AF outcome(s) (<i>1st = primary outcome</i>):</p> <p>The incidence of postoperative AF defined as AF at rates greater than 120 bpm</p> <p>Other outcomes (data not extracted):</p> <p>Premature ventricular contractions, ventricular tachycardia, ventricular fibrillation, plasma magnesium ion concentration, haemodynamic variables</p> <p>Adverse symptoms: There was one operative death in intervention group (low cardiac output and multiple organ failure). Postoperative hypertension: one in magnesium group, 14 in placebo group ($p < 0.05$). Postoperative ventricular tachycardia, 0 in magnesium group, one in placebo group. Frequent premature ventricular beats requiring lidocaine: one in magnesium group, 12 in placebo group ($p < 0.05$). No other differences between groups reported by authors</p> <p>Length of follow-up: 24 h after operation</p> <p>Recruitment dates: NR</p>
Results			
AF outcome	Intervention (n = 50)	Control (n = 48)	Difference
Incidence of postoperative AF	22 (44.0%)	18 (37.5%)	NR
<p>Authors reported only the numbers of patients with AF in intervention and placebo groups. Other outcomes are given with <i>p</i>-value or standard deviation.</p> <p>Methodological comments</p> <p><i>Allocation to treatment groups:</i> It is stated that patients were randomly assigned to receive treatment or placebo in coded ampules. The means by which an allocation sequence was generated is not reported and the allocation concealment is unclear.</p> <p><i>Blinding:</i> Patients were blinded (received treatment or placebo in coded ampules) and cardiologists (who interpreted the ECG) were also blinded to whether the patient received the magnesium sulphate.</p> <p><i>Comparability of treatment groups:</i> No notable differences; authors report that the groups were similar in baseline characteristics and other preoperative and operative characteristics. No statistical tests reported.</p>			
<i>continued</i>			

Method of data analysis: Patients' parameters were compared between two groups using an unpaired t-test and χ^2 test, as required. However, no statistical method used in the analysis of the outcome concerning AF.

Sample size/power calculation: The power calculations and variance estimates are not provided.

Attrition/dropout: NR.

General comments

Generalisability: No details of ethnicity or social background; single location in Israel.

Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: NR.

Reference and design	Intervention	Participants	Outcome measures
<p><i>Authors:</i> Dagdelen et al.^{17,18a}</p> <p><i>Year:</i> 2002 (2003^a)</p> <p><i>Country:</i> Turkey</p> <p><i>Study design:</i> RCT</p> <p><i>Number of centres:</i> NR (assumed one)</p> <p><i>Funding:</i> Not stated</p>	<p><i>Interventions:</i></p> <p>1: Intravenous magnesium sulphate</p> <p>2: Control (not reported as placebo)</p> <p><i>Intervention details:</i></p> <p>1: 1.5 g day⁻¹ magnesium sulphate in 100 ml of 0.9% sodium chloride (25 ml h⁻¹) 1 day before CPB, just after CPB surgery and then once daily for 4 days after CPB</p> <p>2: 100 ml of 0.9% sodium chloride (25 ml h⁻¹) at the same time-points as in the magnesium sulphate group</p>	<p><i>Number of participants:</i></p> <p>Intervention: 93</p> <p>Control: 55</p> <p><i>Sample attrition/dropout:</i> NR</p> <p><i>Inclusion criteria for study entry:</i> Patients scheduled for elective, first time, isolated CABG (no other details)</p> <p><i>Exclusion criteria for study entry:</i> Patients with AF, a past history of AF, heart valve disease, diabetes, chronic renal disease, thyroid disorders and/or COPD; cardiomyopathy, congenital heart disease, congestive heart failure, pericarditis, pulmonary embolism, pre-excitation syndromes, sick sinus syndrome, complete bundle branch block, atrioventricular block, and patients receiving antiarrhythmic drugs, digoxin and/or β-blocking agents</p> <p><i>Baseline characteristics of subjects:</i></p> <p><i>Mean \pm SD age (years):</i></p> <p>Intervention: 62.5 \pm 7.0</p> <p>Control: 61.3 \pm 7.0; difference ns</p> <p><i>Gender (M/F) (n):</i> Intervention: 93/21</p> <p>Control: 43/12; difference: ns</p> <p>Preoperative heart rate, magnesium concentration, QRS interval duration, PQ interval duration, P wave maximum, P wave minimum, and P wave dispersion did not differ significantly between groups (data not extracted)</p>	<p><i>AF outcome(s):</i> Incidence of AF after CABG</p> <p><i>Other outcomes (data not extracted):</i> Heart rate; magnesium concentration. Also ECG characteristics of QRS, PQ intervals and P wave</p> <p><i>Adverse symptoms:</i> NR</p> <p><i>Length of follow-up:</i> 4 days after CABG surgery</p> <p><i>Recruitment dates:</i> NR</p>
<p>^a Dagdelen et al. (2003)¹⁸ is a brief summary of the same data.</p>			
			<i>continued</i>

Results			
AF outcome	Intervention (n = 93)	Control (n = 55)	Difference
Frequency of AF occurrence (n) (%)	2 (2%)	20 (36%)	$p < 0.001$
Other outcomes	Outside scope of assessment: data not extracted		
Methodological comments			
<i>Allocation to treatment groups:</i> Stated that this was randomised, but no details provided.			
<i>Blinding:</i> Control group patients received the same doses and times of saline solution administration as the intervention group, but without magnesium; cardiologist reading ECGs was blinded.			
<i>Comparability of treatment groups:</i> No notable differences.			
<i>Method of data analysis:</i>			
<ul style="list-style-type: none"> • Stated that demographic and clinical variables as well as the incidence of AF were analysed using the χ^2-test or Fisher's exact test for categorical variables and paired or unpaired <i>t</i>-tests for continuous variables. However, the results presented do not indicate in which cases each of these tests was applied, nor whether the data satisfied assumptions of normality. • Overall, statistical tests were not supported by any hypotheses. 			
<i>Sample size/power calculation:</i> NR.			
<i>Attrition/dropout:</i> NR.			
General comments			
<i>Generalisability:</i> No details of ethnicity; limited information on social background; single-location study in Turkey.			
<i>Outcome measures:</i> Appropriate, but note that 12-lead Holter ECG was only used for reactive, not continuous, monitoring (authors discuss whether this might have missed some cases of AF).			
<i>Intercentre variability:</i> NA.			
<i>Conflict of interests:</i> None reported.			

Reference and design	Intervention	Participants	Outcome measures	
<p>Authors: Fanning et al.¹⁹</p> <p>Year: 1991</p> <p>Country: USA</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: NR</p>	<p>Interventions:</p> <p>1. Intravenous magnesium sulphate</p> <p>2. Placebo</p> <p>Intervention details:</p> <p>1. Study group maintenance fluids of 5% dextrose in water with 20 mEq of potassium chloride I⁻¹ plus 40 mEq magnesium sulphate I⁻¹, given as a continuous infusion at 100 ml h⁻¹ over the first 24 h postoperatively, followed by 25 ml h⁻¹ over postoperative hours 25–96. Magnesium sulphate given was 96 mEq in first 24 h (4 mEq h⁻¹) followed by 72 mEq over hours 25–96 (1 mEq h⁻¹)</p> <p>2. Placebo maintenance fluids of 5% dextrose in water with 20 mEq of potassium chloride I⁻¹, given as a continuous infusion according to the same protocol as group 1</p>	<p>Number of participants:</p> <p>Intervention: 49</p> <p>Control: 50</p> <p>Sample attrition/dropout:</p> <p>NR</p> <p>Inclusion criteria for study entry: Patients undergoing elective first time CABG</p> <p>Exclusion criteria for study entry: Serum creatinine >2.5 mg l⁻¹; history of second or third degree heart block; with a permanent pacemaker; any documented or suspected supraventricular or ventricular arrhythmias, including isolated atrial or ventricular premature depolarisations noted on preoperative surface ECG; patients requiring additional procedures such as valve replacement or left ventricular aneurysmectomy; patients refusing to participate</p> <p>Baseline characteristics of subjects:</p> <p>Age (years) (range):^b</p> <p>Intervention: 59 (43–75)</p> <p>Control: 62 (42–79)</p> <p>Gender (M/F):</p> <p>Intervention: 35/14</p> <p>Control: 39/11</p> <p>The authors stated that the two groups were well matched with regard to age, gender and history of MI, and that there was no significant difference in β-blocking or calcium channel-blocking medication, although preoperative digoxin use was more common in the control group (digitalis 0 vs 7 in magnesium and control groups; no <i>p</i>-value provided). Also stated that left ventricular systolic function, ischaemic time, number of grafts and the method of myocardial protection were similar between the groups (no <i>p</i>-values provided). LVEF, plasma magnesium concentration, number of grafts and duration of cross-clamp were also reported (no comments or <i>p</i>-values provided)</p>	<p>AF outcome:</p> <p>Incidence of AF (primary outcome)^a</p> <p>Other outcomes (data not extracted):</p> <p>Postoperative MI; serum magnesium concentration; ventricular arrhythmias</p> <p>Adverse symptoms: Stated that no side-effects could be attributed to the magnesium therapy. One death occurred in the control group during the study</p> <p>Length of follow-up:</p> <p>4 days post-CABG</p> <p>Recruitment dates:</p> <p>NR</p>	
<p>^a The primary article does not clearly define whether AF includes atrial flutter and/or SVT.</p> <p>^b Not stated whether mean or median.</p>				
Results				
AF outcomes		Intervention (n = 49)	Control (n = 50)	Difference
No. of patients with AF (no. of AF episodes)		7 (12)	14 (42)	<i>p</i> < 0.02 ^c
No. of patients with persistent and/or recurrent (> 1 episode) AF		2	9	NR
<p>^c Unclear whether statistical test refers to the number of patients or the number of episodes.</p>				
				<i>continued</i>

Other outcomes **Outside scope of assessment: data not extracted****Methodological comments**

Allocation to treatment groups: Stated that patients were randomised, but no details were provided.

Blinding: Stated that: (1) all patients received the intervention or placebo in a double-blind fashion (no details); (2) the patient's physicians were not blinded to the results of serum magnesium testing but no patient received additional magnesium therapy.

Comparability of treatment groups: Preoperative digoxin and diuretic use was more frequent in the control group; other baseline variables were similar between the groups (reported narratively; no *p*-values provided).

Method of data analysis:

- The incidence of postoperative arrhythmias was analysed for group differences using a two-sided χ^2 analysis.
- Mean serum magnesium levels and the mean number of episodes of arrhythmias between groups were analysed using a paired or two-sample *t*-test where appropriate.
- Presentation of results did not indicate which tests were actually used for each outcome; normality of outcomes was not reported.

Sample size/power calculation: NR.

Attrition/dropout: NR.

General comments

Generalisability: No details of ethnicity or social background; single-location study in USA.

Outcome measures: Appropriate, but note ambiguous definition of AF (see above).

Intercentre variability: NA.

Conflict of interests: None reported.

