# Message Passing-Aided Joint Data Detection and Estimation of Nonlinear Satellite Channels 

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#### Abstract

Satellite communication is capable of supporting seamless global coverage. However, owing to the reliance on limited-duration solar power, the high power amplifier (HPA) is often driven close to its saturation point, which leads to severe nonlinear distortion in satellite channels. Thus, mitigating the effect of the nonlinear distortion becomes essential for reliable communications. In this paper, we propose an efficient joint channel estimation and data detection method based on message passing within the associated factor graph modelling the HPA employed in nonlinear satellite channels. Then, we develop a combined belief propagation and mean field (BP-MF) method to cope with the hard constraints and dense short loops on the factor graph. In particular, the parametric message updating expressions relying on the canonical parameters are derived in the symbol detection part. To alleviate the impact of dense loops, we reformulate the system model into a compact form within the channel estimation part and then reconstruct a loop-free subgraph associated with vector-valued nodes to guarantee convergence. Furthermore, the proposed BP-MF method is also extended to the realistic scenario of having unknown noise variance. To further reduce the computational complexity of the largescale matrix inversion of channel estimation, the generalized approximate message passing (GAMP) algorithm is employed to decouple the vector of channel coefficient estimation into a series of scalar estimations. Simulation results show that the proposed methods outperform the state-of-the-art benchmarks both in terms of bit error rate performance and channel estimation accuracy.


Index Terms-Nonlinear satellite channel, Volterra series, joint channel estimation and data detection, mean field approximation, generalized approximate message passing.

## I. Introduction

Satellite communication plays an essential role in next generation wireless networks by providing seamless global

[^0]coverage [1]-[5]. The onboard high power amplifier (HPA) is a key component of the transponder, which relies on limitedduration solar charging. To achieve a high power efficiency, the HPA is typically operated near its saturation point, which leads to nonlinear distortion.

The nonlinearity of the HPA has been popularly characterized by a frequency-independent memoryless model [6], [7]. The most widely used memoryless HPA model is described by the AM/AM and AM/PM functions [8]. Nonetheless, the input multiplexing (IMUX) and output multiplexing (OMUX) filters placed before and after the HPA introduce memory into these nonlinear satellite channels. Several models have been proposed for characterizing nonlinear channels exhibiting memory, such as the polynomial model [9], the Wiener model [10], the Hammerstein model [10] and Volterra series [11]. The Volterra series based model has a generic expression and hence it is widely applied for characterizing nonlinear satellite channels.
Numerous authors have investigated the design of equalizers for channels modeled by Volterra series [12]-[17]. The optimal maximum a posteriori (MAP) equalizer is developed in [12]. However, the complexity of the optimal MAP equalizer increases exponentially with the channel's memory length. To reduce the computational complexity, a nonlinear soft interference canceller (NL-SIC) is proposed in [13], which is an extension of the classic SIC designed for linear channels in [18]. In [14], a linear minimum mean square error (LMMSE) criterion-based equalizer is used for canceling the nonlinear distortion. However, the complexity of the LMMSEbased methods still grows cubically with the length of their processing window.

Several low-complexity iterative equalizers have also been developed based on the factor graph (FG) modelling nonlinear channels [15]-[17]. In [15], a Markov Chain Monte Carlo (MCMC) equalizer relying on different sampling methods is employed for reducing the complexity of the sum-product algorithm (SPA)-based forward-backward (FB) equalizer, while achieves the comparable performance. However, the MCMC method may suffer from convergence problems resulting from the sensitivity to the noise fluctuations [19]. In [16], a softinput soft-output (SISO) equalizer is proposed to attain a linear computational complexity, where a simplified Volterra series model is employed for characterizing nonlinear channels, resulting in a performance erosion. In [17], a different message passing algorithm is designed based on FG, which gleans extrinsic information from the associated equalizer. By avoiding any matrix inversion operation, the computational complexity

TABLE I
EXPLICITLY CONTRASTING OUR CONTRIBUTIONS TO THE LITERATURE.

| Feature | This Paper | [13], [14] | [15] | [16] | [17] | [20] | [25] | [26] |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Nonlinear Channels | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |  | $\checkmark$ |
| Factor Graph Framework | $\checkmark$ |  | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |  |
| Joint Channel Estimation and Equalization | $\checkmark$ |  |  |  |  |  | $\checkmark$ | $\checkmark$ |
| Parametric Nonlinear Channel Estimation | $\checkmark$ |  |  |  |  |  |  |  |
| Valuable Use for Estimation Uncertainty | $\checkmark$ | $\checkmark$ |  |  | $\checkmark$ | $\checkmark$ | $\checkmark$ |  |
| Without Linearization | $\checkmark$ | $\checkmark$ | $\checkmark$ |  | $\checkmark$ |  |  | $\checkmark$ |
| Unknown Noise Variance | $\checkmark$ |  |  |  |  |  |  |  |

of the equalizer can be significantly reduced compared to the LMMSE equalizer [14] and the FB equalizer [15]. In [20], a Gaussian message passing receiver is developed for faster-than-Nyquist signaling over nonlinear channels, where the channel model is linearized using Taylor expansion, and thus the problem of symbol detection is reformulated into a linear state-space model.

Nevertheless, both of the above-mentioned algorithms only considered the idealized simplifying assumption of perfectly known channel state information (CSI). In practice, the CSI has to be estimated by relying on sufficiently long training sequences. However, due to the inevitable radiation, ageing, temperature variations or different mismatch effects, the characteristics of HPAs are time-varying, which requires reestimation of CSI and results in an extra training overhead. Also, with the increase of the memory length and nonlinear order of channels, the loss of spectral efficiency imposed by pilots is even higher. It has been shown [21]-[26] that joint channel estimation and data detection not only significantly reduces the number of pilots, but also improves the CSI estimation accuracy in linear channels. In [27], joint channel estimation and equalization is proposed for nonlinear MIMO channels characterized by the Volterra series, where the alternating least squares (ALS) algorithm is employed for both channel estimation and data detection in an iterative way. However, this ALS algorithm fails to reach its full potential due to failing to rely on soft information during its iterations.

Against this backdrop, we propose a joint channel estimation and data detection method for nonlinear satellite channels based on FGs. The novel contributions of this paper are explicitly itemized next and they are also clearly contrasted to the literature in Table I. Our main contributions are summarized as follows:

- We construct a FG to represent communications over nonlinear channels exhibiting memory characterized by Volterra series. To tackle the problems of message passing caused by hard constraints and dense short loops, the FG is intentionally split into belief propagation (BP) and mean field (MF) regions. Then, we invoke a combined BP-MF method for deriving messages on a FG. Due to the nonlinear terms in the system model, the messages are in non-Gaussian forms. To this end, a parametric message update with canonical parameters are derived. Moreover, by reformulating the system model, we conceive a loopfree subgraph in terms of vectors for performing channel
estimation. Furthermore, the noise variance estimation using a Gamma a priori distribution is embedded into the proposed method.
- To circumvent the high computational complexity imposed by matrix inversion, we further resort to the generalized approximate message passing (GAMP) algorithm [28] for estimating the channel coefficients. Accordingly, the vector estimation process is transformed into scalar operations. The iterations of the GAMP algorithm may be incorporated into the iterations between the equalizer and decoder. The complexity of channel estimation based on the GAMP algorithm only grows linearly with the number of channel coefficients.
The rest of this paper is organized as follows. Section II presents the system model and problem formulation over nonlinear satellite channels. In Section III, the proposed BPMF method of joint channel estimation and data detection over nonlinear satellite channels is derived in detail, and extended to the scenario of realistic unknown noise variance. In Section IV, our GAMP-aided joint channel estimation and data detection algorithm is developed for reducing the computational complexity. The performance of the proposed algorithms is evaluated by Monte Carlo simulations in Section V, while our conclusions are offered in Section VI.
Notations: Boldface capital and lower case letters denote matrices and vectors, respectively; The operations $(\cdot)^{*},(\cdot)^{T}$ $(\cdot)^{H}$ and $(\cdot)^{-1}$ represent the complex conjugate, transpose, conjugate transpose, and matrix inverse, respectively; $|\cdot|$ denotes the modulus of a complex number; $\Re(x)$ denotes the real part of complex number $x$; The $(n, m)$ th element of the matrix $\mathbf{X}$ is written as $X_{n m} ; \mathbb{E}\{\cdot\}$ and $\mathbb{V}\{\cdot\}$ represent expectation and variance operations, respectively; $\operatorname{Tr}\{\cdot\}$ denotes the trace operation; The symbol $\propto$ denotes equality up to a constant; $\mathcal{C N}\left(x ; m_{x}, V_{x}\right)$ denotes a complex Gaussian probability density function (pdf) of the random variable $x$ with mean $m_{x}$ and variance $V_{x}$.


