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# Long-term outcomes of bilateral salpingo-oophorectomy in women with personal history of breast cancer

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# ABSTRACT

**Objectives** To investigate the association between bilateral salpingo-oophorectomy (BSO) and long-term health outcomes in women with a personal history of breast cancer.

Methods and analysis We used data on women diagnosed with invasive breast cancer between 1995 and 2019 from the National Cancer Registration Dataset (NCRD) in England. The data were linked to the Hospital Episode Statistics-Admitted Patient Care dataset to identify BSO delivery. Long-term health outcomes were selected from both datasets. Multivariable Cox regression was used to examine the associations, with BSO modelled as a time-dependent covariate. The associations were investigated separately by age at BSO.

Results We identified 568 883 women, 23 401 of whom had BSO after the breast cancer diagnosis. There was an increased risk of total cardiovascular diseases with an HR of 1.10 (95% Cl 1.04 to 1.16) in women who had BSO<55 years and 1.07 (95% Cl 1.01 to 1.13) for women who had BSO≥55 years. There was an increased risk of ischaemic heart diseases, but there was no association with cerebrovascular diseases. BSO at any age was associated with an increased risk of depression (HR 1.20, 95% Cl 1.12 to 1.28) and increased risk of second nonbreast cancer in older women (HR 1.21, 95%Cl 1.08 to 1.35). BSO in older women was associated with reduced risk of all-cause mortality (HR 0.92, 95% Cl 0.87 to 096), but not in women who had BSO<55 years.

**Conclusion** In women with a personal history of breast cancer, BSO before and after the age of 55 years is associated with an increased risk of long-term outcomes. BSO after 55 years is associated with reduced all-cause mortality. Family history or genetic predisposition may confound these associations.

#### INTRODUCTION

Women with a personal history of breast cancer are at increased risk of developing second cancers including ovarian cancer. The cumulative 20-year risk of developing ovarian cancer after a breast cancer diagnosis has been estimated to be 1.4% for women diagnosed with breast cancer before the age

#### WHAT IS ALREADY KNOWN ON THIS TOPIC

⇒ Prior research has shown that bilateral salpingooophorectomy (BSO) is associated with a reduction in ovarian cancer risk but an increase in the risk of adverse long-term health outcomes. Though, there is a gap in examining the long-term outcomes of BSO in women with a personal history of breast cancer.

#### WHAT THIS STUDY ADDS

⇒ Our findings indicate an increased risk of total cardiovascular diseases and depression in women who underwent BSO. However, there is a reduction in allcause mortality in women who had BSO after the age of 55 years.

# HOW THIS STUDY MIGHT AFFECT RESEARCH, PRACTICE OR POLICY

⇒ We hope that our study will aid in creating personalised counselling and enhancing decision-making for women with a personal history of breast cancer who opt for BSO.

of 50 and 1.9% for those diagnosed at the age of 50 years or older.<sup>2</sup> There is no effective modality for screening for ovarian cancer. The only recommended method for prevention in women at elevated risk for ovarian cancer is bilateral salpingo-oophorectomy (BSO). BSO is associated with more than a 90% reduction in the risk of ovarian cancer.<sup>34</sup> However, the benefit of ovarian cancer risk reduction should be balanced against the health sequelae caused by the premature loss of oestrogen. In women with no history of cancer BSO with hysterectomy for benign indications has been found to be associated with reduced risk of ovarian cancer and breast cancer and increased risk of cardiovascular diseases (CVD), depression, dementia, cancer and all-cause mortality.<sup>5</sup>



BSO is also performed to suppress ovarian function in women with oestrogen receptor-positive breast cancer. Ovarian function suppression (OFS) could be achieved by surgical removal of both ovaries or radiation-induced ablation both of which are permanent, or by using gonadotropin-releasing hormone agonists which results in temporary suppression of ovarian function. Oestrogen cessation might lower the risk of recurrence, contralateral breast cancer (CBC) and mortality. The present study focuses on other long-term outcomes of BSO after breast cancer diagnosis, including CVD, neuropsychiatric outcomes and second cancer occurrence. This has not been studied before at a population scale using electronic health records. Evidence is limited and focuses only on mortality outcomes<sup>7 8</sup> or on women who had the BSO prior to the diagnosis of breast cancer. 9 Also, evidence derived from the general population is often used to counsel high-risk women who are opting for BSO.

Using general population data to counsel women with a personal history of breast cancer presents challenges, as these women may exhibit different benefit-risk profile, for example, reproductive history, family history of cancer, weight and alcohol consumption. 10 These in turn may influence their baseline risk of developing longterm health sequelae. Tumour characteristics and treatment choices could confound the association between BSO and the long-term outcomes. 11 12 Moreover, in the general population, BSO is often performed at the time of hysterectomy which means that the evidence could be confounded by the indication for the hysterectomy or by the potential long-term outcomes of hysterectomy alone in the comparison groups used.<sup>5</sup> Hence, there is a need for guidance to be based on studies specifically conducted on these women. Women should be able to make decisions based on accurate knowledge of the risks and benefits of the BSO. This information is particularly relevant in women with a personal history of breast cancer who may have additional concerns about using hormonal replacement therapy.

