A modified version of this manuscript has been accepted by the Journal of Survey Statistics and Methodology on 17th March 2015. Please refer to this paper for citation.

To cite:

Durrant, G.B., Maslovskaya, O., Smith, P.W. (2015) Modelling Final Outcome and Length of Call to Improve Efficiency in Call Scheduling, *Journal of Survey Statistics and Methodology*, forthcoming.

Modelling Final Outcome and Length of Call Sequence to Improve Efficiency in Interviewer Call Scheduling

Gabriele B. Durrant, Olga Maslovskaya and Peter W.F. Smith

Southampton Statistical Sciences Research Institute
University of Southampton, United Kingdom
(g.durrant@southampton.ac.uk)

Address for correspondence:

Gabriele Durrant

Southampton Statistical Sciences Research Institute

University of Southampton

SO17 1BJ, Southampton

United Kingdom

g.durrant@southampton.ac.uk

Abstract:

Survey practitioners are increasingly interested in how best to use paradata to improve data collection processes. One particular question is if it is possible to identify early on during fieldwork sample cases that may require a long time, and therefore a lot of financial and staff resources, until interviewing is completed. More specifically, we aim to identify cases with long unsuccessful call sequences. This paper models call record data predicting final call outcome and length of a call sequence. Separate binary and joint multinomial logistic models for the two outcomes are presented, accounting for the clustering of households within interviewers. Of particular interest is to identify explanatory variables that predict final outcome and length of a call sequence. The study uses data from Understanding Society, a large-scale UK longitudinal survey. The work has implications for responsive and adaptive survey designs. The results indicate that modelling outcome and length of a call sequence jointly improves the fit of the model. Outcomes of previous calls, in particular from the most recent call, are highly predictive. The timing of calls and interviewer observation variables, although significant in the models, only slightly improve the predictive power.

Key Words: survey non-response, interviewer call record data, paradata, call sequence, responsive and adaptive survey designs.

Acknowledgement

This work was supported by the UK Economic and Social Research Council (ESRC), 'The Use of Paradata in Cross-Sectional and Longitudinal Research' [grant number: RES-062-23-2997], grant holders: Gabriele B. Durrant, Peter W. F. Smith and Frauke Kreuter.

Data Statement

This study uses Wave 1 data from Understanding Society, the United Kingdom Household Longitudinal Study (UKHLS). The data were obtained from the UK Data Archive (http://www.data-archive.ac.uk/). Data reference: University of Essex. Institute for Social and Economic Research and National Centre for Social Research, Understanding Society: Wave 1-3, 2009-2012. 5th Edition. Colchester, Essex: UK Data Archive, November 2013. SN: 6614, http://dx.doi.org/10.5255/UKDA-SN-6614-5

1. Introduction

For interviewer administered surveys many statistical agencies nowadays routinely collect call record data. Examples of such data are recordings of the day, time and outcome of each call or visit and any observations made about the person talked to at the call. For face-to-face surveys a range of interviewer observation variables, such as physical and social characteristics of the selected household and neighbourhood, may also be recorded. Researchers have become increasingly interested in how best to use and analyse such information. It is hoped that a better understanding of the calling patterns and the mechanisms leading to particular call sequences help improve data collection through reducing both costs and non-sampling errors. For statistical agencies, investigating time and effort into repeated calls and follow-ups to a sample unit is very resource- and cost-intensive. It is therefore desirable to avoid long unsuccessful call sequences to improve efficiency of call scheduling. The aim is to identify cases prone to long and unsuccessful call sequences.

This paper presents models for call record data predicting final call outcome and length of a call sequence early on in the data collection process, say after the first three calls. Separate binary logistic models and a multinomial logistic model for the two outcomes combined are considered. The clustering of sample cases within interviewers is taken into account by using robust standard error estimation. To assess the accuracy of the models, a range of methods are employed, including the widely used R² statistic and classification tables (Agresti 2013). Concepts from epidemiology are introduced here into the context of survey methodology including discrimination and prediction (Plewis, Ketende, and Calderwood 2012; Pepe 2003). A particular focus is the identification of explanatory variables that predict final outcome and/or length, especially those characterising long unsuccessful call sequences. Further research

questions that we aim to address in this study are: how can call record predictors best be incorporated into the model(s): as summary statistics or as individual outcomes?; how predictive are the models?; does their ability to predict improve once more and more call record data are available?; can we achieve similarly performing models by using fully observed auxiliary variables available prior to any data collection?; how can these models best be used in adaptive and responsive survey designs?

Past research mostly aimed to predict final non-response, often did not include paradata and relied on fully observed frame data or socio-demographic variables. However, such models have not been found to predict well the outcomes of calls (Groves and Couper 1996; Groves and Couper 1998; Bates, Dahlhamer, and Singer 2008). In recent years, researchers have explored the potentials of including newly available paradata, such as call record data and interviewer observation variables, in models for non-response outcomes with some success (Potthoff, Manton, and Woodbury 1993; Groves and Couper 1996; Sinibaldi, Durrant, and Kreuter 2013; Sinibaldi, Trappmann, and Kreuter 2014; Bates, Dahlhamer, and Singer 2008; Wagner 2013a and 2013b; Kreuter et al. 2010), but typically still with low predictive power (R² values between 3%-8%) (Olson, Smyth, and Wood 2012; Olson and Groves 2012; West and Groves 2013). Also, the length of a call sequence to obtain a response has not been the focus and has been, if at all, used only as an explanatory variable in models. Furthermore, the use of discrete time event history analysis to predict the outcome at the next call rather than the final outcome has been advocated, exploiting call record data as explanatory variables (Groves and Heeringa 2006; Durrant, D'Arrigo, and Steele 2011 and 2013; Durrant, D'Arrigo, and Müller 2013; Sinibaldi 2014; Hanly 2014). However, specialised software and knowledge of such modelling techniques

are required. Also the outcome over the next few calls may be of greater interest than the outcome of the next call only.

One novel aspect of this paper is to consider modelling *both* length and outcome of call sequences simultaneously. For comparison, separate and joint models for the two outcomes are developed. The method proposed employs standard modelling techniques, which is an advantage. The study uses data from a large-scale face-to-face longitudinal survey in the UK, Understanding Society. The research is motivated by findings from a recent sequence analysis of the same data (Durrant, Maslovskaya and Smith 2014), which highlights the importance of sequence length and outcome as two key characteristics of the large number of different sequences, supporting findings from earlier work (Kreuter and Kohler 2009). This analysis identified a significant number of long call sequences (up to 30 calls) for some households and a distinct grouping of sequences into short and long, successful and unsuccessful call sequences.

The research has implications for survey practice, in particular for adaptive and responsive survey designs. The method developed may be of particular benefit for longitudinal surveys, where the same or similar auxiliary variables, are available for every wave. The models could then be used to predict final outcome and sequence length for future waves.

The remainder of the paper is structured as follows. First, the Understanding Society survey and the analysis sample are discussed. The methods section presents the logistic regression models, the modelling strategy and methods to compare and assess different models. Then, the results are presented. The paper concludes with a summary of the main findings, implications for survey practice, limitations and further work.

2. Data

2.1 The Understanding Society Survey

This paper uses call record data and interviewer observations from the first wave of the UK Understanding Society survey, which is the Household Longitudinal Study in the United Kingdom. The survey covers topics of health, work, education, income, family and social life to help understand the long term effects of social and economic change, as well as policy interventions. The study has many advantages over previously existing datasets in the UK by being exceptionally large and comprehensive. In particular, as part of the study a range of paradata were collected, including call record data and a wide range of interviewer observation variables. Also, only interviewers with above average experience and ability were selected for the study.

Data collection for each wave is scheduled across a 24-months period, with interviews taking place annually. Wave 1 data collection took place between January 2009 and March 2011. All Wave 1 interviews were carried out face-to-face in respondents' homes by trained interviewers using computer-assisted personal interviewing (CAPI). Households are therefore clustered within interviewers. All adult household members (age 16 and older) are asked to respond. In addition to individual interviews, a member of the household needs to respond to a household questionnaire. Interviewers have one month to contact households. A minimum of six calls are made at each sampled address before it is considered unproductive, but interviewers are encouraged to make further calls if possible (McFall 2012).

During the first visit to a household, interviewers collect a wide range of *interviewer observations* capturing characteristics about each household and surrounding neighbourhood.

(see online Appendix Table A1 for wording of all variables considered). In addition, *call record data* are available, which capture information about each call, including outcome, date and time of each call. The outcome of a call in the Understanding Society survey is defined as non-contact, contact, appointment, interview, and 'any other status' (which includes ineligibles and refusals). For this study, call record data, including final response outcome, are combined with interviewer observation data, variables about the study design and geographic information, using a unique household identifier. All variables are available for both respondents and non-respondents.

2.2 Analysis sample and definition of response and explanatory variables

The analysis was initially carried out on all households from Wave 1, excluding cases from an Ethnic Minority Boost sample as rules for the selection of this sample differ from the main sample. A separate analysis could be undertaken for this group. This distinction is, however, not of interest here. Ineligibles are included in our analysis since it is a true outcome and survey agencies may be interested in mechanisms for identifying ineligibles as early as possible to reduce survey costs. This initial analysis sample contained 47,913 households (including both responding and non-responding households). First, we carried out preliminary work using this initial analysis sample, conducting separate analyses for all cases with at least 1, 2, 3 etc. calls. The aim was to evaluate after every call, how well we can predict final outcome and length as more information about each case becomes available. This preliminary analysis suggested that 3 calls may be sufficient to reach an acceptable level of predictability. To more formally assess this and to be able to compare model fit and prediction across the different models -since measures of predictability are dependent on sample size- we restrict our final analysis sample to all households from Wave 1 that received four or more calls (27,995 households). This approach enables us to employ call record information from the beginning of the field period (here the first three calls) to predict final length and outcome. For these final models, we also excluded a small number of cases with missing values in the explanatory variables, as follows: there are no missing cases in any of the geographic information and design variables since these are derived from administrative data. Date and time of a call are automatically captured using computer assisted methods leading to no missing cases in these variables. Recordings of the call outcome, contained a very small amount of missing data, and such calls were excluded (311 cases). Households with missing items in the interviewer observation variables were deleted (2,015 cases). The *final analysis sample*, including only cases with four or more calls, contains 25,358 households within 734 interviewers.

To analyse length and call outcome (and to ensure the comparability of the different models), the following three response variables are considered. Their distributions are presented in Table 1.

- 1.) *length of call sequence* (binary), distinguishing short sequences (up to 6 calls) and long sequences (more than 6 calls). The cut-off point at 6 calls was intentionally selected to fit with the survey protocol requirements of conducting a minimum of 6 call attempts if contact was not established earlier in the process.
- 2.) (final) outcome of call sequence (binary), distinguishing successful call sequences with at least one interview conducted in a household (after call 3) and unsuccessful call sequences with no interviews achieved (after call 3).
- 3.) a variable combining both length and final outcome, distinguishing 4 categories short unsuccessful (up to 6 calls, no interview), short successful (up to 6 calls, at least one interview after call 3), long unsuccessful (more than 6 calls, no interview) and long successful (more than 6 calls, at least one interview after call 3).

