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3 What I don't know can hurt you: Collateral combat damage seems more acceptable
4 when bystander victims are unidentified

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21 **Author Contributions**

22 Conception, design, and data collection for study 1: PC and AV. Conception for
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27

28 **Abstract**

29 Five experiments ($N=2,204$) examined responses to a realistic moral dilemma: a military pilot
30 must decide whether to bomb a dangerous enemy target, also killing a bystander. Few people
31 endorsed bombing when the bystander was an innocent civilian; however, when the
32 bystander's identity was unknown, over twice as many people endorsed the bombing.
33 Follow-up studies tested boundary conditions and found the effect to extend beyond modern-
34 day conflicts in the Middle East, showing a similar pattern of judgment for a fictional war.
35 Bombing endorsement was predicted by attitudes towards *total war*, the theory that there
36 should be no distinction between military and civilian targets in wartime conflict. Bombing
37 endorsement was lower for UK compared to US participants due to differences in total war
38 attitudes. This work has implications for conflicts where unidentified bystanders are common
39 by revealing a potentially deadly bias: people often assume unidentified bystanders are guilty
40 unless proven innocent.

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42 **Keywords:** Morality, war, judgment, decision-making

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48 **Introduction**

49 About as many civilians as soldiers die in war each year [1, 2], some during strikes
50 targeted at enemy combatants. For instance, US airstrikes targeting ISIS militants reportedly
51 caused a steep civilian death toll across multiple incidents [3, 4]. Some of these may be
52 calculated sacrifices, but there have been many reported cases of mistaking innocent civilians
53 for enemy combatants, with the possibility of many more being unreported [4, 5]. Why do
54 military strikes so often strike innocent non-combatants? The current research offers five
55 studies testing a potential explanation why: people tend to assume unknown bystanders in a
56 combat zone are enemies rather than civilians, reducing concerns about collateral damage.

57 Many factors influence sacrificial decisions, such as emotional impact of harm,
58 number of lives saved, and self-presentation [6-10]. Although most dilemma research ignores
59 social relationships [11, 12], sacrifices increase when targets appear different from decision-
60 makers [13] or evil and blameworthy [14]. Accordingly, war increases willingness to
61 sacrifice outgroups to achieve utilitarian outcomes [15].

62 In war, danger is prevalent. People are generally sensitive to threats [16] and bad
63 intentions [17], but wartime may amplify such processing. Threatening environments
64 increase detection of harmful agents [18], categorization of others as outgroups [19], and
65 perception of outgroups as threatening [20] and untrustworthy [21]. Hence, wartime threats
66 may lead people to assume that unknown targets are enemies, justifying attack.

67 Moreover, war entails uncertainty. Theoretically, uncertainty could either amplify or
68 decrease hesitation to sacrifice. For example, sacrificing one person who *might* die to save
69 five who *might* live seems worse than definitely sacrificing one to save five [22, 23]. On the
70 other hand, uncertainty can also absolve deciders of responsibility for their decisions [24],
71 suggesting more unknowns in war might make wartime killing seem less risky. For instance,
72 people may be more willing to sacrifice bystanders of uncertain identities, due to the

73 possibility they are enemies. Prospect theory argues uncertainty can be appealing or
74 unappealing depending on whether s decision-maker is trying to minimize losses or maximize
75 gains [25]. Loss consideration might be salient in war, but wartime conflict could also be
76 framed in terms of strategic gains, making it unclear what prospect theory would predict in
77 terms of wartime moral judgments. Taken together, this work shows that uncertainty can
78 impact judgments, especially concerning losses, moral elements, and utilitarian choices.
79 Despite a large body of evidence and theory however, it remains unclear what these findings
80 predict for wartime moral judgments.

81 Killing non-combatants in an enemy nation during war is not always considered a
82 sacrifice. *Total war* approaches to combat conceptualize civilians as legitimate military
83 targets [26, 27] because they conceptualize the struggle as between entire nations. This
84 thinking was common during the Second World War and despite it being less common in
85 modern warfare, it is still present today. People who endorse total war principles may view
86 enemy civilian deaths as acceptable or even desirable, compared to people who instead favor
87 *limited warfare* between military combatants only. For supporters of total war, civilian deaths
88 are not seen as a sacrifice or an error in judgments, but an inevitable feature of war.

89 Therefore, across five studies, ($N=2204$), we investigated to what extent people would
90 endorse a military bombing that would as a side effect kill an unidentified bystander.
91 Participants considered a dilemma set in the real-world conflict between the United States
92 and Islamic State (ISIS) (Tables S1-S3 for full text). Studies 1 and 2 explored the extent to
93 which participants endorsed bombing which would kill an unidentified bystander. Using
94 experimental vignettes, we examined how a lack of information about a bystander impacts an
95 individual's endorsement of the bombing as compared to a known bystander (innocent or
96 otherwise) or non-innocent combatant. We also examine whether participants given no
97 information about a wartime bystander make any assumptions about their allegiance, i.e.,

98 whether they are combatants or civilians. In Study 3 we extended this paradigm outside the
99 real-world setting of the ISIS conflict into a fictional conflict, where individual judgments
100 would not favor one side of the conflict over another. Study 4 investigated how attitudes
101 toward total war impact endorsement of bombing civilian bystanders, and Study 5 replicated
102 this and Study 1 among UK, rather than US, participants. Study 1 was exploratory; we
103 preregistered Studies 2-5.

104 Studies were approved by the Florida State University Human Subjects Ethics
105 Committee, HSC No. 2016.18130 and The University of Canterbury Human Research Ethics
106 Committee, HEC 2020/16/LR-PS. We obtained consent for all studies by asking participants
107 to select a mandatory consent item before proceeding with the study. The dates of running
108 each study were as follows: October 1-2, 2016, Study 1b: May 31- June 3 2020, Study 1c:
109 January 14 2021, Study 2: September 17 2020, Study 3: August 17-20, 2020, Study 4:
110 February 28, 2022, Study 5: April 11, 2022, Supplemental Study 1: March 18-19, 2021. All
111 materials are reported between the main text and supplemental materials. Datasets and code
112 can be found at: [https://osf.io/c5hjq/?view_only=71191a9d5fe14c6fbf548418d1a04a51].

113 **Study 1a**

114 Study 1a examined whether the identity of a sacrificial bystander influences decisions
115 to bomb a military target. We compare how many participants endorse a military strike that
116 would kill an unknown bystander to a strike that would kill an either innocent bystander,
117 another combatant, or one of several other bystanders tied in some way to the enemy.

118 **Method**

119 In 2016, we recruited 464 adult Americans via Amazon's Mechanical Turk (MTurk).
120 We analyzed all completed responses; no exclusions. In a sensitivity analysis, the pwr
121 package [28] for R Statistics revealed 80% power to detect effect sizes of Cohen's $w=0.17$.

122 We report demographics for all studies in Table 1. Participants read about a US pilot in Iraq
 123 who must choose whether to a) fire a missile on a farmhouse where there was an ISIS
 124 combatant known for making chemical weapons and bystander thus killing them both, or b)
 125 refuse to bomb, letting both individuals live (S1 Table for full text).

126

127 **Table 1. Sample Sizes and Demographics for all Studies.**

Study	Study Manipulation	<i>N</i>	Nationality	Mean Age (<i>SD</i>)	Gender (% Female)
1a	Bystander's association to the enemy target (7 levels)	464	United States	36.70 (12.00)	57.97% (0.28% other)
1b	Bystander's association to the enemy target (7 levels)	271	United States	30.90 (10.78)	47.23% (0.00% other)
1c	Unidentified bystander only	93	United States	31.23 (11.19)	52.68% (0.02% other)
2	Percent chance the bystander is an enemy combatant (11 levels)	297	United States	37.76 (11.97)	41.75% (1.01% other)
3	Real world wars versus fictional wars (2 levels)	203	United States	33.15 (11.29)	59.70% (1.00% other)
4	Foreign/local and innocent/guilty bystander; total war attitudes	579	United States	38.4 (14.4)	50.26% (0.52% other)
5	UK replication of Study 1 (7 levels)	297	United Kingdom	37.99 (13.22)	49.83% (0.68% other)

128

129

130 We varied the bystander's identity across seven conditions. We manipulated the
131 bystander's combatant status and degree of support and involvement in the conflict. We
132 described the bystander as either a) another ISIS soldier, b) a devout Muslim who agrees with
133 ISIS's extremism, c) a civilian farmer who profits off the war, d) a devout Muslim who feels
134 conflicted about extremism, e) a devout Muslim who rejects extremism, f) an innocent
135 civilian farmer, or g) an unidentified person about whom nothing is known. These conditions
136 aim to create a spectrum of involvement ranging between an enemy combatant and an
137 uninvolved civilian bystander.