This study aims to investigate the association between BSO and long-term health outcomes, compared with receiving standard treatment after a breast cancer diagnosis. This is the largest study to date to examine these associations, using population-scale linked data from the National Cancer Registration Dataset (NCRD) and Hospital Episode Statistics-Admitted Patient Care (HES-APC) in England.

#### **METHODS**

#### **Data sources**

National Cancer Registration Dataset

The NCRD is collected and managed by the National Cancer Registration and Analysis Service (NCRAS), a population-based cancer registry for England with national coverage since 1971. NCRAS collects data from multiple sources including multidisciplinary team meetings, pathology reports, molecular testing results,

treatment records and hospital activity records. The NCRD contains demographic data including the age at diagnosis, gender, deprivation index, tumour characteristics data including the cancer stage using the TNM (tumour, node, metastasis) staging system, grade, hormonal receptor status, tumour morphology, tumour size, number of nodes excised and treatment data on the receipt of radiotherapy, chemotherapy, immunotherapy and surgery. The Office for National Statistics provides NCRAS with data on the date and cause of death. Section 254 of the Health and Social Care Act 2012 allows NCRAS to collect individual-level data on patients with cancer without consent.

#### Hospital Episode Statistics-Admitted Patient Care

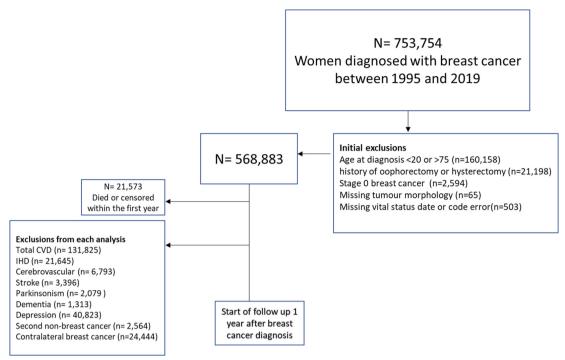
HES-APC collects data on all NHS hospital admissions in England. HES-APC also collects data on admissions to independent providers funded by the NHS.14 The NHS funds more than 98% of the hospital activity in England. 14 HES-APC includes all hospital care episodes since the financial year 1989/1990. Data fields include diagnoses, procedures, patient demographics and admission and discharge dates. Diagnoses are coded using the International Classification of Diseases version 10 (ICD-10). Each admission could have up to 20 diagnoses. We considered any of the first three diagnosis fields as the primary diagnosis/es and the other diagnosis fields as comorbidities. Procedures are coded using OPCS4 codes (Office of Population, Census and Surveys Classification of Interventions and Procedures, fourth Revision). Pseudonymised patient identifiers allow linkage of HES APC to NCRD.

#### Inclusion and exclusion criteria

We included women diagnosed with invasive breast cancer between 1995 and 2019, diagnosed between the ages 20 and 75 and who had no history of previous cancer diagnosis. We excluded women with a history of hysterectomy or oophorectomy before the date of breast cancer diagnosis. The number of women excluded for different reasons is summarised in figure 1. For each outcome investigated, women diagnosed with the specific outcome (in any of the 20 HES diagnosis fields) before or within the first year of the breast cancer diagnosis were excluded from the analysis. For the association with CBC, women with bilateral tumours or unknown laterality of the first tumour were excluded from the analysis.

#### **Identifying BSO procedures**

BSO procedures were identified using the OPCS4 Q221 and Q223 codes in HES-APC. Hysterectomy procedures were identified using the codes such as Q073, Q074, Q075, Q078, Q079, Q081, Q082, Q083, Q088 and Q089. BSO was considered to be for a malignant indication if the patient was diagnosed with a gynaecological malignancy 1 year before or 1 year after the procedure date.



**Figure 1** Exclusion process among women diagnosed with breast cancer between 1995 and 2019. Patients diagnosed with a particular outcome before the start of the follow-up were excluded from the analysis of that outcome. CVD: Cardiovascular diseases; IHD: Ischaemic heart diseases.

#### **Identifying long-term outcomes**

Outcomes of interest were all-cause mortality, total CVD, ischaemic heart diseases (IHD), cerebrovascular diseases, dementia, depression and parkinsonism (extrapyramidal movement disorders), breast cancer mortality, non-breast cancer mortality, CBC and second non-breast cancer (excluding non-melanoma skin cancer). Cancer outcomes and mortality data were identified from the NCRD. Breast cancer-specific mortality was defined as death where breast cancer was listed as a cause on part I of the death certificate. Non-cancer outcomes were identified from HES admissions using the first three diagnosis fields (primary diagnosis) or from part Ia of the death certificate. The ICD-10 codes used are summarised in online supplemental table S1. The associations with the severity of CVD outcomes (fatal/non-fatal) were assessed separately. Fatal outcomes were outcomes identified either from the cause of death or from HES admissions followed by fatality within the first 48 hours after the admission. Non-fatal were outcomes identified from HES admissions which were not followed by death within the first 48 hours.