Reference and design	Intervention	Participants	Outcome measures
<p><i>Authors:</i> Forlani et al.^{20,21a}</p> <p><i>Year:</i> 2002 (2003^a)</p> <p><i>Country:</i> Italy</p> <p><i>Study design:</i> RCT</p> <p><i>Number of centres:</i> NR (assumed one)</p> <p><i>Funding:</i> NR</p>	<p><i>Interventions:</i></p> <ol style="list-style-type: none"> 1. Intravenous magnesium sulphate 2. Control 3. Oral sotalol 4. Intravenous magnesium sulphate + oral sotalol <p><i>Intervention details:</i></p> <ol style="list-style-type: none"> 1. 1.5 g (12 mEq) i.v. magnesium sulphate daily, starting just before CPB until 5 days postoperation 2. No drug 3. 80 mg sotalol given orally twice daily from 1 day to 5 days postoperation, then 40 mg orally twice daily for 4 weeks 4. A combination of magnesium sulphate and sotalol in the same dosages as groups 1 and 3 	<p><i>Number of participants:</i></p> <ol style="list-style-type: none"> 1. Mg: 54 2. Control: 50 3. Sotalol: 51 4. Mg + sotalol: 52 <p><i>Sample attrition/dropout:</i> Excluded from study:</p> <ol style="list-style-type: none"> 1. Magnesium: three patients: one due to intraoperative MI with reduced cardiac index; two due to AF on day of operation 2. Control: no withdrawals reported 3. Sotalol: one patient due to prolongation of QTc 4. Magnesium + sotalol: two patients, one due to intraoperative MI with reduced cardiac index; one due to bradycardia <p>In addition, patients were excluded on the morning of the first postoperative day (i.e. after commencement of intervention) if heart rate <60 bpm (beats per minute), systolic arterial pressure <100 mmHg, wedge pulmonary pressure >15 mmHg, cardiac index <2.5 l minute⁻¹ m⁻² with inotropic support (dopamine dose >3 µg kg⁻¹ minute⁻¹). However, the group(s) from which these patients were excluded were not reported</p>	<p><i>AF outcomes:</i> Incidence, and time to onset, of AF</p> <p><i>Other outcomes extracted:</i> Postoperative length of stay</p> <p><i>Other outcomes (data not extracted):</i> Heart rate during AF; serum magnesium concentration</p> <p><i>Adverse symptoms:</i> One patient in the magnesium group died of a pulmonary embolism 5 days after operation. One patient in the magnesium group and one in the magnesium + sotalol group (both excluded) had intraoperative MI with reduced cardiac index. Sotalol was cautiously discontinued</p>

continued

Reference and design	Intervention	Participants	Outcome measures		
		<p><i>Inclusion criteria for study entry:</i> Patients undergoing first time isolated CABG with CPB</p> <p><i>Exclusion criteria for study entry:</i> Patients with preoperative ejection fraction <0.40, sick sinus syndrome and atrioventricular node disease, a corrected QT interval >440 ms; preoperative use of antiarrhythmic drugs, with the exception of β-blockers; history of SVA, severe COPD, serum creatinine levels >2.0 mg dl⁻¹</p> <p><i>Baseline characteristics of subjects:</i></p> <p><i>Mean \pm SD age (years):</i></p> <ol style="list-style-type: none"> 1. Mg: 64 \pm 7 2. Control: 64 \pm 9 3. Sotalol: 64 \pm 10 4. Mg+sotalol: 62 \pm 11 <p><i>Gender (% M):</i></p> <ol style="list-style-type: none"> 1. Mg: 85% 2. Control: 88% 3. Sotalol: 82% 4. Mg+sotalol: 91% <p>With reference to 18 listed variables, the authors stated that preoperative and intraoperative variables did not differ significantly among the four groups (SD provided for six of these variables)</p>	<p>3 days postoperation in two patients (both excluded) due to bradycardia (one patient in magnesium + sotalol group) and prolongation of QTc (one patient in sotalol group). The authors reported that there were no noted proarrhythmic effects of magnesium, sotalol or magnesium + sotalol</p> <p><i>Length of follow-up:</i> End-point of study was considered any AF episode that required treatment for symptoms or haemodynamic deterioration from 1 day to 1 month postoperation. Interventions were administered for 5 days postoperation</p> <p><i>Recruitment dates:</i> January to July 2001</p>		
^a Forlani et al. (2003) ²¹ is a brief summary of the same data.					
Results					
AF Outcomes	1. Mg (n = 54)	2. Control (n = 50)	3. Sotalol (n = 51)	4. Mg + sotalol (n = 52)	Difference
Incidence of AF, n (%)	8 (14.8%)	19 (38%)	6 (11.8%)	1 (1.9%)	<p>Versus control:</p> <ol style="list-style-type: none"> 1. $p = 0.007$ 3. $p = 0.002$ 4. $p < 0.0001$ <p>Versus Mg + sotalol:</p> <ol style="list-style-type: none"> 1. $p = 0.01$ 3. $p = 0.04$ <p>Mg versus sotalol: $p > 0.05$ (ns)^a</p>
Mean \pm SD time (days postoperation) to onset of AF	2.4 \pm 1.4	2.3 \pm 0.8	2.0 \pm 1.6	2.0 \pm 0.0	$p > 0.05$ (ns) ^a
^a Inferred by reviewers (not directly reported).					
					<i>continued</i>

Other outcomes	1. Mg (n = 54)	2. Control (n = 50)	3. Sotalol (n = 51)	4. Mg + sotalol (n = 52)	Difference
Mean \pm SD length of stay (days postoperation)	5.7 \pm 0.9	5.9 \pm 1.7	5.6 \pm 1.4	5.2 \pm 1.3	Versus control: 4. $p = 0.02$ All other pairwise comparisons: $p > 0.05$ (ns) ^b
^b Inferred by reviewers (not directly reported).					
<p>Methodological comments</p> <p><i>Allocation to treatment groups:</i> Stated that patients were randomised to the study groups according to a computer-generated random code.</p> <p><i>Blinding:</i> NR.</p> <p><i>Comparability of treatment groups:</i> Authors stated that preoperative and intraoperative measurements did not differ significantly among the four groups of patients.</p> <p><i>Method of data analysis:</i></p> <ul style="list-style-type: none"> • Comparisons of continuous or discrete variables between the four groups were performed using an unpaired Student's t-test or a χ^2 test, respectively. • Other analyses were also carried out for comparisons outside the scope of this assessment, including stepwise logistic regression to investigate predictors of AF (data not extracted). <p><i>Sample size/power calculation:</i> NR.</p> <p><i>Attrition/dropout:</i> Stated that all analyses were performed according to the ITT principle (ITT not defined). Patients excluded owing to AF, MI or related arrhythmia were described, but it is unclear whether these patients were also excluded from the analyses (if so, also unclear how the missing data were treated). Other patients were excluded after CABG if haemodynamic criteria were not met, but the numbers and identities of these patients were not reported. Accordingly, there is considerable doubt as to the final composition of the analysed populations.</p> <p>General comments</p> <p><i>Generalisability:</i> No details of ethnicity or social background; assumed single-location study in Italy. Following the oral presentation of this paper at the 38th Annual Meeting of the Society of Thoracic Surgeons, limited generalisability of the findings due to the highly-selected patient population was noted.</p> <p><i>Outcome measures:</i> Appropriate.</p> <p><i>Intercentre variability:</i> NA</p> <p><i>Conflict of interests:</i> None reported.</p>					
SVA, supraventricular arrhythmia.					

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Hazelrigg et al.²²</p> <p>Year: 2004</p> <p>Country: USA</p> <p>Study design: RCT</p> <p>Number of centres: one (comprising two hospitals)</p> <p>Funding: Part-funded by Southern Illinois University School of Medicine, Central Research Committee grant</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Control <p>Intervention details:</p> <ol style="list-style-type: none"> 80 mg kg⁻¹ (ideal body weight) of magnesium sulphate administered in 100 ml D₅W over a 30-minute period before CPB. Thereafter, 8 mg kg⁻¹ (ideal body weight) per hour magnesium sulphate i.v. infusion in 100 ml D₅W continued for 48 h 100 ml of D₅W pre-CPB and 100 ml of D₅W for 48 h post-CPB according to same schedule as intervention group 	<p>Number of participants:</p> <p>Intervention: 105</p> <p>Control: 97</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry: Patients undergoing elective isolated CABG; age ≥ 18 years; lack of chronic arrhythmia history; ejection fraction > 25%; normal renal function (creatinine < 1.5 mg dl⁻¹)</p> <p>Exclusion criteria for study entry: Elevated liver functions [sic]; hypotension (systolic BP < 90 mmHg); postoperative creatinine > 1.8 (units not stated); patients who required cardiac assist</p> <p>Baseline characteristics of subjects</p> <p>Mean ± SD age (years):</p> <p>Intervention: 62.1 ± 9.5</p> <p>Control: 63.7 ± 11.1</p> <p>Difference: <i>p</i> ≥ 0.05 (ns)</p> <p>Gender (M/F):</p> <p>Intervention: 78/27 (26% F)</p> <p>Control: 66/31 (32% F)</p> <p>Difference: <i>p</i> ≥ 0.05 (ns)</p> <p>Differences for the mean ± SD duration of cross-clamping (minutes) were significant (<i>p</i> = 0.04):</p> <p>Intervention: 61.11 ± 21.1</p> <p>Control: 55 ± 21.89</p> <p>Differences between groups were not statistically significant for 21 other preoperative and perioperative variables listed</p>	<p>AF outcome: Incidence of AF (primary outcome)</p> <p>Method of assessing AF outcome: NR</p> <p>Other extracted outcomes: Length of stay in ICU; length of stay in hospital</p> <p>Other outcomes (data not extracted): Incidence of atrial flutter; incidence of frequent/multifocal premature ventricular contraction; ventricular tachycardia; ventricular fibrillation; mean arterial pressure, BP and other haemodynamic variables; serum, urine and tissue electrolyte concentrations</p> <p>Adverse symptoms: Ventricular tachycardia occurred in significantly more control patients (details below); urine magnesium and calcium excretion were significantly higher in the intervention group (details below); death occurred in one and two patients in intervention and control groups, respectively (<i>p</i> ≥ 0.05, ns)</p> <p>Length of follow-up: 5 days after operation</p> <p>Recruitment dates: Over a 5-year period (dates not reported)</p>

continued

Results			
AF outcomes	Intervention (n = 105)	Control (n = 97)	Difference
No. of patients with AF			
Operative	1	2	$p \geq 0.05$ (ns)
0 days postoperative	1	1	$p \geq 0.05$ (ns)
1 day postoperative	4	10	$p < 0.05$
2 days postoperative	12	15	$p \geq 0.05$ (ns)
3 days postoperative	13	10	$p \geq 0.05$ (ns)
4 days postoperative	1	3	$p \geq 0.05$ (ns)
5 days postoperative	0	0	$p \geq 0.05$ (ns)
Total	32	41	$p \geq 0.05$ (ns)
Specific p -values were not provided.			
Other outcomes	Intervention (n = 105)	Control (n = 97)	Difference
Mean \pm SD length of ICU stay (days)	1.33 \pm 0.72	1.36 \pm 1.4	$p \geq 0.05$ (ns)
Mean \pm SD length of hospital stay (days)	6.65 \pm 3.27	6.96 \pm 4.98	$p \geq 0.05$ (ns)
Specific p -values were not provided.			
Methodological comments			
<i>Allocation to treatment groups:</i> Pharmacy staff randomised patients to the interventions using a randomisation table created by a statistician (no other details).			
<i>Blinding:</i> The patients and those caring for them were blinded to the randomisation arm [sic].			
<i>Comparability of treatment groups:</i> No notable differences, except that mean cross-clamp time was significantly (~ 5 minutes) longer in the magnesium sulphate group ($p = 0.04$).			
<i>Method of data analysis:</i>			
<ul style="list-style-type: none"> • A 2×2 χ^2 test was used to compare incidence of AF between the groups. • For other comparisons between the groups, χ^2 tests, a Fisher's exact test for categorised variables, and an independent-groups t-test for continuous variables were used. However, the results presented do not indicate in which cases each of these tests was applied, nor whether the data satisfied assumptions of normality. • Overall, statistical tests were not supported by any hypotheses. 			
<i>Sample size/power calculation:</i> Stated that, based on the incidence of arrhythmias after CABG (data source not reported), 100 patients per group would assure detecting an effect of supplemental magnesium that reduced arrhythmias by half, with power 0.80 and $\alpha = 0.05$ (two-sided).			
<i>Attrition/dropout:</i> NR.			
General comments			
<i>Generalisability:</i> No details of ethnicity or social background; single-location study in USA.			
<i>Outcome measures:</i> Appropriate.			
<i>Intercentre variability:</i> NA.			
<i>Conflict of interests:</i> None reported.			
BP, blood pressure.			