138 To measure bombing endorsement participants decided, "Should the pilot fire upon
139 the building to kill both people inside?" (*Yes/No*) and "How acceptable is it for the pilot to
140 fire upon the building, killing both people inside?" on a scale from 1 (*Not at all*) to 7 (*Very*
141 *much*). We predicted that people would endorse sacrificing the innocent civilian bystander
142 less compared to another enemy combatant, and all other conditions would fall between these
143 extremes. However, participants may endorse the bombing wholeheartedly or reluctantly. To
144 assess this aspect of the dilemma we included three questions: "How comfortable are you
145 with your decision?" (1-7, *Not very-Very*), "How much does the ISIS Operative who makes
146 chemical weapons deserve to die?" (1-7, *Not at all-Very much*), and "How much does the
147 second person who was already in the farmhouse deserve to die? (1-7, *Not at all-Very much*).

148 **Results**

149 We anticipated that people would be sensitive to the identity of a bystander in the line
150 of fire when judging a potential military strike. As predicted, the bystander's identity
151 significantly impacted endorsement of bombing for the binary yes/no question,
152 $\chi^2(6,464)=100.5, p<.001, w=0.47$ (Fig 1). The Likert style question which asked about the
153 acceptability of firing also showed a significant impact of bystander identity with a similar
154 pattern of results in a non-parametric Kruskal-Wallis ANOVA, $\chi^2(6,456)=84.98, p<.001,$

155 $\epsilon^2=0.18$ (Fig S1). In the binary yes/no question people were most willing to bomb when the
 156 bystander was an enemy (77.9%) and least willing (24.2%) when they were an innocent
 157 civilian. Willingness to bomb decreased relative to how much the bystander was involved in
 158 the conflict. For example, 53.7% agreed to bomb a Muslim extremist who supported ISIS and
 159 only 31.8% endorsed the bombing when the bystander was a farmer who benefitted
 160 financially from the ISIS takeover.

161

162 **Fig 1. Percent of “Yes” Responses to “Should The Pilot Fire?” Depending on Bystander**

163 **Identity, Study 1a.** Error bars reflect 95% CIs. Asterisks denote Bonferroni corrected p-
 164 values compared with the unidentified bystander condition (* $p<.05$, ** $p<.01$, *** $p<.001$).

165 More innocent and uninvolved bystanders see a reduction in firing compared to bystanders
 166 who were guilty (e.g., ISIS soldier) or tied to ISIS (e.g., Extremist Muslim, civilian war
 167 profiteer). This pattern does not hold for the unidentified bystander condition, for which most
 168 (58.5%) participants said the pilot should fire.

169

170 Unexpectedly, over half of participants (58.5%) endorsed bombing unidentified
 171 bystanders despite no concrete evidence they were enemies. This was roughly twice the
 172 bombing rate of civilian bystanders, similar to bombing Muslim extremists. Bonferroni
 173 corrected pairwise comparisons in a generalized linear model using a log link transform
 174 showed that people were significantly more willing to bomb an unidentified than innocent
 175 civilian ($exp(B)=1.48$, $SE=.382$, $p=.002$), anti-ISIS Muslim ($exp(B)=2.21$, $SE=.438$, $p<.001$),
 176 a Muslim conflicted about ISIS ($exp(B)=1.78$, $SE=.398$, $p<.001$), or a war profiteer
 177 ($exp(B)=1.11$, $SE=.085$, $p=.058$). Conversely, unidentified bystander bombing rates did not
 178 significantly differ from an ISIS soldier ($exp(B)=-.920$, $SE=.386$, $p=.358$), or the Muslim
 179 extremist ($exp(B)=.192$, $SE=.351$, $p=1.00$, Fig 1).

180 Several more items assessed the manipulation of the different bystander identities.

181 When participants were asked how comfortable they were with their decision, there was no

182 impact of condition in a non-parametric Kruskal-Wallis one way ANOVA, $\chi^2(6)=10.08$,

183 $p=.121$, $\epsilon^2=0.02$. However, an exploratory analysis found an effect of participant's decisions

184 to endorse bombing: those who opted not to bomb were significantly more comfortable with

185 their decision overall ($M=5.38$, $SD=1.75$) than those who opted to bomb ($M=4.63$, $SD=1.96$),

186 $U(N_{\text{Endorse}}=185, N_{\text{Not endorse}}=279)=20085$, $p<.001$, $r=0.22$. Participants were also asked in

187 separate questions how much each person—the targeted ISIS fighter and the bystander—

188 deserved to die. Participants showed high agreement that the ISIS fighter deserved to die

189 across conditions ($M=5.08$, $SD=1.87$) with no difference between conditions which changed

190 the bystander, $\chi^2(6)=2.38$, $p=.881$, $\epsilon^2=0.01$. However, ratings of bystander's deservingness to

191 die varied significantly across the different bystander identities, $\chi^2(6)=173.46$, $p<.001$,

192 $\epsilon^2=0.38$. This broadly followed the same pattern as firing acceptability and firing

193 endorsement (Fig 2). However, participants rated the unidentified bystander as significantly

194 more deserving of death than any other non-combatant bystanders—even the ISIS

195 sympathizer and war profiteer—according to a Mann-Whitney test, $U(N_{\text{non-combatants}}=330$,

196 $N_{\text{unidentified}}=65)=5040.50$, $p<.001$, $M_{\text{Diff}}=-2.00$, 95%CI[-3.00, -2.00], rank biserial

197 correlation=0.53 (coded unidentified bystander 1, other civilian bystanders 0). However,

198 participants rated the unknown bystander as less deserving than the ISIS soldier bystander,

199 $U(N_{\text{Soldier}}=68, N_{\text{unidentified}}=65)=1274.00$, $p<.001$, $M_{\text{Diff}}=-1.00$, 95%CI[1.00, 2.00], biserial rank

200 correlation=0.42. This suggests that not only do many participants endorse a bombing that

201 would kill an unidentified bystander, many also judge that the unidentified bystander

202 deserved it.

203

204 **Fig 2. Rating of Bystander Deservingness to Die Depending on Identity, Study 1a.** Error
205 bars represent a 95% confidence interval. Asterisks denote Games-Howell corrected p-values
206 compared with the unidentified bystander condition (* $p < .05$, ** $p < .01$, *** $p < .001$).

207 Participants answered the question “How much does the second person [the bystander]
208 deserve to die?” with a Likert scale between 1 (*Not at all*) and 7 (*Very much*). The bystander
209 rated most deserving of death was the ISIS soldier ($M=4.49$, $SD=1.93$), the second highest
210 was the unidentified bystander ($M=3.1$, $SD=1.55$).

211

212 **Discussion**

213 To kill a dangerous enemy, people were more willing to sacrifice an unknown
214 bystander than any known bystander except enemy combatant or Muslim extremist.
215 Likewise, people rated the unknown bystander more deserving of death than any known
216 bystander except an enemy combatant. These results suggest people treated unidentified
217 bystanders more like enemy combatants than innocent (or even morally compromised)
218 civilians.

219 **Study 1b**

220 Study 1a surveyed participants in 2016 during the height of the US-ISIS conflict.
221 Shifting public sentiment about this conflict may limit its generalizability. Study 1b
222 replicated Study 1a in 2020 after much of the threat and news coverage of the ISIS conflict
223 had abated. Preregistered: https://aspredicted.org/EJU_BYG

224 **Method**

225 We recruited 317 American participants from mTurk after excluding incomplete
226 responses and participants who failed a competence check we were left with a final sample of
227 271. A sensitivity analysis showed 80% power to detect effect sizes of Cohen’s $w=0.22$. We

228 presented the same materials and measures as study 1a and added an item asking participants
229 to explain their reasoning for their decision about the bombing in a free response format.

