# **Sleep in youth with Autism Spectrum Disorders: systematic review and meta-analysis**

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**ABSTRACT**

**QUESTIONS:** Sleep problems are common and impairing in individuals withautism spectrum disorders (ASD). Evidence synthesis including both subjective (i.e., measured with questionnaires) and objective (i.e., quantified with neurophysiologic tools) sleep alterations in youth with ASD is currently lacking. Therefore, we conducted a systematic review and meta-analysis of subjective and objective studies sleep studies in youth with ASD. **STUDY SELECTION AND ANALYSIS:** We searched the following electronic databases with no language, date, or type of document restriction, up to May 23rd, 2018: Pubmed, PsycInfo, Embase+Embase Classic, Ovid Medline, and Web of Knowledge. Random-effects models were used. Heterogeneity was assessed with Cochran's Q and I2 statistics. Publication (small studies) bias was assessed with final plots and the Egger’s test. Study quality was evaluated with the Newcastle Ottawa Scale. Analyses were conducted using *Review Manager* and *Comprehensive meta-analysis.* **FINDINGS:** From a pool of 3,359 non-duplicate potentially relevant references, 47 datasets were included in the meta-analyses. Subjective and objective sleep outcome measures were extracted from 37 and 15 studies, respectively. Only five studies were based on comorbidity free, medication-naïve participants. Compared to typically developing controls, youth with ASD significantly differed in 10/14 subjective parameters and in 7/14 objective sleep parameters. The average quality score in the Newcastle-Ottawa scale was 5.9/9. **CONCLUSIONS:** A number of subjective and, to a less extent, objective sleep alterations might characterise youth with ASD, but future studies should assess the impact of pharmacological treatment and psychiatric comorbidities.

**Summary box**

What is already known about this subject?

* Autism spectrum disorders (ASD) represent common and impairing neurodevelopmental conditions.
* Sleep problems are commonly reported in individuals with ASD and have a negative impact on their daily functioning.
* Although a previous meta-analysis explored objective (i.e., measured with physiological tools) sleep problems in children with ASD, updated evidence synthesis including both subjective (i.e., assessed via questionnaires) and objective sleep problem is lacking.

What are the new findings?

* Our meta-analysis of 47 datasets showed that, compared to typically developing controls, youth with ASD significantly differed in 10/14 subjective parameters and in 7/14 objective sleep parameters.
* Comorbid psychiatric conditions and the pharmacological treatment may contribute to sleep impairment in ASD.
* Our findings allow a fine-grain characterisation of sleep alterations in ASD which is helpful for their daily clinical management.

How might it impact on clinical practice in the foreseeable future?

* Presence of sleep alterations in youths with ASD should be systematically screened in the clinical practice in order to reduce their impact on daytime functioning.

**BACKGROUND**

Autism spectrum disorders encompass a wide range of neurodevelopmental conditions characterized by a deficit in social communication, together with restricted, repetitive and stereotyped behaviours, interests, or activities. 1 Although not formally part of the diagnostic criteria, 1,2 sleep problems are frequently reported in individuals with ASD (e.g., 3–5) and contribute to their functional impairment. Sleep difficulties are associated with a significant amount of distress for the patients and their families, 6 but also negatively impact on cognitive abilities and self-regulation of disruptive behaviours during the daytime. 7–9

In order to appropriately manage them, it is necessary to characterise the profile of sleep problems in children and adolescents with ASD. Whilst a number of individual studies have been conducted, we are aware of only one meta-analysis that summarised the available body of evidence. 10 However, this meta-analysis was limited to objective sleep studies, i.e., studies relying on actigraphic or polysomnographic measures. Whilst these (in particular, polysomnography) are considered rigorous measures of sleep, it is important to also consider sleep measures subjectively reported by patient and/or their parents via questionnaires, as they are arguably more “ecological” and they reflect the subjective perception, which is important in the management process of the disorder. Furthermore, the meta-analysis by Elrod and Hood 10 was published in 2015 and, as such, an update is warranted.

**OBJECTIVE**

To conduct a systematic review and meta-analysis of subjective and objective studies of sleep in children and adolescents with ASD compared to typically developing controls.

**STUDY SELECTION AND ANALYSIS**

We followed the recommendations of the Meta-Analysis of Observational Studies in Epidemiology group 11 and the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement. 12 The protocol of this systematic review was registered in PROSPERO (CRD42018100016).

**Type of studies**

We included case-control studies comparing children with ASD to typically developing individuals on subjective and/or objective sleep parameters.

**Type of participants**

We included studies on children/youth (≤ 20 years) diagnosed with ASD according to DSM III to 5 criteria or ICD-9 to 10 criteria, or according to a clinical diagnosis of ASD, compared to typically developing participants. Definition of ASD based on cut-off on questionnaires targeting ASD symptoms was not considered rigorous and as such was exclusionary. Psychiatric comorbidities were not an exclusionary criterion.

