**Vitamin D and COVID-19 disease: don’t believe everything you read in the papers!**

**Reply to Dr William B. Grant**

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To the editor:

We read with interest the letter(1) from William B Grant, director of the Sunlight, Nutrition and Health Research Centre, San Francisco, commenting on our narrative review of the evidence around vitamin D and COVID-19 disease(2). Dr Grant suggests that confounding and reverse causality, at least as addressed in studies to date, do not account for associations between vitamin D and COVID-19 disease, where these have been observed. His conclusion is that the majority of the Bradford Hill criteria(3) for causality have been met. In support of these arguments, he cites a recent systematic review and meta-analysis bringing together studies of vitamin deficiency, 25(OH)-vitamin D [25(OH)D] levels, COVID-19 susceptibility, severity and sequelae including mortality. In this paper, Kazemi et al.(4) demonstrate a number of associations between lower 25(OH)D levels and greater risk of these outcomes. Although meta-analysis of observational studies may be viewed as the highest level of observational evidence, the synthesised output can clearly only be as good as the source studies. i.e. a confounded relationship which is consistent will only be more apparent (but of course not more causal) in a meta-analysis(5).

In our review(2) we attempted to critically appraise the source studies available rather than simply report, or indeed synthesize, their results. It is clear from the characteristics of the component studies in Kazemi et al.’s analysis that they are highly heterogeneous in term of setting, design and methodology, generally of small sample size and also highly variable in ability to account for confounding and bias(4). The notion that one can simply talk about “studies adjusted for confounding” is somewhat optimistic given that, for example, in the 3 such adjusted studies of associations between vitamin D and COVID-19 positivity, confounding structures were variable and did not consistently include factors such as adiposity, physical activity, ethnicity or sun exposure, the key determinants of 25(OH)D level, factors likely linked in turn with general health, socioeconomic status, lifestyle, comorbidity, and thus potentially with COVID-19 susceptibility, severity or sequelae. That the effect size appears (for this outcome in this study at least) similar to that from the unadjusted studies, suggests that the covariates incorporated did not explain the effect, but of course, by definition, does not remove the possibility of residual confounding as a result of unmeasured factors(5). Thus comparable effect sizes between “adjusted” and “unadjusted” studies may just as readily indicate an inadequate confounding structure as unconfounded associations.

Dr Grant also cites a recent study by Smolders et al.(6), which elegantly demonstrates the temporal change in 25(OH)D, TNF and IL-8 concentrations in response to an infusion of lipopolysaccharide (LPS) in human volunteers. Consistent with the studies cited in our review, during the infusion the inflammatory cytokines rose and 25(OH)D levels decreased. Following cessation of the infusion the concentrations tended to return to pre-infusion levels over the next 24-hours. Dr Grant suggests that because this is much quicker than the gap between admission and vitamin D testing in the majority of studies, the association between low 25(OH)D and COVID-19 cannot be due to reverse causation. However, COVID-19 infection, as is sadly all too apparent, produces an inflammatory reaction of rather longer duration than the 3 hours’ LPS infusion in this study and therefore will be expected to result in longer term reduction in 25(OH)D, given ample scope for reverse causation(7).

A superficial review of the observational data clearly can support associations between low vitamin D status and increased risk of COVID-19 disease and its sequelae. A detailed, critical appraisal of the data, as we undertook in our article, reveals a much more complex picture with the vast majority of these studies small, of low quality and prone to bias and reverse causation(8). The largest observational studies in UK Biobank, are limited by the 25(OH)D measures being obtained around a decade prior to COVID-19 diagnosis(9). Definitive randomised controlled trials are lacking and the one well conducted trial undertaken to date has shown no effect, albeit in a relatively small population(10). We do not view the evidence-base as conclusive one way or the other - the underlying biology, in terms of 1,25(OH)2D, as Dr Grant suggests, and as we documented in our review, would be compatible with an effect of vitamin D on COVID-19(2). Additionally a large meta-analysis of supplementation for reduction in (non-COVID-19) respiratory infections also supports the possibility of an effect on this outcome(11). We are not convinced however that the criteria of Bradford Hill are sufficiently met(3): there is marked variation in effect size and inconsistency between studies; specificity is absolutely not met given the vast number of conditions purported to be associated with 25(OH)D; temporality is largely uncharacterised; a biological gradient/dose response uncertain; biological plausibility and coherence are partly supported in the context of 1,25(OH)2D but with uncertainty as to how much supplementation with inactive vitamin D will actually change this system; and convincing experimental evidence from randomised trials is awaited(2, 4, 12).

The last Bradford Hill criterion of analogy is also problematic given the plethora of examples where low 25(OH)D has been associated with outcomes and then not substantiated in large randomised controlled trials or Mendelian randomisation studies(12-14). Indeed a recent genetic causal analyses in 443,734 individuals of European ancestry demonstrated no association between genetically determined 25(OH)D and COVID-19 susceptibility, hospitalisation or severe disease(15). A limitation of this approach is the inability to take account of overt vitamin D deficiency, reinforcing the need for well conducted randomised controlled trials or at least large-scale high quality observational studies using state-of-the-art analytical techniques such as propensity score matching and instrumental variable analyses to reduce the risk of residual confounding(16).

In conclusion we would strongly counsel against a superficial appraisal of the evidence which accepts source studies without critical review and ignores issues of small sample size, heterogeneity, residual confounding (or no adjustment at all) and reverse causality, which the arguments of Dr Grant notwithstanding, are patently evident as caveats within the current literature.

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**Compliance with Ethical Standards**

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**Conflict of Interest**

NCH reports personal fees, consultancy, lecture fees and honoraria from Alliance for Better Bone Health, AMGEN, MSD, Eli Lilly, Servier, Shire, UCB, Consilient Healthcare, Kyowa Kirin and Internis Pharma, outside the submitted work. CC reports personal fees, consultancy, lecture fees and honoraria from Alliance for Better Bone Health, Amgen, Eli Lilly, GSK, Medtronic, Merck, Novartis, Pfizer, Roche, Servier, Takeda and UCB. ZRE reports no conflicts of interest.

**Ethical approval/ informed consent**

There was no direct involvement of patients or participants in this letter. Ethics review and informed consent are not applicable.

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