230 **Results**

231 As in Study 1a which was run in 2016, the identity of the bystander again mattered for
232 endorsing the bombing, $\chi^2(6)=49.42, p<.001$. The patterns of results for endorsing the
233 bombing were similar to Study 1a (Fig 3). When both datasets were combined and Study was
234 included as a factor along with condition to predict bombing endorsement, there was no main
235 effect of study, $\chi^2(1)=1.78, p=.183, \exp(B)=0.79$, but a small significant interaction between
236 condition and study, $\chi^2(6)=14.43, p=.025$. However, no Bonferroni corrected pairwise
237 comparisons between conditions across studies 1a and 1b reached significance. The greatest
238 difference was in the Muslim extremist condition, which saw less endorsement of firing in
239 2020 (23.1%) than it did in 2016 (48.1%), but nevertheless the difference was nonsignificant,
240 $\exp(B)=3.87, p=.251$. Additionally, as in Study 1a, ratings of acceptability of the bombing
241 showed a significant effect of condition in a Kruskal-Wallis ANOVA, $\chi^2(6, 264)=54.91,$
242 $p<.001, \epsilon^2=0.20$. The pattern of results was also similar to that of Study 1a (Fig 3).

243

244 **Fig 3. A) Percent of “Yes” Responses to “Should The Pilot Fire?” and B) Rating of**
245 **firing acceptability Depending on Bystander Identity in 2016 (Study 1a) and 2020**
246 **(Study 1b).** Error bars represent a 95% confidence interval. A) Bonferroni-corrected post-hoc
247 comparisons showed no significant differences in firing endorsement by study for any
248 condition. B) Dwass-Steel-Critchlow-Fligner pairwise comparisons showed no significant
249 differences in bombing acceptability by study for any condition.

250

251

252

253 Participants were asked how comfortable they were with their decision, which unlike
 254 study 1a showed a medium effect of condition in a Kruskal-Wallis ANOVA, $\chi^2(6)=16.62$,
 255 $p=.011$, $\varepsilon^2=0.06$. Pairwise comparisons showed significant differences only in the conflicted
 256 Muslim condition, in which participants rated their comfort as significantly lower than
 257 participants in the Innocent condition ($W=4.78$, $p=.013$) and the anti-ISIS Muslim condition
 258 ($W=4.46$, $p=.027$). However, this effect is likely due to differences in firing rates between the
 259 conditions as similar to study 1a, participants who endorsed bombing rated their comfort as
 260 significantly lower ($M=4.49$, $SD=1.82$) than participants who did not endorse bombing
 261 ($M=5.45$, $SD=1.74$), $U(N_{\text{Endorse}}=94, N_{\text{Not endorse}}=177)=5681$, $p<.001$, $r=0.32$.

262 When asked to rate how much the ISIS fighter and the bystander each deserved to die,
 263 participants were again high in condemnation of the target ISIS soldier ($M=5.43$, $SD=1.48$),
 264 an effect which did not differ by condition, $\chi^2(6)=5.19$, $p=.520$, $\varepsilon^2=0.02$. However, as in
 265 Study 1a, there was a strong impact of condition on ratings of whether the bystander deserved
 266 to die, $\chi^2(6)=66.37$, $p<.001$, $\varepsilon^2=0.25$. Pairwise comparisons showed the unidentified
 267 bystander condition ($M=2.95$, $SD=1.67$) was significantly lower than the ISIS soldier
 268 condition ($M=4.41$, $SD=1.83$, $W=5.11$, $p=.006$), and significantly higher than the conflicted
 269 Muslim ($M=1.89$, $SD=1.47$, $W=-4.26$, $p=.041$), the innocent bystander ($M=1.88$, $SD=1.62$,
 270 $W=-4.53$, $p=.023$), the anti-ISIS Muslim ($M=1.76$, $SD=1.45$, $W=-5.04$, $p=.007$). The
 271 unidentified bystander condition was marginally higher than the extremist Muslim ($M=1.95$,
 272 $SD=1.64$, $W=-4.17$, $p=.050$), and no different from the war profiteer bystander ($M=1.94$,
 273 $SD=1.47$, $W=-3.80$, $p=.102$). These findings largely mirror the findings of Study 1a.

274 **Qualitative justifications**

275 We asked participants to justify their bombing decision with a free response of a few
 276 sentences. Participant answers were grouped by the bystander's identity and the participant's
 277 decision to fire. A coding scheme was developed post hoc that included 14 categories (Table

278 S4). Across conditions, the most common type of justification for endorsing the bombing was
279 *The greater good* ($N=54$ out of 94, 57%) which included statements like “The death of the
280 innocent goat farmer, however tragic, is one death compared to the possible hundreds of
281 deaths that could result from the manufacture of chemical weapons from the ISIS operative
282 also killed in the missile attack. One innocent's death compared to the death of hundreds of
283 innocents,” and “I would personally hate myself for pulling the trigger and killing both
284 people, but in this scenario it's the one for the many. If the chemical weapons guy gets to
285 make chemicals it could lead to greater loss of life. Not happy about my decision, but that's
286 the one that likely has to be made as is war.” For participants who did not endorse the
287 bombing the most common kind of justification was *Appeal to innocence* ($N=124$ out of 177,
288 70%) which included statements like “He should not - would never condone the killing of an
289 innocent person,” and “In America we do not believe in killing innocent people.” Because in
290 the majority of our conditions the bystander held some degree of innocence in the conflict,
291 this kind of justification being so common is reasonable.

292 **Justifying bombing the unidentified bystander**

293 Study 1a showed participants were surprisingly accepting of the bombing that would
294 kill the unidentified bystander, so we looked specifically at the justifications of participants
295 who opted to fire in the unidentified bystander condition. In the unidentified condition 18 out
296 of 37 (49%) chose to fire. Of those, 8 (44%) cited *The greater good* as a justification with
297 answers like “Can't afford to risk the people being able to make more weapons,” and “If the
298 pilot doesn't fire the missile and kill the operative then many more than just one other
299 innocent person may die.” This is reflective of the data overall across conditions. However,
300 unique to this condition was 5 (28%) participants who justified the bombing on the
301 assumption that the bystander was part of ISIS including statements like “I said yes because
302 more than likely he ran into someone he knew at home (sic) so they probably are with him

303 making them more than likely ISIS also,” and “Chances are that the person the operative ran
304 into the house with is also an operative of ISIS.” This sample is quite small, but suggests that
305 ambiguity about the identity of a bystander may lead around a quarter of participants to
306 assume they are an enemy and endorse a bombing they otherwise may have not.

307 **Discussion**

308 Overall, the main finding that endorsement of the bombing was much higher for the
309 unidentified bystander compared to the innocent bystander was replicated from Study 1a.
310 This suggests the finding was not dependent on public sentiment at the height of the armed
311 conflict in 2016 and extended into 2020. Additionally, exploratory analysis into how
312 participants justified their decision showed that around a quarter (28%) of participants in the
313 unidentified bystander condition justified the bombing because of the possibility or the
314 assumption that the unidentified bystander was an enemy combatant rather than an innocent
315 civilian. It remains unclear to what extent this assumption impacts participant’s reasoning
316 about the dilemma, so in a follow-up study we asked directly.

317 **Study 1c**

318 When bystanders to a military target are unidentified, there is a possibility they are a
319 civilian, but also a possibility they are another enemy combatant. Studies 1a and 1b showed
320 participants endorse bombing an unidentified bystander at rates much higher than an
321 innocent. Many participants even reported they thought the unidentified bystander deserved
322 to die. Furthermore, when given a chance to explain their decision, around a quarter of
323 participants who endorsed bombing the unidentified bystander believed the bystander was
324 likely to be a combatant rather than a civilian. A possible explanation for the high rate of
325 bombing the unidentified bystander is a tendency to assume they are an enemy. To
326 investigate how participants are thinking about unidentified bystanders we asked directly.

327 Given participants have no information about this person, if participants are unbiased, we
328 would expect no consistent pattern of answers, i.e., the average response should be around the
329 midpoint. If, however, participants tend to assume guilt, we would expect significantly more
330 “yes” answers and a higher probability the bystander is an enemy, especially for those who
331 endorse the bombing. Preregistered: https://aspredicted.org/TPK_HTT

332 **Method**

333 To better understand the responses to the unidentified bystander condition from
334 Studies 1a and 1b, we ran a small-scale replication of only the unidentified bystander
335 condition ($N=93$) as part of a larger study run in 2021 and not reported here. We asked
336 questions identical to Study 1a and added two added additional questions: “The pilot
337 followed the ISIS operative to a farmhouse where a second person already was. Do you think
338 this second person is another member of ISIS?” (*Yes/No*) and “Please rate how likely it is the
339 second person is another ISIS member” on a scale from 1 (*Extremely unlikely*) to 100
340 (*Extremely likely*). The former, binary measure was analyzed with a proportion test while the
341 latter continuous measure was analyzed with a one sample t test. For a proportion test, a
342 sensitivity analysis showed 80% power to detect effect sizes of $h=0.29$. For a one sample t
343 test, a sensitivity analysis showed 80% power to detect effect sizes of *Cohen’s* $d=0.29$.