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Kaplan et al.²³</p> <p>Year: 2003</p> <p>Country: Turkey</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: NR</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Placebo control <p>Intervention details:</p> <ol style="list-style-type: none"> 24.34 mEq (3 g) magnesium sulphate in 100 ml saline solution administered over 2 h (50 ml h⁻¹) preoperatively (-12 h), perioperatively and postoperatively (0, 1, 2, 3 days) 100 ml saline according to same administration schedule as group 1 	<p>Number of participants:</p> <p>Intervention: 100</p> <p>Control: 100</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry:</p> <p>Patients undergoing elective and initial CABG (no other details)</p> <p>Exclusion criteria for study entry:</p> <p>History of AF (or paroxysmal AF with subsequent sinus rhythm); preoperative heart rate <50 bpm; concomitant valve surgery; redo coronary artery surgery; BP < 100 mmHg; history of renal failure (serum creatinine >2.0 mg dl⁻¹); severe respiratory function disorder</p> <p>Baseline characteristics of subjects:</p> <p>Mean ± SD age (range) (years):</p> <p>Intervention: 57.63 ± 9.68 (41–76)</p> <p>Control: 59.56 ± 9.29 (44–80)</p> <p>Gender (M/F):</p> <p>Intervention: 76/24</p> <p>Control: 74/26</p> <p>Difference: <i>p</i> = 0.183 (ns)</p> <p>Differences between groups were not statistically significant (<i>p</i> > 0.05) for previous MI, hypertension or diabetes; preoperative frequency of β-blockers, calcium channel blockers, ACE inhibitors or digoxin; LVEF; duration of CPB or aortic cross-clamping; number of blood transfusion units</p> <p>Groups appeared similar (statistical significance not reported) for preoperative magnesium sulphate concentration^a and number of bypass grafts; preoperative risk (EuroSCORE %): intervention 1.75, control 2.15</p>	<p>AF outcomes:</p> <p>Postoperative development time of AF (primary outcome); day of onset of AF; incidence of AF</p> <p>Other extracted outcomes:</p> <p>Duration in ICU; duration in hospital</p> <p>Other outcomes (data not extracted):</p> <p>Perioperative and postoperative arrhythmias (extrasystole and SVT attacks); ventricular rate (LVEF) at onset of AF; duration of intubation; plasma magnesium sulphate concentration</p> <p>Adverse symptoms:</p> <p>One death in each group. Stated that magnesium sulphate did not cause any cases of severe bradycardia or hypotension</p> <p>Length of follow-up:</p> <p>2 days after operation</p> <p>Recruitment dates:</p> <p>NR</p>
^a Units not stated.			
Results			
AF outcomes	Intervention	Control	Difference
No. of patients with AF	15 (<i>n</i> = 100)	16 (<i>n</i> = 100)	<i>p</i> = 0.845
Mean ±SD time to onset of AF (h) (primary outcome)	37.87 ± 12.76 (<i>n</i> = 15)	45.26 ± 15.27 (<i>n</i> = 16)	<i>p</i> = 0.140
% of patients with AF whose age was >60 years	62.5 (<i>n</i> = 15)	77.7 (<i>n</i> = 15)	<i>p</i> = 0.174 ^b
No. of patients with AF onset ~0 days/+1 day/+2 days after operation	1/2/12 (<i>n</i> = 15)	2/2/12 (<i>n</i> = 15)	NR
<i>continued</i>			

Other outcomes	Intervention (n = 100)	Control (n = 100)	Difference
Mean \pm SD length of ICU stay (h)	22.40 \pm 4.79	23 \pm 4.46	$p = 0.650$
Mean \pm SD length of hospital stay (days) ^c	5.16 \pm 1.18 ^c	5.67 \pm 1.31 ^c	$p = 0.004^c$
Mean \pm SD length of hospital stay (days) in patients with AF	6.0 \pm 1.2 (n = 15 ^d)	6.31 \pm 0.87 (n = 16 ^d)	$p = 0.410$

^b This p -value is assumed by the reviewers to refer to the difference between interventions (meaning not stated in the primary article).

^c It is unclear whether these data refer to the total population (with AF + without AF) or the cohort without AF (accordingly, the appropriate n -value is also unclear).

^d n -Values were not reported here, but are assumed by the reviewers to include all patients with AF.

Methodological comments

Allocation to treatment groups: stated random; no further details.

Blinding: Stated that only 100 ml of saline was administered to the control group for the placebo effect (no other details reported).

Comparability of treatment groups: No notable significant differences.

Method of data analysis:

- Independent samples t -test for comparing mean durations of hospitalisation.
- Fisher's exact test for comparing ratios of atrial extrasystole and SVT in treatment and control groups (interpretation ambiguous).
- Mann-Whitney U -test for comparing numbers of patients with/without AF (unequal n precluded t -test).
- Pearson χ^2 test for effects of interventions on time to onset of AF and for age differences in incidence of AF.
- Binary logistic regression used to evaluate the importance of baseline risk factors on incidence of AF (data not extracted; all p -values ns).
- Multivariate linear regression to investigate correlation between heart rate and serum magnesium sulphate concentration (data not extracted; $p = 0.158$).

Sample size/power calculation: NR.

Attrition/dropout: NR.

General comments

Generalisability: No details of ethnicity or social background; single-location study in USA.

Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: None reported.

ACE, angiotensin-converting enzyme.

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Karmy-Jones <i>et al.</i>²⁴</p> <p>Year: 1995</p> <p>Country: Canada</p> <p>Study design: RCT</p> <p>Number of centres: One (assumed by the reviewers)</p> <p>Funding: NR</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Placebo control <p>Intervention details:</p> <ol style="list-style-type: none"> Six doses of 2.4 g (19.2 mEq) magnesium sulphate i.v. in the first 24 h after the cardiac operation. Total dosage: 14.4 g (115 mEq). First dose: magnesium sulphate in 50 ml of D₅W over 20 minutes at the termination of CPB; further 5 doses: the same dosage every 4 h 50 ml D₅W at the same time-points as group 1 	<p>Number of participants:</p> <p>Intervention: 46</p> <p>Control: 54</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry:</p> <p>Patients who were undergoing elective coronary artery bypass, valve replacement/repair, or a combination of these; enrolled after informed consent was obtained. 39 patients (84.8%) in intervention group and 47 patients in control group (87.0%) were undergoing CABG. One patient in intervention group and two patients in control were undergoing combined procedure</p> <p>Exclusion criteria for study entry:</p> <p>Abnormal renal function, reoperation, emergency operation, evidence of ongoing ischaemia (angina, ST changes), use of medications to control dysrhythmias, and inability to obtain informed consent</p> <p>Baseline characteristics of subjects:</p> <p>Mean \pm SD age (years):</p> <p>Intervention: 64.5 \pm 7.9</p> <p>Control: 60.2 \pm 11.9</p> <p>Gender (M/F) (n):</p> <p>Intervention: 28/18</p> <p>Control: 38/16</p> <p>Authors noted that the magnesium group was significantly older ($p = 0.036$ reported) and that there were no significant differences in other baseline characteristics between groups (mean \pm SD but not p-values reported): preoperative calcium channel blocker, β-blocker, diuretics, digoxin; COPD; diabetes; prior MI; history of dysrhythmia: supraventricular or ventricular, palpitations; NYHA class; Canadian Cardiovascular Society class; ejection fraction; number of vein grafts; cross-clamp time; CPB time</p>	<p>AF outcome:</p> <p>Not reported directly; however, incidence of AF inferred from the incidence of SVAs (of which 92% overall were AF)</p> <p>Other extracted outcomes:</p> <p>Length of hospital stay (days); length of stay at ICU (h)</p> <p>Other outcomes (<i>data not extracted</i>):</p> <p>Ventilator hours (mean \pm SD); number of patients with ventilator hours >24 h; cardiac performance after CABG (cardiac index, left ventricular stroke work index, stroke volume index); patients with inotropes, intra-aortic balloon pump, pacing and/or ischaemia; ventricular and supraventricular arrhythmias (and Lown grades); MB isoenzyme of creatine kinase and serum magnesium concentrations</p> <p>Adverse symptoms:</p> <p>There were two deaths in the placebo group, one caused by multiple organ failure and the other caused by ventricular fibrillation. Over the whole hospital stay, six types of ventricular tachycardia and two classes of ischaemia were each significantly more frequent in placebo than magnesium group (various p-values). Postoperative ST segment elevation was significantly less frequent in the magnesium group. Postoperative creatinine kinase concentration was significantly lower in the magnesium group</p>

continued

Reference and design	Intervention	Participants	Outcome measures	
			<p><i>Length of follow-up:</i> Entire hospital stay (see results on other outcomes). Continuous monitoring at bedside ≥ 24 h, then subsequently SVT identified on clinical signs^a</p> <p><i>Recruitment dates:</i> NR</p>	
^a ECGs were obtained postoperatively and on the first two mornings after operation. After this period, when there was a question of dysrhythmia either on clinical grounds or by telemetry, a 12-lead ECG was obtained.				
Results				
AF outcomes		Intervention (n = 46)	Control (n = 54)	Difference
Incidence of SVT episodes, n (%) ^b		12 (26%)	13 (24%)	p = 0.83 (ns)
Mean \pm SD SVT ^b episodes per patient		0.46 \pm 0.98	0.51 \pm 1.1	p = 0.80 (ns)
^b 92% of all SVT episodes were AF (this implies that 23 out of the 25 patients with SVT had AF).				
Other outcomes		Intervention (n = 46)	Control (n = 54)	Difference
Mean \pm SD length of hospital stay (days)		6.0 \pm 1.9	8.3 \pm 11.5	p = 0.35 (ns)
Mean \pm SD length of stay at ICU (h)		25.6 \pm 16.6	38.0 \pm 73.4	p = 0.68 (ns)
Methodological comments				
<i>Allocation to treatment groups:</i> It was only mentioned that patients were randomised into placebo and magnesium groups.				
<i>Blinding:</i> Patients received magnesium sulphate in solution or the solution without magnesium at the same time points (no further details; blinding of patients was not stated explicitly). Individuals recording data and designating dysrhythmias were not involved in patient management and were blinded to serum magnesium concentrations. Physicians directing therapy were also blinded to all but the initial postoperative serum magnesium results.				
<i>Comparability of treatment groups:</i> The magnesium group was significantly older (p = 0.036 reported); stated by authors that there were no significant differences between groups in other baseline characteristics (p-values not reported).				
<i>Method of data analysis:</i>				
Discrete variables were analysed using χ^2 test. Continuous variables were studied using two-tailed t-test. Kruskal-Wallis [sic] one-way analysis of variance was used to perform rank analysis.				
<i>Sample size/power calculation:</i> Assuming a 30% incidence of ventricular tachyarrhythmia and hoping to demonstrate a 50% reduction, the authors calculated that a sample size of 154 would be required, using an α error of 0.05 and a β error of 0.1. They mentioned that a review of the literature had suggested that significant differences could be identified with a sample of 100 patients.				
<i>Attrition/dropout:</i> NR.				
General comments				
<i>Generalisability:</i> No details of ethnicity or social background; single location in Canada.				
<i>Outcome measures:</i> Authors noted only the numbers of patients with SVT episodes. The incidence of AF in the intervention and the control groups was not reported, but is inferred here from the SVT data (of which 92% overall comprised AF). Data on the length of hospital stay and stay at ICU appropriate.				
<i>Intercentre variability:</i> NA.				
<i>Conflict of interests:</i> Not reported.				