**Outcomes**

Any subjective sleep parameters from any sleep questionnaire and/or any objective sleep parameters measured using polysomnography (PSG), actigraphy or multiple sleep latency test (MSLT), which were presented in at least two studies, were meta-analysed. We selected the following subjective parameters: bedtime resistance, sleep onset delay, sleep duration, sleep anxiety, night awakenings, parasomnias, sleep-disordered breathing, daytime sleepiness, general sleep problems, sleep quality, sleep efficiency, sleep onset latency (min), sleep duration (min), and restorative value of sleep (*i.e.,* feeling well-rested after waking up). For PSG, we considered total sleep time, sleep onset latency, time spent in each sleep stage, REM latency, sleep efficiency, and wake time after sleep onset. As for actigraphic parameters, we selected: sleep onset latency, true sleep, assumed sleep time, actual wake time, and sleep efficiency. For MST, we considered latency to falling asleep.

**Search strategy/syntax**

We searched the following electronic databases: Pubmed (Medline), OVID databases (PsycInfo, EMBASE+EMBASE classic, OVID Medline), and WEB OF KNOWLEDGE Databases (Web of science (Science Citation Index Expanded), Biological abstracts, Biosis, Food science and technology abstracts), up to May 23rd, 2018 with no language/date/type of document restrictions. Further details on the search strategy/syntax, including search terms for each database, are reported in the Supplemental Material 1. References of included studies and of reviews conducted on this topic were also hand-searched to find potential pertinent studies undetected with the electronic search strategy.

**Screening and data extraction**

*Screening*

Title and abstracts of all non-duplicated papers were independently screened by two of the authors (JZ, AD). Potential pertinent papers were retained and assessed for eligibility by screening the full-text. A third senior author (SC) acted as arbitrator when disagreement in any screening stage. If needed, corresponding authors of retained studies were also contacted to request further information.

*Data extraction*

Data extraction was independently performed by two of the authors (JZ, AD), and any discrepancy between them was resolved by consensus. The following data were extracted from each study: first author and publication year, country where the study was conducted, study participants’ details (number, percentage of males, mean age and SD, ASD diagnostic criteria, medication status, and comorbidities), mean and SD for each outcome measure (subjective and/or objective sleep parameters), and nights recorded for sleep assessment.

**Risk of bias assessment**

Two authors (JZ, AD) independently assessed the methodological quality or risk of bias of included studies using the Newcastle-Ottawa Scale for case-control studies (<http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp>). This scale includes the following domains: case definition, representativeness of the cases, selection of controls, definition of controls, comparability of cases and controls on the basis of the design or analysis, ascertainment of exposure, and non-response rate. Disagreements between both authors were resolved by consensus.

**Statistical analysis**

Analyses were performed with Review Manager 5.3 (<http://community.cochrane.org/tools/reviewproduction-tools/revman-5>) and Comprehensive Meta-Analysis, CMA (<http://www.meta-analysis.com/index.php>). Random-effects models were used to compute standardized mean difference (SMD) for each sleep parameter, with 95% confidence interval and the Hedges’ correction 13 to avoid sample size bias. The inverse variance method and the Z statistic were used to calculate the pooled SMD and assess its statistical significance. Heterogeneity degree between studies was measured with Cochran's Q and I2 statistics. 14 Publication bias were explored using the Egger’s test and the funnel plots. 15 We also conducted a post hoc analysis including only studies based on comorbidity free, medication-naïve participants.

**FINDINGS**

From a pool of 3,359 non-duplicate potentially relevant references, 47 datasets (reported in 48 references) were included in our meta-analysis (Figure 1). The list of excluded reports (with reasons for exclusions) and included studies are provided in the Supplemental Materials 2 and 3, respectively. Table 1 shows the main characteristics of the studies included in the meta-analysis. All studies were cross-sectional and the average quality score in the Newcastle-Ottawa scale was 5.9/9 (scores ranged from 3 to 8; Supplemental Material 4).

Subjective outcome measures were extracted from 37 studies, whilst objective outcome measures were obtained from 15 studies (eight studies using PSG,16-23 six using actigraphy,24-29 and one using both 30). Overall, the number of participants ranged from 75 to 5,430 for studies reporting subjective sleep parameters, and from 144 to 312 for sleep objective studies. Two studies 31, 32 reported sleep data of two different samples and we included both samples in the meta-analysis independently.

*Subjective measures of sleep difficulties*

Compared to control individuals, participants with ASD, showed significantly higher bedtime resistance (SMD = 1.00, 95% CI: 0.67 to 1.33), sleep onset delay (0.98, 0.66 to 1.29), sleep anxiety (0.96, 0.61 to 1.32), night awakenings (0.72, 0.44 to 1.01), parasomnias (0.88, 0.60 to 1.15), sleep-disordered breathing (0.48, 0.28 to 0.67), daytime sleepiness (0.34, 0.16 to 0.52), sleep onset latency (in min) (0.81, 0.59 to 1.02), restorative value of sleep (0.81, 0.59 to 1.02), and general sleep problems (0.93, 0.67 to 1.20). They also showed lower sleep duration (-0.88, -1.18 to -0.57). In contrast, children with ASD did not significantly differ from control individuals in sleep quality, sleep efficiency, or sleep duration in min (Table 2 and the Supplemental Material 5). As shown in Table 2, the heterogeneity between studies was statistically significant for almost all subjective sleep parameters (I2 ranged from 81% to 95%), except for sleep efficiency and sleep onset latency (in min). There was also evidence for publication bias for 5 out of 14 subjective sleep parameters: sleep duration (t = 2.19, p = 0.040), sleep anxiety (t = 2.69, p = 0.014), parasomnias (t = 3.30, p = 0.003), daytime sleepiness (t = 2.26, p = 0.032) and general sleep problems (t = 2.31, p = 0.028). The results of the Egger’s test and the funnel plots are reported in Table 2 and the Supplemental Material 6, respectively.

