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Common variants in breast cancer risk loci predispose to distinct tumor subtypes

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36 **Abstract**

37 **Background:** Genome-wide association studies (GWAS) have identified multiple common breast cancer suscepti-
 38 bility variants. Many of these variants have differential associations by estrogen receptor (ER) status, but how these
 39 variants relate with other tumor features and intrinsic molecular subtypes is unclear.

40 **Methods:** Among 106,571 invasive breast cancer cases and 95,762 controls of European ancestry with data on 173
 41 breast cancer variants identified in previous GWAS, we used novel two-stage polytomous logistic regression models
 42 to evaluate variants in relation to multiple tumor features (ER, progesterone receptor (PR), human epidermal growth
 43 factor receptor 2 (HER2) and grade) adjusting for each other, and to intrinsic-like subtypes.

44 **Results:** Eighty-five of 173 variants were associated with at least one tumor feature (false discovery rate < 5%), most
 45 commonly ER and grade, followed by PR and HER2. Models for intrinsic-like subtypes found nearly all of these variants
 46 (83 of 85) associated at $p < 0.05$ with risk for at least one luminal-like subtype, and approximately half (41 of 85) of the
 47 variants were associated with risk of at least one non-luminal subtype, including 32 variants associated with triple-
 48 negative (TN) disease. Ten variants were associated with risk of all subtypes in different magnitude. Five variants were
 49 associated with risk of luminal A-like and TN subtypes in opposite directions.

50 **Conclusion:** This report demonstrates a high level of complexity in the etiology heterogeneity of breast cancer sus-
 51 ceptibility variants and can inform investigations of subtype-specific risk prediction.

52 **Keywords:** Breast cancer, Etiologic heterogeneity, Genetic predisposition, Common breast cancer susceptibility
 53 variants

54 **Introduction**

55 Breast cancer represents a heterogeneous group of diseases
 56 with different molecular and clinical features[1]. Clinical
 57 assessment of estrogen receptor (ER), progesterone
 58 receptor (PR), human epidermal growth factor receptor
 59 2 (HER2) and histological grade are routinely determined
 60 to inform treatment strategies and prognostication[2].
 61 Combined, these tumor features define five intrinsic-
 62 like subtypes (i.e., luminal A-like, luminal B-like/HER2-
 63 negative, luminal B-like/HER2-positive, HER2-positive/
 64 non-luminal, and triple negative) that are correlated with
 65 intrinsic subtypes defined by gene expression panels[2, 3].
 66 Most known breast cancer risk or protective factors are
 67 related to luminal or hormone receptor (ER or PR) posi-
 68 tive tumors, whereas less is known about the etiology of
 69 triple-negative (TN) tumors, an aggressive subtype[4, 5].

70 Breast cancer genome-wide association studies
 71 (GWAS) have identified over 170 common susceptibility
 72 variants, most of them single nucleotide polymorphisms
 73 (SNPs), of which many are differentially associated with
 74 ER-positive than ER-negative disease[6–8]. These include
 75 20 variants that primarily predispose to ER-negative or
 76 TN disease[7, 8]. However, few studies have evaluated

variant associations with other tumor features, or simul- 77
 78 taneously studied multiple, correlated tumor markers
 79 to identify source(s) of etiologic heterogeneity[7, 9–13].
 80 We recently developed a two-stage polytomous logistic
 81 regression method that efficiently characterizes etiologic
 82 heterogeneity while accounting for tumor marker corre-
 83 lations and missing tumor data[14, 15]. This method can
 84 help describe complex relationships between susceptibil-
 85 ity variants and multiple tumor features, helping to clar-
 86 ify breast cancer subtype etiologies and increasing the
 87 power to generate more accurate risk estimates between
 88 susceptibility variants and less common subtypes. We
 89 recently demonstrated the power of this method in a
 90 GWAS to identify novel breast cancer susceptibility
 91 accounting for tumor heterogeneity[15].

92 In this report, we sought to expand our understanding
 93 of etiologic heterogeneity among breast cancer subtypes,
 94 by applying the two-stage polytomous logistic regression
 95 methodology to a large study population from the Breast
 96 Cancer Association Consortium (BCAC) for detailed
 97 characterization of risk associations with 173 breast can-
 98 cer risk variants identified by GWAS[6, 7] by tumor sub-
 99 types defined by ER, PR, HER2 and tumor grade.

100 Methods

101 Study population and genotyping

102 The study population and genotyping are described in
 103 previous publications[6, 7] and in the Additional file 3:
 104 Methods. We included invasive cases and controls from
 105 81 BCAC studies with genotyping data from two Illumina
 106 genome-wide custom arrays, the iCOGS and OncoArray
 107 (106,571 cases (OncoArray: 71,788; iCOGS: 34,783) and
 108 95,762 controls (OncoArray: 58,134; iCOGS: 37,628);
 109 Additional file 1: Table S1). All subjects in the study
 110 population were female and of European ancestry, with
 111 European ancestry determined by ancestry informative
 112 GWAS markers as previously described [6]. We evaluated
 113 173 breast cancer risk variants that were identified in or
 114 replicated by prior BCAC analyses to be associated with
 115 breast cancer risk at a p-value threshold $p < 5.0 \times 10^{-8}$
 116 [6, 7]. Most of these variants (n=153) were identified
 117 because of their association with risk of overall breast
 118 cancer, and a small number of variants (n=20) were
 119 identified because of their association specific to ER-neg-
 120 ative breast cancer (Additional file 1: Table S2). These 173
 121 variants have not previously been simultaneously investi-
 122 gated for evidence of tumor heterogeneity with multiple
 123 tumor markers[6, 7, 15, 16]. Genotypes for the variants
 124 marking the 173 susceptibility loci were determined by
 125 genotyping with the iCOGS and the OncoArray arrays
 126 and imputation to the 1000 Genomes Project (Phase 3)
 127 reference panel.