## II. System Model and Problem Formulation

We consider a coded system communicating over nonlinear satellite channel exhibiting memory, as depicted in Fig. 1. At the transmitter side, the information bit sequence $\mathbf{b} \triangleq\left[b_{1}, b_{2}, \ldots, b_{N_{b}}\right]^{T}$ is encoded into a coded bit sequence $\mathbf{c} \triangleq\left[c_{1}, c_{2}, \ldots, c_{N_{c}}\right]^{T}$, where $N_{b}$ and $N_{c}$ are the lengths of information bit sequence and coded bit sequence, respectively. The sequence $\mathbf{c}$ is further mapped to a data symbol vector


Fig. 1. System model of the coded system communicating over nonlinear satellite channels.
$\mathbf{x} \triangleq\left[x_{1}, x_{2}, \ldots, x_{N}\right]^{T}$ of length $N$, where we have $x_{n} \in \chi$ and $\chi$ denotes the $2^{M}$-ary symbol alphabet. Then, $\mathbf{x}$ is passed through the shaping filter, whose output signal is transmitted over the nonlinear satellite channel corrupted by additive white Gaussian noise (AWGN). The satellite transponder consists of an IMUX filter, an HPA and an OMUX filter, as shown in Fig. 1. The IMUX filter aims for removing the adjacent-channel interferences, while the OMUX filter is intended for reducing the out-of-band emissions of spectral expansion. Volterra series are employed for modelling the satellite channel's nonlinearity. Due to the bandpass nature of the satellite communication channel, only the odd-order Volterra terms are considered [12]. In practice, employing a third-order Volterra series is sufficient for modelling nonlinear satellite channels [13]-[17], [20].

For a third-order Volterra series [11], the $n$-th discretetime received signal at the output of the matched filter can be represented as

$$
\begin{equation*}
y_{n}=\sum_{l=0}^{L} h_{l} x_{n-l}+\sum_{i=0}^{L} \sum_{j \geq i}^{L} \sum_{k=0}^{L} h_{i j k} x_{n-i} x_{n-j} x_{n-k}^{*}+w_{n} \tag{1}
\end{equation*}
$$

where $L$ describes the dispersion of the nonlinear channels, $h_{l}$ and $h_{i j k}$ represent the Volterra channel coefficients, and $w_{n}$ is the AWGN having zero mean and variance of $\sigma^{2}$.

The joint a posteriori distribution of channel coefficients and information bits can be written as

$$
\begin{align*}
p(\mathbf{b}, \mathbf{h} \mid \mathbf{y}) \propto & p(\mathbf{y} \mid \mathbf{h}, \mathbf{x}) p(\mathbf{h}) p(\mathbf{x} \mid \mathbf{c}) p(\mathbf{c} \mid \mathbf{b}) p(\mathbf{b})  \tag{2}\\
\propto & \prod_{n=1}^{N} p\left(y_{n} \mid \mathbf{h}, x_{n}, \ldots, x_{n-L}\right) p\left(x_{n} \mid \mathbf{c}_{n}\right) \\
& \times \prod_{l=0}^{L} p\left(h_{l}\right) \prod_{i, j, k} p\left(h_{i j k}\right) \times p(\mathbf{c} \mid \mathbf{b}) \times \prod_{n=1}^{N_{b}} p\left(b_{n}\right), \tag{3}
\end{align*}
$$

where $\mathbf{h} \triangleq\left[h_{0}, h_{1}, \ldots, h_{L}, h_{000}, h_{001}, \ldots, h_{00 L}, \ldots, h_{L L L}\right]^{T}$ and $\mathbf{c}_{n} \triangleq\left[c_{n, 1}, \ldots, c_{n, M}\right]^{T}$ with $c_{n, k} \triangleq c_{(n-1) M+k}$. The $a$ priori probability $p\left(b_{n}\right)$ of the information bit is assumed to be uniformly distributed. Furthermore, $p(\mathbf{c} \mid \mathbf{b})$ and $p\left(x_{n} \mid \mathbf{c}_{n}\right)$ denote the deterministic coding function and symbol mapping


Fig. 2. Factor graph representation of joint a posteriori distribution in (3).
function, respectively, while $p\left(h_{l}\right)$ and $p\left(h_{i j k}\right)$ are the a priori probabilities of the Volterra channel coefficients $h_{l}$ and $h_{i j k}$, respectively. The likelihood function $p\left(y_{n} \mid \mathbf{h}, x_{n}, \ldots, x_{n-L}\right)$ can be expressed as


For the joint a posteriori distribution in (3), the corresponding FG representation is shown in Fig. 2, where the squares denote the factor nodes (FN) and the circles denote the variable nodes (VN). In Fig. 2, the FN $f_{\mathcal{M}, n}$ corresponding to $p\left(x_{n} \mid \mathbf{c}_{n}\right)$ represents symbol mapping, i.e., $f_{\mathcal{M}, n}\left(x_{n}, \mathbf{c}_{n}\right) \triangleq$ $\delta\left[x_{n}-\mathcal{M}\left(\mathbf{c}_{n}\right)\right]$ with $\mathcal{M}$ being the symbol mapping function, and the FN $f_{r, n}$ corresponding to $p\left(y_{n} \mid \mathbf{h}, x_{n}, \ldots, x_{n-L}\right)$ stands for the likelihood function.

To obtain the marginal distributions of channel coefficients and information bits, BP may be adopted directly. However, the FG in Fig. 2 contains short loops, which may result in the divergence of BP [29]. Although the convergence-guaranteed MF can be employed, it is not compatible with the hard constraints involved in the coding and modulation functions. To this end, we propose a combined BP-MF method for joint channel estimation and data detection in nonlinear satellite channels.

## III. BP-MF for Joint Channel Estimation and Decoding in Nonlinear channels

This section presents the proposed message passing algorithm conceived for jointly estimating the nonlinear channel's coefficients and detecting the transmitted information. Given the associated hard constraints and the FG's loopy nature, we split the FG into two regions, i.e., the BP region and the MF region. For the sake of reducing the computational complexity, a series of parametric message updating expressions are derived to obtain the estimates of data symbols. By introducing


Fig. 3. Factor graphs to illustrate (a) the factor nodes classified into BP region and MF region and (b) the non-loopy subgraph with vector-valued nodes for channel estimation using the MF approximation.
the vector-valued nodes for eliminating the dense loops of the FG, a loop-free subgraph is constructed for performing channel estimation. Furthermore, the proposed algorithm is also extended to the case of realistic unknown noise variance.

## A. The Proposed Joint Channel Estimation and Detection Algorithm Based on the BP-MF Method

Let us now define the BP region $\mathcal{A}_{\mathrm{BP}}$ and the MF region $\mathcal{A}_{\mathrm{MF}}$, where we have $\mathcal{A}_{\mathrm{BP}} \triangleq\left\{f_{\mathcal{M}, n} \mid n=1,2, \ldots, N\right\}$ and $\mathcal{A}_{\mathrm{MF}} \triangleq\left\{f_{r, n} \mid n=1,2, \ldots, N\right\}$, as shown in Fig. 3(a). Then, we can derive the a posteriori beliefs of data symbols and channel coefficients using the BP-MF method.

First, we consider the message update related to the FNs in the BP region. According to [30], the message emerging from $f_{\mathcal{M}, n}$ to $x_{n}$ is updated by
$\mu_{f_{\mathcal{M}, n} \rightarrow x_{n}}\left(x_{n}\right)=\sum_{\mathbf{c}_{n} \in\{0,1\}^{M}} f_{\mathcal{M}, n}\left(x_{n}, \mathbf{c}_{n}\right) \prod_{k=1}^{M} \mu_{c_{n, k} \rightarrow f_{\mathcal{M}, n}}\left(c_{n, k}\right)$.

Since $x_{n}$ is a discrete random variable, the message in (5) is the probability mass function (PMF) of $x_{n}$. Thus, the message $\mu_{f_{\mathcal{M}, n} \rightarrow x_{n}}\left(x_{n}\right)$ has the form

$$
\begin{equation*}
\mu_{f_{\mathcal{M}, n} \rightarrow x_{n}}\left(x_{n}\right)=\sum_{j=1}^{2^{M}} P\left(x_{n}=s^{(j)}\right) \delta\left(x_{n}-s^{(j)}\right) \tag{6}
\end{equation*}
$$

where $s^{(j)}$ is the $j$-th symbol in $\chi$.
Then, the message passed from $x_{n}$ to $f_{r, i}$ is updated by

$$
\begin{equation*}
\mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}\right)=\mu_{f_{\mathcal{M}, n} \rightarrow x_{n}}\left(x_{n}\right) \prod_{k=n}^{n+L} \mu_{f_{r, k} \rightarrow x_{n}}\left(x_{n}\right) . \tag{7}
\end{equation*}
$$

In order to obtain $\mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}\right)$ in (7), we have to derive the expression of $\mu_{f_{r, k} \rightarrow x_{n}}\left(x_{n}\right)$ for $k=n, \ldots, n+L$.