(Different definitions of the dependent variables - such as complete household outcome and different categorisations of sequence length - were also explored. The overall conclusions were very similar to the ones presented in this paper.)

Table 1: Distributions of the three response variables in the final analysis sample (25,358 households).

Variables with categories	Frequencies	Percentages
Length		
Short sequence (up to 6 calls)	12353	48.7
Long sequence (7-30 calls)	13005	51.3
Final outcome		
No single interview in a sequence	13565	53.5
At least one interview in a sequence	11793	46.5
Combined response		
Short unsuccessful	4962	19.6
Short successful	7391	29.1
Long unsuccessful	8603	33.9
Long successful	4402	17.4

The explanatory variables used in the analysis can be split into three main groups. The distributions of the explanatory variables broken down by the categories of the response variables are presented in the online Appendix (Table A2).

- 1.) *geographic information and design variables* (4 variables: urban/rural indicator, government office region, low density area for ethnic minorities and month of household issue);
- 2.) interviewer observation variables (14 variables, e.g. indicators of entry barriers, conditions of surrounding area such as litter in street, abandoned buildings, heavy traffic, type of accommodation, presence of children in a household, relative condition of the property, garden);

3.) *call record variables* (20 variables, e.g. date, time of day, day of week, call outcome; also derived variables including time between calls, number of previous non-contacts, contacts, appointments and broken appointments).

3. Methods

3.1 Binary and multinomial regression models

To model length and final outcome binary logistic regression analysis is used. Multinomial logistic regression is employed to model the combined response variable as it contains four categories. To correctly control for the clustering of households within interviewers robust standard error estimation is employed (Huber 1967; White 1980, 1984 and 1994). As an alternative modelling approach, multilevel models could have been employed also allowing for the nesting of households within interviewers. However, since interviewer effects are not of substantive interest here (e.g. we are not interested in explaining effects of contextual or interviewer-level variables), a multilevel modelling approach is not necessary. To assess predictability of different models R² statistics and classification tables are obtained.

We denote by \boldsymbol{y}_i a binary response variable of household i. The dependent variable 'length of call sequence' is coded

$$y_i = \begin{cases} 1 & \text{short call sequence (up to 6 calls)} \\ 0 & \text{long call sequence (more than 6 calls)} \end{cases}$$

and the dependent variable 'final outcome of call sequence' is coded

$$y_i = \begin{cases} 1 & \text{successful call sequence (at least one interview after call 3)} \\ 0 & \text{unsuccessful call sequence (no interview after call 3).} \end{cases}$$

The response probabilities are denoted by π_i =Pr(y_i =1) and are related to the explanatory variables using logistic regression (e.g. Agresti 2013):

$$logit(\pi_i) = log(\frac{\pi_i}{1-\pi_i}) = \boldsymbol{\beta}^T \mathbf{x}_i,$$

where \mathbf{x}_i is a vector of household-level covariates including intercept and interactions, and $\boldsymbol{\beta}$ is a vector of coefficients.

Multinomial logistic regression is an extension of binary logistic regression. If the response variable has S categories, then the multinomial logistic regression model can be expressed as a set of S-1 non-redundant logistic model equations. We denote by y_i the dependent variable which combines length and final outcome in household i, coded

$$y_i = \begin{cases} 1 & \text{short successful} \\ 2 & \text{short unsuccessful} \\ 3 & \text{long successful} \\ 4 & \text{long unsuccessful}. \end{cases}$$

The response probabilities are denoted by $\pi_i^{(s)} = \Pr(y_i = s)$, s = 1, 2, 3, 4. Taking 'long unsuccessful' as the reference category, the multinomial logistic regression model can be expressed as

$$\log\left(\frac{\pi_{i}^{(s)}}{\pi_{i}^{(4)}}\right) = \boldsymbol{\beta}^{(s)^{T}} \mathbf{x}_{i}^{(s)}, s=1, 2, 3,$$

where $\mathbf{x}_i^{(s)}$ is a vector of covariates including intercept and interactions, and $\boldsymbol{\beta}^{(s)}$ is a vector of coefficients.

3.2 Modelling strategy and comparison of models

Likelihood ratio tests (using the change in the L^2 goodness-of-fit statistic) are used to test the significance of a term in the model. For data preparation and analysis we used SPSS version

20 and STATA 12. STATA can estimate robust standard errors to control for the non-independence of observations (*ro* function; Long and Freese 2006). To simplify the interpretation of the modelling results predicted probabilities are obtained for specific variables, holding all other variables constant at their means.

Modelling Strategy

A forward stepwise model selection procedure is employed. Explanatory variables, together with selected interaction effects, are included into the models: first, only geographic and design variables are included; then interviewer observations are added. Information about the timing of the first call attempt follows, then timing of second and third call (including time between calls). Information about the outcome for the first three calls are added one by one and then as combinations, either as raw outcomes (i.e. outcome of first, second and third call, interactions between outcomes) or as summary information (i.e. number of non-contacts, contacts, appointment and interviews across the first three calls). Allowing for interaction effects between outcomes accounts for the sequence of outcomes as a whole rather than simply as individual outcomes. For example, a non-contact call after an appointment should be interpreted differently from a non-contact after a previous non-contact. The former may reflect an indirect refusal, whereas the latter indicates a longer period of absence. At the last stage of the modelling strategy, time of day and day of the week of the fourth call is added to the models, but without call outcome of call 4. Here, we would like to know if we can predict length and outcome after say three calls when the fourth call is about to start, assuming the timing of the call the interviewer would have chosen. In total, 19 models for each response variable were fitted to the final analysis sample (57 models overall). Only selected models will be discussed further in this paper. It should be noted

that ideally survey researchers would want to control for household size in the models. Here, this was not directly possible since this variable is only observed for responding households. Although there is no direct measure of household size from the interviewer observation variables in Understanding Society, a number of proxy measures are used, such as type of accommodation, floor level and presence of children.

Comparison of model performance

To compare the different models and to assess the quality of model predictions and model fits, we employ the R² statistic and classification tables. The R² coefficient focuses on prediction and is a measure of the goodness-of-fit. It represents the proportion of variation in the dependent variable that is explained by the model, ranging from 0 to 1. The closer it is to 1, the greater the proportion of variation explained by the model. The better the fit of the model, the more confidence it provides that the regression-based predictions are accurate. A number of pseudo-R² measures exist for logistic regression models, e.g. Cox and Snell, McFadden and Nagelkerke R². Different pseudo-R² coefficients can produce different values. The Cox and Snell and the McFadden R² shows the improvement of the model over the null model. The main limitation of these coefficients is that they do not reach the maximum value of 1 which makes the interpretation of the coefficient more difficult (Field 2009; Long and Freese 2006). The Nagelkerke R² coefficient adjusts the Cox and Snell R² so that the range of possible values extends to 1 for ease of interpretation (Field 2009). Here, the Nagelkerke R² coefficient is employed.

Although the R² statistic is widely used to assess the prediction performance of nonresponse models (Groves and Couper 1998; Bates, Dahlhamer, and Singer 2008), it is designed to assess the overall fit of the model and therefore does not distinguish between the accuracy of the model for respondents and non-respondents separately (Plewis, Ketende, and Calderwood 2012). To achieve this we borrow ideas from concepts widely used in epidemiology. Accuracy may be determined by two related concepts: discrimination and prediction (Pepe 2003). Discrimination refers to the conditional probability that a household is predicted to be a respondent (non-respondent) given that a household is indeed a respondent (non-respondent). Prediction refers to the conditional probability of being a respondent (non-respondent) given a household is predicted to be a respondent (non-respondent) (Plewis, Ketende, and Calderwood 2012). To evaluate both concepts classification tables are derived cross-classifying the observed binary response with a prediction of whether $y_i=0$ or $y_i=1$ (Agresti 2013). The prediction for an observation will be obtained depending on a cut-off π_0 . The prediction for observation i is $\hat{y}_i=1$ if $\hat{\pi}_i > \pi_0$, and $\hat{y}_i = 0$ if $\hat{\pi}_i \le \pi_0$, where $\hat{\pi}_i$, denotes the predicted probability from the model. The discrimination power may be summarised by $\textit{sensitivity} = P(\hat{y}_i = 1 | y_i = 1)$ (or true positive fraction) and $\textit{specificity} = P(\hat{y}_i = 0 | y_i = 0)$ (1 minus the false positive fraction) (Agresti 2013; Plewis, Ketende, and Calderwood 2012; Altman 1991). As an overall summary measure of model performance, the percentage of observations correctly classified may be used (i.e. a summary of the diagonal of the classification table), which is a weighted average of sensitivity and specificity:

P(correctly classified) = P(
$$y_i$$
=1 and \hat{y}_i =1)+ P(y_i =0 and \hat{y}_i =0)
$$= P(\hat{y}_i$$
=1| y_i =1) P(y_i =1)+ P(\hat{y}_i =0| y_i =0) P(y_i =0).

The concept of prediction is useful, since it tackles the problem from the other direction. In practice, survey researchers would not know who will be a respondent or a non-respondent before the end of data collection. This means we are interested in the probability of the prediction being correct whether the household is in reality a respondent or a non-respondent. These are measured by the positive predictive value, $P(y_i=1|\hat{y}_i=1)$, and the negative predictive *value*, $P(y_i=0|\hat{y}_i=0)$. The concepts of classification table, discrimination and prediction can be extended to the multinomial case, allowing for several groups of misclassifications. For a 4 category variable this results in a 4×4 classification table, allowing for 4 correctly and 12 incorrectly classified groups. In this analysis we are particularly interested in predicting nonresponse and in discriminating between respondents and non-respondents. In the results section we therefore refer to sensitivity and positive predicted values with respect to modelling long and unsuccessful call sequences. In our analysis we use the default option of π_0 =0.50 for the binary and π_0 =0.25 for the multinomial case, although in practice different values can be explored to optimise the balance between discrimination and prediction (Plewis, Ketende, and Calderwood 2012). Since both the R² statistics and the classification tables allows the comparison of different models for the same data, the final analysis sample was restricted to households with 4 or more calls, although in survey practice the models can be fitted to either all cases or restricted samples, such as all households with at least 1, 2, 3 calls etc.