# Follow-up and censoring

Follow-up started 1 year after the date of breast cancer diagnosis to avoid misclassifying bilateral breast cancers as second primary cancers and to allow time for exclusion of cases with any comorbidities identified after the breast cancer diagnosis. Follow-up ended at the first of the following: the outcome of interest, censoring or end of data collection. Data collection ended on 30 December 2020 for cancer outcomes and 1 May 2022 for mortality

and non-cancer outcomes. Follow-up time was censored at the time of unilateral oophorectomy, pelvic clearance or bilateral oophorectomy for malignant indications (except for ovarian cancer and second non-breast cancers analyses). In addition, for the CBC analysis women were censored 1 year after a breast surgery on the opposite side or a breast surgery with unknown laterality occurring more than 1 year after the breast cancer diagnosis. In each separate analysis for an outcome of interest, women who developed any of the other outcomes studied here were still followed for the outcomes of interest. Follow-up for BSO started 1 year after the surgery date, to allow time for biological plausibility (allow time for the cessation of oestrogen to cause a pathological effect) and to minimise detection bias. We performed separate analyses by the age at surgery, BSO<55 years in younger women and BSO≥55 years in older women. The age 55 was chosen as a proxy for the age at menopause. According to the British Menopause Society, it is estimated that more than 80% of women will be menopausal by the age of 54.15 In the younger women analysis, we only included women diagnosed with breast cancer before the age of 55. Women who were diagnosed before age 55 and had BSO after age of 55 years were included in the non-BSO group and censored at the date of their surgery, the occurrence of events, whichever occurred first.

#### Multiple imputation of missing data

We used multivariate imputations by chained equations to impute the missing TNM stage, grade, oestrogen receptor (ER) status, human epidermal growth factor receptor 2 (HER2) status, tumour size, number of lymph nodes excised, ethnicity and Charlson Comorbidity Index (online supplemental materials).

#### Statistical analysis

We used Cox regression to calculate HRs for the association between BSO and the long-term outcomes, with BSO modelled as a time-dependent covariate. The models were adjusted for age at breast cancer diagnosis, year of diagnosis, tumour size, number of excised lymph nodes, M-stage, grade, ER status, HER2 status, ethnicity (white, Asian, black, mixed and other), deprivation index (1-least deprived, 2, 3, 4 and 5-most deprived) and Charlson Comorbidity Index. The Charlson Comorbidity Index is a weighted scale that predicts the risk of mortality within 1 year of hospitalisation. 16 The index was derived from HES-APC records for all the patients 6 years prior to the breast cancer diagnosis and calculated using the method described by Quan et al. 17 The association with second cancer was further adjusted for hysterectomy. A regression model was fitted using each of the imputed datasets and the log (HRs) were combined using Rubin's rule. 18 We fitted the models for the CBC and second non-breast cancer analyses with an interaction term between BSO and M-stage (0/1). This was done to provide estimates for the association in women with M-stage 0. The Coxregression proportional hazard assumption was assessed by plotting the Schoenfeld residuals (supplementary materials 2: online supplemental figure 2S-17S). All analyses were carried out using R (V.4.1.0; R Foundation for Statistical Computing, Vienna, Austria).

#### **Sensitivity analyses**

We assessed whether the findings were influenced after adjustment for hysterectomy. We also examined the associations between the long-term outcome and three types of procedures (hysterectomy alone, hysterectomy and BSO, BSO alone).

We conducted sensitivity analyses for the associations between the long-term outcomes and BSO in older women by restricting the analyses to women diagnosed with breast cancer at age 55 years or older.

To examine whether the association with breast cancer mortality is explained by CBC diagnosis, we conducted a sensitivity analysis where CBC diagnosis served as a censoring event. Given that the initial cohort analysis for assessing the association with breast cancer mortality included women with bilateral tumours or unknown literalities of the first tumour, we repeated the analysis after excluding women with unknown first tumour literalities or bilateral tumours. This was done while retaining the initial censoring methodology, and subsequently employing CBC as a censoring event to allow comparability of the results.

# Patients and public involvement

The CanGene-CanVare programme includes a Patient Reference Panel (PRP) made up of patients, carers and members of the public who are, or have been, affected by cancer. The PRP is involved in programme governance and oversight, contributing personal perspectives and experiences as well as in the communication and dissemination of programme output.