*Objective parameters of sleep alterations*

As reported in Table 3, children with ASD significantly differed from control individuals in several objective parameters measuring sleep patterns using PSG. Specifically, children with ASD showed lower total sleep time (-0.90, -1.51 to -0.30), longer sleep onset latency (0.53, 0.21 to 0.86), higher time spent in stage 1 sleep (0.48, 0.06 to 0.90), lower time of REM sleep (-0.88, -1.56 to -0.21), lower sleep efficiency (-1.20, -1.98 to -0.41), and higher time awake after sleep onset (0.49, 0.11 to 0.87). However, no significant differences were observed between children with ASD and control individuals in stage 2 sleep, slow wave sleep, and REM latency (Table 3 and the Supplemental Material 5). In relation to actigraphy, we found differences between both groups only in sleep onset latency (Table 3 and the Supplemental Material 5). Children with ASD displayed significantly longer sleep onset latency than control individuals (0.80, 0.55 to 1.05). Evidence of heterogeneity was found for almost all polysomnographic sleep parameters (I2 ranged from 55% to 85%), with the exception of sleep onset latency and wake time, but only for a single actigraphic sleep parameter (sleep efficiency, I2 = 62%). No evidence of publication bias was detected in the Egger’s test (Table 3) and the funnel plots (Supplemental Material 6).

*Post hoc analysis*

The post hoc analysis based on studies including only comorbidity-free and medication naïve participants was limited to PSG studies as only two studies for subjective measures and one study for actigraphic measures, respectively, provided usable data. As shown in Table 4 the post hoc analysis of PSG studies replicated the results of the main analysis (excepted for the parameter duration of sleep stage 1, that was not more significant between participants with ASD and controls).

**CONCLUSIONS AND CLINICAL IMPLICATIONS**

To our knowledge, this is the first meta-analysis including both subjective and objective measures of sleep in children with ASD. We found that, compared to typically developing children, those with ASD presented with a number of significant sleep impairments, quantified both by subjective and objective parameters.

Our results were in accordance with the findings of a previous meta-analysis in children with ASD, 10 in which these children also showed significantly lower total sleep time, increased sleep onset latency, and worse sleep efficiency compared to typically developing children. However, these differences between groups observed during PSG were not consistent with actigraphy-defined measures since only sleep time or sleep efficiency statistically differed.

It should be noted that, although the previous meta-analysis by Elrod and Hood 10 pooled both PSG and actigraphy sleep outcomes together, their analyses of moderating factors revealed a significant impact of sleep assessment method on sleep efficiency. Specifically, their results suggested no difference in children with ASD and controls in actigraphic sleep efficiency, being this consistent with our results to a greater extent.

Our work adds meta-analytic evidence to the Elrod and Hood study, 10 extending to subjective measures of sleep disturbances in ASD. Our results stress further that children with ASD displayed a considerable burden of sleep problems. These children seemed to experience greater bedtime resistance, sleep anxiety, sleep-disordered breathing, and parasomnias, as well as longer sleep onset latency and higher daytime sleepiness. However, there was less consistency about total sleep time, depending on how it was estimated (i.e., with a score in a sleep questionnaire or a length in min). Finally, despite children with ASD showed significantly higher scores in general sleep problems than typically developing children, they did not differ in terms of subjectively reported sleep quality and sleep efficiency, although the limited number of included subjective studies reporting sleep quality (n = 3) and sleep efficiency (n = 2) suggest that this conclusion should be considered with caution.

Our results based on subjective measures were generally not consistent with those obtained with objective parameters. For instance, children with ASD and control individuals did not significantly differ in terms of sleep duration based on parents’ report. By contrast, children with ASD showed a significantly lower total sleep time compared to typically developing children according to PSG measures. This is not surprising and reflects the well-known mismatch between subjective and objective measures.33 Indeed, discrepancies between subjective and objective sleep measures have been reported in earlier studies in both children with ASD (e.g., 30,34) and children with other neurodevelopmental disorders (e.g., 35,36). For example, objective measures were usually taken on one or two nights, while subjective measures reflected the perception of the parent over several nights. Taking into account the advantages and limitations of both subjective and objective sleep measures, rather than considering these two types of measures as exclusionary, we would suggest they should be seen as providing complementary information. Additionally, even within objective studies, there were some discrepancies among apparently similar parameters. Of note, sleep efficiency was significantly different between participants with and without ASD when measured with PSG but not when assessed via actigraphy. The different degree of ecological validity of PSG (usually implemented in a lab) and actigraphy (in the home environment) may contribute to explain these discrepancies.