According to the MF rules, the message passed from the likelihood node $f_{r, n}$ to the symbol node $x_{n}$ is updated by

$$
\begin{align*}
& \mu_{f_{r, n} \rightarrow x_{n}}\left(x_{n}\right) \propto \exp \left\{\sum_{\sim\left\{x_{n}\right\}} \int_{\mathbf{h}} \ln f_{r, n}\left(\mathbf{h}, x_{n}, \ldots, x_{n-L}\right)\right. \\
& \left.\quad \times \prod_{x_{n}^{\prime} \in N\left(f_{r, n}\right) \backslash x_{n}} \mu_{x_{n}^{\prime} \rightarrow f_{r, n}}\left(x_{n}^{\prime}\right) \prod_{m} \mu_{h_{m} \rightarrow f_{r, n}}\left(h_{m}\right)\right\} \tag{8}
\end{align*}
$$

Substituting the expressions of $f_{r, n}$ in (4) into (8), $\mu_{f_{r, n} \rightarrow x_{n}}\left(x_{n}\right)$ can be rewritten as

$$
\begin{align*}
& \mu_{f_{r, n} \rightarrow x_{n}}\left(x_{n}\right) \propto \exp \left\{\mathbb { E } _ { x _ { n } ^ { \prime } \in N ( f _ { r , n } ) \backslash x _ { n } , h _ { m } } \left[\left.-\frac{1}{\sigma^{2}} \right\rvert\, y_{n}\right.\right. \\
& \left.\left.\quad-\sum_{l=0}^{L} h_{l} x_{n-l}-\left.\sum_{i=0}^{L} \sum_{j \geq i}^{L} \sum_{k=0}^{L} h_{i j k} x_{n-i} x_{n-j} x_{n-k}^{*}\right|^{2}\right]\right\} . \tag{9}
\end{align*}
$$

Through the detailed derivations in Appendix A, we have

$$
\begin{align*}
& \mu_{f_{r, n} \rightarrow x_{n}}\left(x_{n}\right) \\
& \quad \propto \exp \left\{-\frac{\sum_{a, b} \varphi_{a, b}^{(n)} x_{n}^{a} x_{n}^{* b}+2 \Re\left(\sum_{m, v} \varphi_{m, v}^{(n)} x_{n}^{m} x_{n}^{* v}\right)}{\sigma^{2}}\right\} \tag{10}
\end{align*}
$$

where $a, b, m, v \in\{1,2,3\}, \varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ denote the canonical parameters corresponding to the sufficient statistics $x_{n}^{a} x_{n}^{* b}$ and $x_{n}^{m} x_{n}^{* v}$. Note that the message in (10) is non-Gaussian distribution, which is different from the Gaussian distribution in linear channels. Hence, for a third-order Volterra series associated with memory length $L$, the corresponding canonical parameters $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ are computed according to Table $\mathrm{II}^{1}$, where the detailed derivations of $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ are given in Appendix B. Note that only the canonical parameters have to be updated in each iteration, which significantly simplifies the updating of messages. Upon substituting (6) and (10) into (7), we have

$$
\begin{align*}
& \mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}\right) \propto \sum_{j=1}^{2^{M}} P\left(x_{n}=s^{(j)}\right) \delta\left(x_{n}-s^{(j)}\right) \\
& \quad \times \exp \left\{-\frac{\sum_{k=n}^{n+L}\left[\sum_{a, b} \varphi_{a, b}^{(k)} x_{n}^{a} x_{n}^{* b}+2 \Re\left(\sum_{m, v} \varphi_{m, v}^{(k)} x_{n}^{m} x_{n}^{* v}\right)\right]}{\sigma^{2}}\right\} \tag{11}
\end{align*}
$$

Since we have $f_{r, i} \in \mathcal{A}_{\mathrm{MF}}$, the belief of $x_{n}$ equals to the message $\mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}\right)$. Thus, we can obtain the a posteriori mean $m_{x_{n}}$ and variance $v_{x_{n}}$ of $x_{n}$ as

$$
\begin{align*}
m_{x_{n}} & =\frac{\sum_{j=1}^{2^{M}} s^{(j)} \mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}=s^{(j)}\right)}{\sum_{j=1}^{2^{M}} \mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}=s^{(j)}\right)}  \tag{12}\\
v_{x_{n}} & =\frac{\sum_{j=1}^{2^{M}}\left|s^{(j)}\right|^{2} \mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}=s^{(j)}\right)}{\sum_{j=1}^{2^{M}} \mu_{x_{n} \rightarrow f_{r, i}}\left(x_{n}=s^{(j)}\right)}-\left|m_{x_{n}}\right|^{2} \tag{13}
\end{align*}
$$

[^1]TABLE II
PARAMETERS $\varphi_{a, b}^{(n)}$ AND $\varphi_{m, v}^{(n)}$ FOR A THIRD-ORDER Volterra Channel WITH MEMORY LENGTH $L$.

| sufficient statistics | canonical parameters |
| :---: | :---: |
| $\left\|x_{n}\right\|^{2}(a=b=1)$ | $\begin{aligned} & \mathbb{E}\left(\left\|h_{0}\right\|^{2}\right)-2 \Re\left(y_{n}^{*} \sum_{i=1}^{L} \mathbb{E}\left(h_{0 i 0}\right) \mathbb{E}\left(x_{n-i}\right)\right)+2 \Re\left(\sum_{i, j=1}^{L} \mathbb{E}\left(h_{0}^{*} h_{0 i j}+h_{i} h_{0 j 0}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*}\right)\right)+\sum_{i, j=1}^{L} \mathbb{E}\left(\left\|h_{0 i j}\right\|^{2}\right) \mathbb{E}\left(\left\|x_{n-i}\right\|^{2}\left\|x_{n-j}\right\|^{2}\right) \\ &++2 \Re\left(\sum_{i, j, p>i, q>j}^{L} \mathbb{E}\left(h_{0 i j}^{*} h_{0 p q}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j} x_{n-p} x_{n-q}^{*}\right)\right)+\sum_{i, j \geq i}^{L} \mathbb{E}\left(\left\|h_{i j 0}\right\|^{2}\right) \mathbb{E}\left(\left\|x_{n-i}\right\|^{2}\left\|x_{n-j}\right\|^{2}\right) \\ &+ 2 \Re\left(\sum_{i, j, p>i, q>j}^{L} \mathbb{E}\left(h_{i j 0}^{*} h_{p q 0}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-p} x_{n-q}\right)\right)+2 \Re\left(\sum_{i, j, p \geq j, q} \mathbb{E}\left(h_{0 i 0}^{*} h_{j p q}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j} x_{n-p} x_{n-q}^{*}\right)\right) \end{aligned}$ |
| $\left\|x_{n}\right\|^{4}(a=b=2)$ | $\begin{aligned} & 2 \Re\left(\mathbb{E}\left(h_{000}^{*}\right)\left(\mathbb{E}\left(h_{0}\right)+\sum_{i, j=1}^{L} \mathbb{E}\left(h_{0 i j}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*}\right)\right)\right)+\sum_{i=1}^{L} \mathbb{E}\left(\left\|h_{00 i}\right\|^{2}\right) \mathbb{E}\left(\left\|x_{n-i}\right\|^{2}\right) \\ & +\sum_{i=1}^{L} \mathbb{E}\left(\left\|h_{0 i 0}\right\|^{2}\right) \mathbb{E}\left(\left\|x_{n-i}\right\|^{2}\right)+2 \Re\left(\sum_{i, j>i} \mathbb{E}\left(h_{00 i}^{*} h_{00 j}+h_{0 i 0} h_{0 j 0}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*}\right)\right) \end{aligned}$ |
| $\left\|x_{n}\right\|^{6}(a=b=3)$ | $\mathbb{E}\left(\left\|h_{000}\right\|^{2}\right)$ |
| $\Re\left(x_{n}\right)(m=1, v=0)$ | $\begin{aligned} & \quad-y_{n}^{*} \mathbb{E}\left(h_{0}\right)-y_{n}^{*} \sum_{i, j=1}^{L} \mathbb{E}\left(h_{0 i j}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*}\right)+\mathbb{E}\left(h_{0}\right) \sum_{j=1}^{L} \mathbb{E}\left(h_{j}^{*}\right) \mathbb{E}\left(x_{n-j}^{*}\right)-y_{n} \sum_{i, j}^{L} \mathbb{E}\left(h_{i j 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*}\right) \\ & +\mathbb{E}\left(h_{0}\right) \sum_{i, j \geq i, k} \mathbb{E}\left(h_{i j k}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-k}\right)+\sum_{i, j, k}^{L} \mathbb{E}\left(h_{i}^{*} h_{0 j k}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j} x_{n-k}^{*}\right)+\sum_{i, j, k}^{L} \mathbb{E}\left(h_{i} h_{j k 0}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*} x_{n-k}^{*}\right) \\ & +\sum_{i, j, k, p, q} \mathbb{E}\left(h_{0 i j} h_{k p q}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*} x_{n-k}^{*} x_{n-p}^{*} x_{n-q}\right)+\sum_{i, j, k, p, q} \mathbb{E}\left(h_{i j 0}^{*} h_{k p q}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-k} x_{n-p} x_{n-q}^{*}\right) \end{aligned}$ |
| $\Re\left(x_{n}^{2}\right)(m=2, v=0)$ | $\begin{aligned} & -y_{n}^{*} \sum_{i=1}^{L} \mathbb{E}\left(h_{00 i}\right) \mathbb{E}\left(x_{n-i}^{*}\right)+\mathbb{E}\left(h_{0}\right) \sum_{i, j \geq i}^{L} \mathbb{E}\left(h_{i j 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*}\right)+\sum_{i, j=1}^{L} \mathbb{E}\left(h_{i}^{*} h_{00 j}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*}\right) \\ + & \sum_{i, j, p, q \geq p} \mathbb{E}\left(h_{0 i j} h_{p q 0}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*} x_{n-p}^{*} x_{n-q}^{*}\right)+\sum_{i, j, p \geq j, q} \mathbb{E}\left(h_{00 i}^{*} h_{j p q}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-p}^{*} x_{n-q}\right) \end{aligned}$ |
| $\Re\left(x_{n}^{3}\right)(m=3, v=0)$ | $\sum_{i, j, k \geq j} \mathbb{E}\left(h_{00 i} h_{j k 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-k}^{*}\right)$ |
| $\Re\left(x_{n}\left\|x_{n}\right\|^{2}\right)(m=2, v=1)$ | $\begin{aligned} &-y_{n}^{*} \mathbb{E}\left(h_{000}\right)+\sum_{i=1}^{L} \mathbb{E}\left(h_{0}^{*} h_{00 i}+h_{0} h_{0 i 0}^{*}+h_{i}^{*} h_{000}\right) \mathbb{E}\left(x_{n-i}^{*}\right)+\sum_{i, j, k=1}^{L} \mathbb{E}\left(h_{00 i} h_{0 j k}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-k}\right) \\ &+\sum_{i, j, k=1}^{L} \mathbb{E}\left(h_{0 i 0}^{*} h_{0 j k}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j} x_{n-k}^{*}\right)+\sum_{i, j, k \geq j}^{L} \mathbb{E}\left(h_{0 i 0} h_{j k 0}^{*}\right) \mathbb{E}\left(x_{n-i} x_{n-j}^{*} x_{n-k}^{*}\right)+\sum_{i, j \geq i, k} \mathbb{E}\left(h_{000} h_{i j k}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*} x_{n-k}\right) \end{aligned}$ |
| $\Re\left(x_{n}\left\|x_{n}\right\|^{4}\right)(m=3, v=2)$ | $\sum_{i=1}^{L} \mathbb{E}\left(h_{000}^{*} h_{00 i}\right) \mathbb{E}\left(x_{n-i}^{*}\right)+\sum_{i=1}^{L} \mathbb{E}\left(h_{000} h_{0 i 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*}\right)$ |
| $\Re\left(x_{n}^{2}\left\|x_{n}\right\|^{2}\right)(m=3, v=1)$ | $\sum_{i, j \geq i} \mathbb{E}\left(h_{000} h_{i j 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*}\right)+\sum_{i, j}^{L} \mathbb{E}\left(h_{00 i} h_{0 j 0}^{*}\right) \mathbb{E}\left(x_{n-i}^{*} x_{n-j}^{*}\right)$ |

