

# Low-stakes accountability and public service turnarounds

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## ABSTRACT

Can accountability to external bodies induce performance turnarounds in struggling public services? And if so, must account-holders use incentives and sanctions to change organizations that have yet to self-correct, or can a gentler, more informational regime suffice? Using recent research into “low-stakes” accountability, we argue that feedback and standard-setting on their own may stimulate subunit turnarounds in complex, multi-service organizations by directing leaders’ scarce attention, control efforts and resource oversight toward previously unnoticed performance deficits. However, given the potential for “tunnel vision” among account-givers, accountability-induced turnarounds may be confined to dimensions of performance most relevant to

the account-holder, to the neglect of others. We test each stage of this theory using quasi-experimental methods on 133 matched local government units and data from the Local Government Ombudsman in England. We show that councils that are notified of maladministration in their social care provision significantly increase leadership attention and control efforts toward this service, and invest more in its core staffing, compared with a matched sample of unnotified councils. On average, the rate of maladministration falls by more than half; though, as predicted, we detect no wider performance gains. The immediacy of the improvement further suggests that responsible subunits may act *in anticipation* of leaders' increased attention.

## KEYWORDS

Accountability; turnaround; difference-in-differences; ombudsmen; tunnel vision

## ACKNOWLEDGEMENTS

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## DATA AVAILABILITY

Data and code can be downloaded from the Harvard Dataverse here:

<https://doi.org/10.7910/DVN/VGIL4G>

## 1. INTRODUCTION

Ensuring and enabling the turnaround of underperforming public services is an important responsibility for policymakers, regulators and public managers alike (Jas & Skelcher, 2005; Meyers & Murphy, 2007; Boyne & Meier, 2009; Rutherford & Favero, 2020; Alonso & Andrews, 2021). At present, most research into organizational turnarounds focuses on troubled firms (see reviews in Trahms et al., 2013; Schweizer & Nienhaus, 2017; Castelló-Sirvent & Roger-Monzó, 2023), although there is a good track record of adapting these insights for the public sector (Boyne, 2006; Harvey et al., 2010; Murphy, 2010). Yet public service turnarounds also differ in that they typically take place under the “glare of public accountability.”

The likely effect of accountability to external bodies on the speed, quality and sustainability of public service turnarounds is at present unclear. Psychologists have shown in lab experiments that judgement tends to improve under pressure of accountability (Schillemans, 2016; Aleksovska et al., 2019). Some econometric studies of school and hospital accountability also report positive effects (Kelman & Friedman, 2009; Propper et al., 2010; Dee & Jacob, 2011; Burgess et al., 2013), although others are more cautious (Deming et al., 2016; Hofer et al., 2020). And the public administration literature on accountability lists many cases of oversight regimes that failed to deliver performance gains, consumed significant organizational resources, and/or encouraged various dysfunctional behaviors (Romzek & Dubnick, 1987; Dubnick, 2005; Bevan & Hood, 2006). Thus, the evidence base relating accountability to public service improvement *in general* is mixed (Dubnick, 2005), even before its applicability to struggling agencies awaiting turnaround is considered (see Rutherford, 2014).

Different *types* of accountability may also affect turnaround prospects differently. Intuitively, it might be reasoned that “the worse the malady, the stronger the medicine,” in which case *high*-stakes accountability will be needed to change organizations that have hitherto failed to self-correct. This means imposing costly sanctions for continued failure and tangible rewards for improvement, as judged by the account-holder. And yet, recent research into gentler performance regimes suggests that feedback and standard-setting alone, without significant economic or reputational consequences, might be more suited to fostering the deep and reflective organizational learning necessary to improve public services (Jakobsen et al., 2018; Andersen & Nielsen, 2020). In particular, dysfunctional behaviors like threat rigidity and gaming may lessen under such *low*-stakes regimes, while open debate and broad search activity may increase. However, empirical tests of low-stakes accountability remain rare (examples are Woo et al., 2015; and Andersen & Nielsen, 2020); and, again, it is unclear whether this approach is suited to controlling agencies that are particularly struggling.

To advance this debate, we synthesize the literatures on turnaround and low-stakes accountability to theorize and then test the effects of feedback and standard-setting alone on underperforming public services. We focus on complex, multi-service organizations (like local governments) in which a variety of activities are undertaken by distinct organizational subunits under common leadership. We argue that initiating turnarounds in subunits will depend on external accountability bodies directing leaders’ limited “budget” of attention toward performance deficits that were hitherto unnoticed or downplayed. Absent the threat of financial or reputational consequences for inaction, leaders will need to receive novel yet credible evidence of service failure sufficient to disconfirm their prior beliefs (see Meier et al., 2015;

Andersen & Nielsen, 2020). When this occurs, they may increase control over the problem areas and divert resources from peripheral to core activities – both of which are recognized turnaround strategies (Boyne & Meier, 2009; Murphy, 2010; O’Kane & Cunningham, 2014). However, given the risk of “tunnel vision” among any account-giver (even in low-stakes regimes), performance recovery in the subunit may still be concentrated on dimensions within the remit of the external account-holder, to the neglect of other “extra-jurisdictional” aspects. That is, accountability-induced turnarounds will tend to be *partial*.

We test these hypotheses quantitatively using quasi-experimental methods and the case of the Local Government Ombudsman in England. This is a prime example of a seemingly underpowered accountability body tasked with informing and advising but not sanctioning or coercing its account-givers (Kirkham, 2005; Thomas, 2024; see also Hertogh, 2001 and Harlow, 2018). (Indeed, “ombudsman or ombudsmouse?” is a question periodically asked of these institutions, so feeble are their powers of enforcement believed to be.)<sup>1</sup> Using coarsened exact matching and staggered difference-in-differences, we show that councils that are notified by the Ombudsman of maladministration in their adult social care provision significantly increase leadership attention and control efforts toward this service, and invest more in its core staffing, compared with a matched sample of untreated councils. The incidence of maladministration also falls in notified councils by, on average, 57% (equivalent to 3.05 fewer Ombudsman-upheld complaints per year). However, consistent with the

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<sup>1</sup> Most notably, this was the title of an article by William Gwyn (1973) about the ombudsman for national government – whose statutory makeup is very similar to the LGSCO, but who at least benefits from the support of a dedicated parliamentary committee (see Harlow & Rawlings, 2021, ch. 13).

expected tunnel vision, we detect no broader performance gains on a wide variety of additional measures. Moreover, the immediacy of the improvement suggests that, once notified of failures, subunits proactively seek to recover performance *in anticipation* of leaders' increased attention, rather than in its aftermath. This is a novel finding in the turnaround literature.

The article proceeds as follows. The second section predicts the effects of low-stakes accountability on performance turnaround, and the third describes our empirical case. The fourth section presents the data and empirical strategy, the fifth provides the results and robustness checks, and the sixth discusses the findings.

## **2. ACCOUNTABILITY AND TURNAROUNDS**

### **2.1 Low-stakes accountability, learning, and tunnel vision**

For many scholars and policymakers, it is imperative that accountability regimes are capable of imposing meaningful consequences on account-givers if they are to induce behavioral change and performance improvement. In the regulation literature, for instance, the principle of the “benign big gun” rests upon regulators having access to credible and costly sanctions against wrongdoers, even if these are only to be used sparingly (Ayres & Braithwaite, 1992). Moreover, one of the main challenges in effectively regulating public-sector entities (compared with firms) is thought to be the reticence of one part of the state to litigate against another (Wilson & Rachal, 1977; Konisky & Teodoro, 2016). In administrative law, too, the dominant view is that “the effectiveness of the courts and the ombudsman is primarily determined by the legal force [or not] of their decisions” (Hertogh, 2001). And similarly, in accountability studies, it is often argued that “accountability cannot flourish if the forums, the account holders, are weak” (Bovens & Wille, 2021, p. 857). As for performance

management research, while there is growing interest in using non-financial means of reward and sanction, the presumption remains that some form of meaningful consequences must be tied to results. As Bevan (2010, p. 36) argues: “whatever system of performance measurement is used, to have an impact it needs to be designed to inflict reputational damage on poor performers.”

On the other hand, there is also some emerging interest in the potential for low-stakes accountability to drive performance improvement in public services (Woo et al., 2015; Jakobsen et al., 2018; Andersen & Nielsen, 2020). “Low-stakes” denotes regimes where the measurement and judgement of performance proceed on the same basis as in high-stakes systems, but without any consequences being tied to those findings (see Heinrich & Marschke, 2010; Verger & Parcerisa, 2017). That is, there are no automatic or probable economic effects (for individuals or organizations), reputational incentives (for instance, use of rankings or league tables, which may have second-order economic effects), or other forms of top-down intervention (like merging or closing inadequate agencies) tied to the judgement of the account-holder.