Our findings could have been impacted by the presence of psychiatric comorbidities and the drug intake of the subjects included in our meta-analysis. For example, psychiatric comorbidities, including epilepsy and Attention-Deficit/Hyperactive Disorder (AHDH) were reported in at least 19% (7/37) of the studies included in our meta-analysis. Drug intake, including stimulants and melatonin, was mentioned in 35% (13/37) of the studies included. In fact, the impact of psychiatric comorbidities on sleep has been consistently reported in research.37,38 The effects of drug on sleep patterns were also well-documented.38 Thus, the high heterogeneity of the results among studies we observed (based on the results of the Egger’s test and the funnel plot) may be related to participants’ comorbid conditions or medication intake. Unfortunately, most of the studied taking account in our analysis did not provide this information which could have been useful to perform a meta-regression to assess the impact of these possible confounders. Additionally, our post hoc analysis based on studies including only comorbidity free and medication naïve participants could only confirm the results of the main meta-analysis of PSG studies, as there were not enough studies for the analysis of subjective and actigraphic parameters.

In addition to the possible role of psychiatric comorbidities and medications, the causes of sleep impairments in children with ASD are likely to be complex and not mutually exclusive. Behavioral factors such as dysfunctional bedtime routines, exacerbated by comorbid anxiety or ADHD, may disrupt sleep, especially sleep onset delay. There is also an increasing body of evidences suggesting the contributing role of biological clock factors (mainly endocrine and genetic) which could be involved in dysregulation of day-night rhythm and sleep patterns in subjects with ASD (see review of 39). For instance, ASD was associated with decreased urinary or blood melatonin level, 40 probably due to genetic and epigenetic abnormalities affecting the enzymes of the melatonin synthesis and degradation pathways.41 Similarly, several studies suggested that BMAL1 or additional clock genes involved in the synchronisation of biological rhythms, may be impaired in ASD. This may affect the ability of patients with ASD to anticipate and adapt their behaviours (including their sleep patterns) to environmental changes.42

The results of our systematic meta-analysis should be considered in the light of its strengths and limitations. As for the strengths, we pre-registered the protocol in a publicly available repository (PROSPERO), reducing the risk of reporting bias. Furthermore, we endeavoured to perform a comprehensive and systematic search in several databases, with no restrictions in terms of language or document type, and we gathered unpublished data from study authors. Additionally, we used a state-of-the-art tool, the Newcastle-Ottawa scale, to assess the quality of the retained studies.

There were also a number of limitations that should be taken into account. First, statistical heterogeneity was significant for the majority of the included measures. Although this did not invalidate the results, it indicated that the pooled effect sizes could not appropriately summarise the results from all datasets. Second, whilst we endeavoured to perform a comprehensive search, there was evidence of publication bias for a number of measures, which suggested that a more transparent report of research findings in the field is needed. Third, as for the quality of individual studies, most of them were rated at overall medium quality or risk of bias using the Newcastle-Ottawa scale, and main concerns were noted in relation to the comparability of groups and exposure-related items. The latter, along with the relatively sparse evidences available for some sleep parameters, calls for more studies assessing sleep impairments in children with ASD and by controlling the mentioned concerns.

Despite these caveats, we deem that our study provides meta-analytic evidence of objective and subjective sleep difficulties in patients with ASD. Clinicians managing children with ASD should systematically query about sleep alterations, during the first assessment and all along the follow-up. Subjective questionnaires such as the scale by Bruni et al. 43 or Owens et al. 44 can be used to screen sleep difficulties at the first assessment and at each follow-up visit with children with ASD. The extent to which these alterations are accounted for by comorbid disorders and/or the effect of pharmacotherapy should be better explored in future studies recruiting only medication naïve and comorbidity free participants, although our post hoc analysis suggest that objective differences inn sleep parameters are detected regardless the effect of comorbidities and medications. Additionally, further research needs to be performed to dissect the dysfunction of biological regulators in ASD. This may offer new promising avenues for early detection and therapeutic intervention in ASD.