#### **RESULTS**

We identified 568883women, 23401 of whom had BSO after the breast cancer diagnosis, 8243 after the age of 55 years and 15158 before the age of 55 years. Baseline demographic, tumour and clinical characteristics of women who had BSO (<55 and ≥55 years) alongside their corresponding reference groups are shown in table 1. The median age at diagnosis for women who had BSO<55 years was 43 years (IQR 38-47 years) and for women in the reference group was 48 years (IQR 43-51 years). The median age at diagnosis in women who had BSO>55 years was 57 years (IQR:52-63 years) and for women in the corresponding reference group 58 years (IQR 50–66). The percentage of white women in the BSO groups was higher than in the non-BSO groups, in the '≥55 years' BSO group 94% were white versus 86% in the reference group and in the '<55 years' BSO group 92% were white versus 84% in their reference group. Among the women who had BSO≥55 years 65% had hysterectomy compared with 2% in the reference group and 43% in those who had BSO<55 years compared with 3% in the reference group.

Summaries of the numbers at risk, person years and number of events in the cohort and the HRs for the associations between BSO and the long-term outcomes are shown in table 2.

#### **Associations with cardiovascular outcomes**

There was an increased risk of total CVD with an HR of 1.10 (95% CI 1.04 to 1.16) in women who had BSO<55 years and an HR of 1.07 (95% CI 1.01 to 1.13) for women who had BSO $\geq$ 55 years. There was an increased risk of IHD for both younger and older women with HRs of 1.20 (95% CI 1.06 to 1.37) and 1.15 (95% CI 1.05 to 1.27), respectively. Further exploration of associations with IHD subtypes, BSO was significantly associated only with angina (unstratified HR 1.27, 95% 1.13 to 1.42), but not with myocardial infarction (unstratified HR 1.03, 95% CI 0.89 to 1.18) or chronic IHD (unstratified HR 1.08, 95% CI 0.99 to 1.19).

There was no association between BSO and cerebrovascular diseases, haemorrhagic stroke or ischaemic stroke with unstratified HR estimates 1.00 (95% CI 0.91 to 1.11), 0.97 (95% CI 0.80 to 1.19) and 1.04 (95% CI 0.90 to 1.19), respectively. Investigating the associations by the severity of the CVD outcomes yielded significant associations only with the non-fatal total CVD and non-fatal IHD outcomes.

#### Association with neuropsychiatric outcomes

BSO was associated with an increased risk of depression both in women who had BSO<55y (HR 1.18, 95% CI 1.09 to 1.28) and ≥55 years (HR 1.18, 95% CI 1.05 to 1.33). BSO was not associated with parkinsonism in women who

**Table 1** Baseline characteristics of women who had BSO above or after the age of 55 with their respective reference groups used in the association analyses

Diagnosis year  Median (IQR)  Age at diagnosis  Median (IQR)  Age at BSO  Median (IQR)  Follow-up years  Median (IQR)	2010 (2004–2014) 43 (38–47) 47 (43–50) 11 (7–17)	2009 (2002–2015) 48 (43–51) 57 (56–61) ‡	2006 (2000–2012) 57 (52–63)	2009 (2003–2015) 58 (50–66)
Median (IQR) Age at diagnosis Median (IQR) Age at BSO Median (IQR) Follow-up years	43 (38–47) 47 (43–50)	48 (43–51)	· · ·	,
Median (IQR)  Age at BSO  Median (IQR)  Follow-up years	47 (43–50)		57 (52–63)	58 (50–66)
Age at BSO  Median (IQR) Follow-up years	47 (43–50)		57 (52–63)	58 (50–66)
Median (IQR) Follow-up years	· ·	57 (56–61) ‡		
Follow-up years	· ·	57 (56–61) ‡		
	11 (7–17)		63 (58–70)	
Median (IQR)	11 (7–17)			
		10 (6–17)	14 (9–20)	9 (5–15)
Deprivation index				
1—least deprived	3491 (23%)	52702 (24%)	2009 (24%)	128 066 (23%)
2	3515 (23%)	50 084 (23%)	2120 (26%)	125 329 (23%)
3	3190 (21%)	45 113 (20%)	1736 (21%)	113116 (21%)
4	2700 (18%)	39413 (18%)	1395 (17%)	96451 (18%)
5-most deprived	2262 (15%)	34580 (16%)	983 (12%)	82 520 (15%)
- Ethnicity			•	
White	13988 (92%)	186939 (84%)	7721 (94%)	471 331 (86%)
Asian	375 (2.5%)	8184 (3.7%)	125 (1.5%)	14913 (2.7%)
Black	199 (1.3%)	4566 (2.1%)	48 (0.6%)	7300 (1.3%)
Mixed	100 (0.7%)	1416 (0.6%)	25 (0.3%)	2240 (0.4%)
Other	163 (1.1%)	3065 (1.4%)	41 (0.5%)	5610 (1.0%)
Missing	333 (2.2%)	17722 (8.0%)	283 (3.4%)	44 088 (8.1%)
NM stage	,	, ,	, ,	,
1	4259 (28%)	62373 (28%)	3063 (37%)	180 007 (33%)
2	5242 (35%)	69784 (31%)	2070 (25%)	154 457 (28%)
3	1331 (8.8%)	16243 (7.3%)	329 (4.0%)	33840 (6.2%)
4	315 (2.1%)	6104 (2.8%)	90 (1.1%)	17064 (3.1%)
Missing	4011 (26%)	67388 (30%)	2691 (33%)	160114 (29%)
Grade	,	,	,	,
1	1649 (11%)	31 674 (14%)	1633 (20%)	86678 (16%)
2	6673 (44%)	92744 (42%)	3787 (46%)	245 097 (45%)
3	5983 (39%)	79 652 (36%)	2217 (27%)	166 880 (31%)
Missing	853 (5.6%)	17 822 (8.0%)	606 (7.4%)	46827 (8.6%)
ER status	()	()		(2.2.2)
Positive	6527 (43%)	79 593 (36%)	2687 (33%)	206476 (38%)
Negative	984 (6.5%)	18 067 (8.1%)	428 (5.2%)	39371 (7.2%)
Missing	7647 (50%)	124232 (56%)	5128 (62%)	299 635 (55%)
HER2 status	(	(==,-,	(- ,-)	(,-)
Positive	1109 (7.3%)	17333 (7.8%)	287 (3.5%)	36024 (6.6%)
Negative	6251 (41%)	80 291 (36%)	2470 (30%)	206536 (38%)
Missing	7798 (51%)	124268 (56%)	5486 (67%)	302 922 (56%)
Morphology	( , - )	(/-)	(,-)	(3070)
Invasive ductal carcinoma	12191 (80%)	174050 (78%)	6075 (74%)	410727 (75%)
Invasive lobular carcinoma	1457 (9.6%)	20 964 (9.4%)	1051 (13%)	62 170 (11%)
Other	1510 (10%)	26878 (12%)	822 (14%)	56611 (13%)