Similarly, we can derive the messages from the likelihood function to the channel coefficient according to the MF rules. To facilitate the calculation of messages, the received signal in (1) is reformulated into the following compact form

$$
\begin{equation*}
\mathbf{y}=\mathbf{X} \mathbf{h}+\mathbf{w} \tag{14}
\end{equation*}
$$

where $\mathbf{w} \triangleq\left[w_{1}, w_{2}, \cdots, w_{N}\right]^{T}$, and

$$
\mathbf{X}=\left[\begin{array}{cc}
\left(\mathbf{x}_{1}^{L}\right)^{T} & \left(\mathbf{x}_{1}^{N L}\right)^{T}  \tag{15}\\
\left(\mathbf{x}_{2}^{L}\right)^{T} & \left(\mathbf{x}_{2}^{N L}\right)^{T} \\
\vdots & \vdots \\
\left(\mathbf{x}_{N}^{L}\right)^{T} & \left(\mathbf{x}_{N}^{N L}\right)^{T}
\end{array}\right]
$$

with $\quad \mathbf{x}_{n}^{L} \triangleq\left[x_{n}, x_{n-1}, \ldots, x_{n-L}\right]^{T}$, and $\mathbf{x}_{n}^{N L} \triangleq$ $\left[x_{n} x_{n} x_{n}^{*}, x_{n} x_{n} x_{n-1}^{*}, \ldots, x_{n} x_{n} x_{n-L}^{*}, \ldots, x_{n-L} x_{n-L} x_{n-L}^{*}\right]^{T}$. In this way, the message passing can be performed in the non-loopy subgraph with vector-valued nodes, as shown in Fig. 3(b). Accordingly, the message passed from the likelihood function $f_{r}$ to the channel coefficient $h$ can be updated by

$$
\begin{equation*}
\mu_{\mathbf{f}_{\mathbf{r}} \rightarrow \mathbf{h}} \propto \exp \left\{\mathbb{E}_{\mu_{\mathbf{x} \rightarrow \mathbf{f}_{\mathbf{r}}}}[\ln p(\mathbf{y} \mid \mathbf{h}, \mathbf{x})]\right\} \tag{16}
\end{equation*}
$$

Since $w_{n} \sim \mathcal{C N}\left(w_{n} ; 0, \sigma^{2}\right), n=1,2, \cdots, N$, we have

$$
\begin{equation*}
p(\mathbf{y} \mid \mathbf{h}, \mathbf{x}) \propto \exp \left\{-\frac{(\mathbf{y}-\mathbf{X h})^{H}(\mathbf{y}-\mathbf{X h})}{\sigma^{2}}\right\} \tag{17}
\end{equation*}
$$

Substituting (17) into (16) yields

$$
\mu_{\mathbf{f}_{\mathbf{r}} \rightarrow \mathbf{h}}
$$

$$
\propto \exp \left\{\mathbb{E}_{\mu_{\mathbf{x} \rightarrow \mathrm{f}_{\mathbf{r}}}}\left[-\frac{(\mathbf{y}-\mathbf{X h})^{H}(\mathbf{y}-\mathbf{X h})}{\sigma^{2}}\right]\right\}
$$

$$
\propto \exp \left\{-\frac{\mathbf{y}^{H} \mathbf{y}-\mathbf{y}^{H} \mathbb{E}[\mathbf{X}] \mathbf{h}-\mathbf{h}^{H} \mathbb{E}\left[\mathbf{X}^{H}\right] \mathbf{y}+\mathbf{h}^{H} \mathbb{E}\left[\mathbf{X}^{H} \mathbf{X}\right] \mathbf{h}}{\sigma^{2}}\right\}
$$

$$
\begin{equation*}
\propto \exp \left\{-\left(\mathbf{h}-\check{\mathbf{m}}_{\mathbf{h}}\right)^{H} \check{\mathbf{C}}_{\mathbf{h}}^{-1}\left(\mathbf{h}-\check{\mathbf{m}}_{\mathbf{h}}\right)\right\} \tag{18}
\end{equation*}
$$

$\propto \mathcal{C N}\left(\mathbf{h} ; \check{\mathbf{m}}_{\mathbf{h}}, \check{\mathbf{C}}_{\mathbf{h}}\right)$,
where the mean vector obeys $\check{\mathbf{m}}_{\mathbf{h}}=\frac{1}{\sigma^{2}} \check{\mathbf{C}}_{\mathbf{h}} \mathbb{E}[\mathbf{X}]^{H} \mathbf{y}$ and the covariance matrix is given by $\check{\mathbf{C}}_{\mathbf{h}}^{-1}=\frac{1}{\sigma^{2}}\left[\mathbb{E}\left(\mathbf{X}^{H} \mathbf{X}\right)\right]$. With the estimated a posteriori mean and variance of $x_{n}$ in (12) and (13), the mean vector $\check{\mathbf{m}}_{\mathbf{h}}$ and the covariance matrix $\check{\mathbf{C}}_{\mathbf{h}}$ can be computed straightforwardly. Assuming that the a priori probability $p_{\mathrm{h}}$ obeys a multivariate complex Gaussian pdf, i.e., $p_{\mathbf{h}} \propto \mathcal{C N}\left(\mathbf{h} ; \mathbf{h}^{0}, \mathbf{V}_{\mathbf{h}}^{0}\right)$, the a posteriori pdf of channel coefficients $\mathbf{h}$ is also a multivariate complex Gaussian pdf $\mathcal{C N}\left(\mathbf{h} ; \mathbf{m}_{\mathbf{h}}, \mathbf{V}_{\mathbf{h}}\right)$ with

$$
\begin{align*}
& \mathbf{m}_{\mathbf{h}}=\mathbf{V}_{\mathbf{h}}\left[\left(\mathbf{V}_{\mathbf{h}}^{0}\right)^{-1} \mathbf{h}^{0}+\check{\mathbf{C}}_{\mathbf{h}}^{-1} \check{\mathbf{m}}_{\mathbf{h}}\right]  \tag{19}\\
& \mathbf{V}_{\mathbf{h}}=\left[\left(\mathbf{V}_{\mathbf{h}}^{0}\right)^{-1}+\check{\mathbf{C}}_{\mathbf{h}}^{-1}\right]^{-1} \tag{20}
\end{align*}
$$

Given (19) and (20), the channel coefficients $h$ can be jointly updated.