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Nurözler et al.²⁵</p> <p>Year: 1996</p> <p>Country: Turkey</p> <p>Study design: RCT</p> <p>Number of centres: One (assumed by reviewers)</p> <p>Funding: NR</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Placebo <p>Intervention details:</p> <ol style="list-style-type: none"> 16 mmol l⁻¹ magnesium in the cardioplegia solution. Then 100 mEq magnesium sulphate i.v. infusion over the 1st postoperative day, followed by 25 mEq day⁻¹ from the 2nd to 5th days as i.v. continuous infusion (in 1000 ml of D₅W over 24 h) Received a cardioplegia solution without any magnesium. No other details of the placebo were reported 	<p>Number of participants:</p> <p>Intervention: 25</p> <p>Control: 25</p> <p>Sample attrition/dropout: NR</p> <p>Inclusion criteria for study entry: Consenting patients undergoing CABG with good left ventricular function and no history of arrhythmias</p> <p>Exclusion criteria for study entry: Other surgical procedures in addition to CABG, documented history of preoperative supraventricular or ventricular arrhythmias, second or third degree heart block, ejection fraction <40%, postoperative low cardiac output requiring inotropic drugs or intra-aortic balloon support, COPD (FEV₁ < 60% predicted) and/or serum creatinine concentration > 2.5 mg dl⁻¹</p> <p>Baseline characteristics of subjects:</p> <p>Mean ± SD age (years):</p> <p>Intervention: 56.3 ± 1.3^a</p> <p>Control: 53.6 ± 2.0^a</p> <p>Difference: <i>p</i> = 0.23 (ns)</p> <p>Gender (M/F):</p> <p>Intervention: 23/2</p> <p>Control: 23/2</p> <p>Differences between groups were not significant (<i>p</i> > 0.05) for: left ventricular function, ejection fraction, left ventricular end-diastolic pressure, number of diseased vessels, number of bypass grafts, CPB time, ischaemic time or haemoglobin on day 3</p> <p>Group values appeared similar (no statistics presented) for: previous infarction, diabetes, hypertension, smoking and postoperative pericarditis</p>	<p>AF outcome:</p> <p>Incidence of AF</p> <p>Other outcomes (data not extracted):</p> <p>Serum magnesium and potassium concentrations</p> <p>Adverse symptoms: Authors stated that no adverse effect of magnesium replacement was noticed</p> <p>Length of follow-up: 5 days after CABG</p> <p>Recruitment dates: 2-month period (dates not specified)</p>
^a Variance measure not stated.			
Results			
AF outcomes	Intervention (<i>n</i> = 25)	Control (<i>n</i> = 25)	Difference
Incidence of AF, <i>n</i> (%) ^b	1 (4%)	5 (20%)	<i>p</i> = 0.02
^b Excludes transient AF that converted to sinus rhythm in < 1 minute or AF that was not confirmed by ECG.			
Other outcomes	Outside scope of assessment: data not extracted		
<i>continued</i>			

Methodological comments

Allocation to treatment groups: Stated that patients were randomised, but no details of the procedure were given.

Blinding: Stated that the design was double-blind, but no details were given.

Comparability of treatment groups: Stated that there were no differences in clinical, angiographic and surgical characteristics between the two groups; where *p*-values were provided (for six continuous variables) they were not significant ($p > 0.05$).

Method of data analysis:

- χ^2 test was used to analyse differences in proportions.
- Multiple logistic regression and analysis of variance were also used, but address questions outside the scope of this assessment (data not extracted).
- Statistical tests were not supported by hypotheses; normality of data was not reported.

Sample size/power calculation: NR.

Attrition/dropout: NR.

General comments

Generalisability: No details of ethnicity or social background; single-location study in Turkey.

Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: None reported.

FEV₁, forced expiratory volume in 1 second.

Reference and design	Intervention	Participants	Outcome measures
<p><i>Authors:</i> Parikka <i>et al.</i>²⁶</p> <p><i>Year:</i> 1993</p> <p><i>Country:</i> Finland</p> <p><i>Study design:</i> RCT</p> <p><i>Number of centres:</i> One (assumed by reviewers)</p> <p><i>Funding:</i> Finnish Heart Foundation, Helsinki</p>	<p><i>Interventions:</i></p> <ol style="list-style-type: none"> 1. Intravenous magnesium sulphate 2. Placebo control <p><i>Intervention details:</i></p> <ol style="list-style-type: none"> 1. Two doses of magnesium sulphate: <ul style="list-style-type: none"> • 40 mmol l⁻¹ of 5% glucose-in-water solution during the first 24 h after the infusion (the infusion was started within 2 h of the operation) • 30 mmol 500 ml⁻¹ of the solution during the next 24 h 2. The same as in the intervention: 5% glucose-water solution without magnesium 	<p><i>Number of participants:</i></p> <p>Intervention: 69 Control: 71</p> <p><i>Sample attrition/dropout:</i></p> <p>In the final postoperative analysis the number of patients in the control group was 70. One patient died on the first postoperative day because of graft occlusion and consequent perioperative infarction</p> <p><i>Inclusion criteria for study entry:</i> Consecutive patients, who had had their first CABG</p> <p><i>Exclusion criteria for study entry:</i> Patients with chronic AF, concomitant valve replacement and antiarrhythmic medication</p> <p><i>Baseline characteristics of subjects:</i></p> <p><i>Mean age (mean + SD) (years):</i></p> <p>Intervention: 57 ± 8 Control: 54 ± 8</p> <p>Authors noted that the magnesium group was older, $p = 0.032$ (ns)</p> <p><i>Gender (proportion of males, %)</i></p> <p>Intervention: 84% Control: 82% Difference: ns</p> <p>Authors stated that prior AF episodes were more common in the magnesium group (9% vs 1%; $p = 0.061$; ns), and (probably) owing to</p>	<p><i>AF outcomes:</i></p> <p><i>Clinically determined:</i></p> <ul style="list-style-type: none"> • incidence of AF; • time of first AF (days) • incidence of AF relapsed <p><i>Holter ECG^a determined 2 days postsurgery:</i></p> <ul style="list-style-type: none"> • incidence of AF • AF episodes per patient • duration of AF (h) • rate of AF (bpm) <p><i>Other outcomes (data not extracted):</i></p> <p>Ventricular ectopic beat; ventricular tachycardia; supraventricular ectopic beat; SVT other than AF or atrial flutter; serum total calcium, magnesium, sodium and potassium concentrations; haematology (creatinine, creatine kinase isoenzyme and C-reactive protein measurements); perioperative MI; sinus rate (bpm)</p>

continued

Reference and design	Intervention	Participants	Outcome measures	
		<p>this, digoxin therapy was more frequent in this group (12% vs 3%; $p = 0.054$; ns). Groups were not statistically different considering other baseline characteristics, e.g. NYHA class, left ventricular and diastolic BP, ejection fraction, incidence of three-vessel disease, prior β-blockers, diuretics (p-values not provided) or left atrial transverse diameter in echocardiography ($p = 0.071$; ns)</p>	<p><i>Adverse symptoms:</i> One patient had postoperative cerebral thrombosis and died of massive pulmonary embolism 4 weeks after CABG. It is not stated in which group the death occurred</p> <p><i>Length of follow-up:</i> 1. Cardiac rhythm was monitored continuously the first 2 postoperative days. Thereafter, for up to 10 days every symptomatic palpitation was recorded by 12-lead ECG (monitoring in total 12 days after operation) 2. A 48-h two-channel Holter recording was started on the second postoperative day</p> <p><i>Recruitment dates:</i> NR</p>	
<p>^a Technically acceptable recording was achieved in 108 patients.</p>				
Results				
Clinical AF outcomes		Intervention (n = 69)	Control (n = 70)	Difference
Incidence of AF, n (%)		20 (29%)	18 (26%)	ns
Mean \pm SD time of first AF (days)		3.5 \pm 1.9	3.8 \pm 2.7	ns
Incidence of AF relapsed, n (% of AF patients)		9 (45% of AF patients)	11 (61% of AF patients)	ns
Holter AF outcomes		Intervention (n = 52)	Control (n = 56)	Difference
Incidence of AF, n (%)		7 (14%)	7 (13%)	ns
Mean \pm SD episodes of AF		1.4 \pm 0.8	1.2 \pm 0.4	ns
Mean \pm SD duration of AF (h)		5.1 \pm 4.3	5.0 \pm 6.3	ns
Mean \pm SD rate of AF (bpm)		115 \pm 26	123 \pm 15	NS ($p = 0.538$)
Methodological comments				
<p><i>Allocation to treatment groups:</i> It was only mentioned that patients were randomised.</p> <p><i>Blinding:</i> Patients were blinded (received magnesium sulphate in solution or the same volume of solution without magnesium). Authors did not provide any information concerning the blinding of doctors or assessors.</p> <p><i>Comparability of treatment groups:</i> In the magnesium group patients were older ($p = 0.032$), prior AF episodes were more common (9% vs 1%; $p = 0.061$, ns) and digoxin therapy was more frequent (12% vs 3%; $p = 0.054$, ns). The groups were not statistically different in terms of other prognostic factors (either p-values were reported or ns stated if the difference was not significant).</p>				
			<i>continued</i>	

Method of data analysis:

- Continuous variables were analysed by two-tailed Student's *t*-test and for non-normal distributions by the Mann–Whitney rank sum test.
- Predictors of AF were investigated (data not extracted).
- However, it was not stated whether data sets conformed to normality.

Sample size/power calculation: NR.

Attrition/dropout: One patient in the control group died on the first postoperative day because of graft occlusion and consequent perioperative infarction.

General comments

Generalisability: No details of ethnicity or social background; single location in Finland.

Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: Not reported.