In part, interest in low-stakes accountability arises from concern in public administration and allied fields (like education and healthcare) over the track record of using high-stakes regimes for public service improvement, particularly in terms of undesirable side-effects. As Jakobsen et al. (2018, p. 127) write, “it is far from certain that the net impact ... has been positive” (although see Gerrish (2016) and Boyne (2010) for more optimistic assessments). The hope is that if gaming or other pathologies stem mainly from the rewards-and-sanctions elements of accountability, then removing high-powered incentives should tip the balance in favor of

performance gains (de Wolf & Janssens, 2007; Dee & Jacob, 2011; Verger & Parcerisa, 2017; Jakobsen et al., 2018; although see Levatino et al., 2024 for evidence refuting the idea of significant differences in opportunism between high- and low-stakes regimes).

Beyond attenuating negative-side effects, low-stakes regimes may also improve the account-giver's capacity for learning. As Jakobsen et al. (2018, p. 135) write, "goal-oriented learning and innovation will become more likely when organizational actors perceive that they work in a setting where acknowledging errors and problems will lead to a dialogue about problem solving, rather than punishment." Similarly, Andersen and Nielsen (2020, p. 416) argue that, with incentives removed, novel performance information allows actors to "update their prior beliefs ... and use these new beliefs to prioritize their efforts." Crucially, Andersen and Nielsen do not argue that *motivation to learn* does not matter, but rather that intrinsic public service motivation may provide sufficient inducement, once the inadequacy of current performance is understood. This contrasts markedly with the "worse malady, stronger medicine" argument, and signals that low-stakes accountability adopts more of a "trust" than a "mistrust" approach to governing public services (Le Grand, 2010).

The mechanisms expected to enhance organizational learning are not well specified in this literature, although there are several possibilities. One is that high-stakes typically also means high-*stress* for staff and managers, which may impede learning processes. For instance, research on "threat rigidity" in organizations suggests that channels of information and means of adaptation become constricted in the face of adversity, triggering old, habitual and potentially ill-fitted behaviors (Staw et al., 1981). Another possibility is that high-stakes regimes divert scarce resources toward



unproductive activities, like temporary fixes or superficial “window dressing” (Smith, 1995), resulting in less capacity for genuine search and innovation. And a third is that high-stakes accountability may restrict the scope of any learning to domains linked to rewards or penalties, contributing to what Moynihan (2005, p. 203) calls “narrow process improvement ... rather than broad understanding of policy changes.”

Finding ways to enhance organizational learning is especially relevant to questions of turnaround, because inattentiveness to informative signals is one of the chief contributors to decline. As McKiernan (2006, p. 771) explains:

“Primary causes of decline are the result of defects in the corporate learning process. ... Companies with good learning routines will normally be expected to survive ... Companies with defective routines are more likely to struggle and so to drift down the spiral of decline.”

Many scholars of public sector turnaround adopt the same view (Boyne, 2006; Harvey et al., 2010; Harvey et al., 2015; Vindrola-Padros et al., 2022). Consequently, if an accountability regime is to help arrest public service decline and initiate turnaround, it needs to foster rather than further deplete the account-giver’s capacity for learning. This means that accountability *without* consequences may be particularly suited for supporting turnaround. Thus, we hypothesize:

**H1:** Performance improves in organizations that are *notified* of failures, compared with similarly-performing organizations that are unnotified.

Nevertheless, lowering the stakes in accountability regimes is unlikely to suppress all undesirable effects of account-giving (Levatino et al., 2024). Low-stakes is not no-

stakes. And even just singling out an activity for measurement can lead to effort prioritization, as implied by both Campbell's Law<sup>2</sup> and Goodhart's Law,<sup>3</sup> and by the familiar dictum that "what gets measured gets managed." This leads to the pathology known as "tunnel vision," where managers "focus on the dimensions of performance that are included in the measurement system, displacing other important but unmeasured aspects" (Mannion & Braithwaite, 2012, p. 570; see also Smith, 1995). In the context of turnarounds, tunnel vision would produce narrowly-focused recoveries, extending only so far as the remit or priorities of the account-holder extend. This may be especially likely for highly-specialized accountability bodies that provide narrow rather than broad-based feedback. Thus, we hypothesize a second main effect:

**H1a:** Among organizations notified of failures, performance improves more on dimensions that are within the accountability body's remit, and less on dimensions that are extra-jurisdictional.

Having derived our primary hypotheses, we now consider some of the behavioral responses we might expect following notification of failure.

## 2.2 Leadership attention

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<sup>2</sup> "The more any quantitative social indicator is used for social decision-making, the more subject it will be to corruption pressures and the more apt it will be to distort and corrupt the social processes it is intended to monitor." (Campbell, 1979)

<sup>3</sup> "Any observed statistical regularity will tend to collapse once pressure is placed upon it for control purposes" (Goodhart, 1981).

Much research into turnarounds suggests that recognizing the need for corrective action is the first hurdle needing to be overcome. Leaders typically have demanding jobs and limited attention spans (Simon, 1997; Tuggle et al., 2010), and frequently suffer from information overload (Walgrave & Dejaeghere, 2017). What is more, transmission of information to the top of hierarchies may be imperfect and delayed (Argyris, 1977), so that even attentive leaders can be poorly informed about performance in organizational subunits. All of this contributes to what Murphy and Jones (2016) call a “lack of self-awareness” among decision-makers, with the implication that that some kind of shock or “trigger point” is needed to initiate problem recognition and recovery efforts (Gopinath, 2005).

The situation is further complicated in the public sector by the contested nature of success and failure (Brewer, 2006; Andersen et al., 2016). Turnaround involves “a recovery in performance after a period of organizational failure” (Boyne & Meier, 2009, p. 835). For firms, this manifests as decline in market-share or financial return (see Pearce & Robbins, 1993). But in government, performance is multidimensional and the moment of outright failure may be unclear (Paton & Mordaunt, 2004; Van de Walle, 2016). Persistent under-performance does not risk bankruptcy, and government services must be provided in even the most unfavorable circumstances (Hargrove & Glidewell, 1990). Consequently, it is not just leader attention that needs to be secured, but also problem acceptance. Leaders may need to be persuaded that the errors reported are serious and distinctive among comparable organizations – rather than being “par for the course,” a product of unusually difficult operating circumstances, or a function of government’s “impossible jobs” (see Hargrove & Glidewell, 1990; Entwistle & Doering, 2024).

One way to accomplish this reset is to provide feedback that is both novel yet highly credible (Andersen & Nielsen, 2020). If the account-holder can convince leaders that prior beliefs about performance in (part of) their organization were misplaced or outdated, the resulting “gap” between expectations and new evidence may be sufficient to initiate remedial actions (Meier et al., 2015). Of course, the more the feedback disconfirms prior beliefs, the more skeptical leaders may be as to its veracity, and the harder the account-holder must work to be persuasive. Leaders may be tempted to dismiss problems as “one-offs” or “noisy data” – perpetuating what McKiernan (2006) calls “ostrich” behavior. Hence, turnaround is more likely when account-holders have strong claims to the credibility of both their data and their interpretation of it as falling below minimum performance standards. When this is the case, we expect:

**H2:** Following notification of failures, leaders pay more attention to the problematic service area than they did prior to notification.

### **2.3 Centralization, core competencies and urgency**

With leader attention and problem acceptance secured, various internal turnaround strategies may be instigated (see reviews in Boyne, 2006; Trahms et al., 2013; Schweizer & Nienhaus, 2017; Castelló-Sirvent & Roger-Monzó, 2023). For now, we focus on centralization of control over the problem, diversion of organizational resources from peripheral to core activities, and prompt implementation.