4. Results

First, a range of models with increasing amount of explanatory variables are fitted to each of the three dependent variables. Table 2 presents pseudo-R² coefficients and the results from the classification tables of the percentage of correctly classified households for eight selected

models. To be able to interpret the results from the classification tables correctly let us consider a random allocation of households into different categories of the response variables. For a binary dependent variable we would expect 50% of cases to be correctly allocated to one of the two groups at random (e.g. at least one interview/no interview; short/long call sequence). For the multinomial case with four categories we would expect this percentage to be 25%. So to do better than random allocation, we aim to find classification values above 50% for the binary and above 25% for the multinomial case. Table 2 suggests that the predictive power of models only including geographic location and design variables (Model 1) is very low with pseudo-R² coefficients all below 3% and only about 55% of cases being correctly classified for both binary models and 36% for the multinomial model. The introduction of interviewer observations to the models improves the predictability of the models in relative terms. The R² value doubles for length, triples for the multinomial model and quadruples for the final outcome variable. However, in absolute terms it is still below 9% across all three models. The classification of cases improves by 2.8-3.6 percentage points. Despite the relatively low improvement in prediction performance, it should be noted that the majority of interviewer observation variables are found to be highly significant across all models, stressing nevertheless the importance of such variables. The introduction of the call record information in the form of timings of calls slightly improves the models' predictions (Models 3 and 4). The R² is now around 10% for the binary and 18% for the multinomial model with the classification tables around 61% and 43% respectively. Including information about call outcomes (Models 5-7) substantially improves the models' prediction. The R² values increase to around 25% for the binary case and to even 37% for the multinomial case. The classification table has increased to 70% for the binary and 52% for the multinomial case, reflecting an improvement in the percentage of correctly allocated

cases by 40% and by above 100% in comparison to a random allocation respectively. Comparing the three models including outcome(s) of previous calls, Model 6, which includes raw outcomes for all three calls, performs better with regards to prediction than one outcome on its own (Model 5) or summaries of separate call outcomes (Model 7). As one may expect, the outcome of the last call (here call 3) is the key variable when predicting final outcome and length.

Table 2: Comparison of different models for length, (final) outcome and the combined dependent variable of length and outcome (Nagelkerke R² and classification table, i.e. the percentage correctly classified households).

	Model	Le	ngth	Partial	Outcome	ne Combined Outcome		
		Nagelkerke R ²	Classification table	Nagelkerke R ²	Classification table	Nagelkerke R ²	Classification table	
1	Just geographic	0.027	55.5%	0.013	54.6%	0.029	36.2%	
2	1+interviewer observations	0.062	58.7%	0.053	58.2%	0.085	39.0%	
3	2+call record for call 1 including call outcome	0.078	59.9%	0.065	59.4%	0.115	40.3%	
4	2+call record for calls 1-3 without call outcomes	0.110	61.1%	0.105	61.2%	0.185	42.9%	
5	4+call outcome for call 3	0.248	69.1%	0.236	67.4%	0.360	51.2%	
6	4+call outcomes for calls 1-3	0.258	69.7%	0.242	67.8%	0.371	51.7%	
7	4+4 sums of call outcomes across the calls 1-3	0.224	68.4%	0.209	66.9%	0.332	50.4%	
8	6+call record for call 4 without call outcome for call 4	0.270	70.7%	0.264	68.5%	0.400	52.9%	

Out of curiosity, we also fit a model that controls for the day and time of call 4 chosen by the interviewer for visiting a household, pretending call 4 has not yet happened. Although this further improves the prediction performance of the models, the level of improvement is not substantial when compared to Model 6 (changes in R² vary between 1-3% and classification improves by a maximum of 1.2%). Table 2 suggests that the best models for prediction is obtained in Model 6 which contains variables from the geographic and design group, interviewer observations and call record data including raw call outcomes for the first three calls and timing of calls. Also, modelling outcome and length of call sequence jointly improves the fit of the model in comparison to the two separate models for either length or final outcome based on the R² value and the classification table.

Table 3: Results of the classification table showing the percentage of correctly classified households by categories of the multinomial dependent variable (combined length and outcome) for each of the 8 modelling stages considered. (Column percentages shown, i.e. percentage of those households which were estimated correctly out of the total observed in the group).

Model	Short Unsuccessful (n=4962)	Short Successful (n=7391)	Long Unsuccessful (n=8603)	Long Successful (n=4402)
1	0.0%	43.2%	69.6%	0.0%
2	6.5%	52.8%	65.9%	0.1%
3	20.4%	49.8%	64.2%	0.1%
4	31.1%	51.6%	63.9%	0.4%
5	44.3%	50.2%	79.5%	5.2%
6	45.1%	51.0%	79.5%	5.6%
7	42.2%	54.4%	75.3%	3.9%
8	50.7%	52.5%	78.2%	6.8%

In the non-response literature researchers are interested in predicting the non-respondents correctly. More specifically for this application, we are interested in predicting households with no interview and long call sequences early on in the data collection process. Table 3 shows the percentage of correctly classified households by the categories of the multinomial dependent

variable for each of the 8 modelling stages considered. Across all stages the percentage of correctly classified households with long unsuccessful call sequences (around 64-80%) is quite high. For Model 6 this value increases to 80%, meaning that of the cases that have long unsuccessful call sequences 80% are correctly classified by the model as being indeed in this category.

Table 4: Complete classification table for the multinomial model (dependent variable is combined length and outcome) for Model 6: (A) column percentages (percentages predicted out of the total observed in the category) reflecting sensitivity of modelling long unsuccessful calls and (B) row percentages (percentages of households observed in the group out of the total predicted in the category) reflecting positive predictive values.

				Obse	erved	
		·	Short Unsuccessful (n=4962)	Short Successful (n=7391)	Long Unsuccessful (n=8603)	Long Successful (n=4402)
A	Predicted	Short Unsuccessful	45.1%	6.8%	9.8%	5.2%
		Short Successful	13.4%	51.0%	8.8%	16.8%
		Long Unsuccessful	40.6%	39.8%	79.5%	72.4%
		Long Successful	0.9%	2.4%	1.9%	5.6%
В	Predicted	Short Unsuccessful	58.7%	13.1%	22.1%	6.1%
		Short Successful	11.2%	63.6%	12.8%	12.4%
		Long Unsuccessful	13.4%	19.6%	45.7%	21.3%
		Long Successful	7.4%	28.3%	25.2%	39.1%

Table 4 breaks this down further for the multinomial model for Model 6, showing marginal summaries of the classification table. The upper panel (A) indicates sensitivity (for a model for long unsuccessful calls) and the lower panel (B) the positive predictive values. For example, we can see that of the cases that are predicted to have long unsuccessful call sequences (panel B) indeed 46% belong correctly to this category, 13.4% and 19.6% would be wrongly

classified and belong in reality to the short successful and short unsuccessful groups respectively.

Misclassification to short sequences would not have in practice negative implications since six calls might be made anyway. Of the cases predicted to have long unsuccessful call sequences 21% turn out to have long successful calls.

Let us now turn to the discussion of the effects of different explanatory variables on the three outcomes. A number of variables are found to be consistently significant across all models. For example, for length these are months of household issue, urban/rural, region, low density ethnic minority area, some interviewer observation variables (accommodation, floor, car/van, child, unkempt garden, relative condition of property) and some call record variables (time of day, time between calls and call outcome variables). Consistently significant predictors for final outcome are region, interviewer observation variables (floor, car/van, child, relative condition of property), time of day (call 1), time between calls and, as might be expected, all call outcome variables, but, unlike in the models for length, they exclude urban/rural, unkempt garden and low density area. Consistently significant predictors for the combined response models are similar to the model for length alone including geographic and design variables (months, urban/rural, region), interviewer observation variables (locked gates, accommodation, floor, car/van, child, unkempt garden, relative condition of property) and call record variables (all time of day variables, time between calls and call outcome variables). Interestingly, day of the week is not significant in most models.

The results of the final models for all three response variables from Model 6 are discussed below. Table 5 presents estimated regression coefficients together with odds ratios from the two binary logistic models (Model 6). The results suggest that the odds of having a short call sequence are higher in the South West (1.267 times) and Scotland (1.199 times) and

reduced in London (1.320 times) when compared to the North East. They are also higher in low density areas for ethnic minorities (1.090 times), in rural areas (1.128 times), for properties with 2 or less floors (1.216-2.247 times), for households which definitely do not have a car (1.947 times), are unlikely to have children or do not have children (1.193 and 1.232 times) and if there is no unkempt garden (1.225 times). The odds are lower in terraced houses (1.230 times) and flats (1.344 times) when compared to detached houses, also lower (1.107 times) when properties are worse than others in the neighbourhood, and lower when the calls are made in the evening compared to morning calls. (It should be noted that we cannot interpret this as a causal effect since calling times are not allocated randomly but are merely determined by the interviewer). The model suggests a positive association for time between calls: the longer the time between calls, the higher the probability of a short sequence. Shorter sequences are also more likely when previous call outcomes are a contact, an appointment, any other status or an interview when compared to non-contact. There is a marked monotone increase in the effect of the call outcome variables across the three calls, indicating that although the outcome of each call is significant, it may be the most recent call that has the highest influence, rather than, for example, the first call. An appointment or interview at the third call increases the odds of a short sequence by between 7 to 10 times compared to having a non-contact at this call attempt.

Table 5: Estimated coefficients for the two logistic regression models for length and (final) outcome including geographic and design variables, interviewer observation variables and call record variables comprising timing and outcome of the call(s) (Model 6).