344 **Results**

345 In a replication of the unidentified bystander condition ($N=93$), when asked, a
346 majority assumed the unidentified bystander was part of ISIS (61%, significantly $>50\%$,
347 $p=.038$, 95%CI [60.62%, 71.22%]), with an average probability judgment above the
348 midpoint, $M=57.60$, $SD=23.52$, student’s $t(92)=3.12$, $p=.002$, 95%CI [52.76, 62.45], $d=0.32$.
349 Participants who endorsed bombing gave a higher percent chance the bystander was a
350 combatant ($M=70.89$, $SD=17.45$) than those who refused to bomb ($M=45.67$, $SD=21.92$),
351 $t(91)=-6.09$, $p<.001$, $M_{Diff}=-25.21$, $SE_{Diff}=4.14$, 95%CI [-33.44,-16.99], $d=-1.26$. These results

352 demonstrate that many people believe an unidentified bystander to be an enemy combatant,
353 despite no evidence this may be the case. In the absence of identity information, participants
354 assuming the bystander was an enemy combatant felt free to endorse the bombing. This
355 pattern suggests either motivated justification of bombing, or increased willingness to bomb
356 given assumptions of enemyship.

357 **Discussion**

358 Participants were significantly more likely to assume the unidentified bystander is an
359 enemy rather than a civilian. Participants may endorse firing because of this assumption
360 about a completely unknown person in the warzone. Perhaps the mere presence in a warzone
361 is taken to suggest that unidentified individuals must be enemies—a dangerous assumption.
362 To test this, Study 2 directly manipulated the probability the unidentified target was civilian
363 and measured endorsement of bombing.

364 **Study 2**

365 In Study 1, participants frequently endorsed sacrificing an unknown bystander to kill
366 an enemy, many assuming the bystander was also an enemy. Different people may have
367 differing prior beliefs about the chances of a random person in a warzone being a combatant.
368 Therefore, we directly manipulated the likelihood that the unknown bystander is a civilian in
369 the vignette by giving a precise percentage the bystander was an enemy or civilian. We again
370 measured participant's endorsement of the bombing with the same measured used in Study
371 1a. Preregistered: aspredicted.org/RDM_ABG.

372 **Method**

373 Participants were 309 adult American MTurkers, with twelve (3.9%) excluded for
374 failing a competence check, leaving 297 (Table 1). We aimed for >250 participants, a
375 minimum for stable correlations [29]. A sensitivity analysis using G*Power 3 [30] showed

376 80% power to detect effect sizes of $OR=0.66$ in a binary logistic regression. Participants read
377 the unidentified bystander vignette from Study 1, except we manipulated the percentage
378 likelihood the bystander was a member of ISIS from 0% to 100% in 10% increments (e.g.,
379 “there is a 70% chance they are ISIS, 30% chance they are a civilian”). Additionally, we
380 changed the ISIS target from a chemical weapons maker to an ISIS leader (see supplement
381 for full text). We predicted that participants would rate sacrifice increasingly acceptable as
382 the chance they are ISIS increased, but a substantial number (50%) would sacrifice even at
383 maximal uncertainty. As in Study 1, we asked about firing endorsement as a binary yes/no
384 question and as a 7-point Likert item for the acceptability of firing. We also asked the same
385 manipulation check questions, with one change: instead of asking how deserving of death
386 each person was, we asked how appropriate each target would be. Specifically, “If the [ISIS
387 leader/unknown person] was alone in the farmhouse, would they be an appropriate target for
388 the pilot to fire on?” with a Likert response between 1 (*Not at all appropriate*) and 7 (*Very*
389 *appropriate*).

390 **Results**

391 When asked whether or not they endorsed the bombing, people were more willing to
392 bomb as the likelihood the bystander was ISIS increased in a binary logistic regression
393 model, odds ratio $\chi^2(1, N=297)=29.5, p<.001, (OR)=1.02, 95\% CI[1.01, 1.03]$, McFadden’s
394 $R^2=.07$ (Fig 4). When probability was 50%, close to half (40.0%) endorsed firing (95%
395 $CI[21.1, 61.3]$); at 40% probability a minority endorsed firing (25.9%, 95% $CI[11.1, 46.3]$).
396 This pattern shows even under uncertainty, many people still choose to bomb at rates similar
397 to Study 1a. Additionally, at 100% probability the bystander was a civilian, nonetheless
398 30.3% agreed to bomb, comparable to the innocent civilian condition in Study 1a (23.1%).
399 When rating the acceptability of firing on a 1-7 scale, bombing was similarly rated more
400 acceptable as the likelihood the bystander was ISIS increased in a linear regression

401 $F(1,295)=25.4, p<.001, R^2=0.08$. Consistent with Study 1a, people who endorsed firing were
 402 less comfortable with their decision ($M=4.47, SD=1.90$) than people who did not ($M=5.48,$
 403 $SD=1.69$), $U(N_{Fire}=117, N_{Don't Fire}=180)=7187.50, p<.001, biserial rank correlation=0.32$.

404

405

406 **Fig 4. Firing Endorsement by Percent Chance the Bystander is also in ISIS, Scatterplot**
 407 **with Applied Jitter, Study 2.** Participants were given a percent chance the bystander is an
 408 enemy combatant (ISIS member) from 0% to 100% and asked “should the pilot fire?
 409 [yes/no]. At 50% when there is an equal chance the bystander is an enemy versus a civilian,
 410 44% of participants say the pilot should fire. This proportion of “yes” answers increases
 411 linearly as the chance the bystander is ISIS increases above 50%. Includes a loess line with
 412 grey area indicating confidence region.

413

414

415 Participants also rated how appropriate it would be to bomb each individual, the ISIS
 416 soldier and the bystander, if they were alone. We computed a linear regression depending on
 417 percent chance the bystander is in ISIS. As expected for the ISIS leader there was no effect of
 418 condition, $F(1,295)=0.88, p=0.348, R^2=.003$, but participants rated the bystander as more
 419 appropriate target to fire on as the percent chance they are ISIS increased, $F(1,195)=22.97,$
 420 $p<.001, R^2=.072$. Overall, participants rated the ISIS leader as a highly appropriate target
 421 ($M=6.02, SD=1.46$) and rated the bystander as a less appropriate target ($M=1.87, SD=1.47$).

422 Discussion

423 These results supported both hypotheses. As the probability that an unidentified
 424 bystander was an enemy increased, so did willingness to kill them; yet, under the highest
 425 uncertainty (50% chance the bystander was ISIS), almost half of participants endorsed

426 bombing. Hence, many participants appear quite willing to risk sacrificing an innocent
427 civilian to kill an enemy. Such findings raise the question of mechanism. Perhaps this effect
428 reflects motivated reasoning by American participants due to their position in the conflict and
429 their reasonable bias against ISIS; if so then a different pattern should emerge for a fictional
430 conflict that avoids motivation to support one side over the other.

431 **Study 3**

432 We examined whether participants would be less willing to sacrifice innocent
433 bystanders in a fictional conflict. Studies 1 and 2 asked Americans about a real-world conflict
434 between the US and ISIS, where motivated reasoning favoring the ingroup may drive
435 assumptions, limiting generalizability. To test this possibility, we manipulated whether the
436 conflict was between America and ISIS or two fictional countries, where motivation and
437 ingroup affiliation should not matter. Specifically, in the real-world conflict condition we
438 replicated the unidentified condition from Study 1 regarding an ISIS fighter and unknown
439 bystander; in the fictional conflict condition participants read about an identical situation
440 about a pilot from “Nibia” contemplating a strike against “Sorovia Federation fighters” (see
441 supplement for full text). Prior studies described the target of the attack as an ISIS soldier
442 known for making chemical weapons; we worried this detail could make readers biased
443 against the target in the fictional conflict condition since chemical weapon use is widely seen
444 as taboo in modern warfare. For this reason, both conditions in Study 3 describe the target
445 more neutrally as “a prominent leader and military strategist for [ISIS/the Federation].” We
446 preregistered the study: aspredicted.org/QFW_ZFZ.