Firstly, when leaders are concerned about performance problems, they may seek greater control over the subunit responsible for the errors (O'Kane & Cunningham, 2014; Meier et al., 2015), at least until improvement takes hold. As O'Kane and Cunningham (2014, p. 965) write, “‘Hard’ leadership is important during turnaround

to centralize command, reduce participation and instill close supervision and control.” McKiernan (2006) summarizes this approach as “more autocracy.” The point is not only that sub-unit discretion declines, but also that enhanced efforts are made to ensure that commands issued at the top are obeyed below – lest recovery be impeded by non-adherence to the turnaround strategy. Thus, we hypothesize:

**H3:** Following notification of failure, leaders will temporarily centralize control over the responsible subunit.

Secondly, successful turnarounds often involve a shift in resources from peripheral to core purposes. The idea is to dispense with inessential (or perhaps not-*imminently* essential) activities, and redirect staff and managers toward “doing what they know best” (McKiernan, 2006, p. 760) and “attending to the basics ... to the work that really counts” (Murphy, 2010, p. 169). Practically, as Boyne and Meier (2009, p. 859) write, this often means “replacing non-core with core staff.” Thus, we hypothesize:

**H4:** Following notification of failures, organizations invest more in core competencies in the responsible subunit.

Finally, just as “time is of the essence” in successful corporate turnarounds (Tangpong et al., 2015; Barbero et al., 2017), we similarly expect that more rapid reactions by public managers will produce larger recoveries. In firms, the aim is to stem decline before the loss of market share and/or investor confidence creates a downward spiral in the organization’s resource base. In government, analogously, the sooner the agency enters a recovery phase, the easier it will be to retain critical staff unnerved at the downturn, recruit new high-quality staff who would be deterred from “joining a sinking ship,” retain good will from both service users and overseers, and escape top-down regulatory intervention. Hence:

**H5:** Following notification of failures, organizations that implement changes more rapidly will experience a greater recovery in performance.

## 2.4 Summary of theory

To summarize, our primary proposition is that external accountability *can* induce performance turnaround in failing organizations (or subunits thereof), but that a low-stakes regime may be sufficient if it enhances organizational learning and lessens unproductive gaming. However, absence of high-powered incentives is unlikely to fully suppress tunnel vision, so we also expect accountability-induced turnarounds to be partial rather than comprehensive, focused on dimensions of performance most relevant to the account-holder. To initiate public service turnarounds, leadership attention needs to be captured through novel but credible performance feedback and authoritative standard setting. When this occurs, internal turnaround strategies will begin, including centralized control over problems and investment in core competencies. The more rapidly this occurs, the greater the recovery.

## 3. EMPIRICAL CASE

### 3.1 The Local Government Ombudsman

The Local Government and Social Care Ombudsman<sup>4</sup> is responsible for investigating complaints from users of local public services in England, and for encouraging lesson learning and service improvement. Led by a named Ombudsman, the LGSCO employs nearly 200 staff and is one of several third-party

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<sup>4</sup> In statute, the Commission for Local Administration.

complaint handlers in the UK's administrative justice system (see Gill, 2018; Harlow & Rawlings, 2021; McBurnie et al., 2023). The LGSCO can only commence investigations after efforts at resolving the complaint locally have failed, and its remit extends only to matters of maladministration and service failure.<sup>5</sup> Broader issues of policy and performance cannot be investigated. Maladministration concerns “the manner in which decisions are reached,” and whether any bias, neglect, inattention, delay, incompetence, ineptitude, perversity, turpitude, arbitrariness, or similar was involved (Parpworth, 2022, p. 391). Service failure typically means failure to provide a service to which the complainant was entitled. For simplicity, we refer to both dimensions as “maladministration,” hereafter.<sup>6</sup>

During 2012-2019, the LGSCO received an average of 11,772 complaints annually, undertook 5,947 investigations and issued 3,483 adverse findings per year in which the complainant was found to have suffered injustice following maladministration. With each upheld complaint, the LGSCO makes recommendations for how the council should respond, such as apologizing, providing services, improving procedures or staff training, and/or offering financial compensation (LGSCO, 2024b). Anonymized decisions and recommendations are published online, but the LGSCO has no powers to compel councils to act upon them (Kirkham, 2005; Thomas, 2024). In addition, in a very small minority of cases, the Ombudsman herself (rather than

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<sup>5</sup> This remit resembles the issues studied by the “service recovery” literature (Thomassen et al., 2017; Döring, 2022; Caillier, 2023), although complaints referred to Ombudsman tend to be more serious and are always subject to some form of dispute between citizen and provider.

<sup>6</sup> Indeed, when service failure was added to the legislation governing the Ombudsman in 2007, this largely formalized the LGSCO's existing approach under the flexible concept of “maladministration” (see Seneviratne, 2008).

the organization at large) issues a “public interest report” to formally notify council leaders of the seriousness of the errors uncovered. This is mostly<sup>7</sup> a discretionary action taken in cases of recurring faults, significant injustices, systemic problems, frequent complaints about a single issue, or significant topical issues (LGSCO, 2024). Again, there are no powers to compel change, although the report must be discussed at a senior decision forum and notified to the local press (Thomas, 2024). Over time, the proportion of cases escalated in this manner has reduced, reflecting resource constraints and a desire to conserve the distinctiveness of this solitary tool of enforcement (Kirkham, 2005). Indeed, as our matching analysis below reveals, a substantial number of councils display observable characteristics associated with the escalation criteria, even though relatively few receive censure in any given year.

Ombudsmen provide an ideal case for exploring the effects of low-stakes accountability. As noted, the advisory nature of their work and lack of enforcement powers are widely recognized (Stacey, 1971; Gwyn, 1973; Harlow, 2018). And yet there is also a small body of qualitative work pointing to the remarkable influence of these “soft power” institutions on both complainant outcomes and broader organizational behavior (Hertogh, 2001; Kirkham et al., 2008; McBurnie et al., 2023; Thomas, 2024). As for the LGSCO specifically, this is clearly a low-stakes accountability regime. As Thomas (2024) writes, “of all UK ombuds, the LGSCO perhaps has the weakest structural position as regards compliance.” It cannot compel changes in councils, nor sanction funding or autonomy, and nor does it call for disciplinary action against individual officials. We have seen no evidence of

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<sup>7</sup> A report is also the legal default outcome where a council refuses to carry out the recommendations and the Ombudsman decides to maintain her position.



LGSCO findings being cited by central government or others to justify regulatory action. And our own analysis (Figure 8 below, and Table A8 in the appendix) shows no effects of Ombudsman notifications on leadership turnover in councils. (Whereas, historically, England's hospital rankings led to several chief executives losing their jobs (Bevan, 2021), and research shows that headteacher retention rates are lowest in schools receiving the worst, "inadequate" judgement from the English education inspectorate (Lynch et al., 2017)).<sup>8</sup>

Most legal scholars regard the combination of publicity and moral authority as the principal means of potential impact for the LGSCO. Yet, while the Ombudsman's work is published, the prospect of significant reputational damage is remote.

Benchmarked against Hibbard et al.'s (2003) four criteria for gauging the strength of reputation-based accountability systems, this regime appears very weak. LGSCO does *not* (i) rank organizations best-to-worst, as happens, for instance, with school league tables; and its findings are *not* (ii) widely disseminated, given the decline in local media outlets and readership (Clark, 2021). Moreover, neither the ombudsman system in general, nor the verdict of "injustice caused by maladministration," is (iii) clearly understood by the general public<sup>9</sup> – quite the reverse, in fact (Kirkham, 2005; Dunleavy et al., 2010). And nor is (iv) performance re-assessed to a regular schedule. Consequently, while external reputation may partially influence council deliberations (see McBurnie et al., 2023), there are far stronger means of leveraging

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<sup>8</sup> Inspection can also be highly traumatic for school leaders, and in recent years contributed to at least one suicide (Perryman et al., 2025).