**Table 1. Descriptive table of the studies included in the meta-analysis.**

|  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  | **ASD group** | | | **Control group** | | |  |  |  |  |
| **First author (year)** | **Country** | **Diagnosis** | ***n* (% male)** | **Age (years)** | **Type** | ***n* (% male)** | **Age** | **Co-occurrent drugs** | **Comorbidities** | **Main sleep measures** | **Nights recorded (PSG/ACT)** |
| Aathira (2017) | India | DSM-IV | 71 (80.28) | 5.3±1.8 | TD | 65 (61.54) | 5.7±1.6 | 22 psychotropic medication, 11 antiepileptics | 11 children with ASD had epilepsy | CSHQ | NA |
| Al-Farsi (2018) | Oman | DSM-V-TR | 122 (82.8) | 3-14 | TD | 90 (64) | 3-14 | NS | NS | CSHQ | NA |
| Allik (2006)a | Sweden | ICD-10 | 32 (87.5) | 10.80±1.25 | TD | 32 (87.5) | 10.9 ± 1.3 | 0 | 0 | ACT, Other subjective | 7 |
| Baker (2013) | Australia | Clinical diagnosis | 27 (81.48) | 15.5 ± 1.3 | TD | 27 (81.48) | 15.5 ± 1.1 | 12 ASD (including anti-depressants, psycho-stimulants, sleeping agents, anti-acne treatments, the contraceptive pill, antipsychotics and anti-inflammatories) and 2 TD (anti-acne medication and thyroxine) | NS | ACT, Other subjective | 7 |
| Bruni (2007) | Italy | ICD-10 and DSM-IV | 18 (88.89) | 12.26±2.50 | TD | 12 (58.33) | 12.6±3.7 | 0 | 0 | PSG | 2 |
| Chou (2012) | China | DSM-IV | 110 (87.3) | 7.6 ± 2.4 | TD | 110 (87.3) | 7.9 ± 2.0 | NS | NS | Other subjective | NA |
| Cotton (2006)a | Australia | Clinical diagnosis | 37 (75.7) | 7.09 ± 2.49 | TD | 55 (58.2) | 8.88 ± 3.95 | NS | NS | Other subjective | NA |
| Cotton (2010)a | Australia | Clinical diagnosis | 34 (73.53) | 7.2 ± 2.5 | TD | 33 (69.70) | 8.2 ± 3.2 | NS | NS | Other subjective | NA |
| Couturier (2005) | Canada | DSM-IV | 23 (95.7) | 9.4 ± 2.0 | TD | 23 (95.7) | 9.5± 2.0 | 75% ASD (56% stimulants, antipsychotics, and antidepressants, and 12.9% others) | NS | CSHQ | NA |
| Elia (1991) | Italy | DSM-III-R | 4 (100) | 13.3 (10.5-15) | TD | 5 (100) | 13.0 (9-17) | NS | 0 | PSG | 3 |
| Elia (2000) | Italy | DSM-IV | 17 (100) | 10.36 ± 3.79 | TD | 5 (100) | 9.22 ± 2.02 | 0 | 0 | PSG | 2 |
| Fletcher (2017) | Australia | DSM-IV-TR | 21 (81.0) | 106.67 ± 26.82 months | TD | 29 (48.3) | 102.10 ± 17.07 months | ASD group: 3 melatonin, 2 stimulant, 2 SSRI, 1 atomoxetine | NS | ACT, CSHQ | 14 |
| Giannotti (2008)a | Italy | DSM-IV-TR | 104 (86.54) | 2.3-7.10 | TD | 162 (62) | 2.2-7.11 | 0 | 19.44% epilepsy (ASD group) | CSHQ | NA |
| Giannotti (2011)a | Italy | DSM-IV-TR | 40 (77.5) | 5.32±3.12 | TD | 12 (75) | 5.8 ± 2.4 | 0 | 0 | PSG | 2 |
| Guler (2016) | Turkey | DSM-5 | 60 (73.3) | 7.10 ± 1.50 | TD | 60 (65) | 6.93 ± 1.59 | NS | NS | CSHQ | NA |
| Han (2017) | China | DSM-5 | 212 (85.4) | 6.0 ± 2.7 | TD | 334 (81.4) | 5.9 ± 2.6 | NS | 0 | CSHQ | NA |
| Harder (2016)a | USA | Clinical diagnosis | 21 (100) | 7.8 ± 1.8 | TD | 23 (78.26) | 8.0 ± 1.9 | 0 | 0 | PSG | 1 |
| Henderson (2011) | USA | Clinical diagnosis | 58 (86.2) | 9.0 ± 2.09 | TD | 57 (52.6) | 8.25 ± 1.98 | NS | NS | Other subjective | NA |
| Hirata (2016) | Japan | DSM-5 | 193 (80.83) | 4.45 ± 1.24 | Community group | 965 (80.83) | 4.51 ± 1.15 | NS | 0 | Other subjective | NA |
| Hodge (2014)a | USA | DSM-IV-TR (questionnaire) | 108 (83.33) | 7.33 ± 3.18 | TD | 108 (83.33) | 7.71 ± 3.13 | NS | NS | CSHQ | NA |
| Hoffman (2006)a | USA | DSM-IV-TR | 106 (84) | 8.2 ± 2.69 | TD | 168 (55) | 8.62 ± 3.28 | NS | ASD group: 14 seizure disorders, 15 ADHD, 6 cerebral palsy | CSHQ | NA |
| Inamuna (1984) (1)b | Japan | ICD-9 (modified version) | 11 (-) |  | TD | 16 (-) | NS | NS | NS | Other subjective | NA |
| Inamuna (1984) (2)b | Japan | ICD-9 (modified version) | 19/20 (-) |  | TD | 17/18 (-) | NS | NS | NS | Other subjective | NA |
| Kelmanson (2018) | Russia | DSM-5 (questionnaire) | 18 (100) | 5 | TD | 54 (100) | 5 | NS | NS | CSHQ | NA |
| Kheirouri (2016) | Iran | DSM-IV-TR | 35 (68.6) | 8.1 ± 4.0 | TD | 31 (58.1) | 7.3 ± 2.6 | 97.1 ASD | 0 | CSHQ | NA |
| Lambert (2016)a | Canada | DSM-IV | 11 (NS) | 10.27 ± 2.24 | TD | 13 (NS) | 10.23 ± 2.01 | 1 ASD (methylphenidate) | 0 | PSG, CSHQ, Other subjective | 2 |