Continued

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Table 1	Continu	ച

Table 1 Continued							
Bilateral salpingo-oophorectomy	Yes Age at BSO<55 n=15158	No Reference* n= <b>221 892</b>	Yes Age at BSO≥55 n=8243	No Reference† n= <b>545482</b>			
No	12576 (83%)	165405 (75%)	6205 (75%)	357 566 (66%)			
Yes	2582 (17%)	56 487 (25%)	2038 (25%)	187916 (34%)			
Hormonal treatment							
No	9761 (64%)	149666 (67%)	4890 (59%)	339772 (62%)			
Yes	5397 (36%)	72 226 (33%)	3353 (41%)	205710 (38%)			
Radiotherapy treatment							
No	5517 (36%)	81 066 (37%)	3162 (38%)	207378 (38%)			
Yes	9641 (64%)	140826 (63%)	5081 (62%)	338 104 (62%)			
Chemotherapy							
No	6061 (40%)	106622 (48%)	5697 (69%)	346 938 (64%)			
Yes	9097 (60%)	115270 (52%)	2546 (31%)	198544 (36%)			
Hysterectomy							
Yes	6478 (43%)	6439 (2.9%)	5321 (65%)	11419 (2.1%)			
No	8680 (57%)	215 453 (97%)	2922 (35%)	534 063 (98%)			

<sup>\*</sup>Women who were diagnosed with invasive breast cancer before the age of 55 and did not have a BSO before the age of 55.

had BSO<55 years (HR 0.98, 95% CI 0.67 to 1.44), but was associated with a reduced risk of parkinsonism in those who had BSO $\geq$ 55 years (HR:0.67, 95% CI 0.48 to 0.94).

#### **Association with cancer outcomes**

BSO was associated with increased risk of second primary non-breast cancers in women having BSO≥55 years (HR:1.21, 95% CI 1.08 to 1.35), but there was no association in women who underwent BSO before age 55 years (HR 1.03, 95% CI 0.91 to 1.17). BSO at any age was not associated with CBC (HR 1.04, 95% CI 0.95 to 1.15).

# **Association with mortality outcomes**

Having BSO at or after the age of 55 years was associated with reduced risk of all-cause mortality with a HR of 0.92 (95% CI 0.87 to 0.96), but not in younger women (HR 1.03, 95% CI 0.98 to 1.08). BSO<55 years was associated with an increased risk of breast cancer mortality 1.09 (1.04–1.15), but no association was observed in women who had BSO after the age of 55 years (HR 0.96, 95% CI 0.89 to 1.05). Finally, BSO was associated with a reduction in the risk of non-breast cancer mortality for both women who had BSO<55 years (HR 0.86, 95% CI 0.78 to 0.96) and BSO $\geq$ 55 years (HR 0.93, 95% CI 0.88 to 0.99).

#### Sensitivity analyses

Sensitivity analysis for the associations in women who underwent the BSO≥55 years and restricted on women diagnosed with breast cancer at or after age of 55 years yielded similar results (supplementary materials 2: online supplemental table 6S).