447 **Method**

448 We recruited 223 adult American MTurkers, excluding 20 (9%) for failing a
449 competence check, leaving 203. The pwr package for R revealed 80% power to detect

450 Cohen's $w=0.20$. The *real-world conflict* condition replicated the Study 1 unidentified
451 bystander condition. The *fictional conflict* condition was identical except involving fictional
452 countries *Nibia* and *Sorovian Federation* (Table S1). Participants reported bombing
453 endorsement as in Study 1a, and rated bombing acceptability 1=*not at all*, 7=*very much*. We
454 described the main target as *leader* rather than *chemical weapons maker* so as to not vilify
455 one side of the conflict. Participants also reported the same manipulation check questions as
456 Study 2. As in Study 1c, participants also reported whether they thought the bystander was an
457 enemy combatant or a civilian, both as a binary yes/no and as a continuous likelihood. For the
458 real-world conflict condition these were the same as Study 1c and the wording was adapted
459 for the fictional conflict condition, e.g. "The pilot followed the Sorovia Federation leader to a
460 farmhouse where a second person already was. Do you think this second person is another
461 member of the Sorovia Federation?"

462 **Results**

463 The real-fictional manipulation revealed no significant effect for the binary choice of
464 bombing or not, $\chi^2(1,203)=2.39$, $p=0.122$, $w=0.11$. Next, we tested whether the two
465 conditions were significantly *equivalent* via a two-sided equivalence test [31]. The TOST
466 procedure for two proportions, with equivalence bounds of the raw score of $\Delta_L=-0.1$ and
467 $\Delta_U=0.1$ (a 10% change in proportion who fired) revealed that the two conditions were not
468 statistically equivalent, as the larger of the two p values is greater than 0.05, $z=.03$, $p=.511$.
469 As both tests were nonconclusive, we ran similar analyses on the continuous measure of
470 bombing acceptability. As with the binary measure of bombing endorsement, the Likert scale
471 bombing acceptability measure also showed no difference between conditions, $t(201)=0.91$,
472 $p=0.363$. We ran a TOST procedure for an independent samples t -test with unequal variances
473 with equivalence bounds of $\Delta_L=-0.7$ and $\Delta_U=0.7$ (a 10% change in ratings of acceptability).

474 This test suggested that the two groups were statistically equivalent, because the larger p
 475 value was still less than .05, $t(196.2)=-4.07$, $p<.001$.

476 As in prior studies, participants who read the real-world war vignette and those who
 477 read the fictional war vignette were no different in their comfort with their decision, $t(201)=-$
 478 $.35$, $p=0.727$. As in previous studies, across groups, participants who opted to fire showed
 479 less comfort with their decision ($M=4.27$, $SD=1.78$) than those who opted not to fire
 480 ($M=5.20$, $SD=1.77$), $t(201)=-3.48$, $p<.001$. Participants did not differ between conditions in
 481 judgments of the appropriateness of bombing the enemy leader, *Welch's* $t(200.56)=1.29$,
 482 $p=0.200$, or the bystander, *Welch's* $t(194.09)=0.04$, $p=0.971$. However, as expected,
 483 participants judged the enemy leader as a significantly more appropriate target for firing a
 484 missile at ($M=5.78$, $SD=1.57$) compared to the bystander ($M=1.72$, $SD=1.41$), $t(202)=27.78$,
 485 $p<.001$).

486 When asked whether they thought the bystander was a combatant, there was no
 487 difference between participants in the real world conflict condition (48.51%, not different
 488 from 50%, $p=.842$) and the fictional conflict condition (50.00%), $\chi^2(1,203)=0.04$, $p=.832$.
 489 Similarly, when asked to give a likelihood that the bystander was a combatant, there was no
 490 difference between the real-world condition ($M=55.81$, $SD=21.95$) and the fictional condition
 491 ($M=56.69$, $SD=20.60$), $U(N_{\text{real-world}}=101, N_{\text{fictional}}=102) = 4973$, $p=.670$. The average
 492 likelihood the bystander was a combatant was significantly above 50% in both the real world
 493 condition, $W(101)=2326$, $p=.008$, 95%CI[52.00,63.00], rank biserial correlation = -.08, and
 494 the fictional condition, $W(102)=2276$, $p=.002$, 95%CI[53.00,65.00], rank biserial correlation
 495 = -.13.

496 Discussion

497 Whether participants considered a conflict between the US and ISIS or two fictional
 498 countries, a similar proportion endorsed killing an unknown bystander to bomb a military

499 target. This pattern emerged on both the dichotomous measure and a more sensitive
500 continuous measure with additional power to detect significant similarity. Hence, results do
501 not seem to reflect motivated reasoning predicated on participant's position in the conflict;
502 rather, they appear to reflect a general tendency to assume that unidentified bystanders are
503 likely enemy combatants.

504 Participants rated the bystander as more likely to be a combatant than a civilian in
505 both the real world and fictional conditions when asked to give a percentage. This is similar
506 to the finding in Study 1c and shows that the tendency to assume guilt over innocence on the
507 battlefield is not confined to the conflict between the US and ISIS. Though significant, this
508 difference was small, and when reporting the binary yes/no measure of whether they thought
509 the bystander was a combatant, participants in both conditions gave answers around 50%,
510 contrary to Study 1c which showed a strong majority favoring combatant. However, both
511 measures showed the pattern of judgment did not change whether participants judged a US
512 ISIS conflict or a conflict between fictional countries.

513 The high rate of bombing unidentified bystanders may seem an error in judgment.
514 Yet, from a "total war" perspective, treating civilians as part of a global struggle between
515 military powers, harming civilians may seem like a rational trade-off. Next, we examined
516 how participants' attitudes toward total war impacts their decisions.

517 **Study 4**

518 In Studies 1-3, participants considered bombings that would kill an enemy and a
519 bystander. Bystanders varied in relationship to enemies but were always described as locals
520 to the combat region. Perhaps the 20-30% of participants endorsing bombing innocent
521 bystanders viewed them as not 'wholly' innocent: rather people may infer enemyship from
522 the target's mere presence in the region. Such beliefs may reflect "total war" beliefs that

523 warfare involves a struggle between nations that extends beyond military combatants to
524 civilian populations who contribute indirectly to conflict.

525 If so, then participants may be more willing to sacrifice innocent local bystanders
526 (i.e., Iraqis) than members of neutral foreign nations (e.g., Sweden), as the latter cannot be
527 construed as an enemy combatant even under total war beliefs. Crucially, this effect should
528 pertain only for innocent targets, not those who join ISIS voluntarily (whether foreign or
529 local). Study 4 therefore manipulated bystander nationality (Iraqi vs Swede) and innocence
530 (documenting ISIS vs aiding ISIS). We also developed an exploratory measure of support for
531 total war using a novel questionnaire. We used this to assess whether total war beliefs predict
532 bombing endorsement; participants higher in such beliefs should be more willing to bomb,
533 especially for local vs foreign targets, preregistered: https://aspredicted.org/M8G_MGC.

534 **Method**

535 We recruited 602 American participants through Prolific, excluding 23 (3.8%) for
536 failing a competence check, leaving 579. The pwr package for R showed 80% power to
537 detect Cohen's $w=0.14$. The study used four versions of the Study 1a vignette, manipulating
538 bystander identity and guilt in a 2 (foreign vs local) by 2 (civilian vs ISIS affiliated) design.
539 The bystander was described as either "a local Iraqi reporter" or "a foreign reporter from
540 Sweden," who either writes articles "informing the world about ISIS activity," or "in support
541 of ISIS and their ideology." Participants answered the same questions as in Study 1a,
542 excluding the comfort with decision question.

543 As an exploratory measure, we developed ten total war belief questions (see
544 Supplement) to assess how much participants consider all citizens and infrastructure in
545 conflict areas legitimate targets for attack. For example, we asked, "During war it is
546 acceptable to bomb cities and other population centers in an enemy nation if it results in a
547 crucial strategic advantage to help end the war" and "In war, everything and anything is fair

548 game.” (1=*strongly disagree* - 7=*strongly agree*, $\alpha=0.92$, see Supplement). We predicted that
 549 people scoring higher on this measure would endorse bombing more often, especially for
 550 local versus foreign targets.