| Levin (2016) | Israel | DSM-IV | 34 (73.5) | 3.28±0.43 | TD | 31 (48.4) | 3.02±0.48 | NS | 0 | CSHQ | NA |
| Li (2012) (1)b | China | DSM-IV | 49 (-) | 4.7 ± 0.7 | TD | 49 (-) | 4.5 ± 0.9 | NS | NS | CSHQ | NA |
| Li (2012) (2)b | China | DSM-IV | 35 (-) | 9.0 ± 2.0 | TD | 42 (-) | 8.3 ± 1.8 | NS | NS | CSHQ | NA |
| Lopez-Wagner (2008)a | USA | DSM-IV-TR | 106 (84) | 8.2 ± 2.69 | TD | 168 (55) | 8.62 ± 3.28 | NS | ASD group: 14 seizure disorders, 15 ADHD, 6 cerebral palsy | CSHQ | NA |
| Malow (2009)a | USA | Clinical diagnosis | 93 (90) | 5.7 ± 2.1 | TD | 64 (59) | 6.8 ± 2.2 | 13% ASD and 5% TD (including atomoxetine, benadryl, citralopram, clonidine, dexmethylphenidate, divalproex sodium, fluoxetine, guanfacine, melatonin, methylphenidate, oxycarbmazapine, risperidone, and sertraline) | 0 | CSHQ | NA |
| Maski (2015) | USA | Clinical diagnosis | 22 (86) | 11.3 ± 2.1 | TD | 20 (90) | 12.3 ± 2.1 | 31.8 % ASD (3 stimulants, 2 antidepressants, 1 guanfacine ± antidepressant, 1 mood stabilizer ± antidepressant) | 0 | PSG, ACT, CSHQ | 1 PSG/7 ACT |
| Matsuoka (2014) | Japan | DSM-IV-TR | 31 (93.55) | 6-12 | TD | 372 (48.9) | 9.4 ± 4.5 | Anti-allergics, antipsychotics, anti-epileptics, methylphenidate (exact percentage for ASD group is not reported) | 2 TD had epilepsy, ASD no reported | CSHQ | NA |
| May (2015) | Australia | DSM-IV-TR | 46 (52.17) | 9.84±1.89 | TD | 38 (63.16) | 9.04±1.62 | 7 ASD (including serotonin-specific reuptake inhibitors, stimulants, risperidone, and melatonin) | NS | CSHQ | NA |
| Miano (2007) | Italy | DSM-IV | 16 (100) | 9.4 ± 2.33 | TD | 18 (50) | 10.2 ± 2.93 | 0 | 0 | PSG | 2 |
| Mutluer (2016) | Turkey | DSM-5 | 64 (79.69) | 11.66 ± 3.8 | TD | 53 (75.47) | 11.75 ± 0.85 | 0 | 0 | Other subjective | NA |
| Paavonen (2008) | Finland | DSM-IV and ICD-10 | 52 (76.9) | 10.1 ± 3.4 | TD | 61 (47.5) | 10.0 ± 1.9 | 0 | NS | Other subjective | NA |
| Pace (2016) | France | DSM-5 | 19 (NS) | 10.7 ± 1.2 | TD | 19 (NS) | 9.9 ± 1.6 | 0 | 0 | ACT | 7 |
| Park (2012) | Korea | DSM-IV-R | 166 (87.3) | 7.49 ± 3.05 | TD (siblings) | 111 (47.7) | 7.94 ± 3.50 | NS | 0 | Other subjective | NA |
| Patzold (1998)a | Australia | DSM-III and DSM-III-R | 38 (81.58) | 7.79±2.63 | Some with developmental disabilities | 36 (80.56) | 8.42±2.58 | 45% ASD (antipsychotics and over-the-counter medications), 20% TD (anticonvulsants and asthma medications) | NS | Other subjective | NA |
| Phung (2017) | USA | Clinical diagnosis | 19 (84.2) | 16.88 ± 2.50 | Neurotypical adolescents | 10 (60) | 15.73 ± 2.00 | NS | NS | ACT, Other subjective | 7 |
| Richdale (1995) | Australia | DSM-III and DSM-III-R | 12 (50) | 9.12±4.99 | Non-ASD | 35 (57.14) | 7.33±2.61 | NS | 2 ASD had epilepsy, 4 TD were asthmatic, 1 had allergy, 1 both | Other subjective | NA |
| Souders (2009) | USA | DSM-IV-TR (checklist) | 59 (81.4) | 7.53 ± 1.92 | TD | 40 (65) | 7.09 ± 2.09 | 56.7% ASD (15 melatonin, 3 catapres, 4 risperidone, 2 aripripazole, 1 hydroxyzine, 1 fluoxetine) | 0 | ACT, CSHQ | 10 |
| Tessier (2015)a | Canada | DSM-IV-TR | 13 (100) | 10.23 ± 2.08 | TD | 13 (100) | 10.23 ± 2.0 | 0 | 0 | PSG | 2 |
| Tzischinsky (2018) | Israel | DSM-5 | 69 (81.16) | 4.94 ± 1.23 | TD | 62 (66.13) | 4.82 ± 1.15 | 22 ASD (including melatonin, risperdal, ritalin, and neuleptil) | NS | CSHQ | NA |
| van der Heijden (2018) | Netherlands | DSM-IV | 68 (89.7) | 9.6 ± 1.9 | TD | 243 (51.9) | 8.7 ± 2.1 | NS | ASD: 16 ADHD, 7 other | Other subjective | NA |
| Yang (2018) | China | DSM-IV | 169 (85.80) | 5.23 ± 2.0 | TD | 172 (84.88) | 5.29 ± 1.58 | NS | NS | CSHQ | NA |
| PSG = polysomnography; ACT = actigraphy; TD = typically development children; NA = not applicable; NS = not specified.  aThe following studies presented some overlapped participants: 1) Allik 2006, Allik 2006, Allik 2008; 2) Cotton 2006, Cotton 2010, Patzold 1998; 3) Giannotti 2008, Giannoti 2011; 4) Harder 2016, Malow 2009; 5) Hodge 2014, Hodge 2013; 6) Hoffman 2006, Lopez-Wagner 2008; and 7) Lambert 2016, Tessier 2015. Studies reported in this table are those studies from which data for meta-analysis were extracted.  bNumbers 1 and 2 within parenthesis denote two different samples were reported in studies. | | | | | | | | | | | |