		Model for L	ength	Model fo		outcome
		Robust			Robust	
Variable	В	SE	OR	β	SE	OR
Constant	-1.087	0.149		0.363	0.149	
Geographic and design variables						
Months						
January-June year 1 (ref)	0.000		1.000	0.000		1.000
July-December year 1	-0.099	0.040	0.905*	0.154	0.040	1.167***
January-June year 2	-0.213	0.039	0.808***	-0.001	0.039	0.999
July-December year 2	-0.319	0.039	0.726***	0.016	0.039	1.017
Low density area for ethnic						
minorities No (ref)	0.000		1 000			
Yes	0.000	0.026	1.000			
Government Office Region (GOR)	0.086	0.036	1.090*			
North East (ref)	0.000		1.000	0.000		1.000
North West	0.000	0.079	1.000	-0.164	0.077	0.848*
Yorkshire and the Humber	0.091	0.079	1.093	-0.164 -0.168	0.077	0.845*
East Midlands	0.130	0.082	1.050	0.072	0.080	1.075
West Midlands	0.048	0.083	1.088	-0.244	0.083	0.783**
East of England						0.783***
London	0.016 -0.282	0.077 0.082	1.016 0.754**	-0.212 -0.449	0.078 0.075	0.638***
South East	0.061	0.082	1.063	-0.449	0.073	0.038***
South East South West	0.001	0.077	1.267**	-0.230	0.074	0.779**
Wales	0.257	0.083	1.169	-0.164 -0.151	0.081	0.859
Scotland	0.130	0.093	1.109	-0.131	0.092	0.839
Urban/rural	0.162	0.063	1.199	-0.554	0.080	0.710
Urban area (ref)	0.000		1.000			
Rural area	0.000	0.040	1.128**			
Interviewer observations variables	0.121	0.040	1.120			
Accommodation						
Detached house/bungalow (ref)	0.000		1.000			
Semi-detached house/bungalow	-0.062	0.042	0.940			
Terraced house/end terraced	-0.002	0.042	0.540			
house/bungalow	-0.207	0.046	0.813***			
Flats/maisonettes purpose built or converted	-0.296	0.061	0.744***			
Bedsitters, dwellings with business and sheltered accommodation	0.257	0.184	1.293			
Floor	0.237	0.104	1.493			
0 floors	0.809	0.236	2.247**	0.806	0.213	2.238***

1 floor	0.318	0.078	1.374***	0.430	0.073	1.538***
2 floors	0.196	0.075	1.216**	0.402	0.065	1.495***
3 floors	0.051	0.079	1.052	0.413	0.077	1.511***
4 floors and above (ref)	0.000		1.000	0.000		1.000
Car/van						
Definitely has a car/van (ref)	0.000		1.000	0.000		1.000
Likely	-0.039	0.041	0.962	-0.189	0.039	0.827***
Unlikely	0.153	0.078	1.165	-0.452	0.075	0.636***
Definitely does not have a car/van	0.666	0.095	1.947***	0.654	0.095	1.924***
Cannot tell from observation	-0.137	0.041	0.872**	-0.400	0.036	0.670***
Child						
Definitely has a child/children aged						
under 10 (ref)	0.000		1.000	0.000		1.000
Likely	0.013	0.075	1.013	-0.099	0.075	0.906
Unlikely	0.177	0.068	1.193**	-0.272	0.065	0.762***
Definitely does not have a						
child/children aged under 10	0.209	0.068	1.232**	-0.163	0.066	0.849*
Cannot tell from observation	0.058	0.062	1.060	-0.312	0.061	0.732***
Unkempt garden						
Yes (ref)	0.000		1.000			
No	0.203	0.056	1.225***			
No obvious garden	0.113	0.062	1.120			
Relative conditions of the address to other residential properties						
Better (ref)	0.000		1.000	0.000		1.000
About the same	-0.047	0.054	0.954	-0.242	0.053	0.785***
Worse	-0.202	0.078	0.903*	-0.440	0.074	0.644***
Unable to obtain information	0.008	0.239	1.008	-1.248	0.260	0.287***
Call Record Variables						
Time of day call 1						
Morning (0.00-12.00) (ref)				0.000		1.000
Afternoon (12.00-17.00)				0.010	0.036	1.011
Evening (17.00-24.00)				-0.130	0.050	0.878**
Time of day call 2						
Morning (0.00-12.00) (ref)	0.000		1.000			
Afternoon (12.00-17.00)	-0.064	0.038	0.937			
Evening (17.00-24.00)	-0.102	0.043	0.903*			
Time of day call 3	0.102	0.0.0	0.502			
Morning (0.00-12.00) (ref)						
Afternoon (12.00-17.00)	-0.015	0.039	0.985			
Evening (17.00-24.00)	-0.094	0.033	0.912*			
Time between call 1 and call 2	0.026	0.002	1.026***	-0.026	0.002	0.974***
Time between call 2 and call 3	0.030	0.002	1.030***	-0.028	0.002	0.973***
Call 1 outcome	0.000	0.001	1.000	0.020	5.55 2	5.7.5

	No contact (ref) Contact made	0.000 0.081	0.037	1.000 1.085*	0.000 -0.098	0.037	1.000 0.907**
	Appointment made	0.025	0.069	1.026	0.242	0.065	1.274***
	Any other status	0.124	0.095	1.133	-0.035	0.096	0.966
	Interview done	1.022	0.240	2.780***	0.372	0.207	1.450
Call 2 outcome							
	No contact (ref)	0.000		1.000	0.000		1.000
	Contact made	0.095	0.038	1.099*	-0.062	0.037	0.940
	Appointment made						
	Appointment made	0.205	0.063	1.228**	0.458	0.062	1.581***
	Any other status	0.293	0.088	1.340**	-0.024	0.092	0.977
	Interview done	1.539	0.143	4.662***	0.517	0.116	1.677***
Call 3 outcome							
	No contact (ref)	0.000		1.000	0.000		1.000
	Contact made	0.499	0.037	1.645***	-0.148	0.037	0.862***
	Appointment made						
	Appointment made	2.000	0.048	7.389***	2.024	0.053	7.568***
	Any other status	1.227	0.065	3.410***	-1.185	0.074	0.305***
	Interview done	2.352	0.110	10.511***	0.860	0.087	2,364***

^{***}p<0.001; **p<0.01; *p<0.05; ref – reference category

In the model for final outcome some of the same variables are significant as in the model for length, however, sometimes with the opposite effect. For example, the odds of having a successful call sequence are lower across all regions, in particular for London and the South-East, when compared to North East. Households with no children are significantly less likely to respond and a longer time in between calls is associated with a reduced likelihood of response. A contact in any of the previous calls is associated with a higher likelihood of non-response as is any other status. Appointment made and any interviewing performed increases the likelihood of a successful call sequence as one would expect. The effect of number of floors and the presence of cars/vans are in the same direction as for length of call sequence.

Table 6 presents estimated coefficients for the multinomial model. To ease interpretation predicted probabilities are computed for selected explanatory variables (Figure 1) holding

constant all other variables at their mean value. Variables from each of the three groups of explanatory variables are chosen: geographic (region), interviewer observations (accommodation type) and call record variables (time of day of first call, third call outcome). Figure 1a suggests that the probability of having long unsuccessful sequences is the highest in London followed by Scotland and East of England. It is also highest in flats compared to other types of accommodation with short successful calls highest in detached houses (Figure 1b). Figure 1c suggests that conducting the first call in the evening is associated with a higher probability of a long unsuccessful call sequence. (Researchers need to be cautious in interpreting this as a causal effect since interviewers may have reasons for calling in the evening, even at the first call). Figure 1d indicates a marked difference between call outcomes, with the likelihood for short successful call sequences clearly highest if the third call is an appointment or an interview. Long unsuccessful call sequences are associated with non-contacts and to a slightly lesser extent with contact. The likelihood of having a short unsuccessful sequence is clearly highest when the third call attempt is recorded as being 'any other status'.

Table 6: Estimated coefficients for the multinomial regression model for the combined dependent variable (length and final outcome) including geographic and design variables, interviewer observation variables and call record variables comprising timing and outcome of the call(s) (Model 6).

	Short Unsuccessful Short Successful		Successful	Long S	Successful	
Variable	β	Robust SE	β	Robust SE	β	Robust SE
Constant	-1.832***	0.216	-0.396*	0.192	-0.018	0.198
Geographic and						
design variables						
Months						
January-June year 1	0.000		0.000		0.000	
(ref)						
July-December	-0.316***	0.058	0.060	0.051	0.032	0.055
year 1						
January-June year 2	-0.354***	0.055	-0.157**	0.051	-0.071	0.054
July-December	-0.424***	0.055	-0.237***	0.051	-0.001	0.053
year 2						
Government						
Office Region						
(GOR)						
North East (ref)	0.000		0.000		0.000	
North West	0.020	0.112	-0.049	0.101	-0.259*	0.101
Yorkshire and the	0.168	0.116	-0.029	0.105	-0.212	0.107
Humber			****		**	
East Midlands	-0.040	0.122	0.105	0.109	0.047	0.110
West Midlands	0.065	0.114	-0.137	0.104	-0.305**	0.104
East of England	-0.130	0.114	-0.149	0.102	-0.358**	0.104
London	-0.323**	0.111	-0.591***	0.100	-0.439***	0.098
South East	0.123	0.108	-0.168	0.098	-0.245*	0.098
South West	0.195	0.118	0.046	0.106	-0.337**	0.110
Wales	0.271*	0.133	-0.007	0.120	-0.125	0.122
Scotland	0.389**	0.117	-0.143	0.120	-0.280**	0.107
Urban/rural	0.307	0.117	0.143	0.107	0.200	0.107
Urban area (ref)	0.000		0.000		0.000	
Rural area	0.000	0.054	0.136**	0.050	0.000	0.057
Interviewer	0.165	0.054	0.130	0.050	0.013	0.037
observations						
variables						
Barrier 2 (locked						
gates)						
Not mentioned (ref)	0.000		0.000		0.000	
Mentioned (1e1)	-0.025	0.151	-0.515***	0.147	-0.147	0.141
Accommodation	-0.023	0.131	-0.515	0.147	-0.147	0.141
Detached	0.000		0.000		0.000	
house/bungalow	0.000		0.000		0.000	
•						
(ref)	0.072	0.050	0.072	0.054	0.011	0.060
Semi-detached house/bungalow	-0.073	0.059	-0.072	0.054	-0.011	0.060
	0.220***	0.066	0.101	0.050	0.000	0.064
Terraced house/end	-0.320***	0.066	-0.101	0.059	0.089	0.064
terraced						
house/bungalow	0.604***	0.006	0.147	0.077	0.111	0.002
Flats/maisonettes	-0.604***	0.086	-0.147	0.077	-0.111	0.082
purpose built or						

converted						
Bedsitters,	0.363	0.242	0.268	0.237	0.163	0.266
dwellings with						
business and						
sheltered						
accommodation						
Floor						
0 floors	0.997**	0.342	1.343***	0.317	1.084**	0.350
1 floor	0.248*	0.111	0.524***	0.101	0.366***	0.104
2 floors	0.135	0.107	0.383***	0.097	0.349***	0.098
3 floors	-0.092	0.114	0.311**	0.102	0.394***	0.100
4 floors and above	0.000		0.000		0.000	
(ref)						
Car/van						
Definitely has a	0.000		0.000		0.000	
car/van (ref)						
Likely	0.067	0.058	-0.206***	0.052	-0.147**	0.056
Unlikely	0.329**	0.105	-0.271**	0.101	-0.500***	0.110
Definitely does not	0.838***	0.157	1.101***	0.133	-0.746***	0.138
have a car/van						
Cannot tell from	0.043	0.058	-0.468***	0.053	0.328***	0.055
observation						
Child						
Definitely has a	0.000		0.000		0.000	
child/children aged						
under 10 (ref)						
Likely	-0.056	0.115	-0.055	0.096	-0.162	0.098
Unlikely	0.237*	0.099	-0.068	0.084	-0.330***	0.086
Definitely does not	0.314**	0.102	0.029	0.086	-0.164	0.087
have a			****			
child/children aged						
under 10						
Cannot tell from	0.095	0.079	-0.185*	0.079	-0.370***	0.080
observation						
Unkempt garden						
Yes (ref)	0.000		0.000		0.000	
No	0.311***	0.077	0.200**	0.072	0.094	0.071
No obvious garden	0.212*	0.085	0.071	0.079	0.047	0.077
Relative	0.212	0.000	0.071	0.075	0.0.7	0.077
conditions of the						
address to other						
residential						
properties						
Better (ref)	0.000		0.000		0.000	
About the same	0.093	0.079	-0.255***	0.067	-0.162*	0.073
Worse	0.014	0.110	-0.525***	0.101	-0.288**	0.105
Unable to obtain	0.337	0.313	-0.872**	0.311	-1.540**	0.480
information						
Call Record						
Variables						
Time of day call 1						
Morning (0.00-	0.000		0.000		0.000	
12.00) (ref)						
Afternoon (12.00-	0.018	0.052	0.034	0.047	-0.021	0.049
17.00)					- · · · -	
Evening (17.00-	-0.081	0.071	-0.180**	0.065	-0.119	0.067
=:					/	