Fig. 4. The subgraph of noise variance estimation.

Finally, the message passed from the mapping function $f_{\mathcal{M}, n}$ to the coded bit $c_{n, k}$ can be computed by

$$
\begin{align*}
\mu_{f_{\mathcal{M}, n} \rightarrow c_{n, k}}\left(c_{n, k}\right) \propto & \sum_{\sim\left\{c_{n, k}\right\}} f_{\mathcal{M}, n}\left(x_{n}, \mathbf{c}_{n}\right) \mu_{x_{n} \rightarrow f_{\mathcal{M}, n}}\left(x_{n}\right) \\
& \times \prod_{k^{\prime} \neq k} \mu_{c_{n, k^{\prime}} \rightarrow f_{\mathcal{M}, n}}\left(c_{n, k^{\prime}}\right) \tag{21}
\end{align*}
$$

where the message passed from $x_{n}$ to $f_{\mathcal{M}, n}$ is updated by

$$
\begin{align*}
& \mu_{x_{n} \rightarrow f_{\mathcal{M}, n}}\left(x_{n}\right)=\prod_{k=n}^{n+L} \mu_{f_{r, k} \rightarrow x_{n}}\left(x_{n}\right) \\
& \propto \exp \left\{-\frac{\sum_{k=n}^{n+L}\left[\sum_{a, b} \varphi_{a, b}^{(k)} x_{n}^{a} x_{n}^{* b}+2 \Re\left(\sum_{m, v} \varphi_{m, v}^{(k)} x_{n}^{m} x_{n}^{* v}\right)\right]}{\sigma^{2}}\right\} \tag{22}
\end{align*}
$$

The message $\mu_{f_{\mathcal{M}, n} \rightarrow c_{n, k}}\left(c_{n, k}\right)$ can be viewed as the extrinsic soft information passed to a SISO decoder and it is treated as the a priori information used for decoding. Once the SISO decoding is terminated, it will produce extrinsic soft information, which is used as the a priori information in (5) to determine the a priori probabilities of the data symbol at the next iteration. Upon using the message $\mu_{f_{\mathcal{M}, n} \rightarrow c_{n, k}}\left(c_{n, k}\right)$, the log-likelihood ratio (LLR) of coded bits can be computed by

$$
\begin{equation*}
L L R\left(c_{n, k}\right)=\ln \frac{\mu_{f_{\mathcal{M}, n} \rightarrow c_{n, k}}\left(c_{n, k}=0\right)}{\mu_{f_{\mathcal{M}, n} \rightarrow c_{n, k}}\left(c_{n, k}=1\right)} \tag{23}
\end{equation*}
$$

The details of the proposed combined BP-MF method of joint channel estimation and decoding are summarized in Algorithm 1.

## B. Extension to the Case of Realistic Unknown Noise Variance

The noise variance is always unknown and has to be estimated in practical wireless communication systems. In this section, we now harness the MF for estimating the noise precision $\lambda$, which is the inverse of noise variance, i.e., $\lambda=1 / \sigma^{2}$. In this case, the joint a posteriori distribution in (2) is represented by $p(\mathbf{b}, \mathbf{h}, \lambda \mid \mathbf{y})$, the likelihood function $p(\mathbf{y} \mid \mathbf{h}, \mathbf{x})$ becomes $p(\mathbf{y} \mid \mathbf{h}, \mathbf{x}, \lambda)$, and the a priori distribution $p(\lambda)$ is added to the products in (2) and (3). On the FG of Fig. 4, the additional VN $\lambda$ and $\mathrm{FN} f_{\lambda}$ denoting the $a$ priori distribution of $\lambda$ are introduced. Then, according to the

## Algorithm 1 The Proposed Combined BP-MF Method for Joint Channel Estimation and Decoding

Initialize $\mu_{c_{n, k} \rightarrow f_{\mathcal{M}, n}}\left(c_{n, k}\right)=1 / 2$ and the a priori distribution of channel coefficients is obtained using pilots based on LMMSE;
for $i_{o}=1$ to $I_{\text {out }}$ do
Calculate the messages from the mapping node to the symbol node according to (5);
for $i_{t}=1$ to $I_{\text {in }}$ do
Calculate the canonical parameters $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ according to Table II;
Calculate the messages from the likelihood node to the symbol node according to (10);
Calculate the belief of variable $x_{n}$ according to (11); Calculate the a posteriori mean and variance $\left\{m_{x_{n}}, v_{x_{n}}\right\}$ according to (12) and (13);

## end for

Calculate the a posteriori mean vector and covariance matrix $\left\{\mathbf{m}_{\mathbf{h}}, \mathbf{V}_{\mathbf{h}}\right\}$ according to (19) and (20);
11: Calculate the messages from the mapping node to the coded bit node according to (21);
Calculate the extrinsic LLRs of the equalizer according to (23) and pass them to the SISO decoder;
Perform channel decoding algorithm and the decoder produces the extrinsic soft information to equalizer;
end for

MF rules, the message from the likelihood node $f_{r}$ to noise precision node $\lambda$ can be updated by
$\mu_{\mathbf{f}_{\mathbf{r}} \rightarrow \lambda}(\lambda) \propto \exp \left\{\int \ln p(\mathbf{y} \mid \mathbf{h}, \mathbf{x}, \lambda) \mu_{\mathbf{x} \rightarrow \mathbf{f}_{\mathbf{r}}}(\mathbf{x}) \mu_{\mathbf{h} \rightarrow \mathbf{f}_{\mathbf{r}}}(\mathbf{h}) d \mathbf{x} d \mathbf{h}\right\}$,
where

$$
\begin{equation*}
p(\mathbf{y} \mid \mathbf{h}, \mathbf{x}, \lambda) \propto \lambda^{N} \exp \left\{-\lambda\|\mathbf{y}-\mathbf{X h}\|^{2}\right\} \tag{25}
\end{equation*}
$$

Upon substituting (25) into (24), we arrive at

$$
\begin{align*}
\mu_{\mathbf{f}_{\mathbf{r}} \rightarrow \lambda}(\lambda) & \propto \exp \left\{N \ln \lambda-\lambda \mathbb{E}\left[\|\mathbf{y}-\mathbf{X h}\|^{2}\right]\right\} \\
& \propto \lambda^{N} \exp \left\{-\lambda \mathbb{E}\left[\|\mathbf{y}-\mathbf{X h}\|^{2}\right]\right\} \\
& \propto \lambda^{N} \exp \left\{-\lambda C_{\lambda}\right\} \tag{26}
\end{align*}
$$

where $C_{\lambda}=\mathbb{E}\left[\|\mathbf{y}-\mathbf{X h}\|^{2}\right]=\mathbf{y}^{H} \mathbf{y}-\mathbf{y}^{H} \mathbb{E}[\mathbf{X}] \mathbb{E}[\mathbf{h}]-$ $\mathbb{E}[\mathbf{h}]^{H} \mathbb{E}[\mathbf{X}]^{H} \mathbf{y}+\operatorname{Tr}\left\{\mathbb{E}\left[\mathbf{X}^{H} \mathbf{X}\right] \mathbb{E}\left[\mathbf{h} \mathbf{h}^{H}\right]\right\}$. The message in (26) is proportional to a one-dimensional complex Wishart distribution with the degree of freedom $N+1$, and covariance of $C_{\lambda}^{-1}$. In this case, we select a Gamma distribution as the conjugate $a$ priori pdf [31], i.e., $p(\lambda) \propto \lambda^{a_{p}-1} \exp \left\{-\lambda A_{p}\right\}$. Due to the fact that the a priori information of the noise precision is usually unavailable at the receiver side, we select a non-informative a priori pdf $p(\lambda)$ associated with $a_{p}=0$ and $A_{p}=0$, i.e., $p(\lambda) \propto 1 / \lambda$. Given $\mu_{\mathbf{f}_{\mathbf{r}} \rightarrow \lambda}(\lambda)$ in (26), we obtain the belief

$$
\begin{equation*}
b_{\lambda}(\lambda) \propto \lambda^{N-1} \exp \left\{-\lambda C_{\lambda}\right\} \tag{27}
\end{equation*}
$$

With the belief of $\lambda$ in (27), the noise precision can be calculated as

$$
\begin{equation*}
\hat{\lambda}=\frac{N}{C_{\lambda}} . \tag{28}
\end{equation*}
$$

Note that for the update of data symbols $\mathbf{x}$ and channel coefficients $\mathbf{h}$, we only have to replace $\sigma^{2}$ by $1 / \hat{\lambda}$.