**Table 2. Summary of the results of the meta-analysis with subjective sleep parameters.**

|  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  |  |  |  |  | **Heterogeneity** | | | **Egger's test** | |
| **Sleep parameter** | **k** | **N** | **SMD (95% CI)** | **Z** | **p** | **Q** | **p** | **I2** | **t** | **p** |
| Bedtime resistance | 21 | 3589 | 1.00 [0.67, 1.33] | 5.98 | <0.00001 | 362.85 | <0.00001 | 94 | 2.00 | 0.059 |
| Sleep onset delay | 22 | 3636 | 0.98 [0.66, 1.29] | 6.03 | <0.00001 | 362.72 | <0.00001 | 94 | 2.09 | 0.049 |
| Sleep duration | 22 | 3376 | -0.88 [-1.18, -0.57] | 5.69 | <0.00001 | 304.14 | <0.00001 | 93 | 2.19 | 0.040 |
| Sleep anxiety | 20 | 3312 | 0.96 [0.61, 1.32] | 5.30 | <0.00001 | 375.97 | <0.00001 | 95 | 2.69 | 0.014 |
| Night awakeningsa | 20 | 3312 | 0.74 [0.44, 1.04] | 4.81 | <0.00001 | 274.60 | <0.00001 | 93 | 1.26 | 0.222 |
| Parasomnias | 22 | 4747 | 0.88 [0.60, 1.15] | 6.24 | <0.00001 | 330.97 | <0.00001 | 94 | 3.30 | 0.003 |
| Sleep-disordered breathing | 24 | 4129 | 0.48 [0.28, 0.67] | 4.80 | <0.00001 | 181.14 | <0.00001 | 87 | 1.06 | 0.299 |
| Daytime sleepiness | 28 | 5430 | 0.34 [0.16, 0.52] | 3.62 | 0.0003 | 226.25 | <0.00001 | 88 | 2.26 | 0.032 |
| General sleep problems | 27 | 5291 | 0.93 [0.67, 1.20] | 6.97 | <0.00001 | 409.23 | <0.00001 | 94 | 2.31 | 0.028 |
| Sleep quality | 3 | 155 | 0.24 [-1.05, 1.52] | 0.36 | 0.72 | 27.38 | <0.00001 | 93 | 0.41 | 0.752 |
| Sleep efficiency | 2 | 75 | -0.28 [-1.07, 0.51] | 0.68 | 0.49 | 2.63 | 0.11 | 62 |  |  |
| Sleep onset latency (min) | 6 | 787 | 0.81 [0.59, 1.02] | 7.33 | <0.00001 | 7.82 | 0.17 | 36 | 0.94 | 0.198 |
| Sleep duration (min) | 6 | 766 | -0.32 [-0.74, 0.11] | 1.47 | 0.14 | 31.78 | <0.00001 | 84 | 0.08 | 0.939 |
| Restorative value of sleep | 2 | 91 | 0.13 [-0.96, 1.23] | 0.24 | 0.81 | 5.23 | 0.02 | 81 |  |  |
| aIn order to reduce the heterogeneity between effect sizes, data already reported in similar measures were not included. This decision led to the exclusion of the data provided in two studies 23,51 for subjective night awakenings, as well as to the division of sleep duration and sleep onset delay in two distinct variables: sleep duration and sleep onset delay reported as a score in questionnaire or as a length (in min). | | | | | | | | | | |