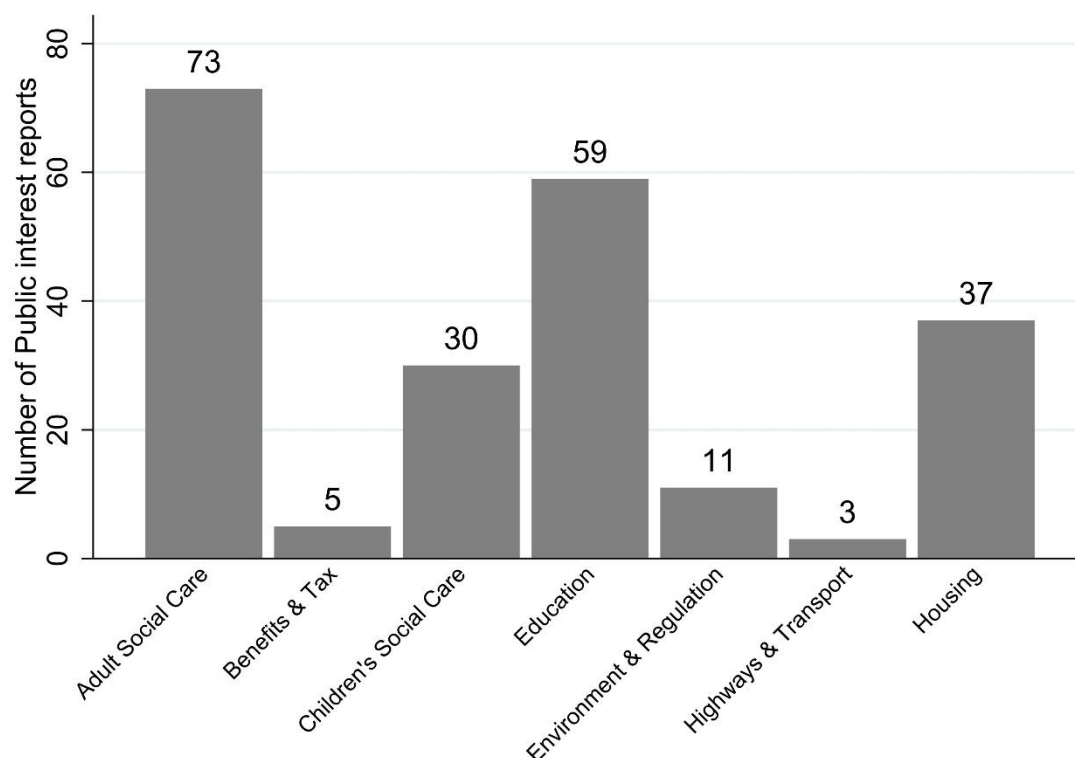
<sup>9</sup> Moreover, Hodges (2018, p. 54) suggests that lawyers and judges are often unfamiliar Ombudsman systems, too.

reputation for purposes of accountability and performance improvement than those available to the LGSCO.

### **3.2 Adult social care**

We focus on Ombudsman oversight of councils' adult social care departments, partly because of the serious consequences of maladministration in this field, and partly because of the empirical opportunity that this provides. As Gulland (2024, p. 203) argues, "much social care decision making happens when people are at their most vulnerable and where outcomes can make life-changing differences to people's health and wellbeing." As such, while we do not claim to match the severity of performance failure examined in some turnaround research, complaints in this domain "range from the petty to the tragic" (Gulland, 2010, p. 457) and at times result in very serious individual and social harms. In addition, adult social care provides an excellent empirical opportunity for quantitative research. As Figure 1 shows, between 2012 and 2019 (our study period), LGSCO issued a total of 237 public interest reports to councils, of which adult social care accounted for the most (73 reports), followed by education (59 reports) and housing (37 reports). This is based on 6,955 upheld complaints about councils' adult social care over the period, amounting 28% of all upheld complaints. This provides ample scope for quantitative analysis.

**Figure 1: Public interest notifications by service category, 2012-2019**



**ALT-TEXT: Bar chart showing number of public interest reports issued for each local government service area**

Adult social care services involve the provision of personal care and domestic assistance to individuals aged 18 and over, including the elderly and others experiencing physical or mental illness, disability, or other challenging circumstances (Glasby et al., 2021; Burn et al., 2024). Unlike healthcare in England, social care is a local government responsibility and is means-tested (Needham & Hall, 2023). It is one of the most high-profile areas of work for English councils, and accounts for about a fifth of local public expenditure. This is rising with increasing life expectancy, an aging population, and the growing proportion of adults with diagnosed disabilities.

In England, care services are almost entirely *produced* by private or third-sector agencies (Goodair et al., 2024). (Complaints made against producers can also be

investigated by the LGSCO, but are excluded from our analysis.) The primary role for councils is to undertake the needs-testing and means-testing required by law to determine what level of support is required and how it should be funded (Needham & Hall, 2023; Burn et al., 2024). Most complaints to the Ombudsman relate to “delays in the assessment of people’s needs, processes that fail to put the individual at the center, care that fits with the system’s offer rather than the person’s needs and preferences, and poor communication” (LGSCO, 2024a, p. 1). In addition, councils have a broad well-being duty to promote quality of life among service users, a responsibility to ensure the availability of quality services in their areas, and a duty to support informal care givers like family members (Burn et al., 2024).

## 4. DATA AND METHOD

### 4.1 Variables and data sources

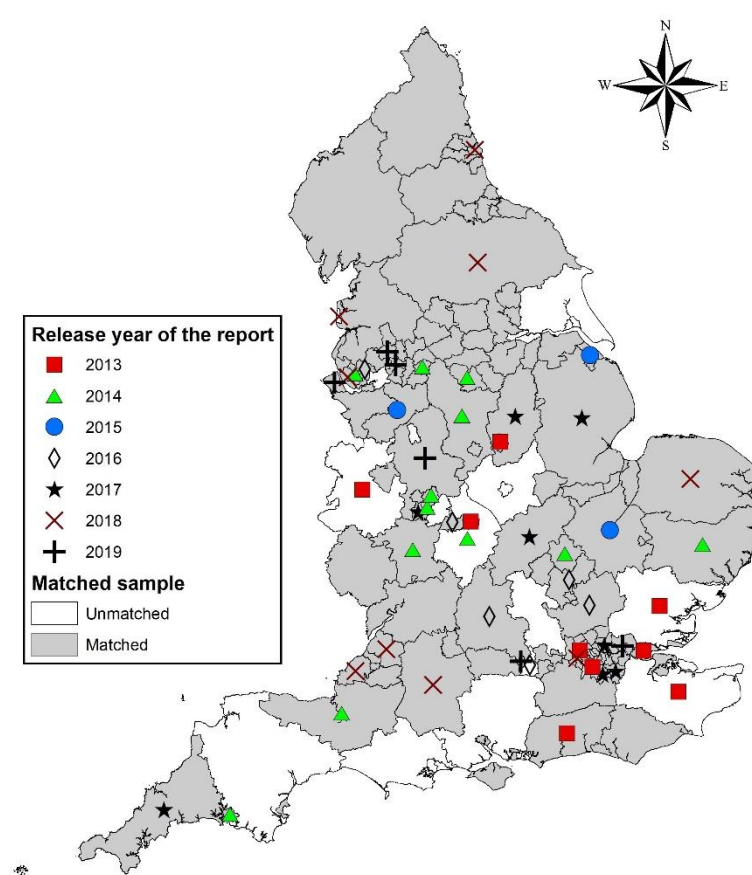
To test the effects of the Ombudsman’s public interest notifications on turnaround outcomes and behaviors in adult social care, we construct a panel dataset at the upper-tier council level from 2012 to 2019.<sup>10</sup> We begin with public interest reports released in 2013, so that all analyses have at least one year of pre-treatment data; and we halt our analysis in 2019 to avoid biases due to COVID-19, which hugely disrupted this sector. Our variable of interest – the release of a public interest notification – is a dummy variable derived from the Ombudsman’s online archive. For

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<sup>10</sup> At the time of our analysis, local government in England consisted of five council types in two vertical arrangements: A two-tier system in predominantly rural areas, comprising 27 county councils (upper tier) and 201 district councils (lower tier); and single-tier authorities elsewhere, including 32 London boroughs, 36 metropolitan districts, and 55 unitary authorities. Social care is a responsibility of all council types except districts.

any given council, the variable is set to 1 for the year the report was issued and all subsequent years, and 0 otherwise. During our study period, 53 out of 151 upper-tier councils received at least one such a notification, with 12 receiving two reports, and four councils receiving three reports. Figure 2 illustrates the timing and geographic distribution of this low-stakes accountability “treatment.”

**Figure 2: Release year of public interest report for each local authority**



**ALT-TEXT: Map showing upper-tier local authorities in England, with treatment and control areas indicated**

Regarding our dependent variables:

We quantify performance turnaround (H1) by examining the **incidence of maladministration** in adult social care departments as determined by the ombudsman's regular casework.<sup>11</sup> To test the breadth of any turnaround (H1a), we further analyze administrative data describing the speed of assessment activity, and two user satisfaction surveys. Specifically, we measure **the number of days of delay** in transferring individuals from hospital to adult social care due to delays in care needs assessment or public funding (separately), or because of disputes between agencies. In addition, we employ **user satisfaction data** from the annual Adult Social Care Survey to measure service users' ease of finding information and advice, satisfaction with care services, and perceptions about the impact of their care on quality of life. Finally, we also adopt two perceptual indicators from the biennial Personal Social Survey of Adult Carers measuring satisfaction of informal carers with council-run services.<sup>12</sup> Overall, these additional measures capture aspects of council performance that extend beyond the Ombudsman's narrow remit of maladministration whilst being potentially amenable to any improvement initiatives that managers attempt after censure.