Reference and design	Intervention	Participants	Outcome measures
<p><i>Authors:</i> Toraman et al.²⁷</p> <p><i>Year:</i> 2001</p> <p><i>Country:</i> Turkey</p> <p><i>Study design:</i> RCT</p> <p><i>Number of centres:</i> One</p> <p><i>Funding:</i> NR</p>	<p><i>Interventions:</i></p> <ol style="list-style-type: none"> 1. Intravenous magnesium sulphate 2. Control (not reported as placebo) <p><i>Intervention details:</i></p> <ol style="list-style-type: none"> 1. 6 mmol magnesium sulphate in 100 ml of 0.9% sodium chloride (25 ml h⁻¹) 1 day before CBP, just after CBP surgery and then once daily for 4 days after CBP 2. 100 ml of 0.9% sodium chloride (25 ml h⁻¹) at the same time-points as in the magnesium sulphate group 	<p><i>Number of participants:</i></p> <p>Intervention: 100 Control: 100</p> <p><i>Sample attrition/dropout:</i> NR</p> <p><i>Inclusion criteria for study entry:</i> Patients scheduled for elective, first time, isolated CABG (no other details)</p> <p><i>Exclusion criteria for study entry:</i> Patients with AF, a past history of AF, heart valve disease, diabetes, chronic renal disease, thyroid disorders and/or COPD; patients receiving antiarrhythmic drugs, digoxin and/or β-blocking agents</p> <p><i>Baseline characteristics of subjects:</i></p> <p><i>Mean \pm SD age (years):</i> Intervention: 62 \pm 6.7 Control: 61.4 \pm 8.7 Difference: $p = 0.56$ (ns)</p> <p><i>Gender (M/F) (n):</i> Intervention: 78/22 Control: 83/17 Difference: $p = 0.48$ (ns)</p> <p>There were no significant differences between the groups in 14 other preoperative and 12 perioperative variables reported (p-values were provided; all $p > 0.05$)</p>	<p><i>AF outcome:</i></p> <p>Incidence of AF after CABG</p> <p><i>Other outcomes extracted:</i></p> <p>Duration in ICU, ICU readmission, length of postoperative hospital stay</p> <p><i>Other outcomes (data not extracted)</i></p> <p>Stroke incidence, duration of extubation, postoperative cardiac output, total chest drainage, serum magnesium concentration</p> <p><i>Adverse symptoms:</i> Stated that all patients in the magnesium group received their scheduled doses of magnesium without any adverse effects such as bradycardia or hypotension (no other details)</p> <p><i>Length of follow-up:</i> 4 days after CABG surgery</p> <p><i>Recruitment dates:</i> February 1999 to March 2000</p>

continued

Results			
AF outcome	Intervention (n = 100)	Control (n = 100)	Difference
Frequency of AF occurrence, n (%)	2 (2%)	21 (21%)	$p < 0.001$
Other outcomes	Intervention (n = 100)	Control (n = 100)	Difference
Mean \pm SD length of stay in ICU (h)	21.6 \pm 5.6	22.6 \pm 6.9	$p = 0.58$
ICU readmission, n	1	3	$p = 0.62$
Mean \pm SD length of postoperative hospital stay (days)	5.4 \pm 0.9	5.8 \pm 4.1	$p = 0.36$
Methodological comments			
<i>Allocation to treatment groups:</i> Stated that this was randomised, but no details provided.			
<i>Blinding:</i> Stated that all ECGs were analysed by a cardiologist who was blinded to the study (no further details provided).			
<i>Comparability of treatment groups:</i> No notable differences.			
<i>Method of data analysis:</i>			
<ul style="list-style-type: none"> Stated that demographic and clinical variables as well as the incidence of AF were analysed using the χ^2 test or Fisher's exact test for categorical variables and <i>t</i>-tests for continuous variables. Kruskal–Wallis one-way analysis of variance was used to perform rank analysis. However, the results presented do not indicate in which cases each of these tests was applied, nor whether the data satisfied assumptions of normality. Overall, statistical tests were not supported by any hypotheses. 			
<i>Sample size/power calculation:</i> NR.			
<i>Attrition/dropout:</i> NR.			
General comments			
<i>Generalisability:</i> No details of ethnicity; limited information on social background; single-location study in Turkey.			
<i>Outcome measures:</i> Appropriate, but note that 12-lead Holter ECG was only used for reactive, not continuous, monitoring (authors discuss whether this might have missed some cases of AF).			
<i>Intercentre variability:</i> NA.			
<i>Conflict of interests:</i> None reported.			

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Treggiari-Venzi <i>et al.</i>²⁸</p> <p>Year: 2000</p> <p>Country: Switzerland</p> <p>Study design: RCT</p> <p>Number of centres: One</p> <p>Funding: Swiss Society of Cardiology</p>	<p>Interventions:</p> <ol style="list-style-type: none"> Intravenous magnesium sulphate Placebo control Amiodarone (data not extracted) <p>Intervention details:</p> <ol style="list-style-type: none"> 16 mmol (32 mEq) (4 g) magnesium sulphate per 24 h (Bichsel, Interlaken, Switzerland), over 72 h, starting within 1 h of arrival in ICU 0.9% sodium chloride, over 72 h, starting within 1 h of arrival in ICU 	<p>Number of participants:</p> <p>Intervention: 47 (49 randomised)</p> <p>Control: 51 (53 randomised)</p> <p>Sample attrition/dropout:</p> <p>Magnesium sulphate and placebo each had two dropouts (magnesium sulphate: one due to cardiac arrest and one incomplete ECG data; placebo: one due to pacemaker dependence and one incomplete data)</p> <p>Inclusion criteria for study entry:</p> <p>Patients scheduled for elective CABG (no other details)</p> <p>Exclusion criteria for study entry:</p> <p>Refusal of consent; chronic AF, second or third degree atrioventricular block, pacemaker dependence, amiodarone treatment < 1 year before operation; thyroid disease, other associated heart surgery, valvular disease, chronic renal failure (creatinine clearance rate < 30 ml minute⁻¹) and liver dysfunction (prothrombin time < 50% and/or bilirubin > 35 µmol l⁻¹ and/or presence of ascites)</p> <p>Baseline characteristics of subjects:</p> <p>Mean age (range) (years):</p> <p>Intervention: 65 (46–81)</p> <p>Control: 65 (37–88)</p> <p>Gender (M/F) (n):</p> <p>Intervention: 42/5</p> <p>Control: 43/8</p> <p>Other baseline values (e.g. previous MI, history of SVA) appear similar between groups. No statistical tests reported</p>	<p>AF outcomes (1st = primary outcome):</p> <p>Prevention of AF first;^a time to onset of AF; frequency of AF > 30s; heart rate during AF</p> <p>Other extracted outcomes:</p> <p>Duration in ICU</p> <p>Other outcomes (data not extracted):</p> <p>Duration of intubation; period of required catecholamine infusion; plasma magnesium sulphate concentration</p> <p>Adverse symptoms:</p> <p>Stated that incidences of MI, cardiac arrest, need for surgical haemostasis and prolonged tracheal intubation (> 72 h) were similar in the two groups</p> <p>Length of follow-up:</p> <p>3 days (72 h) after start of intervention</p> <p>Recruitment dates:</p> <p>NR</p>
<p>^a Primary end-point (outcome) definition does not precisely correspond with the reported data.</p>			
Results			
AF outcomes	Intervention (n = 47)	Control (n = 51)	Difference
Frequency of AF occurrence > 30s n (%)	11 (23%)	14 (27%)	p = 0.82
Time to onset of AF (h) after start of intervention from log-rank test	Given in chart, but not clearly extractable	Given in chart, but not clearly extractable	p = 0.5
Mean ± SD delay in AF onset (h) after start of intervention ^b	45 ± 14	42 ± 12	p = 0.37
No. of AF episodes (maximum) per patient ^c	17 (3)	14 (1)	p = 0.88
Mean ± SD heart rate (bpm) during AF	146 ± 9	153 ± 18	p = 0.26
<p>^b It is not clear which statistical analysis was used here; this could have been an unpaired t-test or a Mann–Whitney test depending upon normality of the variable (not stated).</p> <p>^c Assumed by reviewers (not stated) to be the number per patient.</p>			
Other outcomes	Intervention (n = 47)	Control (n = 51)	Difference
Median (range) length of stay in ICU (days)	3 (2–21)	3 (2–7)	NR
<i>continued</i>			

Adverse events: One patient in the placebo group died 5 days after surgery

Methodological comments

Allocation to treatment groups: Random assignment of patients to interventions was based on colour-coded spheres drawn from an opaque container (no other details). Additional patients were added to compensate for 'technical' dropouts, but no indication is given of whether this procedure was also randomised. Precise timing of the start of interventions not reported; assumed to be immediately post-CABG operation in ICU.

Blinding: Study drugs were prepared in an opaque syringe with opaque tubing by an independent observer; study described as 'double blind' (no other details given).

Comparability of treatment groups: No notable differences; authors report that the groups were similar in baseline characteristics and surgical procedure.

Method of data analysis:

- For continuous data, either an unpaired *t*-test (normally distributed data) or a Mann–Whitney test (non-normal data) was used. However, it is not stated which data sets conformed to normality, hence it is unclear which tests were used for each comparison.
- For categorical variables a χ^2 test was used. However, no χ^2 results were presented, only *p*-values were given, so it is unclear where this test was applied.
- Kaplan–Meier analysis was used to analyse the delay in onset and the duration of AF (no other details provided).
- Overall, statistical tests were not supported by any hypotheses.

Sample size/power calculation: Reported for detecting a 50% reduction of frequency of AF using Holter ECG (80% power with $\alpha = 0.5$), but the power calculation is unclear (no variance estimate provided).

Attrition/dropout: Reported, but not stated whether accounted for in analysis. Probably not analysed according to ITT (the authors remarked for the magnesium group that "the final number of patients in the interim analysis was 47").

General comments

Generalisability: No details of ethnicity or social background; single-location study in Switzerland.

Outcome measures: Appropriate.

Inter-centre variability: NA.

Conflict of interests: None reported.

Reference and design	Intervention	Participants	Outcome measures
<p>Authors: Yilmaz <i>et al.</i>¹⁴</p> <p>Year: 2000</p> <p>Country: Turkey</p> <p>Study design: RCT</p> <p>Number of centres: One (assumed by reviewers)</p> <p>Funding: NR</p>	<p>Interventions:</p> <ol style="list-style-type: none"> 1. Magnesium sulphate bolus^a 2. Control <p>Intervention details:</p> <ol style="list-style-type: none"> 1. 17 mmol l⁻¹ magnesium sulphate in cardioplegia infusion and 0.4 mmol kg⁻¹ bolus of magnesium sulphate administered at the start of CPB 2. 17 mmol l⁻¹ magnesium sulphate in cardioplegia infusion only 	<p>Number of participants:</p> <p>Intervention: 15</p> <p>Control: 15</p> <p>Sample attrition/dropout:</p> <p>NR</p> <p>Inclusion criteria for study entry: Patients undergoing elective CABG, with ejection fraction of >50% and no arrhythmias before surgery</p> <p>Exclusion criteria for study entry:</p> <p>NR</p> <p>Baseline characteristics of subjects:</p> <p>NR</p>	<p>AF outcome:</p> <p>Incidence of AF</p> <p>Other outcomes (data not extracted):</p> <p>Defibrillation, total postoperative arrhythmias, undefined arrhythmia subgroup ('VES'), total lidocaine</p> <p>Adverse symptoms:</p> <p>NR</p> <p>Length of follow-up:</p> <p>Stated that an ECG was analysed for each patient 24 h postoperation (no other details reported)</p> <p>Recruitment dates:</p> <p>NR</p>
<p>^a It was not stated whether the bolus was administered intravenously; the composition and dilution of the infusion were not reported.</p>			
			<i>continued</i>

Results			
AF outcome	Intervention (n = 15)	Control (n = 15)	Difference
Incidence of AF, n	1	3	ns ^b
^b Lack of statistical significance inferred by reviewers (the difference was not marked as significant); threshold <i>p</i> -value not stated. Note that the incidence of all arrhythmias differed significantly between the groups (<i>p</i> < 0.01).			
Other outcomes	Outside scope of assessment: data not extracted		
Methodological comments			
<i>Allocation to treatment groups:</i> Stated that patients were randomised, but no details were provided.			
<i>Blinding:</i> NR.			
<i>Comparability of treatment groups:</i> Unclear, as few baseline characteristics were reported.			
<i>Method of data analysis:</i>			
<ul style="list-style-type: none"> • An unpaired <i>t</i>-test was used for data analysis. • Statistical testing was not supported by hypotheses; normality of data was not reported. 			
<i>Sample size/power calculation:</i> NR.			
<i>Attrition/dropout:</i> NR.			
General comments			
<i>Generalisability:</i> No details of ethnicity or social background; single-location study in Turkey.			
<i>Outcome measures:</i> Appropriate but described only briefly.			
<i>Intercentre variability:</i> NA.			
<i>Conflict of interests:</i> None reported.			