128 Statistical analysis

129 An overview of the analytic strategy is shown in Fig. 1 and
 130 a detailed discussion of the statistical methods, including
 131 the two-stage polytomous logistic regression, are pro-
 132 vided in the Additional file 3: Methods and elsewhere[14,
 133 15]. Briefly, we used two-stage polytomous regression
 134 models that allow modelling of genetic association of
 135 breast cancer accounting for underlying heterogeneity in
 136 associations by combinations of multiple tumor markers
 137 using a parsimonious decomposition of subtype-specific
 138 case-control odds-ratio parameters in terms of marker-
 139 specific case-case odd-ratio parameters[14, 15]. We
 140 introduced further parsimony by using mixed-effect for-
 141 mulation of the model that allows ER-specific case-case
 142 parameters to be treated as fixed and similar parameters
 143 for other markers (PR, HER2 and grade (as an ordinal
 144 variable)) as random. We used an expectation-maxi-
 145 mization (EM) algorithm[17] for parameter estimation
 146 under this model to account for missing data in tumor
 147 characteristics.

148 Our primary aim was to identify which of 173 known
 149 breast cancer susceptibility variants showed heteroge-
 150 nous risk associations by ER-, PR- and HER2-status and
 151 tumor grade. This was tested using a global heterogeneity

152 test by ER, PR, HER2 and/or grade, with a mixed-effect
 153 two-stage polytomous model (model 1), fitted separately
 154 for each variant. The global null hypothesis was that there
 155 was no difference in risk of breast cancer associated with
 156 the variant genotype across any of the tumor features
 157 being evaluated. We accounted for multiple testing (173
 158 tests, one for each of variant) of the global heterogene-
 159 ity test using a false discovery rate (FDR) <5% under the
 160 Benjamini-Hochberg procedure[18].

161 For the variants showing evidence of global heteroge-
 162 neity after FDR adjustment, we further evaluated which
 163 of the tumor features contributed to the heterogeneity
 164 by fitting a fixed-effects two-stage model (model 2) that
 165 simultaneously tested for associations with each tumor
 166 feature (this model was fitted for each variant separately).
 167 We used a threshold of $p < 0.05$ for marker-specific tumor
 168 heterogeneity tests to describe which specific tumor
 169 marker(s) contributed to the observed heterogeneity,
 170 adjusting for the other tumor markers in the model. This
 171 p-value threshold was used only for descriptive purposes,
 172 as the primary hypotheses were tested using the FDR-
 173 adjusted global test for heterogeneity described above.

174 We conducted additional analyses to explore evidence
 175 of heterogeneity. We fitted a fixed-effect two-stage model
 176 (model 3) to estimate case-control odd ratios (ORs) and
 177 95% confidence intervals (CI) between the variants and
 178 five intrinsic-like subtypes defined by combinations of
 179 ER, PR, HER2 and grade: (1) luminal A-like (ER+ and/
 180 or PR+, HER2-, grade 1 or 2); (2) luminal B-like/HER2-
 181 negative (ER+ and/or PR+, HER2-, grade 3); (3) lumi-
 182 nal B-like/HER2-positive (ER+ and/or PR+, HER2+);
 183 (4) HER2-positive/non-luminal (ER- and PR-, HER2+),
 184 and (5) TN (ER-, PR-, HER2-). We also fitted a fixed-
 185 effect two-stage model to estimate case-control ORs and
 186 95% confidence intervals (CI) with tumor grade (model
 187 4; defined ordinally as grade 1, grade 2, and grade 3) for
 188 the variants associated at $p < 0.05$ only with grade in case-
 189 case comparisons from model 2.

190 To help describe sources of heterogeneity from dif-
 191 ferent tumor characteristics in models 2 and 3, we per-
 192 formed cluster analyses based on Euclidean distance
 193 calculated from the absolute z-statistics that were
 194 estimated by the individual marker-specific tumor
 195 heterogeneity tests (model 2) and the case-control
 196 associations with risk of intrinsic-like subtypes (model
 197 3). The clusters were used only for presentation pur-
 198 poses and were not intended to suggest strictly defined
 199 categories, nor are they intended to suggest the vari-
 200 ants are associated with tumor markers through simi-
 201 lar biological mechanisms. Clustering was performed
 202 in R using the function Heatmap as implemented
 203 by the package “Complex Heatmap” version 3.1[19].
 204 Additional details for calculating Euclidean distance