Turning to the expected behaviors during subunit turnarounds:

To capture leadership attention (H2), we record the **number of formal decisions issued each year** by each council about its adult social care services. Councils

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<sup>11</sup> This was supplied to us directly from the LGSCO as a case-level database, although all individual decisions can be inspected online for several years after determination. All survey data reported in this paragraph are also publicly available via government or NHS websites.

<sup>12</sup> As this survey is conducted biennially, we assume that the situation in non-survey years remains consistent with the previous survey year, to align with our annual panel analysis.

publish lists of decisions taken at cabinet meetings and other committees, typically designating these as “key” and “non-key” decisions, respectively. We obtained and coded every decision related to adult social care during our study period, creating a panel dataset to track committee attention toward this service over time. Additionally, to measure the degree of centralized control over adult social care departments (H3), we obtained internal audit plans from council websites and recorded the **number of audit days assigned to adult social care each year**. As Power (1999, p. 6) explains, audit committees function as “internal agents of control,” and internal audit plans are documents that specify where an organization’s control risks lie, and how the internal audit staff is to be deployed to provide internal assurance against those risks. Thus, a higher allocation of audit time signals an increased concern to secure compliance in that subunit, consistent with O’Kane and Cunningham’s advice to “instill close supervision” during turnarounds (already cited).

To observe the rate of investment in core competencies (H4), we measure **staff recruitment patterns** using the Adult Social Care Workforce Dataset, which is collated annually by central government from local statistical returns. This includes information on over 110,000 roles within 151 councils in categories including frontline, professional and supervisory roles. Frontline includes staff interacting directly with the public, other than regulated professionals; while supervisors manage and coordinate frontline and professional teams. In addition, we use data from Municipal Yearbooks to track **turnover in director-level** leadership of adult social care in each council.

Finally, because performance may be influenced by local demographic conditions, we control for the proportion of adults in the local population and the proportion with long-term health conditions, using Office for National Statistics data.

Variable definitions and data sources are provided in Appendix Tables A1 and A2.

## 4.2 Empirical strategy

Our study aims to provide a credible causal explanation of the impact of the Ombudsman's public interest notifications on local government performance and behaviors. The key challenge is ensuring that our treatment and control groups are fully comparable, given their non-random assignment. The Ombudsman can only notify councils about which it has both received and upheld a complaint from the public. (That is, it has no right to self-initiated investigations.)<sup>13</sup> In addition, escalation to the public interest report stage is largely a discretionary choice, albeit justified *internally* according to the aforementioned criteria. To ensure the accuracy of our Average Treatment Effect on the Treated estimates, we first employ matching to maximize comparability between treatment and control groups, and then apply staggered difference-in-differences to infer the effects of the Ombudsman's report releases over time.

We use the Coarsened Exact Matching (CEM) approach proposed by Iacus et al. (2012) and widely used in economics (see Aneja & Xu, 2021; Kotsogiannis et al., 2024). CEM is an exact matching algorithm that partitions data into strata based on

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<sup>13</sup> If, during the course of an investigation, a potential further injustice is uncovered, LGSCO may also investigate it without receiving a further complaint (Thomas, 2024).



all possible combinations of predefined observable bins. Compared to standard matching methods, CEM enhances balance, reduces model dependence, and decreases estimation error. The matching variables are first coarsened (divided into discrete categories) and then exact matches are made based on these categories. This process yields counterfactuals that are comparable in terms of the joint distribution of observable baseline characteristics. (We recognize that the Ombudsman's discretionary decision-making may also reflect unobservable factors that cannot be accounted for in our data.) We use the Ombudsman's case management data to quantify the four aforementioned criteria for regulatory escalation: recurring faults, significant injustices, systemic problems, and repeated complaints about a single issue. To operationalize these criteria, we create seven indicators: the number of complaints received, investigated, and upheld; and the major categories of recommendations – apologies, financial redress, procedure reviews, and staff training. In addition, we include the proportion of the population with long-term health conditions and the proportion of adults in the population.

Table 1 shows that, before matching, there were significant differences between the treated and control groups in total complaints, investigated complaints, financial redress cases, and the rate of health problems in the population (Columns 1 and 2). After matching, we found suitable counterparts for 48 out of 53 treated councils, and for 85 out of 98 control councils. Column 4 reports the mean differences between matched treated and control councils in 2012. Notified and unnotified councils are now highly comparable, with no statistically significant differences in any of the Ombudsman complaints and recommendations or demographic covariates.

In matching councils using 2012 data, one further concern is that these attributes may fluctuate during the remainder of the study period, reducing the quality of the match. Therefore, in Figure A1 in the Appendix, we examine correlations between the values of these covariates in 2012 and their average values from 2013 to 2019. We find these variables to exhibit dynamic stability (close to a 45-degree linear correlation), indicating that our approach is sufficient to capture the long-term characteristics of local councils. (We refrain from implementing a dynamic matching strategy because the staggered timing of treatment necessitates stable untreated controls over time, and re-matching would distort the control group composition and induce post-treatment bias.)

**Table 1: Descriptive statistics of whole and matched samples in 2012**

	(1) Whole sample mean	(2) Treated- control diff.	(3) Matched sample mean	(4) Matched treated- control diff.
<b><i>Panel A: Ombudsman complaints and recommendations</i></b>				
number of total complaints (logged)	1.051	0.591*** (0.139) [0.000]	0.939	0.222 (0.142) [0.118]
number of investigated complaints (logged)	0.461	0.309*** (0.110) [0.005]	0.377	-0.006 (0.105) [0.958]
number of upheld complaints (logged)	0.197	0.114 (0.070) [0.101]	0.140	-0.026 (0.066) [0.692]
number of apologies (logged)	0.124	0.063 (0.053) [0.232]	0.075	0.036 (0.045) [0.420]
number of financial redress cases (logged)	0.151	0.112** (0.054) [0.039]	0.098	0.000 (0.053) [0.998]
number of procedure reviews (logged)	0.068	0.023 (0.039) [0.556]	0.010	0.000 (0.017) [0.995]
number of staff	0.028	0.011	0.015	0.013

training (logged)		(0.023)		(0.020)
		[0.638]		[0.506]
<b>Panel B: Population demographics</b>				
rate of health problems in the population (%)	0.132	0.006** (0.003) [0.048]	0.132	0.004 (0.003) [0.282]
adult rate in the population (%)	0.801	0.000 (0.004) [0.912]	0.802	0.000 (0.004) [0.978]
Sample size		151		133
of which treated upper-tier local authorities		53		48
of which control upper-tier local authorities		98		85

*Notes:* The table compares the whole sample and matched local authorities in 2012. Column 1 shows mean characteristics for the whole sample, and Column 2 reports mean differences between treated and control groups. Columns 3 and 4 present the same metrics for the matched sample (using coarsened exact matching). \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Standard errors are reported in parentheses, and p-values in square brackets.

Using the trimmed sample, we estimate the dynamic effects of the Ombudsman's public interest notifications using a staggered difference-in-differences (DiD) event study framework (Sun & Abraham, 2021). This extends conventional DiD by accounting for staggered treatment and addressing biases from heterogeneous treatment timing, yielding robust causal estimates. It captures both immediate reactions and longer-term trends in councils' behavior. We implement Callaway and Sant'Anna's (2021) group-time average treatment effect estimator to accommodate these features.

For council  $i$  in year  $t$ , let  $E_i$  be the period when council  $i$  is first treated. Then let  $K_{it} = t - E_i$  be the number of years before or after the event. We regress outcome  $Y_{it}$  on 1 ( $K_{it} = k$ ) relative year indicators for the fully saturated set of indicators going from the beginning to end of the sample:

$$Y_{it} = \alpha + \sum_k \beta_k [1(K_{it} = k)] + \gamma_i + \delta_t + \varepsilon_{it} \quad (1)$$

In this specification,  $\beta_k$  is the Average Treatment Effect on the Treated (ATT) estimate at length of exposure  $k$  from the first report releasing. We control for council and year fixed effects, denoted by  $\gamma_i$  and  $\delta_t$  respectively. We also control for demographic conditions. Standard errors are clustered at the council level.

## 5. RESULTS

Using the combined CEM and staggered DiD strategies, we begin by testing the main effects of low-stakes accountability on future council performance, and then turn to the possible mechanisms through which performance improvements occur.