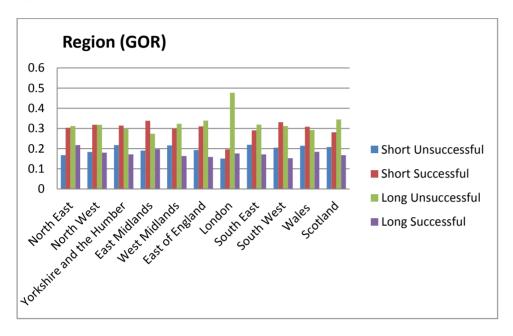
24.00)						
Time of day call 2						
Morning (0.00-	0.000		0.000		0.000	
12.00) (ref)						
Afternoon (12.00-	-0.023	0.053	-0.012	0.048	0.141**	0.052
17.00)						
Evening (17.00-	-0.126*	0.059	-0.072	0.054	0.011	0.059
24.00)						
Time of day call 3						
Morning (0.00-	0.000		0.000		0.000	
12.00) (ref)						
Afternoon (12.00-	-0.025	0.054	0.028	0.050	0.058	0.053
17.00)						
Evening (17.00-	-0.127*	0.057	-0.083	0.052	-0.034	0.055
24.00)						
Time between call	0.034***	0.002	-0.002	0.002	-0.024***	0.003
1 and call 2						
Time between call	0.038***	0.002	-0.001	0.002	-0.028***	0.003
2 and call 3						
Call 1 outcome						
No contact (ref)	0.000		0.000		0.000	
Contact made	0.279***	0.050	-0.038	0.048	0.035	0.052
Appointment made	0.124	0.094	0.219*	0.088	0.337***	0.087
Any other status	0.292*	0.118	0.055	0.130	0.138	0.145
Interview done	1.510***	0.335	1.320***	0.327	0.821*	0.350
Call 2 outcome						
No contact (ref)	0.000		0.000		0.000	
Contact made	0.198***	0.052	0.012	0.049	-0.024	0.052
Appointment made	0.032	0.089	0.516***	0.081	0.316***	0.086
Any other status	0.436***	0.107	0.202	0.124	0.228	0.147
Interview done	2.056	0.221	2.043***	0.210	1.012***	0.238
Call 3 outcome						
No contact (ref)	0.000		0.000		0.000	
Contact made	0.618***	0.050	0.273***	0.049	-0.203***	0.052
Appointment made	0.187	0.102	2.683***	0.063	0.568***	0.080
Any other status	1.780***	0.077	-0.125	0.102	-0.661***	0.119
Interview done	2.344***	0.169	2.704***	0.157	0.556**	0.195
	. ** <0.01. *					

***p<0.001; **p<0.01; *p<0.05; ref – reference category.

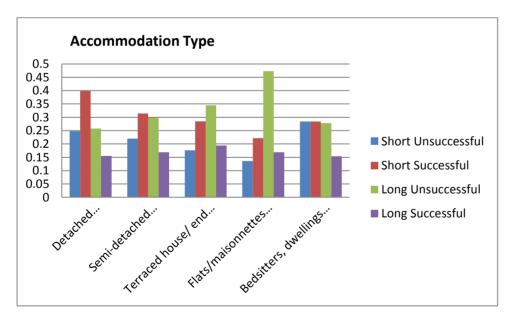
Reference category for the response variable is Long Unsuccessful.

Figure 1: Predicted probabilities for the final combined model (final outcome and length) for selected explanatory variables (at least one variable from each group of explanatory variables in Model 6)

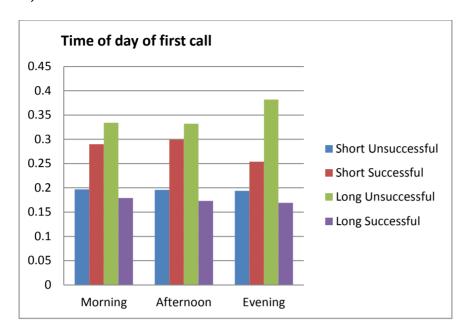
a.)



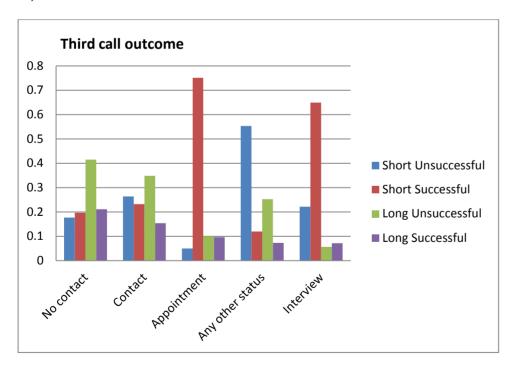
b.)



c.)



d.)



5. Conclusions and implications for survey practice

This paper presents an example of how to use paradata for interviewer-administered surveys, a topic often discussed in the literature. The paper employs a range of interviewer observation and call record variables to predict both final outcome and length of a call sequence early on in the data collection process. Models are developed prior to data collection and after the first, second and third call respectively to see if predictions of the models improve once more and more call record data are available. The models are developed both separately and jointly for the two outcomes of interest. The research was motivated by a recent sequence analysis (Durrant et al. 2014) which identified a categorisation of sequences based on length and outcome. Survey researchers have focussed on predicting outcome, either final outcome or outcome at each call (Groves and Couper 1998; Durrant, D'Arrigo, and Steele 2013). Here, this is extended to also include length of call sequence, a variable that so far has only been considered as an explanatory variable in non-response models.

The key findings from this analysis are:

- Modelling outcome and length of call sequence jointly improves the fit of the model in comparison to the two separate models based on the pseudo-R² values and the classification tables.
- 2. The models proposed have the ability to predict the outcomes of interest reasonably well (length, final call outcome and the combined model) with a pseudo-R² of around 26% for the binary cases and 36% for the combined outcome. This is very high in a social science context and compares to values of around 3-8% for non-response models in the literature (Olson, Smyth, and Wood 2012; Olson and Groves 2012; West and Groves 2013). Classification tables show about 70% correctly classified cases for the binary cases

- (expected baseline is 50%) and even 52% for the multinomial case (expected baseline is 25%).
- 3. A number of variables are significant for both the model for length and the model for outcome. However, their effect may have opposite signs. For example, a household unlikely to have children has a higher probability of a shorter call sequence but a lower probability of response; a longer time in between calls may be associated with a shorter call sequence but with a lower probability of response.
- 4. Call record variables: The outcomes of previous calls, in particular the most recent call, are highly significant for both final response outcome and sequence length and their inclusion greatly improves the predictive power, with the R² increasing from less than 9% to above 26%. As one might expect, the prediction improves significantly when more and more call outcomes are available. We found evidence that it is the most recent call outcome that may have the biggest influence, rather than the first call, for example. We found some indication that entering the raw outcome variables from each call (i.e. outcome of call 1, 2 and 3) as opposed to entering a summary measure, such as total number of non-contacts, improves the fit slightly. Often discussed in the literature as important variables, controlling for timing of calls has less impact on the performance of the model (R²), although some of these variables are significant predictors throughout. Time between calls was often highly significant. Time of day was either not significant or only marginally significant. Interestingly, the day of the week was not found to be significant in the majority of models.
- 5. *Interviewer observation variables:* A number of interviewer observation variables (such as floor, car/van, child, relative condition of property) are found to be significant in all models. However, their contribution to the improvement of the model prediction is limited having

controlled for previous call outcome(s). Adding interviewer observation variables to a model with, for example, only basic geographical information doubles or even quadruples the R^2 value.

6. *Basic geographic information*, available prior to the first call, are found to be significant variables for most models but are not found very effective in predicting final outcome and length of call sequence.

The findings highlight benefits and implications for survey practice, especially for adaptive and responsive survey designs, monitoring continually the streams of process data to alter the survey design during the course of data collection (Groves and Heeringa 2006). The paper proposes a methodology that can be used and adapted by survey managers of other datasets, for both cross-sectional and longitudinal surveys. First, we strongly encourage researchers to use sensitivity and the positive predicted value rather than just the R² statistic to assess non-response models. The predictions from the two separate or the combined model provide *one* way of informing survey managers which households to follow up or where to stop calling. In this paper the focus is on efficiency to avoid very long call sequences. This guidance should, in practice, be supplemented with information about the impact of cases on nonresponse bias during data collection and further work in this area is currently being undertaken. Survey managers may wish to weigh up between the probability of a successful outcome versus sequence length. Ideally survey agencies may want to follow up those with potentially relatively short sequences and successful outcomes. Cases that are predicted to have long call sequences with many interviewer visits to a household may also be contacted by other means first (e.g. telephone) to avoid successive non-contacts at the doorstep. Often call record and interviewer observation

variables can be supplemented further by additional linked datasources, such as from administrative or census data. Here, we only control for linked basic geographic information. Particular benefits of this work exist for data from longitudinal surveys since here the models can be, in principle, fitted for prediction in subsequent waves. Models may then be extended to include call record data and survey information from previous waves. Both of these further work strands are currently being explored.

The analysis indirectly assumes that all calls carry the same costs, e.g. staff and time resources and financial cost of a call. However, in practice some calls may be relatively inexpensive and may imply little extra burden, if, for example, a call can be made on the way to another household. The analysis could be extended to include such cost considerations. However, given the data we have available to us, we are unable to carry this out. The current work assumes a cut-off value π_0 to identify a household as a respondent or non-respondent based on their predicted probabilities. In practice, different values may be explored to optimise sensitivity and specificity in the specific survey setting considered. Other definitions of the dependent variables may also be considered (e.g. complete household response rather than 'at least one interview'). Due to the observational nature of the data we cannot establish causal links but merely associations between the response and the explanatory variables. However, this restriction is not a limitation as the main interest of the analysis is not to establish causal links but to compare performance of different models and to identify significant explanatory variables.

6. References

Agresti, A. (2013), Categorical Data Analysis (3rd edition), New Jersey: John Wiley & Sons. Altman, D. G. (1991), Practical Statistics for Medical Research, London: Champan & Hall.