551 **Results**

552 **Endorsing bombing**

553 A generalized linear model predicting bombing depending on bystander guilt and
 554 nationality showed that only guilt predicted firing rates, $\chi^2(1,579)=133.30$, $p<.001$, $w=0.48$
 555 (Fig. 5). Neither nationality, $\chi^2(1,579)=0.55$, $p=.46$, $w=0.03$, nor the interaction were
 556 significant, $\chi^2(1,579)=0.13$, $p=.72$, $w=0.01$. Planned contrasts of Bonferroni corrected
 557 pairwise comparisons in a generalized linear model using a log link transform likewise
 558 showed sensitivity to bystander guilt, but not nationality. Specifically, we found significant
 559 differences between the following conditions: foreign-guilty vs. foreign-innocent
 560 ($Exp(B)=0.11$, $SE=.03$, $p<.001$), foreign-guilty vs. local innocent ($Exp(B)=0.13$, $SE=.04$,
 561 $p<.001$), foreign-innocent vs. local-guilty ($Exp(B)=10.59$, $SE=3.19$, $p<.001$), and local-guilty
 562 vs. local-innocent ($Exp(B)=0.12$, $SE=.03$, $p<.001$). However, there were no differences
 563 between the foreign-guilty vs. local-guilty ($Exp(B)=1.11$, $SE=.27$, $p=1.00$) or foreign-
 564 innocent vs. local-innocent conditions ($Exp(B)=1.22$, $SE=.42$, $p=1.00$). Hence, bystander
 565 nationality did not influence bombing; all significant effects were driven only by bystander
 566 affiliation with ISIS.

567

568 **Fig 5. The Percentage of Participants willing to Fire Depending on Whether the**
 569 **Bystander is Innocent vs. Guilty and Foreign (Swedish) vs. Local (Iraqi), Study 4.** Error
 570 bars reflect 95% CIs. Asterisks denote Bonferroni corrected p-values compared with the
 571 unidentified bystander condition ($***p<.001$). There is a significant increase in firing when

572 the bystander is guilty rather than innocent and no impact when they are foreign rather than
573 local.

574

575 **Acceptability of bombing**

576 Participants reported “*How acceptable is it for the pilot to fire upon the building,*
577 *killing both people inside?*” on a scale from 1 (*Not at all*) to 7 (*Very much*). A 2×2 between-
578 subjects analysis of variance showed higher scores in the guilt vs innocence conditions,
579 $F(1,567)=140.03, p<.001, \eta^2=0.20$, but no significant difference between local vs foreign
580 conditions, $F(1,567)=0.36, p=.551, \eta^2=0.00$, and no significant interaction, $F(1,567)=0.42,$
581 $p=0.515, \eta^2=0.00$. This pattern of continuous firing acceptability matched the pattern of
582 dichotomous yes/no bombing endorsement.

583 **Deservingness to die**

584 As a manipulation check of perceived bystander innocence vs guilt, participants rated
585 how much both the bystander and the target (ISIS operative) deserved to die on scales from 1
586 (*Not at all*) to 7 (*Very much*). We computed a linear regression on bystander’s deservingness
587 to die depending on total war attitudes, guilt, and nationality, plus all interactions. The overall
588 model was significant, $F(7,562)=82.93, p<.001, R^2=.51$. However, only the interaction
589 between total war beliefs and bystander guilt emerged as significant, $t(579)=5.25, p<.001$:
590 when the bystander was guilty, people high in total war beliefs rated them more deserving of
591 death than people low in total war beliefs. When the bystander was known to be innocent,
592 total war beliefs had no impact on deservingness ratings (Fig 7). No other effects were
593 significant (Table 2). As expected, participants consistently rated the ISIS operative high in
594 deservingness to die with no difference between conditions, $F(3,318)=0.77, p=.51 (M=5.44,$
595 $SD=1.60)$.

596 **Total war attitudes questionnaire**

597 We conducted a principle components analysis using oblimin rotation with 500
 598 iterations before convergence and 500 for rotation, retaining all factors with an eigenvalue
 599 greater than 1 [32]. Results showed a single factor with an eigenvalue of 5.96 accounting for
 600 59.57% of the variance. Therefore, we treated these items as a single reliable measure
 601 ($\alpha=.92$).

602 Next, we conducted a logistic regression on the decision to fire depending on total
 603 war attitudes, guilt, and nationality. The overall model was significant, $\chi^2(4,579)=287.52$,
 604 $p<.001$, $w=0.70$, with significant effects of total war attitudes, $\chi^2(1,579)=152.91$, $p<.001$,
 605 $w=0.51$, and guilt, $\chi^2(1,579)=137.66$, $p<.001$, $w=0.49$ (Fig 6), but not nationality,
 606 $\chi^2(1,579)=.50$, $p=.48$, $w=0.03$, nor the interaction, $\chi^2(1,579)=.03$, $p=.87$, $w=0.01$.

607 The pattern was similar for scale ratings of bombing acceptability. We computed a
 608 linear regression on acceptability depending on total war attitudes, bystander guilt, and
 609 bystander nationality. The overall model was significant, $F(3,566)=153.29$, $p<.001$, $R^2=.45$,
 610 showing a small to moderate relationship (Table 2). Both total war attitudes $t(579)=15.93$,
 611 $\beta=.50$, $SE=.06$, $p<.001$, and bystander guilt, $t(579)=12.31$, $\beta=.77$, $SE=.13$, $p<.001$, but not
 612 bystander nationality, $t(579)=.62$, $\beta=.04$, $SE=.13$ $p=.54$, significantly predicted firing
 613 acceptability. We also ran a regression model including all 2- and 3-way interactions. This
 614 model similar in predictive power to the above model, $F(3,562)=67.32$, $p<.001$, $R^2=0.46$, and
 615 no interaction term was significant, so we report only the first model excluding interactions
 616 above [33].

617

618 **Fig 6. Firing Endorsement by Bystander's Guilt and Total War Beliefs, Scatterplot**

619 **With Applied Jitters Study 4.** Includes a loess line with a grey area indicating confidence
 620 region.

621

622 **Table 2. Regressing Ratings of How Much the Bystander Deserves To Die Depending on**
 623 **their Guilt, Nationality, and Participant's Support For Total War, Study 4**

Predictor	Estimate	SE	<i>t</i>	<i>p</i>
Intercept ^a	0.95790	0.31846	3.0080	0.0027
Guilty (0=Innocent, 1=Guilty):				
1 – 0	-0.27293	0.49038	-0.5566	0.5780
Nationality (0=Iraqi, 1=Swedish):				
1 – 0	-0.22693	0.46627	-0.4867	0.6267
Total War Beliefs	0.09340	0.09978	0.9360	0.3497
Total War Beliefs * Guilty:				
Total War Beliefs * (1 – 0)	0.78179	0.14878	5.2548	<.0001
Total War Beliefs * Nationality:				
Total War Beliefs * (1 – 0)	0.10334	0.14703	0.7028	0.4824
Guilty * Nationality:				
(1 – 0) * (1 – 0)	1.00641	0.68739	1.4641	0.1437
Total War Beliefs * Guilty * Nationality:				
Total War Beliefs * (1 – 0) * (1 – 0)	-0.27870	0.20845	-1.3370	0.1818

624 ^a Represents reference level; Bold indicates significance

625

626

627

628 **Fig 7. Firing acceptability ratings (Left) and ratings of how much the bystander**
 629 **deserves to die (Right) by total war beliefs and bystander's guilt, Study 4.** Scatterplot
 630 with applied jitter includes linear regression line with grey area indicating confidence region.

631

632

633 **Discussion**

634 This study clarified the role of bystander guilty, nationality, and total war beliefs on
 635 willingness to sacrifice bystanders to destroy enemies. We predicted support for total war
 636 would permit harming an innocent local Iraqi but not an innocent foreigner. Conversely, we
 637 expected that both local and foreign civilians who support the enemy would be perceived as
 638 equally available for harm.

639 However, results did not entirely support predictions. Consistent with prior studies,
640 bombing rates were higher for guilty bystanders associated with ISIS than innocent
641 bystanders reporting on ISIS. However, bombing rates were no different whether bystanders
642 were local or foreign. Participants treated local and foreign enemies similarly—but also
643 unexpectedly treated local and foreign innocent targets similarly, contrary to our predictions.
644 Instead, total war attitudes predicted overall increased willingness to bomb, and especially
645 high perceptions that guilty targets (foreign or local) deserved to die.

646 Hence, people who endorse total war appear to admit that some bystanders are
647 innocent, yet nonetheless accept bombing them—they appear to view sacrificing innocents
648 worthwhile to damage a known enemy. Intriguingly, it made no difference whether the
649 innocent bystander to be sacrificed was local (thus presumably part total war conflict) or
650 foreign (thus presumably not part). This pattern suggests that people who endorse total war
651 beliefs may view any innocent targets as justified sacrifices in pursuit of damaging enemies,
652 rather than only local civilians embroiled in total war.