### 5.1 Turnaround outcomes

Our main indicator of turnaround is whether the future incidence of maladministration reduces. Table 2, Panel A, Column 1 presents the baseline specification for Equation (1), with a basic set of fixed effects without additional controls. This indicates that, on average, councils significantly reduced upheld complaints following receipt of a public interest report. The estimates in Column 2, which add demographic controls, remain largely unchanged in both significance and magnitude — a 57% reduction in maladministration after notification, corresponding to approximately 3.05 fewer Ombudsman-upheld cases.<sup>14</sup> Columns 3 and 4 present a placebo test using volume of complaints *not* accepted for investigation as the outcome variable. This tests whether our findings reflect broader reporting trends or unrelated factors, rather than

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<sup>14</sup> Percentage change is calculated as  $(e^{(-0.844)} - 1) * 100 = -57\%$ . Given the mean of 5.347 upheld cases, the estimated treatment is  $5.347 * e^{(-0.844)} = 2.301$ . This corresponds to an absolute reduction of 3.05 cases on average. The 99% confidence interval implies that the percentage reduction may range from approximately 4% to 81%, representing an absolute reduction between 0.21 and 4.32 cases on average.

genuine performance gains. The analysis shows no significant change in unaccepted complaints after notification, reinforcing our interpretation of H1.

**Table 2: Effects of low-stakes accountability on performance turnaround**

	(1)	(2)	(3)	(4)
	Investigated upheld cases		Not accepted for investigation	
<b>Panel A</b>				
Post report releasing	-0.820*** (0.305) [0.007]	-0.844*** (0.312) [0.007]	0.275 (0.254) [0.279]	0.256 (0.245) [0.295]
<b>Panel B</b>				
Year 0	-0.637** (0.284) [0.025]	-0.614** (0.300) [0.041]	0.324* (0.193) [0.093]	0.308* (0.185) [0.096]
Year 1	-0.907*** (0.315) [0.004]	-0.970*** (0.310) [0.002]	0.206 (0.244) [0.397]	0.190 (0.246) [0.440]
Year 2	-0.839** (0.347) [0.016]	-0.856** (0.361) [0.018]	0.373 (0.263) [0.157]	0.370 (0.247) [0.135]
Year 3	-0.898 (0.549) [0.102]	-0.938* (0.561) [0.094]	0.198 (0.407) [0.626]	0.156 (0.401) [0.697]
Local authority FEs	✓	✓	✓	✓
Year FEs	✓	✓	✓	✓
Demographic control		✓		✓
Empirical method	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD
Observations	1,064	1,048	1,064	1,048

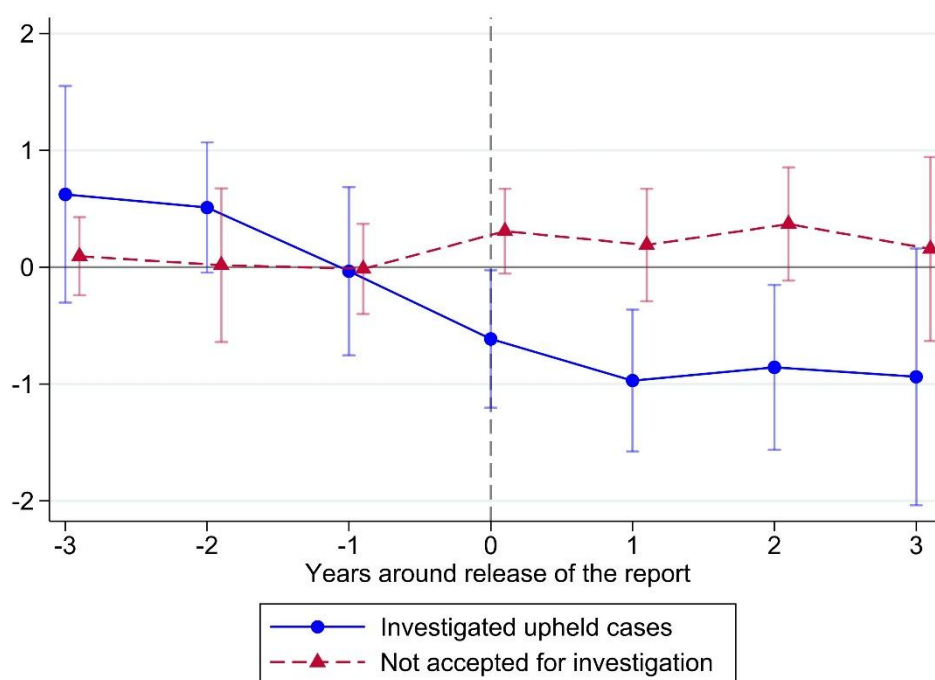
*Notes:* The unit of observation is the local authority × year. Post report releasing is a dummy that is 1 if local authority has received the report, and 0 otherwise. Post is a dummy that is 1 if the year is after the report releasing. Standard errors clustered at the local authority level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors are reported in parentheses, and p-values in square brackets.

The DiD design assumes that, absent the shock event (notification), treated and control councils would have experienced similar trends in outcome variables.

Although this is not directly testable, as usual we test its validity for pre-treatment trends. Figure 3 depicts the impact of notification on both upheld complaints and

cases not accepted for investigation, using year-by-year estimates. Before receiving the Ombudsman's report (year  $-1$  and earlier), there is no significant difference at the 5% level between treated and control councils in the number of upheld complaints, indicating parallel trends. However, after notification, treated councils show a clear reduction in upheld cases relative to the matched control group. This trend is corroborated by the year-by-year estimates presented in Table 2, Panel B. The effect becomes apparent in the year of notification (year 0), with an approximate 47.1% decrease, reaching a peak in the following year (year 1) with a decline of around 59.6%. Conversely, in the placebo test, the number of cases *not* accepted for investigation remains stable before and after notification, again suggesting that the observed improvement in maladministration is not being driven by general changes in complaint reporting.

**Figure 3: Event study - Performance turnaround in censured vs. uncensored councils**



**ALT-TEXT: Event study graph showing effect of notification on upheld complaints**

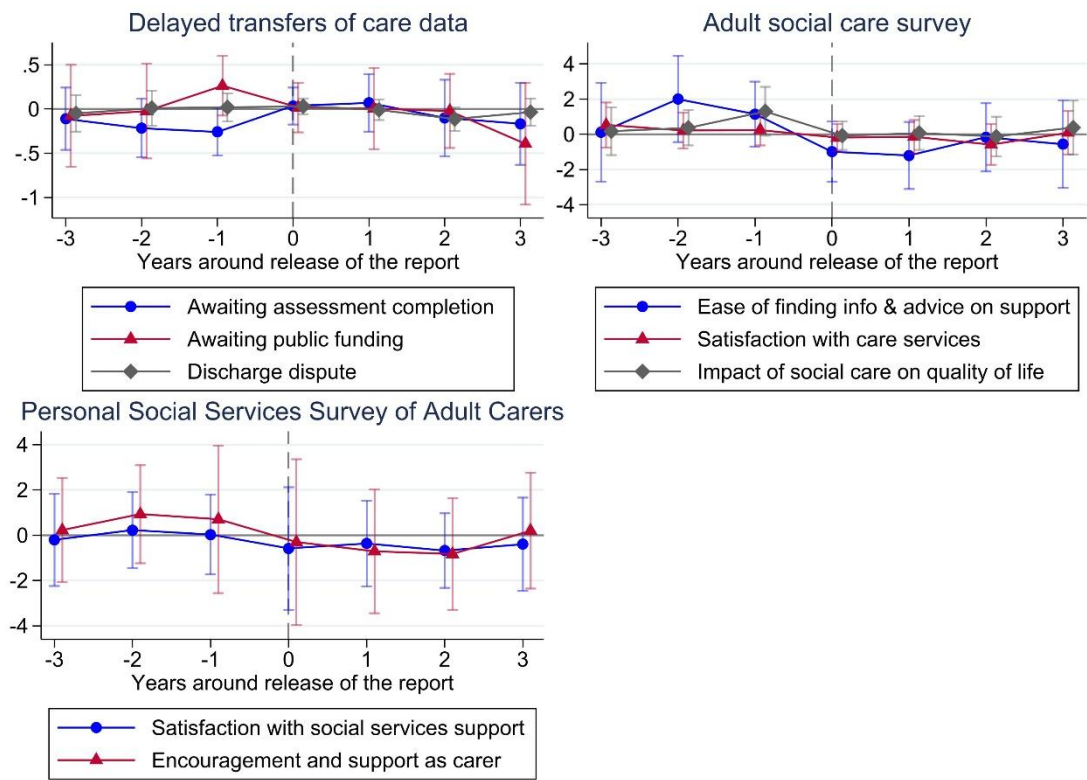
A further concern is that our results might be driven entirely by the influence of never-treated councils in the staggered design, rather than reflecting the temporal differences between councils receiving reports. We therefore also adopt the most conservative strategy by conducting staggered DiD analysis using only councils that received public interest reports during the study period, excluding the never-treated group. Even with this smaller sample, our baseline findings remain supported (see Table A5 and Figure A2 in the appendix).