Reference and design	Intervention	Participants	Outcome measures
<p><i>Authors:</i> Zangrillo et al.²⁹</p> <p><i>Year:</i> 2005</p> <p><i>Country:</i> Italy</p> <p><i>Study design:</i> RCT</p> <p><i>Number of centres:</i> One</p> <p><i>Funding:</i> NR</p>	<p><i>Interventions:</i></p> <ol style="list-style-type: none"> 1. Intravenous magnesium sulphate 2. Placebo <p><i>Intervention details:</i></p> <ol style="list-style-type: none"> 1. 2.5g (20 mEq) magnesium sulphate diluted in 100 ml normal saline solution over a 30-minute period, immediately after central venous cannulation 2. Placebo of 100 ml normal saline solution 	<p><i>Number of participants:</i></p> <p>Intervention: 80 Control: 80</p> <p><i>Sample attrition/dropout:</i></p> <p>The analysis focused on patients receiving off-pump CABG; patients who required conversion to on-pump CABG were discontinued from the intervention (eight in magnesium group, nine in placebo group) but were included in the analysis on an ITT basis (all randomised patients were analysed in their allocated groups)</p> <p><i>Inclusion criteria for study entry:</i></p> <p>Patients referred for isolated elective coronary artery surgery; age ≥ 18 years; in sinus rhythm and for whom off-pump CBP was deemed technically feasible</p> <p><i>Exclusion criteria for study entry:</i></p> <p>History of AF; any other surgical procedure during the current admission; Q-wave MI in the preceding 6 weeks; ongoing treatment with amiodarone, digoxin or warfarin; permanent pacemaker implanted; or valvular regurgitation</p>	<p><i>AF outcomes:</i></p> <p>Incidence of AF (primary outcome); AF at hospital discharge</p> <p><i>Other outcomes extracted:</i></p> <p>Length of stay in ICU; length of stay in hospital</p> <p><i>Other outcomes (data not extracted):</i></p> <p>Ventricular tachycardia; low cardiac output; mechanical ventilation; biochemical data; oral amiodarone at hospital discharge; pacemaker; inotropes and transfusion frequency</p> <p><i>Adverse symptoms:</i> Stated that there were no serious adverse drug events and no patients died in the hospital</p>

continued

Reference and design	Intervention	Participants	Outcome measures	
		<p><i>Baseline characteristics of subjects</i></p> <p><i>Mean ± SD age (years):</i> Intervention: 65 ± 9.8 Control: 66 ± 9.7</p> <p><i>Gender (M/F):</i> Intervention: 74/6 (7.5% F) Control: 69/11 (13.8% F)</p> <p>Preoperative data were presented for 21 other variables; mean ± SD were reported for nine of these, but only <i>n</i> and % were reported for the remaining 12. The authors stated that baseline clinical characteristics and demographic details of the two groups were similar, but statistical significance was not mentioned and no <i>p</i>-values were provided</p> <p>There were no significant intraoperative differences (<i>p</i> > 0.05) between the groups in: number of grafts, pacemakers, ischaemia (>2 minutes), perioperative AF^a or patients with conversion from off-pump to on-pump CABG</p>	<p><i>Length of follow-up:</i> Not reported but mean length of hospital stay was ~1 week</p> <p><i>Recruitment dates:</i> May 2002 to March 2003</p>	
<p>^a Perioperative AF was restored to sinus rhythm in all cases by direct synchronised electric shock (20 J). Patients (<i>n</i>, %) with perioperative AF were: magnesium 7 (8.8%), placebo 4 (5.0%), <i>p</i> = 0.5.</p>				
Results^b				
AF outcomes		Intervention (<i>n</i> = 80)	Control (<i>n</i> = 80)	Difference
Postoperative AF, <i>n</i> (%) ^c		16 (20%) ^c	18 (22.5%)	<i>p</i> = 0.8 (ns)
AF at hospital discharge, <i>n</i> (%)		2 (2.5%)	2 (2.5%)	<i>p</i> = 0.7 (ns)
Other outcomes		Intervention (<i>n</i> = 80)	Control (<i>n</i> = 80)	Difference
Mean ± SD length of ICU stay (h)		26 ± 20.9	31 ± 24.4	<i>p</i> = 0.2 (ns)
Mean ± SD length of hospital stay (days) ^d		7 ± 3.8 ^d	6 ± 2.8 ^d	<i>p</i> = 0.004 ^d
<p>^b No indication was given about the timing of the results relative to patient admission, surgery or intervention.</p> <p>^c Four patients who had perioperative AF went on to also develop postoperative AF; these four were all in the magnesium group and their data are included here.</p> <p>^d The data given here are credited to two different comparisons in the primary article: (1) a comparison between magnesium and placebo groups (reported in Table 4), and (2) a comparison between AF and non-AF groups (reported in the text). The reviewers have assumed that the tabulated data (as extracted here) are correct (those data reported in the text appear inconsistent, suggesting a misprint).</p>				
Methodological comments				
<p><i>Allocation to treatment groups:</i> Randomisation was stratified in blocks of 40 patients; independent nurses dispensed either magnesium or placebo in the operating room according to a computer-generated list (no other details).</p> <p><i>Blinding:</i> Stated that all study personnel and participants were blinded to treatment assignment for the duration of the study (no details provided).</p> <p><i>Comparability of treatment groups:</i> Some slight preoperative differences in numbers in magnesium/placebo groups for diabetes (30/24), COPD (3/6), MI (44/49), β-blocker use (54/42), calcium channel blockers (24/37) and ACE inhibitors (31/36), but no variance measures or <i>p</i>-values were provided for these variables. The authors considered that the baseline characteristics were similar between the groups, but did not robustly test for differences.</p>				
				<i>continued</i>

Method of data analysis:

- The study tested the hypothesis that prophylactic magnesium supplementation as compared with placebo would reduce the occurrence of postoperative AF. Perioperative AF was noted, but not included in the analysis. All significance tests were two-sided.
- Analysis focused on patients receiving off-pump CABG; patients who required conversion to on-pump CABG were discontinued from the intervention (eight in magnesium group, nine in placebo group), but were included in the analysis on an ITT basis (no definition of ITT given).
- Dichotomous data were compared using a two-tailed χ^2 test with Yates' correction, or a Fisher's exact test when appropriate.
- Continuous measures were compared with analysis of variance or the Mann-Whitney *U*-test when appropriate.
- Presentation of results did not indicate which tests were actually used for each outcome; normality of outcomes was not reported.
- Stepwise multivariate logistic regression was used to investigate predictors of AF (analysis outside the scope of this assessment; data not extracted).

Sample size/power calculation: With two-sided $\alpha = 0.05$ and 80% power, based on previous publications, frequency of arrhythmia in intervention and placebo groups was assumed to be 30% and 50%, respectively, which would require 120 patients per group to detect. The actual number of patients recruited was 80 per group, suggesting that statistical power would have been <80%. However, the actual statistical power is unclear (interim analyses were included in the protocol for analyses with 40 patients per group or 80 patients per group, but the statistical power of the interim analyses was not stated; the study was terminated after interim analysis with 80 patients per group, as recruitment of a further 80 per group was considered unlikely to detect a benefit in either treatment group, but no explanation was given).

Attrition/dropout: All randomised patients were analysed; there were no losses to follow-up. Patients who withdrew from intervention (due to conversion to on-pump CABG) were included according to the ITT principle (ITT not defined).

General comments

Generalisability: No details of ethnicity or social background; single-location study in Italy.

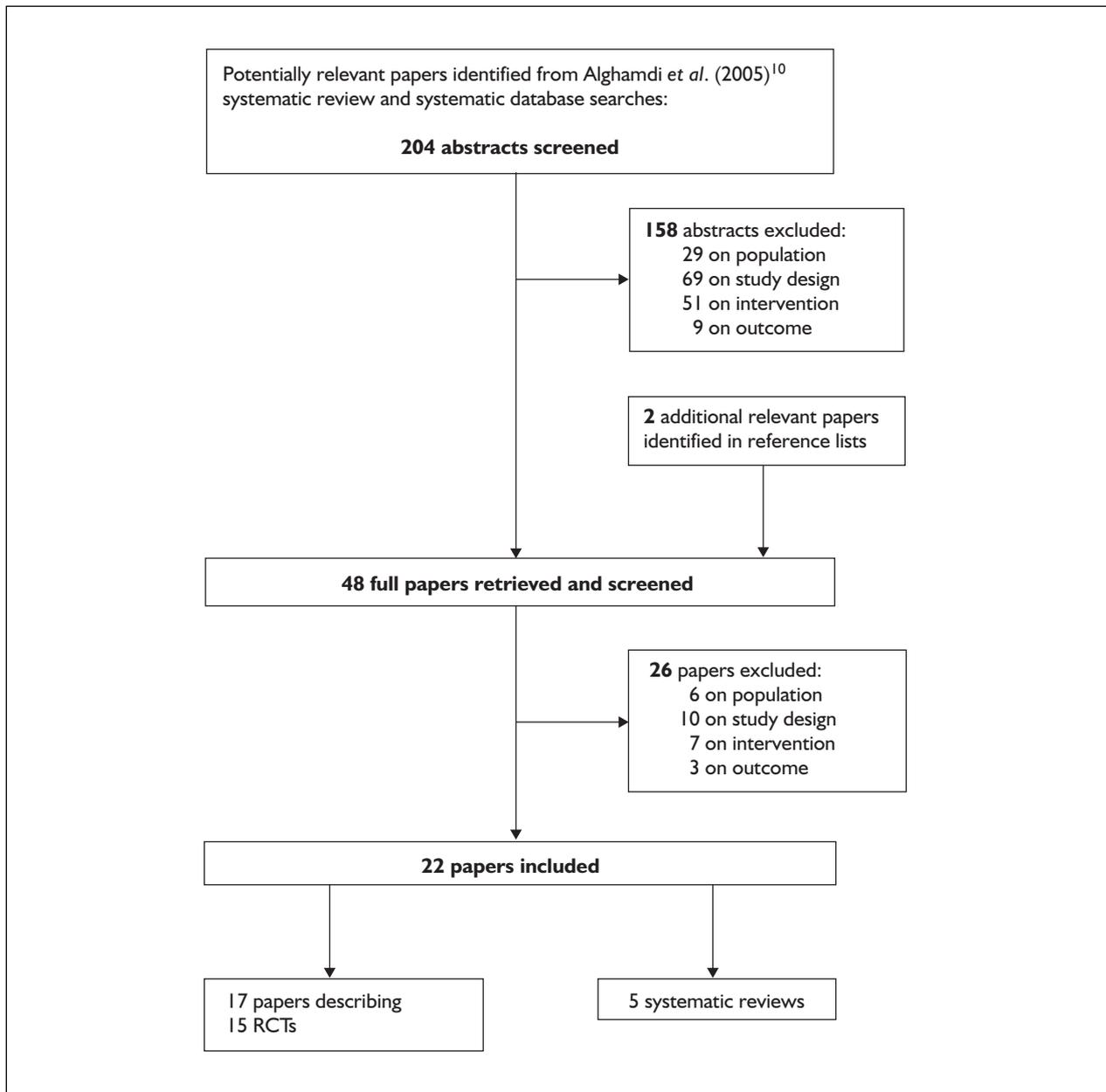
Outcome measures: Appropriate.