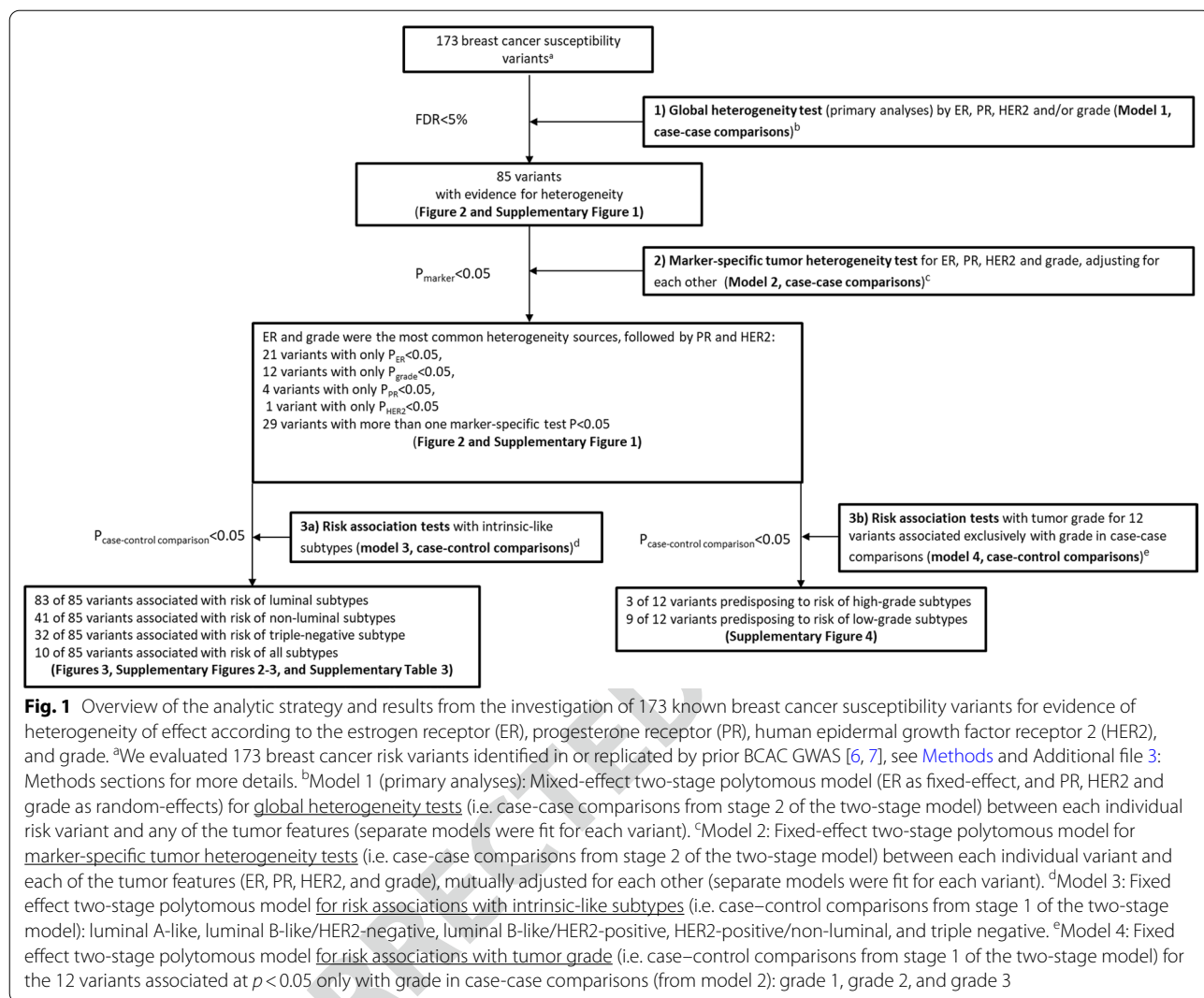


Fig. 1 Overview of the analytic strategy and results from the investigation of 173 known breast cancer susceptibility variants for evidence of heterogeneity of effect according to the estrogen receptor (ER), progesterone receptor (PR), human epidermal growth factor receptor 2 (HER2), and grade. ^aWe evaluated 173 breast cancer risk variants identified in or replicated by prior BCAC GWAS [6, 7], see Methods and Additional file 3: Methods sections for more details. ^bModel 1 (primary analyses): Mixed-effect two-stage polytomous model (ER as fixed-effect, and PR, HER2 and grade as random-effects) for global heterogeneity tests (i.e. case-case comparisons from stage 2 of the two-stage model) between each individual risk variant and any of the tumor features (separate models were fit for each variant). ^cModel 2: Fixed-effect two-stage polytomous model for marker-specific tumor heterogeneity tests (i.e. case-case comparisons from stage 2 of the two-stage model) between each individual variant and each of the tumor features (ER, PR, HER2, and grade), mutually adjusted for each other (separate models were fit for each variant). ^dModel 3: Fixed effect two-stage polytomous model for risk associations with intrinsic-like subtypes (i.e. case-control comparisons from stage 1 of the two-stage model): luminal A-like, luminal B-like/HER2-negative, luminal B-like/HER2-positive, HER2-positive/non-luminal, and triple negative. ^eModel 4: Fixed effect two-stage polytomous model for risk associations with tumor grade (i.e. case-control comparisons from stage 1 of the two-stage model) for the 12 variants associated at $p < 0.05$ only with grade in case-case comparisons (from model 2): grade 1, grade 2, and grade 3

205 using absolute z-statistics are provided in Additional
 206 file 3: Methods.

207 We performed sensitivity analyses, in which we esti-
 208 mated the ORs and 95% CI between the variants and
 209 the intrinsic-like subtypes by implementing a standard
 210 polytomous model that defined the intrinsic-like sub-
 211 types using only the available tumor markers data (not
 212 using the EM algorithm to account for missing data in
 213 tumor markers). For all analyses we analyzed OncoAr-
 214 ray and iCOGS array data separately, adjusting for the
 215 first 10 principal components for ancestry-informative
 216 variants, and then meta-analyzed the results.