- Bates, N., Dahlhamer, J., and Singer, E. (2008), "Privacy Concerns, Too Busy, or Just Not Interested: Using Doorstep Concerns to Predict Survey Nonresponse," *Journal of Official Statistics*, 24, 591–612.
- Durrant, G. B., D'Arrigo, J., and Müller, G. (2013), "Modeling Call Record Data: Examples from Cross-Sectional and Longitudinal Surveys," in *Improving Surveys with Paradata:*Analytic Use of Process Information, ed. Kreuter, F., New Jersey: John Wiley and Sons, pp. 281-308.
- Durrant, G.B., D'Arrigo, J., and Steele, F. (2011), "Using Field Process Data to Predict Best Times of Contact Conditioning on Household and Interviewer Influences," *Journal of the Royal Statistical Society, Series A*, 174, 4, 1029-1049.
- Durrant, G. B., D'Arrigo, J., and Steele, F. (2013), "Analysing Interviewer Call Record Data by Using a Multilevel Discrete-Time Event History Modelling Approach," *Journal of the Royal Statistical Society, Series A, Special issue: The Use of Paradata in Social Survey Research*, 176, 1, 251-269.
- Durrant, G. B., Maslovskaya, O., and Smith, P. W. F. (2014), "Sequence Analysis as a Tool for Investigating Call Record Data," Working paper, University of Southampton.
- Field, A. (2009), *Discovering statistics using SPSS* (3rd edition), Los Angeles: SAGE.
- Groves, R. M., and Couper, M. P. (1996), "Contact-level Influences on Cooperation in Face-to-face Surveys," *Journal of Official Statistics*, 12, 63–83.
- Groves, R. M., and Couper, M. P. (1998), *Nonresponse in Household Interview Surveys*, New York: John Wiley and Sons.

- Groves, R. M., and Heeringa, S. G. (2006), "Responsive Design for Household Surveys: Tools for Actively Controlling Survey Errors and Costs," *Journal of the Royal Statistical Society, Series A*, 169, 439–457.
- Hanly, M. (2014), "Improving Nonresponse Bias Adjustments with Call Record Data," Paper to be presented at 25th International Workshop on Household Survey Nonresponse, Iceland.
- Huber, P. J. (1967), "The Behavior of Maximum Likelihood Estimates under Nonstandard Conditions," in *Proceedings of the Fifth Berkeley Symposium on Mathematical Statistics and Probability*, Berkeley, CA: University of California Press, vol. 1, pp. 221–233.
- Kreuter, F., and Kohler, U. (2009), "Analyzing Contact Sequences in Call Record Data. Potential and Limitations of Sequence Indicators for Nonresponse Adjustments in the European Social Survey," *Journal of Official Statistics*, 25, 2, 203-226.
- Kreuter, F., Olson, K., Wagner, J., Yan, T., Ezzati-Rice, T. M., Casas-Cordero, C., Lemay, M.,
 Peytchev, A., Groves, R. M., and Raghunathan, T. E. (2010), "Using Proxy Measures and
 Other Correlates of Survey Outcomes to Adjust for Non-Response: Examples from
 Multiple Surveys," *Journal of the Royal Statistical Society, Series A*, 173, 389–407.
- Long, J. S., and Freese, J. (2006), Regression Models for Categorical Dependent Variables

 Using Stata (2nd edition), Texas: STATA Press.
- McFall, S. L. (ed.) (2012), *Understanding Society: Findings 2012*, Colchester: Institute for Social and Economic Research, University of Essex.
- Olson, K., and Groves, R. M. (2012), "An Examination of Within-Person Variation in Response Propensity over the Data Collection Field Period," *Journal of Official Statistics*, 28, 29-51.

- Olson, K., Smyth, J. D., and Wood, H. M. (2012), "Does Giving People Their Preferred Survey Mode Actually Increase Survey Participation Rates? An Experimental Examination," *Public Opinion Quarterly*, 76, 611 635.
- Pepe, M.S (2003), *The Statistical Evaluation of Medical Tests for Classification and Prediction*, Oxford: Oxford University Press.
- Plewis, I., Ketende, S., and Calderwood, L. (2012), "Assessing the Accuracy of Response Propensities in Longitudinal Studies," *Survey Methodology*, 38, 2, 167-171.
- Potthoff, R. F., Manton, K. G., and Woodbury, M. A. (1993), "Correcting for Nonavailability Bias in Surveys by Weighting Based on Number of Callbacks," *Journal of the American Statistical Association, Applications and Case Studies*, 88, 424, 1197-1207.
- Sinibaldi, J. (2014), Evaluating the Quality of Interviewer Observed Paradata for Nonresponse Applications, PhD Thesis, München: Ludwig-Maximilian-Universität.
- Sinibaldi, J., Durrant, G. B., and Kreuter, F. (2013), "Evaluating the Measurement Error of Interviewer Observed Paradata," *Public Opinion Quarterly, Special issue: Topics in Survey Measurement and Public Opinion*, 77, 1, 173-193.
- Sinibaldi, J., Trappmann, M., and Kreuter, F. (2014), "Which is the Better Investment for Nonresponse Adjustment: Purchasing Commercial Auxiliary Data or Collecting Interviewer Observations?" *Public Opinion Quarterly* (forthcoming).
- Wagner, J. (2013a), "Adaptive Contact Strategies in Telephone and Face-to-Face Surveys," Survey Research Methods, 7, 1, 45-55.
- Wagner, J. (2013b), "Using Paradata-Driven Models to Improve Contact Rates in Telephone and Face-to-Face Surveys," in *Improving Surveys with Paradata: Analytic Use of Process Information*, ed. Kreuter, F., New Jersey: John Wiley and Sons, pp. 145-170.

- West, B. T., and Groves, R. M. (2013), "A Propensity-Adjusted Interviewer Performance Indicator," *Public Opinion Quarterly*, 77, 352-74.
- White, H. (1980), "A Heteroskedasticity-consistent Covariance Matrix Estimator and a Direct Test for Heteroskedasticity," *Econometrica*, 48, 817–830.
- White, H. (1984), Asymptotic Theory for Econometricians, Orlando, FL: Academic Press.
- White, H. (1994), *Estimation, Inference and Specification Analysis*, New York: Cambridge University Press.

Online Appendix

Table A1: Exact wording of all variables used in the analysis

	Question
Used in the fir	nal models
Months	Month of sample issue
Low density area for ethnic minorities	Low density area for ethnic minorities
Government Office Region (GOR)	Government Office Region of the
	address
Urban/rural	Is the address located in urban or rural
	area?
Barrier 2	Are any of these physical barriers to
	entry present at the address? Locked
	gates
Accommodation	Address dwelling type code
Floor	How many floors are there at the
	address?
Car/van	Based on your observation, is it likely
	that this address has a car or van?
Child	Based on your observation, is it likely
Ciniu	that this address contains one or more
	children aged under 10 including babies?
Unkempt garden	Does the address have an unkempt
onkempt garden	garden?
Relative conditions of the address to other	How is the external condition of the
residential properties	address relative to other residential
residential properties	properties in the area?
Time of day call 1	Time call started
Time of day call 2	Time call started
Time of day call 3	Time call started
Call 1 outcome	Call status
Call 2 outcome	Call status
Call 3 outcome	Call status
Can 3 outcome	Can status
Used for modelling but not s	ignificant in final models
Barrier 1	Are any of these physical barriers to
	entry present at the address? Locked
	common entrance
Barrier 3	Are any of these physical barriers to
	entry present at the address? Security
	staff or gatekeeper
Barrier 4	Are any of these physical barriers to
	entry present at the address? Entry
	phone access
Vicinity 1	Are any of the following present or
, 1011110, 1	The any of the following present of

	within sight or hearing of the address? Boarded houses, abandoned buildings,
	demolished houses or demolished
	buildings
Vicinity 2	Are any of the following present or
	within sight or hearing of the address?
	Trash, litter or junk in street/road
Vicinity 3	Are any of the following present or
	within sight or hearing of the address?
	Heavy traffic on street/road
Conditions of residential property	Which of these best describes the
	condition of residential properties in the
	area?
Start day 1	Day call started
Start day 2	Day call started
Start day 3	Day call started
Start month 1	Month call started
Start month 2	Month call started
Start month 3	Month call started
Start year 1	Year call started
Start year 2	Year call started
Start year 3	Year call started

Table A2: Distribution of explanatory variables by the dependent variables (a.) for length and final outcome and b.) for the combined dependent variable of length and outcome).

a.)

Leng	th	(Final) Outcome	
Short (up to 6 calls)	Long (more than 6 calls)	No interviews after call 3	At least one interview after call 3
3,209 (51.3%)	3,044 (48.7%)	3,253 (52.0%)	3,000 (48.0%)
2,897 (49.6%)	2,946 (50.4%)	2,978 (51.0%)	2,865 (49.0%)
3,095 (47.9%)	3,371 (52.1%)	3,537 (54.7%)	2,929 (45.3%)
3,152 (46.4%)	3,644 (53.6%)	3,797 (55.9%)	2,999 (44.1%)
	(up to 6 calls) 3,209 (51.3%) 2,897 (49.6%) 3,095 (47.9%)	(up to 6 calls) than 6 calls) 3,209 (51.3%) 3,044 (48.7%) 2,897 (49.6%) 2,946 (50.4%) 3,095 (47.9%) 3,371 (52.1%)	(up to 6 calls) than 6 calls) after call 3 3,209 (51.3%) 3,044 (48.7%) 3,253 (52.0%) 2,897 (49.6%) 2,946 (50.4%) 2,978 (51.0%) 3,095 (47.9%) 3,371 (52.1%) 3,537 (54.7%)