Intercentre variability: NA.

Conflict of interests: None reported.

Appendix 4

Studies screened and included in the current review



Appendix 5

Detection and definitions of atrial fibrillation given in the included RCTs

Study	Method of detecting AF	Definition of AF
Bert, 2001 ¹⁵	Continuous bedside ECG monitoring with automated alarmed arrhythmia detection and recall (Hewlett-Packard) in the ICU. On hospital wards, monitoring was performed by telemetry to a central nursing station to ≥ 4 days postoperation. Thereafter, clinical observation and a daily ECG study were performed. If a suspected arrhythmia occurred after discontinuation of continuous ECG monitoring, a 12-lead ECG study was obtained to confirm arrhythmia classification	AF was not defined separately from atrial flutter and SVT. These arrhythmias together were referred to as postoperative atrial tachycardias (POAT). All episodes of POAT had an initial ventricular rate > 110 bpm, were sustained for > 5 minutes and warranted pharmacological therapy. POAT classification excluded premature atrial contractions
Bhudia, 2006 ¹⁶	NR	NR (AF was not a primary outcome)
Caspi, 1995 ¹³	12-Lead ECG was performed before operation and at 6, 12 and 24 h after operation. All patients underwent ECG monitoring to 3 days postoperation using a bedside arrhythmia detection system (Mennen Horizon 2000)	Reported only that AF was at rates greater than 120 bpm
Dagdelen, 2002 ¹⁷	Standard 12-lead surface ECG was obtained on the day before operation, just after operation, and once daily for 4 days postoperation. After ICU discharge, trained nurses performed clinical observation every hour. Patients were monitored routinely with an alarm-triggered seven-lead telemetry system (Siemens Infinity) until 4 days postoperation under the close attention of a monitor technician. When there was a question of AF on clinical grounds or by telemetry, a 12-lead ECG was obtained	AF was defined as the absence of consistent P waves before each QRS complex, irregular QRS complexes and appearance of F waves
Fanning, 1991 ¹⁹	Continuous ECG monitoring to 4 days postoperation using a bedside arrhythmia detection system (Mennen Horizon 2000)	AF was not defined separately from atrial flutter and SVT with rates > 110 bpm and duration > 30 s
Forlani, 2003 ²¹	Cardiac rhythm was continuously monitored in the ICU until the morning of the second postoperative day. For the remaining days until discharge, 12-lead ECG was recorded every 8 h	NR
Hazelrigg, 2004 ²²	NR	NR
Kaplan, 2003 ²³	Rhythm was monitored continuously during the operation and first 2 days postoperation (Datascopie, Datascopie 2001A). In the wards patients were monitored with a 12-lead ECG and telemetry system (Fukuda Denshi DynaScope) if physical examination revealed a tachycardia attack or the development of an arrhythmia or if a patient had a palpitation or any rhythm-related complaint	NR

continued

Study	Method of detecting AF	Definition of AF
Karmy-Jones, 1995 ²⁴	Continuous monitoring for minimum 24 h using alarm-triggered bedside monitors (Hewlett-Packard) capable of automated recall, observed constantly by monitor nurses. ECG were obtained postoperatively and on the first 2 mornings postoperation. Thereafter, a 12-lead ECG was obtained if there was a question of dysrhythmia either on clinical grounds or by telemetry	NR
Nurözler, 1996 ²⁵	Continuous ECG over 5 days postoperation with 12-h memorised arrhythmia detection monitor (Drager U-M 3). AF identified on the monitor confirmed by a reading of a rhythm strip. AF excluded if not confirmed by ECG or if <1 minute duration	NR (except for duration \geq 1 minute)
Parikka, 1993 ²⁶	Continuous monitoring of cardiac rhythm with bedside monitors on the first 2 days postoperation in the ICU. Thereafter, for up to 10 days, every symptomatic palpitation was recorded by 12-lead ECG. A 48-h two-channel Holter recording was started 2 days postoperation. The recording was analysed by one of the investigators using a Marquette electrocardioscanner	Clinical AF was defined as irregular QRS complexes without detectable regular atrial activity continuing for 1 h, or less if treatment was necessary because of intolerable symptoms or haemodynamic deterioration. Episodes of this type of arrhythmia seen at Holter were defined as AF episodes if they lasted for 15 s and were at least 1 minute apart. Holter recording did not reveal any AF patient who had not been identified by clinical symptoms
Toraman, 2001 ²⁷	ECGs were obtained preoperatively and 0–5 days postoperatively. Patients were continuously monitored while in the ICU using alarm-triggered bedside monitors (Siemens). After discharge from the ICU, trained nurses performed clinical observations every 4 h and all patients were monitored with an alarm-triggered seven-lead telemetry system (Siemens Infinity), under close attention of a monitor technician, until the morning of day 5 postoperation. When AF was suspected clinically or by telemetry a 12-lead ECG was obtained	AF was defined as the absence of consistent P waves before each QRS complex and an irregular rate lasting for >10 minutes or requiring therapy as a result of haemodynamic compromise
Treggiari-Venzi, 2000 ²⁸	A Holter ECG (Delmar Avionics three-channel Cardiocorder) recording was obtained throughout the 72-h infusion period. Additional 12-lead ECGs were recorded every 12 h. The Holter recording was analysed on completion of the 72-h study period or earlier if the study was terminated because of arrhythmia. SVA episodes were detected visually and printed for accurate diagnosis by two investigators	SVT was defined as an arrhythmia of more than three narrow QRS complexes at a rate greater than 100 bpm and lasting for more than 30 s. AF was defined as a totally irregular atrial rhythm leading to irregular ventricular rhythm. Unclear whether AF definition was independent of the definition of SVT; note that the authors mention both SVA and SVT as extracted here
Yilmaz, 2000 ¹⁴	A 12-lead standard ECG was analysed for each patient during 30 minutes and 24 h postoperation	NR
Zangrillo, 2005 ²⁹	Continuous monitoring until hospital discharge through telemetry with continuous display of the ECGs on multiple oscilloscopes simultaneously in the ICU, high-dependency unit (HDU) or main ward. 12-Lead ECGs were obtained routinely every 6 h in ICU and daily in HDU, at hospital discharge, and when AF was detected on telemetry. In addition, ECG was performed whenever AF was suspected clinically	Any documented episode of AF (12-lead ECG) defined as an irregular rhythm with an irregular fluctuating baseline, without well-defined P waves and irregular RR intervals. AF excluded if \leq 10 minutes or requiring medical attention owing to patient instability

Appendix 6

Exclusion criteria reported for the selection of patients in RCTs

Exclusion criterion	Study														
	Bert ¹⁵	Bhudia ¹⁶	Caspi ¹³	Dagdelen ^{17,18}	Fanning ¹⁹	Forlani ^{20,21}	Hazelrigg ²²	Kaplan ²³	Karmy-Jones ²⁴	Nurözler ²⁵	Parikka ²⁶	Toraman ²⁷	Treggiari-Venzi ²⁸	Yilmaz ¹⁴	Zangrillo ²⁹
Previous AF		×		×				×				×	×		×
Atrial arrhythmias	×														
Atrioventricular block				×									×		
Angina								×							
Atrioventricular node disorder						×									
Bradycardia								×							
Bronchospastic	×														
Bundle branch block				×											
COPD (defined or undefined)				×		×				×		×			
Diabetes												×			
Heart block, heart disease				×	×					×					
Liver dysfunction								×					×		
Hypotension							×	×							
Inotropic drugs required										×					
Intra-aortic balloon required										×					
Ischaemia ongoing									×						
LV function or LVEF abnormal	×					×				×					
Non-sinus rhythm	×														
Pacemaker/cardiac assistance					×		×						×		×
Pericarditis				×											
Pre-excitation syndrome				×											
Pulmonary embolism				×											
QT interval increased						×									
Q wave MI ≤6 weeks presurgery															×
Creatinine elevated					×	×	×	×		×					
Creatinine clearance decreased													×		
Reoperation								×	×						
Severe respiratory disorder								×							
Sick sinus syndrome				×		×									
Surgery, cardiac	×				×								×		
Surgery, concomitant										×					×
Surgery, emergency									×						
SVA or SVT					×	×				×					
Valve disease or valve surgery				×				×			×	×	×		×
Ventricular arrhythmias	×				×					×					
Non-consent or unavailable consent					×				×				×		
Renal disease, impairment or failure	×	×		×				×	×			×	×	×	
Thyroid disorders				×								×	×		
Antiarrhythmics				×		×			×		×	×	×		
Digoxin				×								×			
β-blockers				×								×			
Amiodarone ≤1 year presurgery													×		
Amiodarone ongoing															×
Digoxin ongoing															×
Warfarin ongoing															×
No exclusion criteria reported			×										×	×	

LV, left ventricular; ×, exclusion criterion reported.

Appendix 7

Search strategy for economic evaluations

Ovid MEDLINE

1950 to May week 1 2007

Searched 15 May 2007

1 "Atrial Fibrillation"/ (19465)
 2 Atrial Flutter/ (3991)
 3 (atrial adj3 fibrillat\$.mp. (25464)
 4 (heart adj3 fibrillat\$.mp. (950)
 5 atrial fibrillat\$.mp. (25294)
 6 (atrium adj3 flutter\$.mp. (23)
 7 (atrium adj3 fibrillat\$r).mp. (0)
 8 (auricular\$ adj3 fibrillat\$.mp. (736)
 9 (auricular\$ adj3 flutter\$.mp. (274)
 10 Tachycardia Supraventricular/ (3781)
 11 or/1-10 (30644)
 12 "Coronary Artery Bypass"/ (31512)
 13 (CABG or coronary artery bypass graft\$.ti,ab.
 (16229)
 14 (coronary artery bypass adj3 surgery).ti,ab.
 (7332)
 15 (coronary adj3 revasculari\$.ti,ab. (4439)
 16 (myocardial adj3 revascular\$.ti,ab. (4063)
 17 (coronary adj6 graft).ti,ab. (6832)
 18 (coronary adj6 bypass).ti,ab. (26085)
 19 (coronary adj6 surgery).ti,ab. (15257)
 20 (aortocoronary adj6 bypass).ti,ab. (2453)
 21 (valve\$ adj6 surgery).ti,ab. (4554)
 22 Cardiac Surgical Procedures/ (23335)
 23 or/12-22 (71271)
 24 MgSO4.mp. (1029)
 25 magnesium sulfate/ (3419)
 26 magnesium sulfate\$.ti,ab. (1735)
 27 exp magnesium compounds/ (10471)
 28 magnesium/ (56302)
 29 (magnesium and (sulphate\$ or sulfate\$)).ti,ab.
 (3192)
 30 magnesium sulphate\$.ti,ab. (729)
 31 magnesium sulfate\$.ti,ab. (1735)
 32 7487-88-9.rn. (3419)
 33 Magnesium Deficiency/ (3387)
 34 or/24-33 (67330)
 35 3930-20-9.rn. (1695)
 36 Sotalol/ (1695)
 37 sotalol.ti,ab. (1997)
 38 or/35-37 (2418)
 39 11 and 23 and 34 (48)
 40 23 and 34 (242)
 41 11 and 34 (151)