217 **Results**

218 The mean (SD) ages at diagnosis (cases) and enroll-
 219 ment (controls) were 56.6 (12.2) and 56.4 (12.2) years,
 220 respectively. Among cases with information on the cor-
 221 responding tumor marker, 81% were ER-positive, 68%

222 PR-positive, 83% HER2-negative and 69% grade 1 or
 223 2 (Table 1; see Additional file 1: Table S1 for details by
 224 study). Additional file 1: Table S3 shows the correlation
 225 between the tumor markers. ER was positively corre-
 226 lated with PR ($r = 0.61$) and inversely correlated with
 227 HER2 ($r = -0.16$) and grade ($r = -0.39$). The most com-
 228 mon intrinsic-like subtype was luminal A-like (54%),
 229 followed by TN (14%), luminal B-like/HER2-negative
 230 (13%), Luminal B-like/HER2-positive (13%) and HER2-
 231 positive/non-luminal (6%; Table 1). These frequencies
 232 varied across BCAC studies due to the studies being
 233 diverse in both design and country of origin (Additional
 234 file 1: Table S1). Notably, there is little population-based
 235 data on the frequencies of intrinsic-like subtypes [20,
 236 21]. The overall frequencies in our study population are
 237 generally similar to those reported by SEER for non-
 238 Hispanic white females and the Scottish cancer registry
 239 [20, 21]; however, given the diverse sources of our data

240 they are not directly comparable to country specific
241 cancer registries.

242 Figure 1 shows an overview of the analytic strategy and
243 results from three main analyses performed separately
244 for each variant: 1) global test for heterogeneity by all
245 tumor markers (model 1; primary hypothesis), 2) marker-
246 specific tumor test for heterogeneity for each marker,
247 adjusting for the others (model 2), and 3) estimation of
248 case-control ORs (95% CIs) by intrinsic-like subtypes
249 (model 3) and by grade (model 4).

250 **Global test for heterogeneity by tumor markers (primary
251 hypothesis)**

252 Mixed-effects two-stage models (model 1) were fitted for
253 each of the 173 variants separately and included terms for
254 ER, PR, HER2 and grade to test for global heterogeneity
255 by any of the tumor features (case-case comparison). This
256 model identified 85 of 173 (49.1%) variants with evidence
257 of heterogeneity by at least one tumor feature (FDR < 5%;
258 Figs. 1, 2; Additional file 1: Fig. S1).

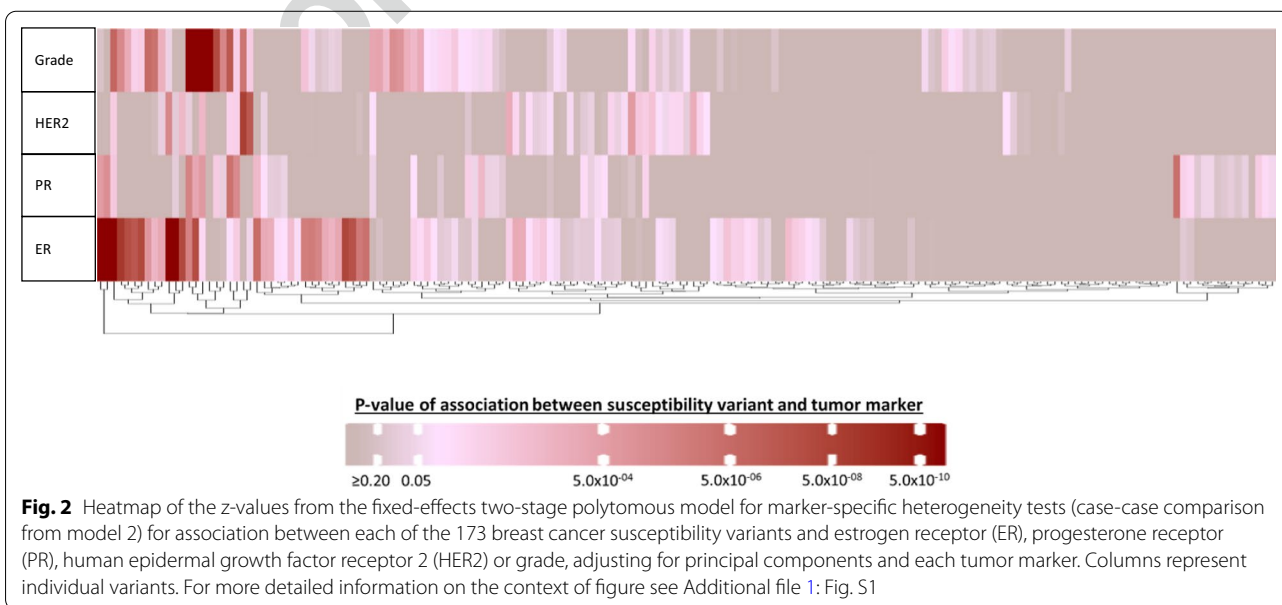
259 **Marker-specific tumor test for heterogeneity for each
260 marker, adjusting for other markers**