Finally, as noted, several councils receive two or even three public interest reports related to adult social care during our study period. So far, we only accounted for the first moment of notification. In the appendix, we modify our approach to explore the impact of subsequent notifications. As explained there, these additional analyses all suggest that the performance improvement is driven largely by the initial notification, with little evidence of incremental effects from subsequent reports.

With the main effect established, we next investigate whether performance improvements are confined narrowly to the Ombudsman's remit or are more broad-based (H1a). In Table 3, we use a combination of administrative and survey data to assess council performance beyond maladministration. In Panel A of Table 3, across the eight columns, we find that censuring by the Ombudsman induced no significant effect on council-attributable delayed discharges from hospital, or satisfaction among either care users or informal carers. In Figure 4 and Panel B of Table 3, we further examine year-by-year estimates. The coefficients for each year following notification show that this maladministration-focused accountability regime *did not* have

significant effects on performance beyond the account-holder’s remit, consistent with H1a.

**Figure 4: Event study - Tunnel vision in censured vs. uncensored councils**



**ALT-TEXT: Event study graphs showing effect of notification on administrative and survey measures of performance**



**Table 3: Effects of low-stakes accountability on performance beyond maladministration**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Delayed Transfers of Care Data			Adult Social Care Survey			Personal Social Services Survey of Adult Carers	
	Awaiting assessment completion	Awaiting public funding	Discharge dispute	Ease of finding info & advice on support	Satisfaction with care services	Impact of social care on quality of life	Satisfaction with social services support	Encouragement and support as carer
<b>Panel A</b>								
Post report releasing	-0.041 (0.158) [0.796]	-0.098 (0.178) [0.582]	-0.031 (0.041) [0.447]	-0.732 (0.802) [0.361]	-0.205 (0.430) [0.633]	0.064 (0.470) [0.892]	-0.506 (0.920) [0.583]	-0.415 (1.183) [0.726]
<b>Panel B</b>								
Year 0	0.034 (0.107) [0.752]	0.018 (0.143) [0.902]	0.031 (0.046) [0.492]	-0.990 (0.874) [0.257]	-0.192 (0.395) [0.627]	-0.072 (0.410) [0.860]	-0.585 (1.380) [0.671]	-0.309 (1.873) [0.869]
Year 1	0.070 (0.165) [0.670]	0.004 (0.234) [0.985]	-0.010 (0.059) [0.872]	-1.208 (0.975) [0.215]	-0.142 (0.488) [0.771]	0.077 (0.490) [0.876]	-0.362 (0.967) [0.708]	-0.714 (1.397) [0.609]
Year 2	-0.100 (0.220) [0.649]	-0.023 (0.214) [0.916]	-0.113* (0.068) [0.096]	-0.171 (0.989) [0.863]	-0.580 (0.593) [0.327]	-0.130 (0.576) [0.822]	-0.679 (0.842) [0.420]	-0.832 (1.255) [0.508]
Year 3	-0.167 (0.236) [0.478]	-0.392 (0.351) [0.264]	-0.034 (0.078) [0.661]	-0.560 (1.270) [0.659]	0.093 (0.627) [0.882]	0.381 (0.775) [0.623]	-0.395 (1.052) [0.707]	0.195 (1.307) [0.882]
Local authority FE	✓	✓	✓	✓	✓	✓	✓	✓
Year FE	✓	✓	✓	✓	✓	✓	✓	✓
Demographic control	✓	✓	✓	✓	✓	✓	✓	✓
Empirical method	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD
Observations	1,040	1,040	1,048	1,026	1,016	1,018	1,048	1,048

*Notes:* The unit of observation is the local authority × year. Post report releasing is a dummy that is 1 if local authority has received the report, and 0 otherwise. Post is a dummy that is 1 if the year is after the report releasing. Standard errors clustered at the local authority level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors are reported in parentheses, and p-values in square brackets.

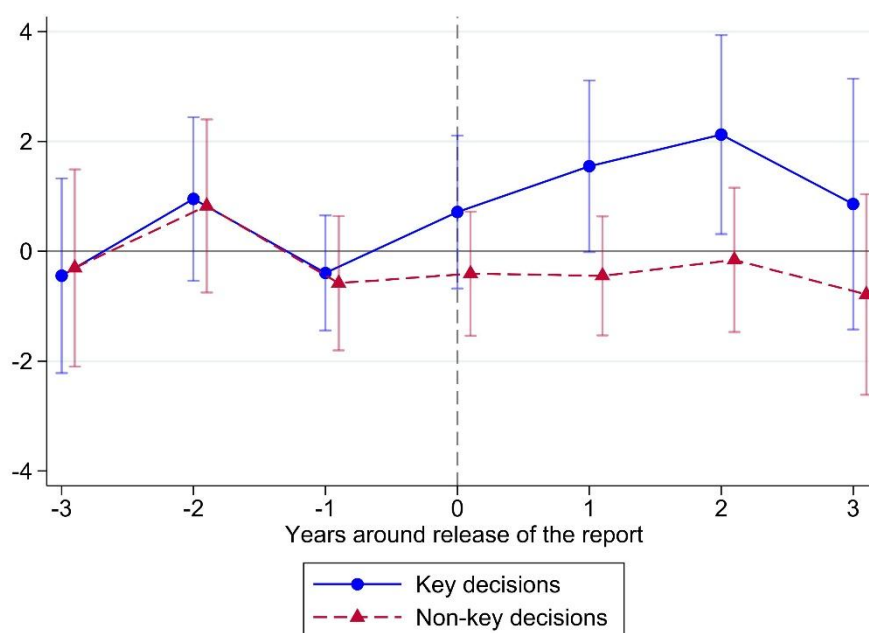
## 5.2 Leadership attention

Having confirmed the main effect of notification on turnaround success (narrowly conceived), we next explore the hypothesized behaviors.

To measure any changes in leadership attention, we quantify the number of key decisions about adult social care taken by council cabinets. If attention is redirected by the low-stakes accountability event, we expect an increase in top-level decision-making related to adult social care. Figure 4 presents an event study where the solid blue line depicts the yearly impact of notification on the number of key decisions. Prior to treatment, the coefficients for key decisions hover around zero, indicating no significant differences between notified and unnotified councils. However, following censure, decision-making on adult social care increases, peaking in the second year with around 9.08 additional top-level decisions on average, relative to the pre-censure period. This suggests that leadership attention is indeed significantly redirected after notification from the Ombudsman is received, supporting H2.<sup>15</sup>

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<sup>15</sup> In the appendix, Table A6 Column 1 provides the specific coefficient estimates illustrating this effect.

**Figure 5: Event study – Leadership attention in censured vs. uncensored councils**

**ALT-TEXT: Event study graph showing effect of notification on key and non-key decisions**

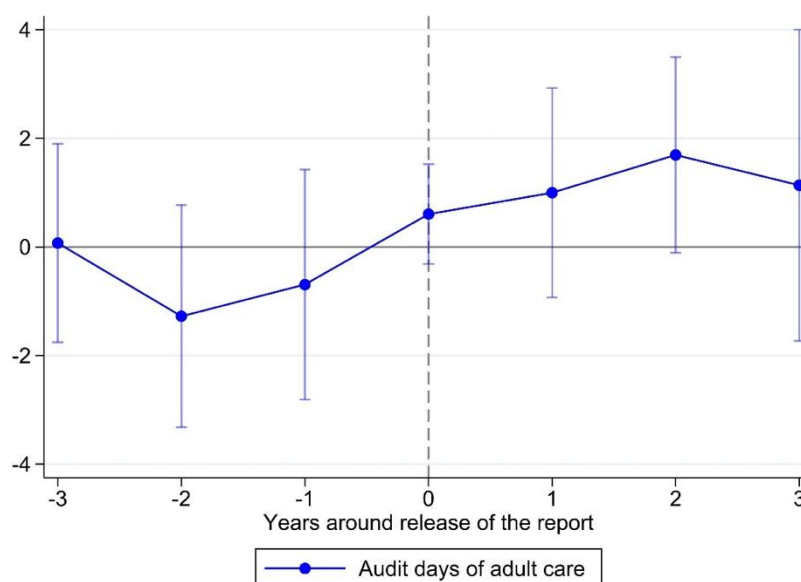
To validate this finding, we again conduct a placebo test, now using “non-key” decisions by more junior and specialist committees as the outcome variable. If the observed increase in key decisions is truly attributable to heightened *leadership* attention, we should not see a similar increase in non-key decisions. The red dashed line in Figure 5 and the results in Table A6 in the appendix confirm this expectation. There is no significant change in non-key decision-making related to adult care following censure.