minorities	4.510.740.000	C 000 (77 10)	5.005 (55.40)	4.505.445.500
No	4,513 (42.9%)	6,000 (57.1%)	5,927 (56.4%)	4,586 (43.6%)
Yes	7,840 (52.8%)	7,005 (47.2%)	7,638 (51.5%)	7,207 (48.5%)
Government				
Office Region (GOR)				
North East	540 (47.0%)	609 (53.0%)	551 (48.0%)	598 (52.0%)
North West	1,405 (50.2%)	1,395 (49.8%)	1,404 (50.1%)	1,396 (49.9%)
Yorkshire and the	1,119 (53.1%)	987 (46.9%)	1,084 (51.5%)	1,022 (48.5%)
Humber	1,117 (55.170)	707 (1 0.770)	1,004 (31.370)	1,022 (40.570)
East Midlands	952 (52.8%)	850 (47.2%)	836 (46.4%)	966 (53.6%)
West Midlands	1,195 (51.4%)	1,129 (48.6%)	1,253 (53.9%)	1,071 (46.1%)
East of England	1,254 (50.3%)	1,241 (49.7%)	1,326 (53.1%)	1,169 (46.9%)
London	1,279 (34.7%)	2,402 (65.3%)	2,312 (62.8%)	1,369 (37.2%)
South East	1,839 (51.0%)	1,770 (49.0%)	1,943 (53.8%)	1,666 (46.2%)
South West	1,062 (53.6%)	919 (46.4%)	1,025 (51.7%)	956 (48.3%)
Wales	594 (52.4%)	540 (47.6%)	574 (50.6%)	560 (49.4%)
Scotland	1,114 (48.9%)	1,163 (51.1%)	1,257 (55.2%)	1,020 (44.8%)
Urban/rural				
Urban area	9,747 (46.9%)	11,045 (53.1%)	11,255 (54.1%)	9,537 (45.9%)
Rural area	2,606 (57.1%)	1,960 (42.9%)	2,310 (50.6%)	2,256 (49.4%)
Interviewer				
observations				
Accommodation	2.016 (50.50)	0.104 (41.00()	2 500 (50 50)	0.540 (40.50()
Detached	3,016 (58.7%)	2,124 (41.3%)	2,598 (50.5%)	2,542 (49.5%)
house/bungalow Semi-detached	2.021 (52.20()	2 420 (46 70)	2 902 (51 70/)	2 5 40 (40 20/)
house/bungalow	3,921 (53.3%)	3,430 (46.7%)	3,803 (51.7%)	3,548 (48.3%)
Terraced	3,483 (46.1%)	4,080 (53.9%)	3,943 (52.1%)	3,620 (47.9%)
house/end	3,403 (40.170)	4,000 (33.770)	3,743 (32.170)	3,020 (47.770)
terraced				
house/bungalow				
Flats/maisonettes	1,841 (35.8%)	3,301 (64.2%)	3,130 (60.9%)	2,012 (39.1%)
purpose built or	-, ()	0,000 (011270)	2,222 (22372)	_,,, (=,,,,,,
converted				
Bedsitters,	92 (56.8%)	70 (43.2%)	91 (56.2%)	71 (43.8%)
dwellings with	•	•	•	
business and				
sheltered				
accommodation				
Floor				
0 floors	76 (67.9%)	36 (32.1%)	49 (43.8%)	63 (56.2%)
1 floor	1,824 (52.5%)	1,651 (47.5%)	1,857 (53.4%)	1,618 (46.6%)
2 floors	9,186 (50.7%)	8,944 (49.3%)	9,415 (51.9%)	8,715 (48.1%)
3 floors	814 (38.0%)	1,328 (62.0%)	1,205 (56.3%)	937 (43.7%)
4 floors and	453 (30.2%)	1,046 (69.8%)	1,039 (69.3%)	460 (30.7%)
above				
Car/van	5 076 (55 10/)	4 122 (44 00/)	A 227 (A7 10/)	4 971 (52 004)
Definitely has a car/van	5,076 (55.1%)	4,132 (44.9%)	4,337 (47.1%)	4,871 (52.9%)
Likely	2,636 (50.2%)	2,610 (49.8%)	2,790 (53.2%)	2,456 (46.8%)
LIKEIY	4,030 (30.4%)	4,010 (47.0%)	4,170 (33.4%)	۷,+JU (40.0%)

Unlikely	527 (50.5%)	516 (49.5%)	647 (62.0%)	396 (38.0%)
Definitely does	417 (62.2%)	253 (37.8%)	206 (30.7%)	464 (69.3%)
not have a	, ,	, ,	,	, ,
car/van				
Cannot tell from	3,697 (40.2%)	5,494 (59.8%)	5,585 (60.8%)	3,606 (39.2%)
observation	3,077 (10.270)	3,171 (37.070)	3,505 (00.070)	3,000 (37.270)
Child				
Definitely has a	853 (51.3%)	811 (48.7%)	693 (41.6%)	971 (58.4%)
child/children	055 (51.570)	011 (40.770)	093 (41.070)	9/1 (30.470)
aged under 10	1 100 (51 00/)	1.072 (40.00()	1 000 (45 00/)	1 102 (54 20/)
Likely	1,128 (51.2%)	1,073 (48.8%)	1,008 (45.8%)	1,193 (54.2%)
Unlikely	2,630 (51.0%)	2,522 (49.0%)	2,866 (55.6%)	2,286 (44.4%)
Definitely does	2,133 (54.9%)	1,755 (45.1%)	1,864 (47.9%)	2,024 (52.1%)
not have a				
child/children				
aged under 10				
Cannot tell from	5,609 (45.0%)	6,844 (55.0%)	7,134 (57.3%)	5,319 (42.7%)
observation	,	, ,	,	, , ,
Unkempt garden				
Yes	1,001 (42.9%)	1,335 (57.1%)	1,249 (53.5%)	1,087 (46.5%)
No	9,007 (52.6%)	8,121 (47.4%)	8,891 (51.9%)	8,237 (48.1%)
No obvious	2,345 (39.8%)	3,549 (60.2%)	3,425 (58.1%)	2,469 (41.9%)
	2,343 (37.670)	3,347 (00.270)	3,423 (30.170)	2,407 (41.770)
garden				
Relative				
conditions of the				
address to other				
residential				
properties _				
Better	1,054 (53.5%)	917 (46.5%)	912 (46.3%)	1,059 (53.7%)
About the same	10,517 (48.9%)	10,981 (51.5%)	11,529 (53.6%)	9,969 (46.4%)
Worse	743 (41.4%)	1,050 (58.6%)	1,047 (58.4%)	746 (41.6%)
Unable to obtain	39 (40.6%)	57 (59.4%)	77 (80.2%)	19 (19.8%)
information				
Call Record				
Variables				
Time of day call				
1				
Morning (0-	2,280 (48.7%)	2,402 (51.3%)	2,485 (53.1%)	2,197 (46.9%)
12.00)	2,200 (10.770)	2,102 (31.370)	2,103 (33.170)	2,177 (10.770)
Afternoon (12.00-	8,485 (49.5%)	8,647 (50.5%)	9,037 (52.7%)	8,095 (47.3%)
17.00)	0,403 (49.370)	0,047 (30.370)	9,037 (32.170)	0,093 (47.370)
	1 500 (44 00/)	1.056 (55.20/)	2.042 (57.60/)	1 501 (42 40/)
Evening (17.00-	1,588 (44.8%)	1,956 (55.2%)	2,043 (57.6%)	1,501 (42.4%)
24.00)				
Time of day call				
2				
Morning (0-	2,377 (50.7%)	2,314 (49.3%)	2,524 (53.8%)	2,167 (46.2%)
12.00)				
Afternoon (12.00-	6,653 (48.7%)	7,016 (51.3%)	7,198 (52.7%)	6,471 (47.3%)
17.00)				
Evening (17.00-	3,323 (47.5%)	3,675 (52.5%)	3,843 (54.9%)	3,155 (45.1%)
24.00)		•	•	•

Time between				
call 1 and call 2				
Time of day call				
3				
Morning(0-12.00)	2,346 (50.4%)	2,313 (49.6%)	2,515 (54.0%)	2,144 (46.0%)
Afternoon (12.00-	5,801 (49.3%)	5,974 (50.7%)	6,325 (53.7%)	5,450 (46.3%)
17.00)				
Evening (17.00-	4,206 (47.1%)	4,718 (52.9%)	4,725 (52.9%)	4,199 (47.1%)
24.00)				
Time between				
call 2 and call 3				
Call 1 outcome				
No contact	7,876 (45.0%)	9,640 (55.0%)	9,135 (52.2%)	8,381 (47.8%)
Contact made	2,957 (55.4%)	2,385 (44.6%)	3,136 (58.7%)	2,206 (41.3%)
Appointment	909 (58.1%)	655 (41.9%)	680 (43.5%)	884 (56.5%)
made				
Any other status	521 (64.2%)	291 (35.8%)	564 (69.5%)	248 (30.5%)
Interview done	90 (72.6%)	34 (27.4%)	50 (40.3%)	74 (59.7%)
Call 2 outcome				
No contact	7,225 (43.2%)	9,513 (56.8%)	8,900 (53.2%)	7,838 (46.8%)
Contact made	2,695 (53.2%)	2,372 (46.8%)	2,909 (57.4%)	2,158 (42.6%)
Appointment	1,215 (62.7%)	722 (37.3%)	750 (38.7%)	1187 (61.3%)
made				
Any other status	808 (71.8%)	317 (28.2%)	843 (74.9%)	282 (25.1%)
Interview done	410 (83.5%)	81 (16.5%)	163 (33.2%)	328 (66.8%)
Call 3 outcome				
No contact	5,706 (37.4%)	9,547 (62.6%)	9,034 (59.2%)	6,219 (40.8%)
Contact made	2,196 (49.6%)	2,230 (50.4%)	2,717 (61.4%)	1,709 (38.6%)
Appointment	2,793 (80.1%)	692 (19.9%)	528 (15.2%)	2,957 (84.8%)
made or interview				
or complete				
Any other status	860 (67.3%)	418 (32.7%)	1,031 (80.7%)	247 (19.3%)
Interview done	798 (87.1%)	118 (12.9%)	2557.8%)	661
				(72.2%)

b.)

Variables	Short	Short Successful	Long	Long Successful
	Unsuccessful		Unsuccessful	·
Geographic and				_
design				
Months				
January-June year	1,279 (20.5%)	1,930 (30.9%)	1,974 (31.6%)	1,070 (17.1%)
1				
July-December	1,061 (18.2%)	1,836 (31.4%)	1,917 (32.8%)	1,029 (17.6%)
year 1				
January-June year	1,271 (19.7%)	1,824 (28.2%)	2,266 (35.0%)	1,105 (17.1%)
2				
July-December	1,351 (19.9%)	1,801 (26.5%)	2,446 (36.0%)	1,198 (17.6%)