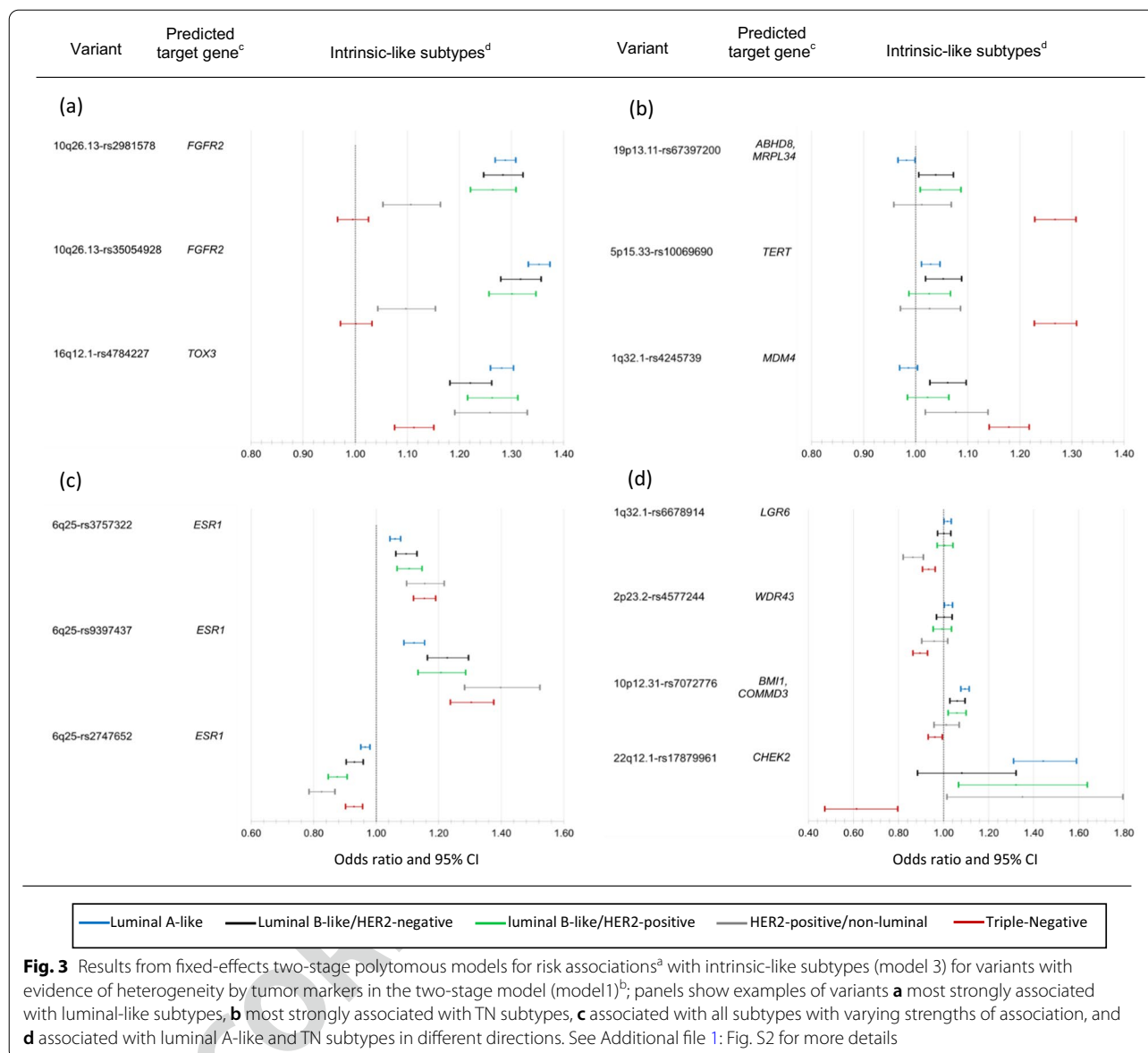
261 Fixed-effects two-stage models (model 2) were used to
262 test which of the correlated tumor markers was respon-
263 sible for the observed global heterogeneity (case-case
264 comparison). Figure 2 and Additional file 1: Fig. S1 show
265 results of these analyses clustered by case-case z-values
266 of associations between susceptibility variants and each
267 tumor marker for the 173 variants. For the 85 variants
268 with observed global heterogeneity, these analyses iden-
269 tified ER and grade as the two features that most often

270 contributed to the observed heterogeneity (45 and 33
271 variants had marker-specific $p < 0.05$ for ER and grade,
272 respectively), and 29 variants were associated with more
273 than one tumor feature (Figs. 1, 2, Additional file 1:
274 Fig. S1). Eighteen of these 85 variants showed no associa-
275 tions with any individual tumor marker at $p < 0.05$ (Fig. 2,
276 Additional file 1: Fig. S1). Twenty-one variants were
277 associated at $p < 0.05$ only with ER, 12 variants only with
278 grade, 4 variants only with PR and one variant only with
279 HER2 (Fig. 2, Additional file 1: Fig. S1, see footnotes).