653 It should be noted that participants were generally less willing to bomb both innocent
654 targets in this study compared to Studies 1 and 2. It may be that this change reflects the edits
655 to the scenario which clarified the conviction of the bystander’s anti-ISIS stance (e.g., ‘he
656 does not support ISIS and never has’), or the description of his position as a reporter rather
657 than a farmer, as this may have increased inferences that the innocent target is not merely
658 neutral but possibly actively working against ISIS in the region. Alternatively, this pattern
659 could reflect shifting public opinion on war in America, as this data was collected during the
660 first days of the Russia-Ukraine conflict. Additionally, this study collected a sample from
661 Prolific instead of mTurk, a platform with a different pool of users. Regardless, despite these
662 changes to the paradigm, a substantial proportion of participants continued to endorse
663 bombing regardless of the innocence of the bystander. It is an open question whether these

664 findings are specific to US participants, so in a fifth study we replicated our method with a
665 non-US sample.

666 **Study 5**

667 Study 5 replicated Study 1 using United Kingdom participants. Pre-registered:
668 https://aspredicted.org/W87_DB2.

669 **Method**

670 We recruited 302 adult participants from the UK through Prolific. We excluded five
671 for failing a competence check leaving a final sample of 297. The pwr package for R revealed
672 80% power to detect effect sizes of $w=0.21$. In addition to items used in Study 1, we also
673 included the total war attitudes questionnaire from Study 4, as well as the additional items for
674 participants assigned to the unidentified bystander condition: “The pilot followed the ISIS
675 operative to a farmhouse where a second person already was. Do you think this second
676 person is another member of ISIS?” (*Yes/No*) and “Please rate how likely it is the second
677 person is another ISIS member” on a scale from 1 (*Extremely unlikely*) to 100 (*Extremely*
678 *likely*). If the effects we observed in previous studies are specific to American participants,
679 we predict that using a sample from the UK will show differences.

680 **Results**

681 As in Studies 1a and 1b, the bystander’s identity significantly impacted bombing
682 endorsement, using a logistic generalized linear model, $\chi^2(6,297)=36.13$, $p<.001$, $w=0.35$.
683 However, bombing rates were generally lower than in the American sample. For example,
684 only 32.5% of UK participants endorsed bombing the unidentified bystander, compared to
685 58.5% of Americans in Study 1 (Fig 8). Furthermore, UK participants did not bomb the
686 unidentified bystander significantly more than other conditions. In addition to being asked
687 whether the pilot should fire, our main dependent measure, participants were also asked

688 “How acceptable is it for the pilot to fire upon the building, killing both people inside?” and
 689 answered on a Likert scale from 1 (*Not at all*) to 7 (*Very much*). A one-way between subjects
 690 analysis of variance showed significant differences between the vignettes across the report
 691 measures, $F(6,128.1)=9.37, p<.001$. However, the unidentified bystander condition did not
 692 differ significantly from any condition except the ISIS soldier condition (Table 3). As in
 693 previous studies, the pattern of acceptability matched the pattern of endorsing firing.

694

695 **Fig 8. Percent of “Yes” Responses to “Should The Pilot Fire” Depending on Bystander**
 696 **Identity, Study 5.** Error bars reflect 95% CIs. Asterisks denote Bonferroni corrected p-values
 697 compared with the unidentified bystander condition ($*p<.05, **p<.01, ***p<.001$). Unlike
 698 prior studies using American participants, bombing rates for the unidentified bystander
 699 condition did not differ significantly from any other condition.

700

701

702 However, UK participants in Study 5 replicated the pattern of American participants
 703 in Study 1 by rating the unidentified bystander as more deserving of death ($M=2.19,$
 704 $SD=1.34, N=43$) than bystanders in all other conditions except a known enemy “ISIS soldier”
 705 condition, ($M=1.34, SD=0.89, N=214$), which a Mann-Whitney U test showed was
 706 significant, $U(N_{unidentified}=43, N_{civilian}=214)=2720, p<.0001, MDiff=-1.00, 95\% CI[-1.00, -0.00],$
 707 $rank\ biserial\ correlation=0.41$ (Fig 9). Thus, although UK participants were more hesitant
 708 than Americans to endorse bombing targets in general, and unidentified targets in particular,
 709 they nonetheless demonstrated a similar pattern of suspicion toward an unknown bystander—
 710 even higher than towards a war profiteer (Table 3).

711

712 **Fig 9. Ratings of Bystander Deservingness of Death Depending on their Identity, Study**
 713 **5.** Error bars represent a 95% confidence interval. Asterisks denote Games-Howell corrected
 714 p-values comparing with the unidentified bystander condition (* $p < .05$, ** $p < .01$, *** $p < .001$).
 715 Participants answered the question “How much does the second person [the bystander]
 716 deserve to die?” with a Likert scale between 1 (*Not at all*) and 7 (*Very much*).

717

718

719 **Table 3. Regressing Ratings of How Much the Bystander Deserves to Die Depending On**
 720 **Bystander Identity and Total War Beliefs, Study 5.**

Predictor	Estimate	SE	t	p	Stand. Estimate
Intercept ^a	1.25585	0.466	2.6968	0.007	
Condition:					
Unidentified – Innocent	-0.77244	0.643	-1.2016	0.231	0.60297
War profiteer – Innocent	-0.69059	0.624	-1.1064	0.270	0.12265
Anti-ISIS Muslim – Innocent	-0.31175	0.640	-0.4872	0.626	-0.01591
Extremist Muslim – Innocent	-1.05885	0.687	-1.5416	0.124	0.17931
ISIS soldier– Innocent	1.66063	0.680	2.4429	0.015	1.93266
Moderate Muslim – Innocent	-0.06990	0.624	-0.1119	0.911	0.00976
Total war attitudes	-0.00744	0.195	-0.0382	0.970	-0.00502
Total war attitudes * Condition:					
Total war attitudes * (Unidentified – Innocent)	0.66642	0.252	2.6476	0.009	0.44986
Total war attitudes * (War profiteer – Innocent)	0.34641	0.250	1.3880	0.166	0.23384
Total war attitudes * (Anti- ISIS Muslim – Innocent)	0.11369	0.267	0.4262	0.670	0.07675
Total war attitudes * (Extremist Muslim – Innocent)	0.52589	0.261	2.0140	0.045	0.35500
Total war attitudes * (ISIS soldier– Innocent)	0.50111	0.256	1.9579	0.051	0.33827
Total war attitudes * (Moderate Muslim – Innocent)	0.03347	0.241	0.1387	0.890	0.02259
^a Represents reference level					
Model Fit Measures					
		Overall Model Test			
R	R²	F	df1	df2	p

0.721	0.520	23.5	13	282	<.001
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721

722 **Total war attitudes**

723 Participants reported total war attitudes as in Study 4. UK participants showed less
724 reliability with *Cronbach's* $\alpha=0.88$. A principle component analysis using oblim rotation
725 based on eigenvalues greater than 1 showed 2 components with eigenvalues of 4.91 and 1.30
726 and accounting for 49% and 13% of the variance respectively. Despite the lower reliability,
727 we again treat the scale as a single factor to compare with findings from study 4. We again
728 conducted a logistic regression on bombing decisions depending on total war attitudes and
729 bystander identity. Results showed significant main effects of total war attitudes,
730 $\chi^2(1,297)=88.09, p<.001, w=0.54$, and bystander identity, $\chi^2(1,297)=24.66, p<.001, w=0.29$
731 but no interaction, $\chi^2(1,297)=7.41, p=.29, w=0.16$. Consistent with Study 4, total war attitudes
732 predicted overall bombing endorsement, regardless of bystander identity.

733 This pattern was not as strong for scale ratings of bombing acceptability where a
734 linear regression which predicted bombing acceptability depending on total war attitudes and
735 bystander identity. Results showed significant main effects of total war attitudes,
736 $F(1,282)=17.02, p<.001$, but no main effect of bystander identity, $F(6,282)=1.70, p=.122$, nor
737 the interaction, $F(6,282)=1.41, p=.209$. It could be that total war attitudes drive decisions of
738 whether bombing is acceptable, but are distinct from actually endorsing whether or not to
739 bomb.

740 We also computed a linear regression on ratings of the bystander's deservingness to
741 die depending on total war attitudes and bystander identity. The overall model was
742 significant, $F(13,282)=23.47, p<.001, R^2=.52$ (Table 3). Similar to Study 4, there was a
743 significant interaction between total war attitudes and bystander identity, $F(6,282)=2.56,$
744 $p=.020$. Total war beliefs were particularly related to judgments of how much some
745 bystanders deserved to die but not others. Total war had little impact on judgments of the war

746 profiteer, anti-ISIS Muslim, moderate Muslim, and ISIS soldier all of which were not
 747 significantly different from the innocent bystander; conversely total war beliefs had a greater
 748 impact on judgments of the unidentified bystander and the extremist Muslim (Fig 10).