## IV. GAMP-Aided Complexity Reduction of BP-MF for Joint Channel Estimation and Decoding

Although the proposed BP-MF algorithm is capable of accurately estimating both channel coefficients and transmitted symbols, it still suffers from a high computational complexity on the order of $\mathcal{O}\left(L_{h}^{3}\right)$ resulting from the matrix inversion operation of (20) involved in channel estimation, where $L_{h}$ denotes the number of channel coefficients. This computational complexity could be a challenging task for communication systems having a large channel dispersion. To circumvent the matrix inversion, we further resort to the computationally efficient GAMP algorithm [27] to develop a low-complexity channel estimation method. With the aid of the central limit theorem and Taylor series expansion, the GAMP decouples the vector-valued channel estimation problem into a sequence of scalar operations, and its asymptotic behavior under i.i.d. transforms can be exactly described by a simple set of state evolution equations.

To commence, we replace $\mathbf{X}$ in (14) by $\hat{\mathbf{X}}$, where $\hat{\mathbf{X}}$ is the result of replacing all the $\left\{x_{n}\right\}$ in $\mathbf{X}$ by its a posteriori mean $m_{x_{n}}$. Then (14) becomes

$$
\begin{equation*}
\mathbf{y}=\hat{\mathbf{X}} \mathbf{h}+\mathbf{w} \tag{29}
\end{equation*}
$$

Let us now define $z_{n} \triangleq \hat{\mathbf{X}}_{n} \mathbf{h}$, where $\hat{\mathbf{X}}_{n}$ denotes the $n$th row of matrix $\hat{\mathbf{X}}$, and the corresponding likelihood function is given by $p\left(y_{n} \mid z_{n}\right)=\mathcal{C N}\left(z_{n} ; y_{n}, \sigma^{2}\right)$. Then, the approximate marginal distribution of $z_{n}$ is expressed as

$$
\begin{equation*}
p\left(z_{n} \mid \mathbf{y} ; \hat{p}_{n}, \gamma_{n}^{p}\right) \triangleq \frac{p\left(y_{n} \mid z_{n}\right) \mathcal{C N}\left(z_{n} ; \hat{p}_{n}, \gamma_{n}^{p}\right)}{\int_{z_{n}^{\prime}} p\left(y_{n} \mid z_{n}^{\prime}\right) \mathcal{C N}\left(z_{n}^{\prime} ; \hat{p}_{n}, \gamma_{n}^{p}\right)} \tag{30}
\end{equation*}
$$

where the quantities $\hat{p}_{n}(t)$ and $\gamma_{n}^{p}(t)$ at the $t$-th iteration are updated by

$$
\begin{align*}
\hat{p}_{n}(t) & =\sum_{m} \hat{X}_{n m} \hat{h}_{m}(t)-\gamma_{n}^{p}(t) \hat{s}_{n}(t-1),  \tag{31}\\
\gamma_{n}^{p}(t) & =\sum_{m}^{m}\left|\hat{X}_{n m}\right|^{2} \gamma_{m}^{h}(t) . \tag{32}
\end{align*}
$$

At the first iteration, $\hat{h}_{m}(t)$ and $\gamma_{m}^{h}(t)$ are initialized based on the pilot-aided LMMSE estimator. The parameter $\hat{s}_{n}(t-1)$ related to the a posteriori mean of $z_{n}$ is set to 0 . Then, they can be updated by exploiting the equation in (31)-(41). Upon using (30), the a posteriori mean and variance of $z_{n}$ are

$$
\begin{align*}
\hat{z}_{n}(t) & =\mathbb{E}\left\{z_{n} \mid \mathbf{y} ; \hat{p}_{n}(t), \gamma_{n}^{p}(t)\right\} \\
& =\hat{p}_{n}(t)+\frac{\gamma_{n}^{p}(t)}{\gamma_{n}^{p}(t)+\sigma^{2}}\left(y_{n}-\hat{p}_{n}(t)\right)  \tag{33}\\
\gamma_{n}^{z}(t) & =\mathbb{V}\left\{z_{n} \mid \mathbf{y} ; \hat{p}_{n}(t), \gamma_{n}^{p}(t)\right\}=\frac{\gamma_{n}^{p}(t) \sigma^{2}}{\gamma_{n}^{p}(t)+\sigma^{2}} \tag{34}
\end{align*}
$$

Based on the GAMP rules, the output scalar estimation function $g_{\text {out }}(\cdot)$ is calculated by

$$
\begin{align*}
g_{\text {out }}\left[\hat{p}_{n}(t), y_{n}, \gamma_{n}^{p}(t)\right] & =\frac{1}{\gamma_{n}^{p}(t)}\left[\mathbb{E}\left\{z_{n} \mid \mathbf{y} ; \hat{p}_{n}(t), \gamma_{n}^{p}(t)\right\}-\hat{p}_{n}(t)\right] \\
& =\frac{y_{n}-\hat{p}_{n}(t)}{\gamma_{n}^{p}(t)+\sigma^{2}} \tag{35}
\end{align*}
$$

and $\hat{s}_{n}(t) \triangleq g_{\text {out }}\left[\hat{p}_{n}(t), y_{n}, \gamma_{n}^{p}(t)\right]$. The negative partial derivative of $g_{\text {out }}(\cdot)$ becomes

$$
\begin{equation*}
\gamma_{n}^{s}(t)=-\frac{\partial}{\partial \hat{p}_{n}(t)} g_{\text {out }}\left[\hat{p}_{n}(t), y_{n}, \gamma_{n}^{p}(t)\right]=\frac{1}{\gamma_{n}^{p}(t)+\sigma^{2}} \tag{36}
\end{equation*}
$$

Similarly, the approximate marginal distribution of $h_{m}$ is updated by

$$
\begin{equation*}
p\left(h_{m} \mid \mathbf{y} ; \hat{r}_{m}, \gamma_{m}^{r}\right) \triangleq \frac{p\left(h_{m}\right) \mathcal{C N}\left(h_{m} ; \hat{r}_{m}, \gamma_{m}^{r}\right)}{\int_{h_{m}^{\prime}} p\left(h_{m}^{\prime}\right) \mathcal{C N}\left(h_{m}^{\prime} ; \hat{r}_{m}, \gamma_{m}^{r}\right)} \tag{37}
\end{equation*}
$$

where the quantities $\hat{r}_{m}(t)$ and $\gamma_{m}^{r}(t)$ at the $t$-th iteration can be formulated as

$$
\begin{align*}
& \hat{r}_{m}(t)=\hat{h}_{m}(t)+\gamma_{m}^{r}(t) \sum_{n} \hat{X}_{n m}^{*} \hat{s}_{n}(t)  \tag{38}\\
& \gamma_{m}^{r}(t)=\left[\sum_{n}\left|\hat{X}_{n m}\right|^{2} \gamma_{n}^{s}(t)\right]^{-1} \tag{39}
\end{align*}
$$

The input scalar estimation function $g_{\text {in }}(\cdot)$ can be written as

$$
\begin{align*}
g_{\text {in }}\left[\hat{r}_{m}(t), \gamma_{m}^{r}(t)\right] & =\mathbb{E}\left\{h_{m} \mid \mathbf{y} ; \hat{r}_{m}(t), \gamma_{m}^{r}(t)\right\} \\
& =\frac{v_{h_{m}}^{0} \hat{r}_{m}(t)+\gamma_{m}^{r}(t) h_{m}^{0}}{v_{h_{m}}^{0}+\gamma_{m}^{r}(t)}, \tag{40}
\end{align*}
$$

and $\hat{h}_{m}(t+1)=g_{\text {in }}\left[\hat{r}_{m}(t), \gamma_{m}^{r}(t)\right]$. Meanwhile, the partial derivative of $g_{\text {in }}(\cdot)$ satisfies

$$
\begin{align*}
\gamma_{m}^{h}(t+1) & =\gamma_{m}^{r}(t) \frac{\partial}{\partial \hat{r}_{m}(t)} g_{\text {in }}\left[\hat{r}_{m}(t), \gamma_{m}^{r}(t)\right] \\
& =\mathbb{V}\left\{h_{m} \mid \mathbf{y} ; \hat{r}_{m}(t), \gamma_{m}^{r}(t)\right\}=\frac{\nu_{h_{m}}^{0} \gamma_{m}^{r}(t)}{\nu_{h_{m}}^{0}+\gamma_{m}^{r}(t)} . \tag{41}
\end{align*}
$$