42 39 or 40 or 41 (345)
 43 limit 42 to (humans and english language and
 yr="2004 - 2007") (45)
 113 exp ECONOMICS/ (374970)
 114 exp ECONOMICS, HOSPITAL/ (14701)
 115 exp ECONOMICS, PHARMACEUTICAL/
 (1761)
 116 exp ECONOMICS, NURSING/ (3739)
 117 exp ECONOMICS, DENTAL/ (3555)
 118 exp ECONOMICS, MEDICAL/ (11345)
 119 exp "Costs and Cost Analysis"/ (129053)
 120 Cost-Benefit Analysis/ (39951)
 121 VALUE OF LIFE/ (4844)
 122 exp MODELS, ECONOMIC/ (5178)
 123 exp FEES/ and CHARGES/ (7124)
 124 exp BUDGETS/ (9947)
 125 (economic\$ or price\$ or pricing or financ\$ or
 fee\$ or pharmacoeconomic\$ or pharma
 economic\$.tw. (325412)
 126 (cost\$ or costly or costing\$ or costed).tw.
 (189830)
 127 (cost\$ adj2 (benefit\$ or utilit\$ or minim\$ or
 effective\$)).tw. (48757)
 128 (expenditure\$ not energy).tw. (10369)
 129 (value adj2 (money or monetary)).tw. (595)
 130 budget\$.tw. (10852)
 131 (economic adj2 burden).tw. (1410)
 132 "resource use".ti,ab. (2055)
 133 or/113-131 (754771)
 134 news.pt. (118759)
 135 letter.pt. (589482)
 136 editorial.pt. (202717)
 137 comment.pt. (329262)
 138 or/134-137 (951676)
 139 133 not 138 (695184)
 140 42 and 139 (8)
 141 from 140 keep 4-6,8 (4)
 142 23 and 38 and 139 (13)
 143 from 142 keep 4-5,8-9,12-13 (6)
 144 11 and 38 and 139 (27)
 145 limit 144 to english language (25)
 146 from 145 keep 1,9,11-15,17,19-20,24 (11)
 147 34 and 38 and 139 (2)

Number of hits (download file)

Costs magnesium: 4

Costs sotalol: 17

Costs magnesium + sotalol: 2

EMBASE

1980 to 2007 week 19

Searched 16 May 2007

- 1 heart atrium fibrillation/ (21875)
- 2 heart atrium flutter/ (3362)
- 3 exp Heart Atrium Arrhythmia/ (35266)
- 4 (atrial adj3 fibrillat\$.mp. (17803)
- 5 (heart adj3 fibrillat\$.mp. (31983)
- 6 atrial fibrillat\$.mp. (17577)
- 7 (atrium adj3 flutter\$.mp. (3380)
- 8 (atrium adj3 fibrillat\$r).mp. (0)
- 9 (auricular\$ adj3 fibrillat\$.mp. (108)
- 10 (auricular\$ adj3 flutter\$.mp. (29)
- 11 Supraventricular Tachycardia/ (5812)
- 12 or/1-11 (47152)
- 13 coronary artery bypass graft/ (22523)
- 14 Coronary Artery Bypass Surgery/ (6944)
- 15 exp coronary artery surgery/ (39530)
- 16 (CABG or coronary artery bypass graft\$.ti,ab. (14959)
- 17 (coronary artery bypass adj3 surgery).ti,ab. (6804)
- 18 (coronary adj3 revasculari\$.ti,ab. (4087)
- 19 (myocardial adj3 revascular\$.ti,ab. (3015)
- 20 (coronary adj6 graft).ti,ab. (6026)
- 21 (coronary adj6 bypass).ti,ab. (23026)
- 22 (coronary adj6 surgery).ti,ab. (13234)
- 23 (aortocoronary adj6 bypass).ti,ab. (1469)
- 24 (valve\$ adj6 surgery).ti,ab. (3618)
- 25 or/13-24 (50320)
- 26 MgSO4.mp. (1011)
- 27 magnesium sulfate/ (6009)
- 28 Magnesium Derivative/ (989)
- 29 magnesium/ (23019)
- 30 (magnesium and (sulphate\$ or sulfate\$)).ti,ab. (2910)
- 31 magnesium sulphate\$.ti,ab. (642)
- 32 magnesium sulfate\$.ti,ab. (1548)
- 33 7487-88-9.rm. (6022)
- 34 Magnesium Deficiency/ (1203)
- 35 or/26-34 (30923)
- 36 3930-20-9.rm. (7170)
- 37 Sotalol/ (7144)
- 38 sotalol.ti,ab. (2048)
- 39 or/36-38 (7279)
- 40 12 and 25 and 35 (88)
- 41 25 and 35 (209)
- 42 12 and 35 (667)
- 43 41 or 42 (788)
- 44 limit 43 to (humans and english language and yr="2004 - 2007") (229)
- 45 12 or 25 (95066)
- 46 35 and 39 and 45 (149)
- 92 (cost\$ adj2 effective\$.ti,ab. (37066)
- 93 (cost\$ adj2 benefit\$.ti,ab. (7616)
- 94 cost effectiveness analysis/ (48505)

- 95 cost benefit analysis/ (26055)
- 96 budget\$.ti,ab. (7837)
- 97 cost\$.ti. (34748)
- 98 (cost\$ adj2 (effective\$ or utilit\$ or benefit\$ or minimi\$)).ab. (40355)
- 99 (economic\$ or pharmacoeconomic\$ or pharmaco economic\$.ti. (13532)
- 100 (price\$ or pricing\$.ti,ab. (10064)
- 101 (financial or finance or finances or financed).ti,ab. (20781)
- 102 (fee or fees).ti,ab. (4745)
- 103 cost/ (19051)
- 104 cost minimization analysis/ (1127)
- 105 cost of illness/ (3798)
- 106 cost utility analysis/ (1908)
- 107 drug cost/ (29425)
- 108 health care cost/ (52431)
- 109 health economics/ (9184)
- 110 economic evaluation/ (3590)
- 111 economics/ (5146)
- 112 pharmacoeconomics/ (884)
- 113 budget/ (6946)
- 114 economic burden.ti,ab. (1415)
- 115 "resource use".ti,ab. (1899)
- 116 or/92-115 (214586)
- 117 (editorial or letter).pt. (525894)
- 118 116 not 117 (191116)
- 119 118 and 35 and (12 or 25) (43)
- 120 from 119 keep 7,15,22,32,34,39 (6)
- 121 118 and 39 and (12 or 25) (137)
- 122 121 not 119 (123)
- 123 limit 122 to english language (118)
- 124 from 123 keep 3,10,18,25-27,31-32,42,47, 49-50,52-56,61,63-64,66,68-69,71,73,77, 79-80,86,88-89,92,94-97,99-100,102-104, 106,111-113,115-118 (49)

Number of hits (download file)

Costs magnesium: 6

Costs sotalol: 49

Ovid MEDLINE In-Process and Other Non-Indexed Citations

14 May 2007

Searched 15 May 2007

- 1 (atrial adj3 fibrillat\$.mp. (728)
- 2 (heart adj3 fibrillat\$.mp. (32)
- 3 atrial fibrillat\$.mp. (727)
- 4 (atrium adj3 flutter\$.mp. (1)
- 5 (atrium adj3 fibrillat\$r).mp. (0)
- 6 (auricular\$ adj3 fibrillat\$.mp. (16)
- 7 (auricular\$ adj3 flutter\$.mp. (0)
- 8 or/1-7 (747)
- 9 (CABG or coronary artery bypass graft\$.ti,ab. (482)

- 10 (coronary artery bypass adj3 surgery).ti,ab.
(220)
- 11 (coronary adj3 revasculari\$.ti,ab. (134)
- 12 (myocardial adj3 revascular\$.ti,ab. (83)
- 13 (coronary adj6 graft).ti,ab. (173)
- 14 (coronary adj6 bypass).ti,ab. (638)
- 15 (coronary adj6 surgery).ti,ab. (353)
- 16 (aortocoronary adj6 bypass).ti,ab. (22)
- 17 (valve\$ adj6 surgery).ti,ab. (117)
- 18 or/9-17 (968)
- 19 MgSO4.mp. (38)
- 20 (magnesium and (sulphate\$ or sulfate\$)).ti,ab.
(98)
- 21 magnesium sulphate\$.ti,ab. (23)
- 22 magnesium sulfate\$.ti,ab. (55)
- 23 19 or 20 or 21 or 22 (129)
- 24 sotalol.ti,ab. (30)
- 25 8 or 18 (1659)
- 26 23 and 25 (3)
- 27 from 26 keep 1-3 (3)
- 28 24 and 25 (14)
- 29 from 28 keep 5,8,10-13 (6)
- 30 (economic\$ or price\$ or pricing or
pharmacoeconomic\$ or pharma
economic\$.tw. (3394)
- 31 (cost\$ or budget\$.tw. (7334)
- 32 (cost\$ adj2 (benefit\$ or utilit\$ or minim\$)).tw.
(373)
- 33 (value adj2 (money or monetary)).tw. (31)
- 34 30 or 31 or 32 or 33 (9858)
- 35 23 and 34 (2)

Number of hits (download file)

Costs magnesium: 3

Costs sotalol: 2

NHS EED (CRD database)

Magnesium and (atrial fibrillate* or coronary)

Sotalol and atrial fibrillat*

Number of hits (download file)

Costs magnesium: 1

Costs sotalol: 6

Econlit

Magnesium 0

Sotalol 0

Number of hits (download file)

0

Total keyworded

Costs magnesium: 8

Costs sotalol: 25

Number of hits (download file)

8

55

Appendix 8

Inclusion criteria for economic review

Q1. Economic evaluation or costing study	Q2. Prophylaxis for AF following CABG	Q3. Comparators	Q4. Include
1 = Full economic evaluation (CBA, CUA, CEA or CMA) 2 = Costing study (no assessment of effectiveness). Reports costings or resource use 3 = No 4 = Unclear	1 = Y 2 = N 4 = Unclear	1 = Mg 2 = Sotalol 3 = Other	1 = Y (primary) 2 = Y (secondary) 3 = N 4 = Unclear
If Q1 = (1 or 2) AND Q2 = 1 AND Q3 = (1 and 2) – then include = 1 If Q1 = (1 or 2) AND Q2 = 1 AND Q3 = (1 or 2 or 3) – then include = 2 If 4 – then include = 4 Otherwise include = 3 CBA, cost–benefit analysis; CEA, cost-effectiveness analysis; CMA, cost-minimisation analysis; CUA, cost–utility analysis.			



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Volume 1, 1997

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Diagnosis, management and screening of early localised prostate cancer.

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Newborn screening for inborn errors of metabolism: a systematic review.

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Antenatal screening for Down's syndrome.

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Informed decision making: an annotated bibliography and systematic review.

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Volume 5, 2001

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Clinical and cost-effectiveness of donepezil, rivastigmine and galantamine for Alzheimer's disease: a rapid and systematic review.

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