749 Hence, similar to Study 4, total war beliefs seemed to especially increase perceptions
 750 that ‘guilty’ bystanders (closely associated to the enemy) deserved to die, rather than
 751 ‘innocent’ bystanders with less association. A possible exception was the ISIS soldier, which
 752 was only marginally ($p=.051$) significantly different from the innocent bystander in how total
 753 war attitudes affected ratings of deservingness to die. Intriguingly, total war beliefs increased
 754 perceptions that the unidentified bystander deserved to die. This finding suggests that people
 755 higher in total war beliefs not only endorse the sacrifice of bystanders more frequently they
 756 may also tend to view unidentified bystanders as guiltier.

757

758 **Fig 10. Ratings of Bystander Deservingness of Death Depending on their Identity, and**
 759 **Total War Beliefs, Study 5.** Includes linear regression line with grey area indicating 95%
 760 confidence region.

761

762 **Assumptions about the unidentified bystander**

763 We asked participants who read about the unidentified bystander ($N=43$) whether they
 764 believed he was part of ISIS, *yes* or *no*, and their probability estimate. In this UK sample,
 765 37% said yes, not significantly different from 50% $p=.13$, and substantially lower than the
 766 American replication of Study 1 where 67% (significantly greater than 50%) said yes.

767 Likewise, the UK probability estimates were not different from 50% ($M=52.4$, $SD=21.6$),
 768 $t(42)=0.73$, $p=.471$, 95% CI[45.76, 59.03], $d=0.11$, unlike American estimates which were
 769 significantly greater. Participants higher in support for total war were more likely to say
 770 “yes” when asked if the unidentified bystander was an ISIS member, $\chi^2(296)= 7.9$, $p=.0049$,

771 $OR=2.55$, 95% CI[1.31, 5.63] and gave a higher probability the bystander was an ISIS
772 member, $F(1,41)=12.70$, $p=.0009$, $R^2=.24$.

773 **Discussion**

774 Replicating Study 1a with a UK sample produced different results from the US
775 sample. Bombing endorsement was lower overall for UK participants, and the proportionally
776 high rate of bombing of the unidentified bystander shown in Study 1 did not appear when
777 using UK participants. Total war attitudes were lower in this UK sample than in the US
778 sample which may explain the lower overall bombing rates. Additionally, UK participants did
779 not display the same tendency to assume the unidentified bystander was more likely to be an
780 enemy combatant than a civilian seen in Study 1. Assuming guilt when the bystander was
781 unidentified was positively correlated to total war attitudes: lower support for total war
782 among UK participants might be a factor underlying the disappearance of the high rate of
783 bombing when the bystander is unidentified. This study suggests the findings from Studies 1-
784 4 may be unique to US participants, possibly due to their relatively high support for total war
785 compared to UK participants.

786 **General Discussion**

787 Five experiments examined how the identity of a wartime bystander influences
788 willingness to sacrifice them to kill a dangerous enemy combatant. We discovered a
789 potentially deadly tendency: when the bystander's identity was unknown, people tended to
790 assume they were an enemy and therefore acceptable collateral damage. Crucially, ~50% of
791 people across studies of American participants endorsed sacrificing unidentified bystanders
792 despite no evidence they were enemies, a sacrificial rate higher than any identified target
793 except a known enemy combatant (though not always significantly higher than some morally
794 compromised targets). This effect emerged even when we explicitly provided probability

795 estimates that the bystander was a civilian, emerged both during the height of the conflict
796 (2016) and after (2020), and emerged for both real and fictional targets.

797 This bias toward sacrificing unknown bystanders appears to stem from assuming the
798 unidentified person is an enemy. Our data suggest this finding is not merely due to ingroup
799 bias—Americans supporting America’s side of a war—as it emerged even when Americans
800 judged a fictional war. Moreover, people endorsed sacrificing the bystander at rates of around
801 50% when the certainty they were an enemy was around 50%. Thus, consistent with other
802 work (Watkins & Laham, 2019), wartime contexts may increase sacrificial acceptance, in
803 part by allowing decision-makers to arrive at unflattering assumptions about unknown
804 targets. This pattern aligns broadly with research showing threatening contexts increase
805 perceptions of harmful agents and outgroup categorization [18, 19].

806 Participants who endorse principles of total war such as “In war, everything and
807 anything is fair game” were also more likely to endorse the bombing. We predicted
808 individual differences in total war attitudes may lead to endorsing bombing when bystanders
809 are civilians of an enemy nation rather than a friendly nation. However, this was not the case:
810 higher support for total war principles predicted endorsing bombing generally and appeared
811 insensitive to the bystander’s nationality. When considering an innocent civilian bystander,
812 participants who were high and those who were low on total war agreed that the bystander
813 did not deserve to die; yet those high in total war attitudes were more likely to support a
814 bombing that would kill that innocent civilian to also kill a dangerous ISIS member. This
815 suggests supporters of total war are more likely to support sacrificing civilians as collateral
816 damage, while still admitting that it is a sacrifice.

817 **Limitations**

818 A replication using participants from the UK showed differences from US
819 participants. Those from the UK were less likely to endorse bombings overall, were lower on

820 support for total war, and were less likely to assume the unidentified bystander was a
821 combatant compared to participants from the US. Although this may put limits on the
822 generalizability of these findings, it is possible that low support for total war in the UK
823 compared to the US is the driving factor for these differences as total war beliefs was
824 positively related to both endorsing bombing and believing the bystander is likely to be a
825 combatant. However, these relationships are correlational and future studies could benefit
826 from more direct comparisons between countries across these measures and future work
827 could generalize this paradigm to other nations and cultures [34].

828 Moreover, all studies recruited from the general population: trained military decision-
829 makers could be either more or less hawkish in these decisions—evidence suggests both
830 directions are possible. Research examining race-bias in police shootings finds trained police
831 are less biased than civilians [35]. On the other hand, we find some evidence to suggest that
832 those with ties to the military are *more* likely to endorse firing on the unidentified bystander
833 (see Supplemental Materials: Study 1a Military Experience). Additionally, all studies wholly
834 or in part referenced an ongoing real-world conflict, so it is possible shifting sentiment about
835 this conflict may change the results of future studies. However, data from Study 1a was
836 collected in 2016 during the conflict, and a 2020 replication (Study 1b) after the conflict
837 reported similar levels of firing on unidentified bystanders.

838 All the studies reported here measure participants' willingness to endorse bombing a
839 bystander they have no information about, and Study 1c shows the high rates in which
840 participants infer the unidentified bystander is an enemy. However, one could argue that our
841 unidentified bystander vignette *does* have circumstantial evidence of the bystander's
842 affiliation with the enemy. Participants may reasonably assume the bystander is likely to be
843 an enemy because of proximity to the enemy. Gestalt psychology [36] demonstrates that in
844 general people are likely to view individuals who are close in proximity to be socially close

845 as well. However, a study which replicated Study 1a while manipulating the bystander's
846 proximity to the enemy showed no impact on endorsing the bombing (Study S1,
847 supplemental materials). Proximity to the enemy is unlikely to fully explain our participants'
848 bombing endorsement since manipulating it directly had no impact on any dependent
849 measures.

850 Choosing when bombing bystanders in war is a choice informed by ideology and can
851 be influenced by an individual's political beliefs. Aside from our measure of support for total
852 war, we did not collect data on political ideology and how this may affect bombing decisions.
853 Future work could compare politics as a possible moderator to these effects.

854 **Conclusion**

855 People often assume unidentified bystanders in a warzone are combatants and
856 acceptable collateral damage. Rather than give bystanders the benefit of the doubt, people
857 tend to treat them as "guilty until proven innocent." These findings have implications for
858 military strategists who must decide whether to attack areas with enemy militants and
859 unidentified bystanders. Our results support a common tendency in people to assume the
860 bystanders are enemies, which can have deadly consequences if they turn out to be innocent
861 civilians. To minimize civilian deaths, future research should investigate how to reduce this
862 bias and get decision-makers to evaluate more carefully who their weapons are targeting.

863

864

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869

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871 **References**

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 964
 965 **Supporting information**

966 **Figure S1.** Rating of firing acceptability from 1 (Not at all) to 7 (Very much) in Study 1a
967 split by bystander condition. Error bars represent a 95% confidence interval. Asterisks denote
968 Dwass-Steel-Critchlow-Fligner pairwise comparison p-values compared with the unidentified
969 bystander condition (* $p < .05$, ** $p < .01$, *** $p < .001$).
970