280 **Estimation of case-control ORs (95% CIs) by intrinsic-like
281 subtypes (model 3)**

282 Fixed-effects two-stage models for intrinsic-like subtypes
283 (model 3) were fitted for each of the 85 variants with evi-
284 dence of global heterogeneity to estimate ORs (95% CIs)
285 for risk associations with each subtype (case-control com-
286 parisons). Additional file 1: Fig. S2 shows a summary of
287 these analyses for the 85 variants, clustered by case-con-
288 trol z-value of association between susceptibility variants
289 and breast cancer intrinsic-like subtypes, and Additional
290 file 2: Fig. S3 shows forest plots for associations with risk
291 by tumor subtypes. Nearly all (83 of 85) variants were
292 associated with risk ($p < 0.05$) for at least one luminal-like
293 subtype, and approximately half (41 of 85) of the variants
294 were associated with risk of at least one non-luminal sub-
295 type, including 32 variants that were associated with TN
296 disease (Fig. 1, Additional file 1: Fig. S2 footnote 'h'). Ten
297 variants were associated with risk of all subtypes (Fig. 1,
298 Additional file 1: Fig. S2 footnote 'j'). Below we describe
299 examples of groups of variants associated with different
300 patterns of associations with intrinsic subtypes (Fig. 3 a-d).





301 Two variants in linkage disequilibrium (LD, $r^2=0.73$)
 302 at 10q26.13 (rs2981578 and rs35054928) and 16q12.1-
 303 rs4784227 had the strongest evidence of association with
 304 risk of luminal-like subtypes (Fig. 3a, Additional file 1:
 305 Fig. S2). The two variants at 10q26.13 showed no evi-
 306 dence of associations with TN subtypes, and a weaker
 307 association with HER2-positive/non-luminal subtype. In
 308 contrast, 16q12.1-rs4784227 was strongly associated with
 309 risk for all luminal-like subtypes and, weaker so, with
 310 risk of HER2-positive/non-luminal and TN subtypes
 311 (Figs. 3a, Additional file 1: Fig. S2).

312 Three variants 19p13.11-rs67397200, 5p15.33-
 313 rs10069690 and 1q32.11-rs4245739 showed the strongest
 314 evidence of associations with risk of TN disease. All three

of these variants showed weaker or no evidence of associ-
 ations with risk of the other subtypes (Fig. 3b, Additional
 file 1: Fig. S2).

Two variants in low LD ($r^2=0.17$) at 6q25, rs9397437
 and rs3757322, and a third variant in 6q25, rs2747652,
 which was not in LD ($r^2<0.01$) with rs9397437 or
 rs3757322, showed strong evidence of being associ-
 ated with risk of all subtypes. rs9397437 and rs3757322
 were most strongly associated with risk of TN disease.
 rs2747652 was most strongly associated with risk of
 HER2-positive subtypes (Figs. 3c, Additional file 1:
 Fig. S2).

Five variants were associated with risk of lumi-
 nal A-like disease in an opposite direction to their

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Table 1 Distribution of estrogen receptor (ER), progesterone receptor (PR), human epidermal growth factor receptor 2 (HER2), and grade and the intrinsic-like subtypes among cases of invasive breast cancer in studies from the Breast Cancer Consortium Association

Tumor marker	N (%)
ER	
Negative	16,900 (19%)
Positive	70,030 (81%)
Unknown	19,641
PR	
Negative	24,283 (32%)
Positive	51,603 (68%)
Unknown	30,685
HER2	
Negative	47,693 (83%)
Positive	9,529 (17%)
Unknown	49,349
Grade	
1	15,583 (20%)
2	37,568 (49%)
3	24,382 (31%)
Unknown	29,038
Intrinsic-like subtypes	
Luminal A-like	27,510 (54%)
Luminal B-like/HER2-negative	6,804 (13%)
Luminal B-like/HER2-positive	6,511 (13%)
HER2-positive/non-luminal	2,797 (6%)
Triple-negative	7,178 (14%)
Unknown	55,771

Luminal A-like (ER + and/or PR +, HER2-, grade 1 & 2); Luminal B-like/HER2-negative (ER + and/or PR +, HER2-, grade 3); Luminal B-like/HER2-positive (ER + and/or PR +, HER2+); HER2-positive/non-luminal (ER- and PR-, HER2+), and triple-negative (ER-, PR-, HER2-)

association with risk of TN disease. 1q32.1-rs6678914, 2p23.2-rs4577244, and 19p13.11-rs67397200 had weaker evidence of associations with risk of luminal A-like disease compared to associations with risk of TN disease, and 10p12.31-rs7072776 and 22q12.1-rs17879961 (I157T) had stronger evidence of an association with risk of luminal A-like disease compared to their association with risk of TN disease (Fig. 3d, Additional file 1: Fig. S2, for rs67397200 see Fig. 3b).

Estimation of case-control ORs (95% CIs) by tumor grade (model 4)

Case-control associations by tumor grade for the 12 variants that were observed associated at $p < 0.05$ only with grade in case-case comparisons are shown in Additional file 2: Fig. S4. 13q13.1-rs11571833, 1p22.3-rs17426269

and 11q24.3-rs11820646 showed stronger evidence for predisposing to risk of high-grade subtypes, and the remaining variants showed stronger evidence for predisposing to risk of low-grade subtypes.

When limiting analyses to cases with intrinsic-like subtypes defined only by available tumor marker data, results from case-control analyses were similar, but less precise than results from the two-stage polytomous regression model using the EM algorithm to account for missing tumor marker data (Additional file 1: Table S4).

Discussion

This study demonstrates the extent and complexity of genetic etiologic heterogeneity among 173 breast cancer risk variants by multiple tumor characteristics, using novel methodology in the largest and the most comprehensive investigation conducted to date. We found compelling evidence that about half of the investigated breast cancer susceptibility loci (85 of 173 variants) predispose to tumors with different characteristics. We identified tumor grade, along with confirming ER status, as important determinants of etiologic heterogeneity. Associations with individual tumor features translated into differential associations with the risk of intrinsic-like subtypes defined by their combinations.