Sensitivity analysis revealed that adjusting for hyster-ectomy primarily influenced the association between BSO performed before age of 55 years and breast cancer and all-cause mortalities. Among the 'BSO<55 years' analysis cohort hysterectomy alone and hysterectomy and BSO were associated with a reduction in the risks of all-cause mortality and breast cancer mortality. While BSO alone was associated with increased risk of breast cancer mortality (HR 1.26, 95% CI 1.19 to 1.34), all-cause mortality (HR 1.18, 95% CI 1.12 to 1.25) and reduced risk of non-breast cancer mortality (HR 0.82, 95% CI 0.71 to 0.96). More details in supplementary materials 2 (online supplemental figure 1S, table 4S, 5S)

When censoring at the date of CBC, BSO alone was not associated with breast cancer mortality with HR estimates of 1.05 (95% CI 0.97 to 1.14) and 0.90 (95% CI 0.73 to 1.11), for women who had BSO<55 years and  $\geq$ 55 years, respectively.

### **DISCUSSION**

We used population-scale electronic health records to assess the association between BSO after breast cancer diagnosis and long-term health outcomes. This is the first time that HES-APC and the NCRD datasets were linked to answer this question in a cohort of patients with breast cancer diagnosed over a 24-year period.

BSO after breast cancer diagnosis was associated with an increased risk of CVD, IHD and depression in women who had BSO at any age and an increased risk of second

<sup>†</sup>Women who were diagnosed with invasive breast cancer at any age and did not have a BSO.

<sup>‡</sup>Reference includes women who were diagnosed before the age of 55 and had a BSO after the age of 55, these women were censored at the date of their BSO≥55, outcome development or occurrence of a censoring event whichever occurred first.

BSO, bilateral salpingo-oophorectomy; ER, oestrogen receptor; HER2, human epidermal growth factor receptor 2; TNM stage, tumour, node, metastasis stage.

O		A + DOO	NI at otal	N	Daws	LID (OFO) ODA
Outcome		Age at BSO	N at risk	N events	Person years	HR (95% CI)*
All-cause mortality		Unstratified	547246	158 447	5108170	1.01 (0.98 to 1.05
	<55	231 281	50 053	2304019	1.03 (0.98 to 1.08	
		≥55	532118	156 083	4947996	0.92 (0.87 to 0.96
Total CVD	Total	Unstratified	415485	104491	3570242	1.09 (1.05 to 1.13
	Non-fatal			100279	3570245	1.10 (1.05 to 1.14
	Fatal			11546	4157063	0.93 (0.82 to 1.06
	Total	<55	206942	33 189	1935037	1.10 (1.04 to 1.16
	Non-fatal			32070	1 935 038	1.10 (1.05 to 1.16
	Fatal			2153	2119670	0.88 (0.70 to 1.11
	Total	≥55	401756	102312	3 434 178	1.07 (1.01 to 1.13
	Non-fatal			98144	3 434 181	1.08 (1.02 to 1.15
	Fatal			11 461	4008443	0.88 (0.75 to 1.03
HD	Total	Unstratified	525 665	28178	4805042	1.15 (1.06 to 1.24
	Non-fatal			25710	4805043	1.17 (1.08 to 1.27
	Fatal			3549	4951005	0.81 (0.61 to 1.06
	Total	<55	229018	5871	2250590	1.20 (1.06 to 1.37
	Non-fatal		220010	5556	2250590	1.21 (1.07 to 1.38
	Fatal			428	2285240	0.93 (0.53 to 1.63
	Total	≥55	510653	27873	4647631	1.15 (1.05 to 1.27
		_ ≥55	510055			
	Non-fatal			25418	4647632	1.20 (1.08 to 1.32
2 1	Fatal			3534	4791818	0.78 (0.56 to 1.06
Cerebrovascular diseases		Unstratified	540517	20141	5 002 065	1.00 (0.91 to 1.11
aiscascs	Non-fatal			18203	5 002 068	1.01 (0.91 to 1.12
	Fatal			4106	5 0 6 8 1 7 6	0.93 (0.73 to 1.18
	Total	<55	230308	3512	2282799	0.92 (0.77 to 1.10
	Non-fatal			3207	2282800	0.94 (0.78 to 1.12
	Fatal			518	2296864	0.94 (0.59 to 1.48
	Total	≥55	525 428	19980	4842903	0.98 (0.87 to 1.10
	Non-fatal			18052	4842906	1.00 (0.88 to 1.13
Fata	Fatal			4085	4908357	0.83 (0.62 to 1.10
Angina		Unstratified	525 664	11760	4872452	1.27 (1.13 to 1.42
		<55	229017	2740	2265645	1.28 (1.06 to 1.54
		≥55	510652	11609	4714269	1.34 (1.16 to 1.54
Myocardial infard	ction	Unstratified	525 664	9074	4917312	1.03 (0.89 to 1.18
		<55	229017	1747	2277193	0.84 (0.65 to 1.09
		≥55	510652	9004	4758434	1.09 (0.92 to 1.29
Chronic IHD		Unstratified	525 664	19686	4849348	1.08 (0.99 to 1.19
		<55	229017	4033	2262022	1.14 (0.98 to 1.34
		≥55	510652	19488	4691219	1.12 (1.00 to 1.25
Haemorrhagic st	roke	Unstratified	543914	4584	5076992	0.97 (0.80 to 1.19
, J. 101		<55	230728	1101	2296233	0.96 (0.71 to 1.29
		≥55	528808	4530	4917201	0.90 (0.69 to 1.18
Ischaemic stroke		Unstratified	543914	9971	5 055 430	1.04 (0.90 to 1.19
		J. ISTI GETTICG	0.0011	507 1		(0.00 to 1.10
ISCHAEITHC SHOKE		<55	230728	1513	2294263	0.92 (0.70 to 1.20