**Table 3. Summary of the results of the meta-analysis with objective sleep parameters.**

|  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  |  |  |  |  | **Heterogeneity** | | | **Egger's test** | |
| **Sleep parameter** | **k** | **N** | **SMD (95% CI)** | **Z** | **p** | **Q** | **p** | **I2** | **t** | **p** |
| ***Polysomnography*** | | | | | | | | | | |
| Total sleep time | 8 | 247 | -0.90 [-1.51, -0.30] | 2.93 | 0.003 | 29.76 | 0.0001 | 76 | 0.35 | 0.734 |
| Sleep onset latency | 7 | 211 | 0.53 [0.21, 0.86] | 3.26 | 0.001 | 7.06 | 0.32 | 15 | 0.24 | 0.818 |
| Stage 1 sleep | 8 | 247 | 0.48 [0.06, 0.90] | 2.25 | 0.02 | 15.67 | 0.03 | 55 | 0.37 | 0.723 |
| Stage 2 sleep | 8 | 247 | 0.12 [-0.50, 0.73] | 0.37 | 0.71 | 33.06 | <0.0001 | 79 | 1.15 | 0.292 |
| Slow wave sleep | 8 | 247 | -0.15 [-0.83, 0.53] | 0.43 | 0.66 | 41.08 | <0.00001 | 83 | 0.56 | 0.596 |
| REM latency | 7 | 211 | -0.03 [-0.48, 0.42] | 0.13 | 0.90 | 13.79 | 0.03 | 56 | 0.42 | 0.689 |
| REM | 9 | 273 | -0.88 [-1.56, -0.21] | 2.56 | 0.01 | 46.07 | <0.00001 | 83 | 1.47 | 0.182 |
| Sleep efficiency | 7 | 238 | -1.20 [-1.98, -0.41] | 2.99 | 0.003 | 40.31 | <0.00001 | 85 | 1.76 | 0.138 |
| Wake time | 7 | 211 | 0.49 [0.11, 0.87] | 2.53 | 0.01 | 9.67 | 0.14 | 38 | 0.75 | 0.485 |
| ***Actigraphy*** | | | | | | | | | | |
| Sleep onset latency | 5 | 276 | 0.80 [0.55, 1.05] | 6.23 | <0.00001 | 2.49 | 0.65 | 0 | 0.36 | 0.745 |
| True sleep | 6 | 301 | -0.04 [-0.37, 0.29] | 0.24 | 0.81 | 9.65 | 0.09 | 48 | 1.74 | 0.156 |
| Assumed sleep time | 2 | 144 | -0.14 [-0.47, 0.20] | 0.80 | 0.42 | 0.05 | 0.83 | 0 |  |  |
| Actual wake time | 4 | 237 | 0.12 [-0.14, 0.38] | 0.92 | 0.36 | 2.03 | 0.57 | 0 | 0.45 | 0.697 |
| Sleep efficiency | 6 | 312 | -0.16 [-0.54, 0.22] | 0.82 | 0.41 | 13.20 | 0.02 | 62 | 0.91 | 0.416 |

**Table 4. Summary of the results of the post hoc analysis with polysomnographic parameters.**

|  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  |  |  |  |  | **Heterogeneity** | | | **Egger's test** | |
| **Sleep parameter** | **k** | **N** | **SMD (95% CI)** | **Z** | **p** | **Q** | **p** | **I2** | **t** | **p** |
| Total sleep time | 5 | 178 | -1.12 [-1.57, -0.66] | 4.76 | <0.00001 | 7.10 | 0.13 | 44 | 1.32 | 0.278 |
| Sleep onset latency | 5 | 178 | 0.56 [0.21, 0.90] | 3.12 | 0.002 | 4.68 | 0.32 | 15 | 0.95 | 0.413 |
| Stage 1 sleep | 5 | 178 | 0.20 [-0.11, 0.52] | 1.28 | 0.20 | 2.25 | 0.69 | 0 | 4.02 | 0.028 |
| Stage 2 sleep | 5 | 178 | 0.03 [-0.68, 0.73] | 0.07 | 0.94 | 18.55 | 0.001 | 78 | 0.35 | 0.747 |
| Slow wave sleep | 5 | 178 | -0.17 [-1.14, 0.80] | 0.35 | 0.73 | 33.61 | <0.00001 | 88 | 0.40 | 0.714 |
| REM latency | 5 | 178 | 0.01 [-0.57, 0.58] | 0.02 | 0.98 | 12.42 | 0.01 | 68 | 0.68 | 0.545 |
| REM | 6 | 204 | -0.58 [-1.13, -0.02] | 2.04 | 0.04 | 16.49 | 0.006 | 70 | 0.04 | 0.972 |
| Sleep efficiency | 5 | 178 | -0.90 [-1.38, -0.42] | 3.67 | 0.0002 | 8.14 | 0.09 | 51 | 0.44 | 0.689 |
| Wake time | 5 | 178 | 0.57 [0.10, 1.04] | 2.38 | 0.02 | 8.30 | 0.08 | 52 | 0.59 | 0.597 |

**FIGURES CAPTIONS**

Figure 1. PRISMA flow chart.

Footnote : \*reasons for exclusion for each paper are reporte in the Supplemental material 2

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