Many of the variants with evidence of global heterogeneity predisposed to risk of multiple subtypes, but with different magnitudes. For example, three variants identified in early GWAS for overall breast cancer, *FGFR2* (rs35054928 and rs2981578)[22, 23] and 8q24.21 (rs13281615)[22], were associated with luminal-like and HER2-positive/non-luminal subtypes, but not with TN disease. rs4784227 located near *TOX3*[22, 24] and rs62355902 located in a *MAP3K1*[22] regulatory element, were associated with risk of all five subtypes. Of the five variants found associated in opposite directions with luminal A-like and TN disease, we previously reported rs6678914 and rs4577244 to have opposite effects between ER-negative and ER-positive tumors[7]. rs17879961 (I157T), a likely causal[16] missense variant located in a *CHEK2* functional domain that reduces or abolishes substrate binding[25], was previously reported to have opposite directions of effects on lung adenocarcinoma and lung squamous cell carcinoma and for lung cancer between smokers and non-smokers[26, 27]. Moreover, the risk association of rs17879961 has been reported to vary across tissue locations/cell-types, as this variant has been associated with a higher risk of pancreatic ductal adenocarcinoma [28], chronic lymphocytic leukemia [29], and colorectal cancer [30], and also associated with a lower risk of aerodigestive squamous cell carcinoma [31] and ovarian cancer [32]. To our knowledge

rs67397200 and rs7072776 have not previously been shown to be associated with subtypes in opposite directions. In a prior breast cancer GWAS that applied the two-stage polytomous model for risk variant discovery we also identified five variants that were associated with risk of luminal A-like and TN disease in opposite directions [15]. Overall, these findings suggest that the same biological pathway has opposite effects on the susceptibility to different tumor types. This interpretation is supported by functional characterization of rs36115365, a variant on 5p15.33 which was found to have similar cis-regulatory effects on TERT in multiple cancers cell lines from different cancers, but was associated with a higher risk of pancreatic and testicular cancer and a lower risk of lung cancer [33]. Alternatively, a causal variant may differently influence cis-gene regulation and/or alter different biological pathways depending on the cell or tissue of origin [34]. Further studies of these variants are required to clarify the biological mechanisms for these apparent cross-over effects.

In prior ER-negative GWAS, we identified 20 variants that predispose to ER-negative disease, of which five variants were only or most strongly associated with risk of TN disease (rs4245739, rs10069690, rs74911261, rs11374964, and rs67397200)[7, 8]. We confirmed these five variants to be most strongly associated with TN disease. The remaining previously identified 15 variants all showed associations with risk of non-luminal subtypes, especially TN disease, and for all but four variants (rs17350191, rs200648189, rs6569648, and rs322144) evidence of global heterogeneity was observed.

Little is known regarding PR and HER2 as sources of etiologic heterogeneity independent of ER status. Of the four variants that showed evidence of heterogeneity only according to PR, rs10759243[6, 35], rs11199914[36] and rs72749841[6] were previously found primarily associated with risk of ER-positive disease, and rs10816625 was found to be associated with risk of ER-positive/PR-positive tumors, but not other ER/PR combinations[12]. rs10995201 was the only variant found in case-case comparisons to be solely associated with HER2 status, although the evidence was not strong, requiring further confirmation. Previously, rs10995201 showed no evidence of being associated with ER status[37]. Most variants associated with PR or HER2, had not been investigated for PR or HER2 heterogeneity while adjusting for ER[9–13]. We previously reported rs10941679 to be associated with PR-status, independent of ER, and also with grade[10]. We also found suggestive evidence of PR-specific heterogeneity for 16q12-rs3803662[13], which is in high LD ($r^2=0.78$) with rs4784227 (*TOX3*), a variant strongly associated with PR status. Our findings for rs2747652 are also consistent with a prior BCAC

fine-mapping analysis across the *ESR1* locus, which found rs2747652 to be associated with risk of the HER2-positive/non-luminal subtype and high grade independent of ER[9]. rs2747652 overlaps an enhancer region and is associated with reduced *ESR1* and *CCDC170* expression[9].

Histologic grade is a composite of multiple tumor characteristics including mitotic count, nuclear pleomorphism, and degree of tubule or gland formation, therefore susceptibility variants associated with tumor grade could affect multiple biological pathways [38]. Evidence from comparisons of tumor morphology and genomic and molecular alterations suggest that tumor grade is likely a ‘stable’ tumor feature and does not progress from low- to high-grade [39–42], thus the variants associated with grade are likely not associated with grade progression. Among the 12 variants identified with evidence of heterogeneity by grade only, rs17426269, rs11820646, and rs11571833 were found to be most strongly associated with risk of grade 3 disease. rs11571833 lies in the *BRCA2* coding region and produces a truncated form of the protein[43] and has been shown to be associated with both risk of TN disease and risk of serous ovarian tumors, both of which tend to be high-grade[44]. To our knowledge, rs17426269 and rs11820646 have not been investigated in relation to grade heterogeneity. The remaining 9 variants were all more strongly associated with grade 1 or grade 2 disease. Six of these variants were previously reported to be associated primarily with ER-positive disease[6, 36, 45, 46], highlighting the importance of accounting for multiple tumor characteristics to better illuminate heterogeneity sources.