Continued

Table 2 Continued						
Outcome	Age at BSO	N at risk	N events	Person years	HR (95% CI)*	
Parkinsonism	Unstratified	545231	3533	5 082 902	0.78 (0.61 to 1.01)	
	<55	230926	648	2298880	0.98 (0.67 to 1.44)	
	≥55	530 135	3493	4923142	0.67 (0.48 to 0.94)	
Dementia	Unstratified	545 933	12036	5072461	0.94 (0.81 to 1.08)	
	<55	231212	625	2301234	0.90 (0.55 to 1.47)	
	≥55	530804	12016	4912385	0.95 (0.82 to 1.11)	
Depression	Unstratified	506 423	21 243	4784915	1.20 (1.12 to 1.28)	
	<55	213588	10491	2150730	1.18 (1.09 to 1.28)	
	≥55	492741	20103	4640259	1.18 (1.05 to 1.33)	
Second non-breast cancer†	Unstratified	544682	39886	4494403	1.05 (0.97 to 1.14)	
	<55	230904	11 049	2044405	1.03 (0.91 to 1.17)	
	≥55	529 581	39362	4353510	1.21 (1.08 to 1.35)	
Contralateral breast cancer	Unstratified	522802	13948	4071952	1.04 (0.95 to 1.15)	
	<55	221 647	6819	1799643	1.01 (0.89 to 1.14)	
	≥55	508 460	13179	3967272	1.13 (0.97 to 1.33)	
Breast cancer mortality	Unstratified	547246	80795	5107400	1.11 (1.06 to 1.16)	
	<55	231 281	36745	2303692	1.09 (1.04 to 1.15)	
	≥55	532 118	78865	4947241	0.96 (0.89 to 1.05)	
Non-breast cancer mortality	Unstratified	547246	77629	5107400	0.91 (0.86 to 0.96)	
	<55	231 281	13300	2303692	0.86 (0.78 to 0.96)	
	≥55	532118	77 196	4947241	0.93 (0.88 to 0.99)	

\*Models adjusted for age at breast cancer diagnosis, year of diagnosis, tumour size, number of excised lymph nodes, M-stage, grade, ER status, HER2 status, ethnicity, deprivation index and Charlson Comorbidity Index. †Additionally adjusted for hysterectomy.

BSO, bilateral salpingo-oophorectomy; CVD, cardiovascular diseases; ER, oestrogen receptor; HER2, human epidermal growth factor receptor 2; IHD, ischaemic heart diseases; N, number.

non-breast cancer among women who had BSO at or after age 55 years. Investigating the CVD associations by the severity of the outcome showed no association with fatal CVD outcomes. The lack of significant association between BSO and fatal CVD outcomes, coupled with the observation that the increased risk of CVD and IHD seems to be primarily driven by an elevated risk of angina, suggests the possibility of detection bias influencing these associations.

BSO before the age of 55 years was associated with increased risk of breast cancer mortality, but reduced risk of non-breast cancer mortality. Although BSO after age of 55 years was not associated with breast cancer mortality, it was associated with a lower risk of death from other causes (non-breast cancer mortality) and lower all-cause mortality. A sensitivity analysis was conducted where CBC was considered a censoring event to assess the robustness of the finding that BSO in younger women was associated with increased breast cancer mortality. This analysis found no association between BSO in younger women and breast cancer mortality. It is therefore possible that the group of women who opted for BSO at a young age may be enriched for women who are genetically susceptible to

breast cancer and who are at an increased risk for CBC. <sup>19</sup> We explored this by stratifying the analysis on the indication for BSO (prophylactic/other benign indication), the indications were derived from the ICD-10 diagnosis codes recorded in HES at the date of having the BSO. The estimated HRs by indication were similar (supplementary materials 2; online supplemental table 7S).