Finally, the outputs $\left\{\hat{h}_{m}(t+1)\right\}$ and $\left\{\gamma_{m}^{h}(t+1)\right\}$ are the mean and variance of the approximate marginal distribution of channel coefficients, respectively, which are invoked in the next iteration. The GAMP algorithm proposed for updating the estimates of channel coefficients can be embedded into the iteration between the equalizer and the decoder. Note that the complexity of the GAMP algorithm is dominated by the matrix-vector multiplications in (31), (32), (38), and (39), and hence its complexity is proportional to $L_{h}$. The proposed joint channel estimation and decoding algorithm based on the GAMP-aided BP-MF method is referred to as BP-MFGAMP, which is summarized in Algorithm 2. Though the GAMP method is computationally efficient, it may diverge in the case of ill-conditioned or correlated matrices. To this end, a series of pioneering work have emerged for mitigating the convergence problem of GAMP [32]-[34]. Specifically, the memory approximate message passing algorithm proposed in [35] can circumvent the strong sensitivity of the illconditioned matrices to small perturbations, while it retain a low computational complexity. Note that benefiting from the

TABLE III
Volterra Channel Coefficients

| $h_{0}$ | $h_{1}$ | $h_{2}$ |
| :---: | :---: | :---: |
| $0.78085+0.41347 i$ | $0.40323-0.0064 i$ | $-0.15361-0.08961 i$ |
| $h_{000}$ | $h_{001}$ | $h_{002}$ |
| $-0.16-0.036 i$ | $-0.14+0.14 i$ | $0.156+0.088 i$ |
| $h_{110}$ | $h_{220}$ | other $h_{i j k}$ |
| $0.004-0.068 i$ | $0.072-0.072 i$ | 0 |

convergence-guaranteed property, the proposed BP-MF algorithm with robust convergence is also appropriate for a more general transform matrix. To verify the convergence behavior of the proposed BP-MF and BP-MF-GAMP algorithms in the considered nonlinear channels, we show the convergence curves of the proposed algorithms in the following Section V.

```
Algorithm 2 The BP-MF-GAMP Method for Joint Channel
Estimation and Decoding
    The same as BP-MF except replacing the step (10) in
    Algorithm 1 with the following steps:
    Inputs: priors \(\left\{p\left(h_{m}\right)\right\}\), likelihood \(\left\{p\left(y_{n} \mid z_{n}\right)\right\}\), maxi-
    mum iterations \(I_{\text {GAMP }},\left\{\hat{X}_{n m}\right\}\), initial channel estimate
    \(\left\{h_{m}^{0}, v_{h_{m}}^{0}\right\}\)
    Initialization: \(\hat{h}_{m}(1)=h_{m}^{0}, \gamma_{m}^{h}(1)=v_{h_{m}}^{0}, \hat{s}_{n}(0)=0\)
    for \(t=1\) to \(I_{\text {GAMP }}\) do
        for each \(n\), do
            Calculate \(\hat{p}_{n}(t)\) and \(\gamma_{n}^{p}(t)\) according to (31) and (32);
            Calculate \(\hat{z}_{n}(t)\) and \(\gamma_{n}^{z}(t)\) according to (33) and (34);
            Calculate \(\hat{s}_{n}(t)\) and \(\gamma_{n}^{s}(t)\) according to (35) and (36);
        end for
        for each \(m\), do
            Calculate \(\hat{r}_{m}(t)\) and \(\gamma_{m}^{r}(t)\) according to (38) and
            (39);
            Calculate the a posteriori mean \(\hat{h}_{m}(t+1)\) and
            variance \(\gamma_{m}^{h}(t+1)\) according to (40) and (41);
        end for
    end for
```


## V. Simulation Results

We consider an LDPC-coded system having the code length of $N_{b}=1440$ and code rate of $R=5 / 7$. The encoded bits are mapped to QPSK symbols. Moreover, the third-order Volterra series model associated with $L=2$ is considered in Table III, the corresponding channel coefficients of which are given in [15] but with stronger linear and nonlinear ISI [17]. The initial estimates of channel coefficients are obtained from the pilots using the LMMSE method, where the ratio of pilots is $5 \%$. In the proposed algorithms, the inner iterations within the symbol detection are embedded into the outer iterations between equalization and LDPC decoding. Unless otherwise specified, the maximum number of iterations of the proposed algorithms are set to $I_{\text {out }}=20$ and $I_{\text {in }}=2$, respectively, for the outer and inner iterations. Furthermore, the number of GAMP iterations is $I_{\text {GAMP }}=20$.
The bit error rate (BER) performance of the proposed joint channel estimation and decoding methods is evaluated in Fig.


Fig. 5. BER performance of different algorithms.
5. For comparison, the performance of ALS, LMMSE-NLSIC and LMMSE-BP-MF are also plotted. In the LMMSE-NL-SIC and LMMSE-BP-MF algorithms, LMMSE is used for channel estimation based on pilots, and the NL-SIC as well as the BP-MF are employed for equalization, respectively. Furthermore, the BER performance of the NL-SIC and BP-MF having perfect CSI (PCSI) is considered as a benchmark. It is observed that, for the PCSI scenario, the BP-MF equalizer outperforms the NL-SIC equalizer. This is because the NL-SIC equalizer uses the Gaussian approximation for calculating the extrinsic soft information gleaned from the equalizer, and the third-order Volterra terms are ignored when calculating the variance of the estimation error. For the unknown CSI scenario, the LMMSE-NL-SIC and LMMSE-BP-MF outperform the ALS as a benefit of exploiting soft information. We can observe that, by gleaning information from both pilots and the transmitted symbols recovered in the previous iterations, the proposed BP-MF and BP-MF-GAMP perform close to the PCSI scenario, and they significantly outperform the pilotbased LMMSE-BP-MF method. The performance gap between BP-MF and BP-MF-GAMP arises due to the fact that the former explicitly takes estimation uncertainty of data detection by updating the second-order moments of transmitted symbols for calculating the parameters $\check{\mathbf{m}}_{\mathbf{h}}$ and $\check{\mathbf{C}}_{\mathbf{h}}$. By contrast, the BP-MF-GAMP directly ignores the variance $v_{x_{n}}$ upon calculating the parameters among the GAMP method.

In Fig. 6, the normalized mean square error (NMSE) of the proposed channel estimation algorithms is evaluated. For comparison, the performance of the ALS and LMMSE methods is also presented. Observe that the NMSE monotonically decreases as the value of $E_{b} / N_{0}$ increases. The proposed BP-MF and BP-MF-GAMP algorithms have a significant NMSE performance gain over the LMMSE and ALS methods. Benefiting from accounting for the symbol estimation uncertainty, the proposed BP-MF algorithm outperforms the BP-MF-GAMP in terms of NMSE performance. Moreover,


Fig. 6. NMSE performance of channel estimation of different algorithms.
with the increase of SNR, the NMSE performance of BP-MFGAMP converges to that of the BP-MF method. This is due to the fact that the symbol detection becomes more accurate at high SNRs, and thus the performance loss imposed by ignoring the uncertainties of the estimated symbols in the BP-MFGAMP method become negligible. Although the proposed BP-MF-GAMP method obtains the similar NMSE performance with the BP-MF at high SNRs, it still suffers from a nonnegligible BER performance loss, as shown in Fig. 5. This is because superior channel estimation cannot always guarantee better BER performance.

In Fig. 7, we show the noise variance estimated by the proposed algorithm at different $E_{b} / N_{0}$. Since the BP-MF and the BP-MF-GAMP have similar noise variance estimation characteristics, only the results of BP-MF are given for brevity. As seen in the figure, after several iterations, the proposed noise variance estimator significantly outperforms its pure pilot-based counterpart.

The NMSE convergence behavior of the proposed BPMF and BP-MF-GAMP algorithms by the SNR is shown in Fig. 8. Observe that both the BP-MF and BP-MF-GAMP can converge fast after about 10 iterations at high SNRs. Moreover, as a benefit of its convergence-guaranteed property as well as due to taking into account the uncertainty of data detection, the BP-MF method converges faster than the BP-MF-GAMP method. The extrinsic information transfer (EXIT) curves of the proposed BP-MF method are illustrated in Fig. 9. As expected, an open tunnel emerges between the curves of the equalizer and channel decoder at $E_{b} / N_{0}=2.4 \mathrm{~dB}$, which demonstrates that the proposed algorithm is expected to converge.