We identified 18 variants with evidence of global heterogeneity (FDR < 5%), but no significant (marker-specific $p < 0.05$) associations with any of the individual tumor characteristic(s). This is likely explained by the fact that the test for association with specific tumor markers using fixed-effects models are less powerful than mixed-effects models used to test the primary hypothesis of global heterogeneity by any tumor marker[14].

To help describe and visualize the strength of the evidence for common heterogeneity patterns, we performed clustered analyses of z-values for tumor marker-specific heterogeneity tests and case-control associations with risk of intrinsic-like subtypes. Because they are based on z-values, these clusters reflect differences in sample size and statistical power to detect associations between variants and specific tumor subtypes. Thus, clusters should not be interpreted as strictly defined categories.

A major strength of our study is our large sample size of over 100,000 breast cancer cases with tumor marker information, and a similar number of controls, making this the largest, most comprehensive breast cancer



501 heterogeneity investigation. Our application of the two-
 502 stage polytomous logistic regression enabled adjusting
 503 for multiple, correlated tumor markers and accounting
 504 for missing tumor marker data. This is a more powerful
 505 and efficient modeling strategy for identifying heteroge-
 506 neity sources among highly correlated tumor markers,
 507 compared with standard polytomous logistic regres-
 508 sion[14, 15]. In simulated and real data analyses, we have
 509 demonstrated that in the presence of heterogenous asso-
 510 ciations across subtypes, the two-stage model is more
 511 powerful than polytomous logistic regression for detect-
 512 ing heterogeneity. Moreover, we have demonstrated that
 513 in the presence of correlated markers, the two-stage
 514 model, incorporating all markers simultaneously, has
 515 much better ability to distinguish the true source(s) of
 516 heterogeneity compared to testing for heterogeneity by
 517 analysis of one marker at a time[14, 15]. In prior analyses,
 518 we showed that the two-stage polytomous regression is a
 519 powerful approach to identify susceptibility variants that
 520 display tumor heterogeneity[15]. Notably, in this prior
 521 investigation we excluded the genomic regions in which
 522 the 173 variants that were investigated in this work are
 523 located[15].

524 Our study also has some limitations. First, many
 525 breast cancer cases from studies included in this report
 526 had missing information on one or more tumor charac-
 527 teristics. ER tumor status data was available for 81%
 528 of cases, but missing data for the other tumor markers
 529 ranged from 27 to 46%. To address this limitation, we
 530 implemented an EM algorithm that allowed a powerful
 531 analysis to incorporate cases with missing tumor charac-
 532 teristics under the assumption that tumor characteristics
 533 are *missing at random* (MAR), i.e., the underlying reason
 534 for missing data may depend on observed tumor mark-
 535 ers or/and covariate values, but not on the missing val-
 536 ues themselves[47]. If this assumption is violated it can
 537 lead to an inflated type-one error[14]. However, in the
 538 context of genetic association testing, the missingness
 539 mechanism would also need to be related to the genetic
 540 variants under study, which is unlikely. The 88 variants
 541 that did not meet the p-value threshold for significant
 542 heterogeneity in the global test, are likely to represent a
 543 combination of variants that are associated with risk of
 544 all investigated tumor subtypes with similar effects and
 545 variants for which we lacked power to detect evidence of
 546 global heterogeneity due to weak effect sizes or uncom-
 547 mon allele frequencies. In addition, our study focused on
 548 investigating ER, PR, HER2, and grade as heterogeneity
 549 sources; future studies with more detailed tumor charac-
 550 terization could reveal additional etiologic heterogeneity
 551 sources.

Conclusion

Our findings provide insights into the complex etiologic heterogeneity patterns of common breast cancer susceptibility loci. These findings may inform future studies, such as fine-mapping and functional analyses to identify the underlying causal variants, clarifying biological mechanisms that drive genetic predisposition to breast cancer subtypes. Moreover, these analyses provide precise estimates of relative risk for different intrinsic-like subtypes that could improve the discriminatory accuracy of subtype-specific polygenic risk scores [48].

Abbreviations

GWAS: Genome-wide association studies; ER: Estrogen receptor; PR: Progesterone receptor; HER2: Human epidermal growth factor receptor 2; SNPs: Single nucleotide polymorphisms; FDR: False discovery rate; TN: Triple-negative; BCAC: Breast Cancer Association Consortium; EM: Expectation-maximization; OR: Odds ratios; 95% CI: 95% Confidence interval; LD: Linkage disequilibrium.

Supplementary Information

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Additional file 1. Figures S1 and S2 and Table S1-S4. This file contains supplementary figures 1-2 and supplementary tables 1-4. In reply to Reviewer #1, we have added the distribution of the intrinsic-like subtypes by studies.

Additional file 2. Figures S3 and S4. This file contains supplementary figures S3 and S4.

Additional file 3. Methods. This file contains the supplementary methods. In reply to Reviewer #2, we have added a more detailed description of the clustering methods.

Additional file 4. Funding and Acknowledgement. This file contains the additional funding not included in the main text, the acknowledgments, and the names of the people in the collaboration groups.

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Availability of data and materials

The datasets generated and/or analyzed during the current study are part of the Breast Cancer Association Consortium and would be available with the appropriate permissions, including an application process and appropriate data transfer agreements.

Declarations**Ethics approval and consent to participate**

All the studies included in these analyses were approved by local IRBs.

Consent for publication

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Competing interests

The authors have no competing interests to declare.

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