year 2				
Low density area				
for ethnic				
minorities				
No	1,821 (17.3%)	2,692 (25.6%)	4,106 (39.1%)	1,894 (18.0%)
Yes	3,141 (21.2%)	4,699 (31.7%)	4,497 (30.3%)	2,508 (16.9%)
Government				
Office Region				
(GOR)				
North East	192 (16.7%)	348 (30.3%)	359 (31.2%)	250 (21.8%)
North West	514 (18.4%)	891 (31.8%)	890 (31.8%)	505 (18.0%)
Yorkshire and the	458 (21.7%)	661 (31.4%)	626 (29.7%)	361 (17.1%)
Humber				
East Midlands	342 (19.0%)	610 (33.9%)	494 (27.4%)	356 (19.8%)
West Midlands	502 (21.6%)	693 (29.8%)	751 (32.3%)	378 (16.3%)
East of England	481 (19.3%)	773 (31.0%)	845 (33.9%)	396 (15.9%)
London	557 (15.1%)	722 (19.6%)	1,755 (47.7%)	647 (17.6%)
South East	792 (21.9%)	1,047 (29.0%)	1,151 (31.9%)	619 (17.2%)
South West	407 (20.5%)	655 (33.1%)	618 (31.2%)	301 (15.2%)
Wales	243 (21.4%)	351 (31.0%)	331 (29.2%)	209 (18.4%)
Scotland	474 (20.8%)	640 (28.1%)	783 (34.4%)	380 (16.7%)
Urban/rural				
Urban area	3,886 (18.7%)	5,861 (28.2%)	7,369 (35.4%)	3,676 (17.7%)
Rural area	1,076 (23.6%)	1,530 (33.5%)	1,234 (27.0%)	726 (15.9%)
Interviewer				
observations				
Barrier 2 (locked				
gates)				
Not mentioned	4,882 (19.6%)	7,311 (29.3%)	8,395 (33.7%)	4,324 (17.4%)
Mentioned	80 (17.9%)	80 (17.9%)	208 (46.6%)	78 (17.5%)
Accommodation				
Detached	1,273 (24.8%)	1,743 (33.9%)	1,325 (25.8%)	799 (15.5%)
house/bungalow				
Semi-detached	1,615 (22.0%)	2,306 (31.4%)	2,188 (29.8%)	1,242 (16.9%)
house/bungalow				
Terraced	1,330 (17.6%)	2,153 (28.5%)	2,613 (34.5%)	1,467 (19.4%)
house/end				
terraced				
house/bungalow	500 (10 50)	1 1 10 (00 00)	2 422 (47 22)	0.50 (4.5.00)
Flats/maisonettes	698 (13.6%)	1,143 (22.2%)	2,432 (47.3%)	869 (16.9%)
purpose built or				
converted	4.5 (2.0 4.5.)			
Bedsitters,	46 (28.4%)	46 (28.4%)	45 (27.8%)	25 (15.4%)
dwellings with				
business and				
sheltered				
accommodation				
Floor	22 (20, 52)	44 (20 20)	15 (15 20()	10 (15 00()
0 floors	32 (38.6%)	44 (39.3%)	17 (15.2%)	19 (17.0%)
1 floor	771 (22.2%)	1,053 (30.3%)	1,086 (31.3%)	565 (16.3%)
2 floors	3,666 (20.2%)	5,520 (30.4%)	5,749 (31.7%)	3,195 (17.6%)

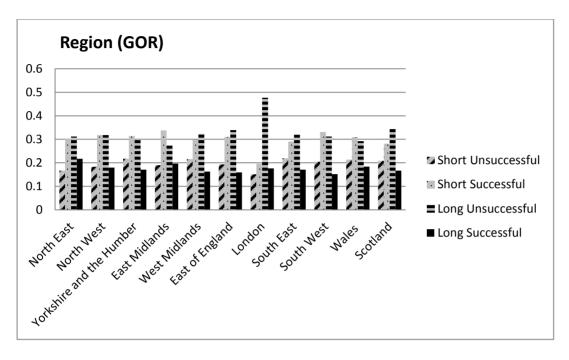
3 floors	288 (13.4%)	526 (24.6%)	917 (42.8%)	411 (19.2%)
4 floors and above	205 (13.7%)	248 (16.5%)	834 (55.6%)	212 (14.1%)
Car/van				
Definitely has a	1,855 (20.1%)	3,221 (35.0%)	2,482 (27.0%)	1,650 (17.9%)
car/van				
Likely	1,092 (20.8%)	1,544 (29.4%)	1,698 (32.4%)	912 (17.4%)
Unlikely	269 (25.8%)	258 (24.7%)	378 (36.2%)	138 (13.2%)
Definitely does	103 (15.4%)	314 (46.9%)	103 (15.4%)	150 (22.4%)
not have a car/van				
Cannot tell from	1,643 (17.9%)	2,054 (22.3%)	3,942 (42.9%)	1,552 (16.9%)
observation				
Child	2.15 (1.1.00)	50.5 (0.5.10)	115 (05 00)	2.57 (24.00)
Definitely has a	247 (14.8%)	606 (36.4%)	446 (26.8%)	365 (21.9%)
child/children				
aged under 10	260 (16 40/)	760 (24.00/)	(40 (00 40/)	425 (10 20/)
Likely	360 (16.4%)	768 (34.9%)	648 (29.4%)	425 (19.3%)
Unlikely Definitely does	1,172 (22.7%)	1,458 (28.3%)	1,694 (32.9%)	828 (16.1%) 714 (18.4%)
not have a	823 (21.2%)	1,310 (33.7%)	1,041 (26.8%)	/14 (18.4%)
child/children				
aged under 10				
Cannot tell from	2,360 (19.0%)	3,249 (26.1%)	4,774 (38.3%)	2,070 (16.6%)
observation	2,300 (19.070)	3,249 (20.170)	4,774 (30.370)	2,070 (10.070)
Unkempt garden				
Yes	359 (15.4%)	642 (27.5%)	890 (38.1%)	445 (19.0%)
No	3,670 (21.4%)	5,337 (31.2%)	5,221 (30.5%)	2,900 (16.9%)
No obvious	933 (15.8%)	1,412 (24.0%)	2,492 (42.3%)	1,057 (17.9%)
garden) (10.070)	1,112 (2.1070)	2, 1, 2 (12.670)	2,007 (27.570)
Relative				
conditions of the				
address to other				
residential				
properties				
Better	351 (17.8%)	703 (35.7%)	561 (28.5%)	356 (18.1%)
About the same	4,272 (19.9%)	6,245 (29.0%)	7,257 (33.8%)	3,724 (17.3%)
Worse	314 (17.5%)	429 (23.9%)	733 (40.9%)	317 (17.7%)
Unable to obtain	25 (26.0%)	14 (14.6%)	52 (54.2%)	5 (5.2%)
information				
Call Record				
Variables				
Time of day call				
1				
Morning (0-	921 (19.7%)	1,359 (29.0%)	1,564 (33.4%)	838 (17.9%)
12.00)	0.050 (10.50)	5.100 (00.00)	T (04 (00 00))	0.040 (17.00)
Afternoon (12.00-	3,353 (19.6%)	5,132 (30.0%)	5,684 (33.2%)	2,963 (17.3%)
17.00)	ZOO (10 40()	000 (05 40)	1.055 (00.000)	(01 (15 00))
Evening (17.00-	688 (19.4%)	900 (25.4%)	1,355 (38.2%)	601 (17.0%)
24.00)				
Time of day call				
Morning (0-	061 (20 50/)	1 /16 (20 20/)	1 563 (22 20/)	751 (16 00/)
Morning (0-	961 (20.5%)	1,416 (30.2%)	1,563 (33.3%)	751 (16.0%)

12.00)				
Afternoon (12.00-	2,682 (19.6%)	3,971 (29.1%)	4,516 (33.0%)	2,500 (18.3%)
17.00)				
Evening (17.00-	1,319 (18.8%)	2,004 (28.6%)	2,524 (36.1%)	1,151 (16.4%)
24.00)				
Time between				
call 1 and call 2				
Time of day call				
3	070 (21 00/)	1.047.400.004	1 50 5 (00 00)	
Morning(0-12.00)	979 (21.0%)	1,367 (29.3%)	1,536 (33.0%)	777 (16.7%)
Afternoon (12.00-	2,455 (20.8%)	3,346 (28.4%)	3,870 (32.9%)	2,104 (17.9%)
17.00)	1 500 (17 10/)	2 (70 (20 00))	2 107 (25 90/)	1 501 (17 00/)
Evening (17.00-	1,528 (17.1%)	2,678 (30.0%)	3,197 (35.8%)	1,521 (17.0%)
24.00)				
Time between call 2 and call 3				
Call 1 outcome				
No contact	2,735 (15.6%)	5,141 (29.4%)	6,400 (36.5%)	3,240 (18.5%)
Contact made	1,527 (28.6%)	1,430 (26.8%)	1,609 (30.1%)	776 (14.5%)
Appointment	310 (19.8%)	599 (38.3%)	370 (23.7%)	285 (18.2%)
made	310 (17.070)	377 (30.370)	370 (23.770)	203 (10.270)
Any other status	355 (43.7%)	166 (20.4%)	209 (25.7%)	82 (10.1%)
Interview done	35 (28.2%)	55 (44.4%)	15 (12.1%)	19 (15.3%)
Call 2 outcome			- (,	()
No contact	2,603 (15.6%)	4,622 (27.6%)	6,297 (37.6%)	3,216 (19.2%)
Contact made	1,283 (25.3%)	1,412 (27.9%)	1,626 (32.1%)	746 (14.7%)
Appointment	332 (17.1%)	883 (45.6%)	418 (21.6%)	304 (15.7%)
made				
Any other status	611 (54.3%)	197 (17.5%)	232 (20.6%)	85 (7.6%)
Interview done	133 (27.1%)	277 (56.4%)	30 (6.1%)	51 (10.4%)
Call 3 outcome				
No contact	2,706 (17.7%)	3,000 (19.7%)	6,328 (41.5%)	3,219 (21.1%)
Contact made	1,170 (26.4%)	1,026 (23.2%)	1,547 (35.0%)	683 (15.4%)
Appointment	176 (5.1%)	2,617 (75.1%)	352 (10.1%)	340 (9.8%)
made				
Any other status	707 (55.3%)	153 (12.0%)	324 (25.4%)	94 (7.4%)
Interview done	203 (22.2%)	595 (65.0%)	52 (5.7%)	66 (7.2%)

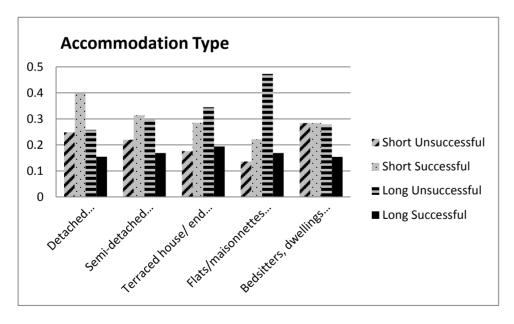
Figure A1: [Figure 1 from text in greyscale.]

Predicted probabilities for the final combined model (final outcome and length) for selected explanatory variables (one variable from each group of explanatory variables in Model 6)

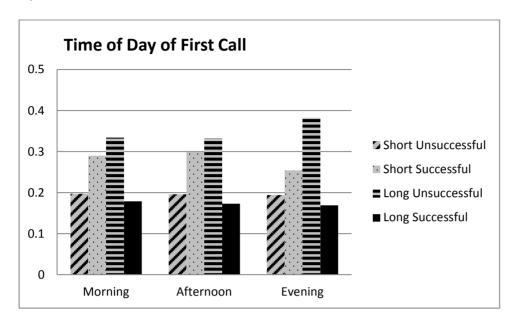
a.)



b.)



c.)



d.)