In Fig. 10, the BER performance of the proposed BPMF method relying on different numbers of inner and outer iterations is evaluated. It may observe that, increasing $I_{\text {out }}$ significantly improves the BER performance. When $I_{\text {out }}$ is higher than 16, however, the performance gain attained by


Fig. 7. Estimated average noise variance versus $E_{b} / N_{0}$.


Fig. 8. NMSE convergence behavior of BP-MF and BP-MF-GAMP algorithms.
increasing $I_{\text {in }}$ and $I_{\text {out }}$ becomes marginal. Hence, the number of inner and outer iterations can be adjusted and optimized to strike a performance versus computational complexity tradeoff.
We now further evaluate the BER and NMSE performance of the proposed BP-MF and BP-MF-GAMP algorithms in the face of stronger nonlinearity. The corresponding channel coefficients are given in Table IV. Fig. 11 shows the BER performance of the aforementioned algorithms. Compared to the results in Fig. 5, the stronger nonlinearity degrades the BER performance of all algorithms. Specifically, the ALS, LMMSE-NL-SIC and NL-SIC with PCSI fail in the face of stronger nonlinearity. The proposed BP-MF and BP-MFGAMP methods perform close to the BP-MF having PCSI, and outperform the LMMSE-BP-MF. Fig. 12 shows the NMSE


Fig. 9. The EXIT chart for the proposed BP-MF method.


Fig. 10. BER performance of the proposed BP-MF method with different numbers of inner iteration and outer iteration.

TABLE IV
Volterra Channel Coefficients with Stronger Nonlinearity

| $h_{0}$ | $h_{1}$ | $h_{2}$ |
| :---: | :---: | :---: |
| $0.78085+0.41347 i$ | $0.40323-0.0064 i$ | $-0.15361-0.08961 i$ |
| $h_{000}$ | $h_{001}$ | $h_{002}$ |
| $-0.2-0.045 i$ | $-0.175+0.175 i$ | $0.195+0.11 i$ |
| $h_{110}$ | $h_{220}$ | other $h_{i j k}$ |
| $-0.005-0.085 i$ | $0.09-0.09 i$ | 0 |

performance of channel estimation at different SNRs. We can see that the channel estimation performance of the proposed BP-MF and BP-MF-GAMP algorithms is not sensitive to the stronger channel nonlinearity, which is a benefit of the superior symbol detection performance. They still significantly outperform the ALS and LMMSE schemes.


Fig. 11. BER performance of different algorithms in system with stronger nonlinearity.


Fig. 12. NMSE performance of channel estimation of different algorithms in system with stronger nonlinearity.

## VI. Conclusions

Message passing-aided iterative receivers were designed for joint channel estimation and data detection in nonlinear satellite channels characterized by the Volterra series. To address the problems of hard constraints and dense short loops on factor graph, we beneficially decomposed the factor graph into the BP and MF regions. Specifically, BP was exploited for calculating the messages harnessed from the factor node associated with hard constraints, while MF was employed at the observation node. The calculation of marginal distributions of the data symbols was significantly simplified by performing parametric message passing. In the channel estimation part, we reformulated the system model into a compact form and transformed the densely connected subgraph into a loop-free
subgraph associated with vector-valued nodes. The proposed BP-MF method was also extended to the realistic scenario of unknown noise variance. To avoid the high complexity of matrix inversion in the BP-MF method, we further employed GAMP to transform the vector estimation problem into a scalar one, and proposed a BP-MF-GAMP algorithm. Simulation results showed that the proposed BP-MF and BP-MF-GAMP algorithms achieved superior BER and NMSE performance compared to the state-of-the-art benchmarks. Moreover, our work provides a theoretical reference for solving the challenging joint channel estimation and data detection problem of other emerging nonlinear systems.

## Appendix A

DERIVATIONS OF (10)
First, we expand the term $\mid y_{n}-\sum_{l=0}^{L} h_{l} x_{n-l}-$

$$
\left.\sum_{i=0}^{L} \sum_{j \geq i}^{L} \sum_{k=0}^{L} h_{i j k} x_{n-i} x_{n-j} x_{n-k}^{*}\right|^{2} \text { as }
$$

$$
\left|y_{n}-\sum_{l=0}^{L} h_{l} x_{n-l}-\sum_{i=0}^{L} \sum_{j \geq i}^{L} \sum_{k=0}^{L} h_{i j k} x_{n-i} x_{n-j} x_{n-k}^{*}\right|^{2}
$$

$$
=y_{n}^{*} y_{n}-2 \Re\left(y_{n}^{*} \sum_{l=0}^{L} h_{l} x_{n-l}+y_{n}^{*} \sum_{i, j \geq i, k} h_{i j k} x_{n-i} x_{n-j} x_{n-k}^{*}\right)
$$

$$
+\sum_{l=0}^{L} h_{l} h_{l}^{*} x_{n-l} x_{n-l}^{*}+2 \Re\left(\sum_{i=0}^{L} \sum_{j>i}^{L} h_{i}^{*} h_{j} x_{n-i}^{*} x_{n-j}\right)
$$

$$
+2 \Re\left(\sum_{l=0}^{L} \sum_{i, j \geq i, k} h_{l}^{*} h_{i j k} x_{n-l}^{*} x_{n-i} x_{n-j} x_{n-k}^{*}\right)
$$

$$
+\sum_{i, j \geq i, k} h_{i j k}^{*} h_{i j k} x_{n-i}^{*} x_{n-j}^{*} x_{n-k} x_{n-i} x_{n-j} x_{n-k}^{*}
$$

$$
\begin{equation*}
+2 \Re\left(\sum_{i, j, k, a, b, c} h_{i j k}^{*} h_{a b c} x_{n-i}^{*} x_{n-j}^{*} x_{n-k} x_{n-a} x_{n-b} x_{n-c}^{*}\right) \tag{42}
\end{equation*}
$$

In (42), $y_{n}^{*} y_{n}$ is a constant, and all the combinations can be split into two parts, namely those with and without real operation. Then, we take the expectation of (42) with respect to the beliefs of all $x_{i}$ for $i \neq n$ and all $h_{l}$ and $h_{i j k}$. Accordingly, these variables can be replaced by their expectations extracted from the a posteriori pdfs. After some further manipulations, the message $\mu_{f_{r, n} \rightarrow x_{n}}\left(x_{n}\right)$ can be written in the form of its sufficient statistics and the canonical parameters in (10), where the canonical parameters $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ consist of the combinations of expectations of the related variables.

## Appendix B

DERIVATIONS OF PARAMETERS $\varphi_{a, b}^{(n)}$ AND $\varphi_{m, v}^{(n)}$
For the derivations of parameters $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$, we take $\Re\left(x_{n}\left|x_{n}\right|^{4}\right)$ as an example. Considering real-valued operations and the form of this term, we focus our attention on the last term $2 \Re\left(\sum_{i, j, k, a, b, c} h_{i j k}^{*} h_{a b c} x_{n-i}^{*} x_{n-j}^{*} x_{n-k} x_{n-a} x_{n-b} x_{n-c}^{*}\right)$ in (42). Then, we have $2 \Re\left(\sum_{i=1}^{L} h_{000}^{*} h_{00 i} x_{n}^{*} x_{n}^{*} x_{n} x_{n} x_{n} x_{n-i}^{*}\right)$ and
$2 \Re\left(\sum_{i=1}^{L} h_{0 i 0}^{*} h_{000} x_{n}^{*} x_{n-i}^{*} x_{n} x_{n} x_{n} x_{n}^{*}\right)$. When the expectations are calculated, all variables except for $x_{n}$ are replaced by their expectations, e.g., $\mathbb{E}\left(h_{000}^{*} h_{00 i} x_{n}^{*} x_{n}^{*} x_{n} x_{n} x_{n} x_{n-i}^{*}\right)=$ $\mathbb{E}\left(h_{000}^{*} h_{00 i} x_{n-i}^{*}\right) x_{n}^{*} x_{n}^{*} x_{n} x_{n} x_{n}$. Since the moment of the product of two statistically independent random variables is equal to the product of the two moments of the corresponding independent random variables, we have $\mathbb{E}\left(h_{000}^{*} h_{00 i} x_{n-i}^{*}\right)=\mathbb{E}\left(h_{000}^{*}\right) \mathbb{E}\left(h_{00 i}\right) \mathbb{E}\left(x_{n-i}^{*}\right)$. Accordingly, the parameters $\varphi_{m, v}^{(n)}(m=3, v=2)$ can be obtained. Given the above derivation, all the $\varphi_{a, b}^{(n)}$ and $\varphi_{m, v}^{(n)}$ can be derived.

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[^1]:    ${ }^{1}$ The term requiring the real part operation in (10) is simplified as $\Re\left(x_{n}^{m} x_{n}^{* v}\right)$ in Table II.