BSO is indicated for various reasons including treatment of ovarian cancer, benign ovarian conditions affecting both ovaries, for example, endometriosis, ovarian cancer risk reduction in high-risk women (eg, *BRCA1* and *BRCA2* PV carriers)<sup>20</sup> or for OFS in patients with breast cancer with ER-positive tumours.<sup>6</sup> In this cohort, we censored women who had BSO for treatment of any gynaecological cancer or as a part of pelvic clearance procedure. Thus, the women in our cohort probably had the BSO for the other listed indications. It is also possible that the cohort included women who are *BRCA1* or *BRCA2* pathogenic variants (PVs) carriers given that the prevalence of both PVs among patients with breast cancer is around 2–3%.<sup>21</sup> <sup>22</sup>

Several studies examined the association between OFS and all-cause mortality, disease-free survival (DFS) and

other outcomes. OFS is recommended to be combined with tamoxifen or aromatase inhibitors in women with hormone-positive breast cancer. However, OFS could be achieved temporarily by GnRH agonist or permanently using radiation-induced ablation or BSO. Only BSO is recommended for ovarian cancer risk reduction. In a recent Cochrane systematic review and meta-analysis, <sup>6</sup> medical OFS was associated with a significant reduction in the risk of all-cause mortality and DFS with HRs 0.80 (95% CI 0.71 to 0.89) and 0.81 (95% CI 0.75 to 0.88), respectively. While OFS by BSO was not associated with all-cause mortality (HR 0.86, 95% CI 0.57 to 1.28) or DFS (HR 0.96, 95% CI 0.70 to 1.30). It is important to note this systematic review answers a different question from our study which compares women who had BSO to women who might have had OFS through GnRH agonist or radiation-induced ablation.

Obermair *et al* examined the association between BSO after breast cancer diagnosis and mortality in two independent studies using the Queensland cancer registry (n=25536)<sup>8</sup> and the Western Australia cancer registry (n=15395).<sup>7</sup> In both studies, hysterectomy and BSO were associated with reduced risk of all-cause mortality and breast cancer mortality. However, BSO alone was not associated with all-cause mortality or breast cancer mortality. These findings are in line with our sensitivity analysis in which both hysterectomy alone and hysterectomy and BSO<55 years were associated with reduced risk of mortality, while BSO alone<55 years was associated with an increased risk of mortality.

The findings for the associations with CVD and all-cause mortality in the age-stratified analyses are in line with the associations in the general population. In a recent systematic review and meta-analysis on the long-term outcomes of BSO at the time of hysterectomy,<sup>5</sup> hysterectomy and BSO were associated with increased risk of IHD in women who had the procedure before or after the age of 50 years and increased risk of all-cause mortality in women who had BSO before the age of 50 years. However, there was no association with all-cause mortality in women who had BSO after the age of 50 years.

These findings are important for counselling women with a personal history of breast cancer who are considering BSO. Guidance based on studies specifically conducted on this group would provide more informative support for decision-making than evidence derived from the general population.

The major strength of this study is the large sample size, up to date this is the largest study to examine the long-term outcomes of BSO after the diagnosis of breast cancer. The NCRD contains data on all patients with cancer diagnosed in England which minimises selection bias. We modelled the BSO and hysterectomy as time-dependent covariates to avoid immortal time bias. Also, we excluded women who had a history of any of the outcomes before the start of the follow-up.

Assessment of the association between BSO and long-term outcomes in electronic health records has

several limitations. Unlike randomised controlled studies several factors (possibly confounders) influence the receipt or uptake of intervention in observational studies. To address this, we adjusted the analysis for a number of confounders including hormonal receptor status, treatment, age at diagnosis, deprivation index and Charlson Comorbidity Index. In addition, we censored women who had BSO for malignant indications to minimise confounding by indication of the BSO. However, there could be residual confounding resulting from the lack of information on confounders like family history of breast cancer, smoking status and body mass index. Also, detection bias could have influenced the results as women with a previous history of surgery might seek more medical attention. Nevertheless, we expect that this has not substantially biased the results as the cohort consists of patients with breast cancer who are likely to receive close medical monitoring. We attempted to minimise detection bias by starting the follow-up for BSO or hysterectomy 1 year following the surgery to avoid adding cases to the surgery groups who were accidentally discovered at the time of the surgery (prevalent cases). Another limitation is that non-cancer long-term outcomes were identified from hospital admissions in HES and from death certificates, which means we were only capable of identifying severe outcomes and this might have limited the power of the study to detect associations with certain outcomes such as dementia. This limitation may also explain the observed reduction in the risk of parkinsonism, which is not in line with previous findings. 23 24 Additionally, the association with breast cancer mortality could be affected by the inaccurate coding of the cause of death. Finally, some tumour characteristic data were missing, for example, 29% of TNM stage and more than 50% of the ER status and HER2 status. To address this, we employed multiple imputation which has been shown to reduce bias in estimates compared with complete case analysis.<sup>25</sup>

#### CONCLUSION

In women with a personal history of breast cancer, BSO before and after the age of 55 years is associated with an increased risk of long-term outcomes including CVD, cancer and depression. However, more work is needed to elucidate the possibility of confounding by family history and genetic susceptibility. Women with BSO above age of 55 years may benefit from a reduction in all-cause mortality.

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**Contributors** HH wrote the manuscript and performed the analyses. IA, LL and MT selected the censoring surgeries for the contralateral breast cancer analysis. TR, AB and CK collected and verified the data. ACA, PP, MT and CT planned the study and raised funding. MT, PP and ACA supervised the study and edited the manuscript. All authors provided feedback to inform the methodology. HH acted as a guarantor.

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