### 5.3 Centralization

Next, we investigate whether notification leads to centralization of control over the responsible subunit, using our dataset of internal audit plans. Figure 6 shows the

annual change in audit days allocated to adult care following the public interest notification. We observe a gradual increase in the year-by-year point estimates, with a significant rise in audit days in the second year after notification. This suggests that councils are deploying control resources to supervise adult social care departments, consistent with H3, although this effect is significant only at the 10% level – likely due to missing data.<sup>16</sup>

**Figure 6: Event study – centralized control in censured vs. uncensored councils**



**ALT-TEXT: Event study graph showing effect of notification on internal audit plans**

<sup>16</sup> Only 61% of council-years have internal audit plans available on a comparable and code-able basis. In the appendix, Table A7 provides the specific coefficient estimates.

## 5.4 Urgent investments in core competencies

Investing in core competencies is a further means of performance turnaround (H4). Following Boyne and Meier (2009), we operationalize this as changes to workforce composition. Table 4, Column 1 shows that censured councils grow their adult social care workforce considerably, compared with the control group. Columns 2 to 4, and Figure 8, present event study results by job groupings, revealing a pronounced increase in frontline roles. Specifically, Panel B of Table 4, Column 2 indicates that these roles increased by in the order of 17.58 posts in the first year after censure, rising to 30.47 in the second. By contrast, Columns 3 and 4 of Table 4 show no significant increases in supervisory or professional roles, with coefficients that are neither economically nor statistically significant at any point. Together, this indicates a targeted investment in public-facing (but unprofessionalized) roles, consistent with turnaround theories predicting a focus on core business.

**Table 4: Effects of low-stakes accountability on core competencies**

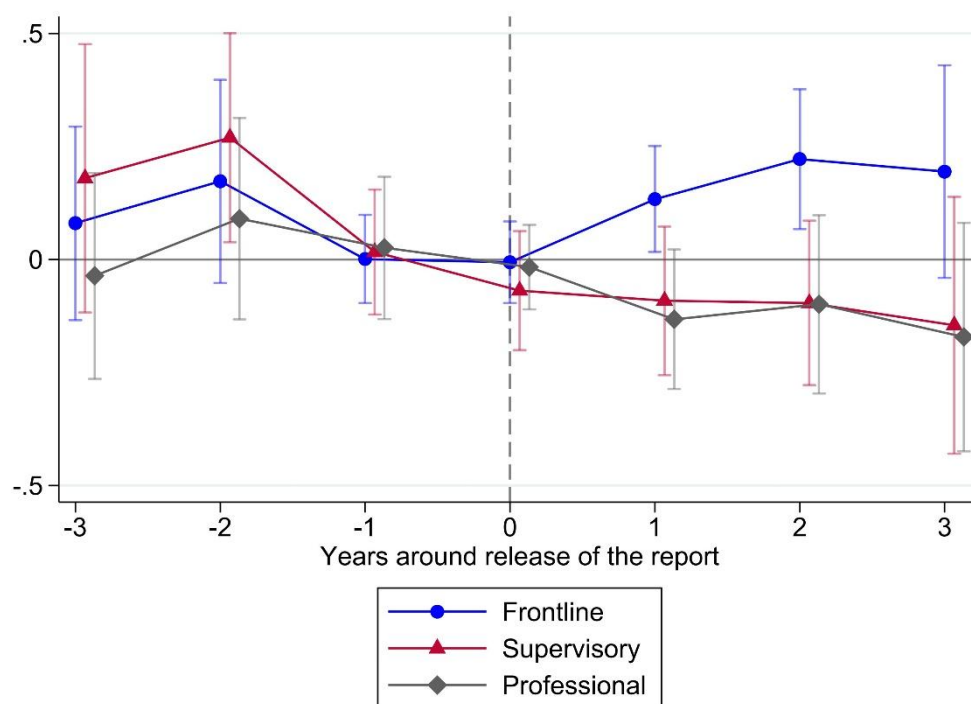
	(1)	(2)	(3)	(4)
	Overall workforce	Job categories		
		frontline	supervisory	professional
<b>Panel A</b>				
Post report releasing	0.032 (0.059) [0.582]	0.136** (0.063) [0.030]	-0.100 (0.082) [0.222]	-0.105 (0.077) [0.171]
<b>Panel B</b>				
Year 0	-0.012 (0.038) [0.744]	-0.006 (0.046) [0.896]	-0.069 (0.067) [0.307]	-0.017 (0.048) [0.724]
Year 1	0.030 (0.054) [0.577]	0.134** (0.060) [0.025]	-0.091 (0.084) [0.277]	-0.132* (0.079) [0.093]
Year 2	0.094 (0.074) [0.202]	0.222*** (0.079) [0.005]	-0.096 (0.093) [0.299]	-0.099 (0.101) [0.325]
Year 3	0.017 (0.104)	0.194 (0.120)	-0.146 (0.145)	-0.171 (0.129)

	[0.867]	[0.105]	[0.315]	[0.184]
Local authority FEs	✓	✓	✓	✓
Year FEs	✓	✓	✓	✓
Demographic control	✓	✓	✓	✓
Empirical method	CEM+DiD	CEM+DiD	CEM+DiD	CEM+DiD
Observations	1,010	998	1,009	1,003

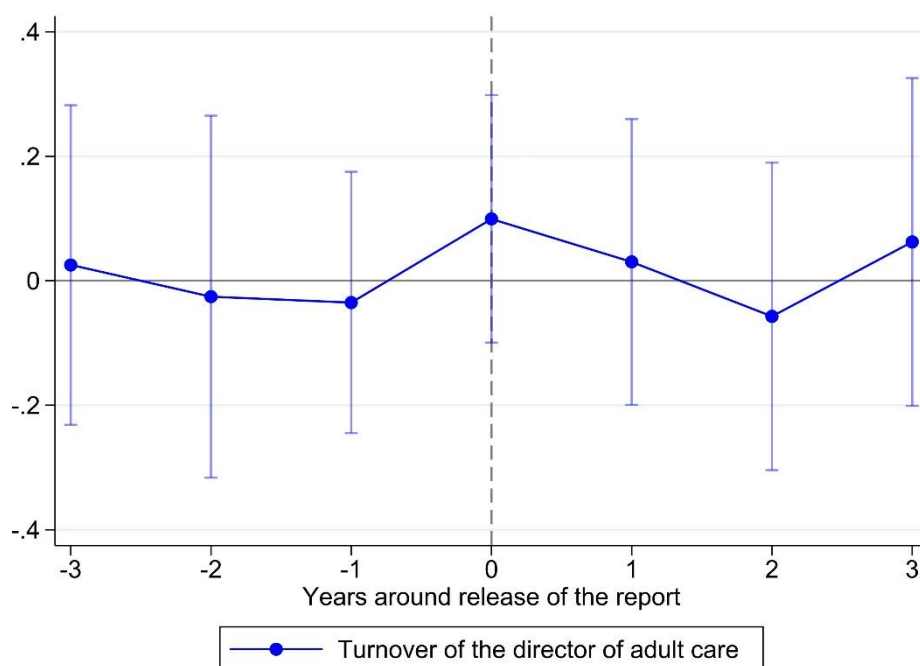
*Notes:* The unit of observation is the local authority × year. Post report releasing is a dummy that is 1 if local authority has received the report, and 0 otherwise. Post is a dummy that is 1 if the year is after the report releasing. Standard errors clustered at the local authority level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Standard errors are reported in parentheses, and p-values in square brackets.

We also examined whether notification influenced turnover in the director position responsible for adult social care within councils' executive teams. Figure 9 shows no discernible effect of censure on turnover, with a flat trend in the three years before and after the report was issued. As already indicated, this is consistent with Ombudsmen providing a low-stakes accountability regime, without threat to job security. But it also demonstrates that the improvement in performance that we observe is not explained by leadership "takeover."

**Figure 7: Event study – Core competencies in censured vs. uncensored councils**



**ALT-TEXT: Event study graph showing effect of notification on hiring practices**

**Figure 8: Event study – Turnover of director in censured vs. uncensured councils**

**ALT-TEXT: Event study graph showing effect of notification on leadership turnover**

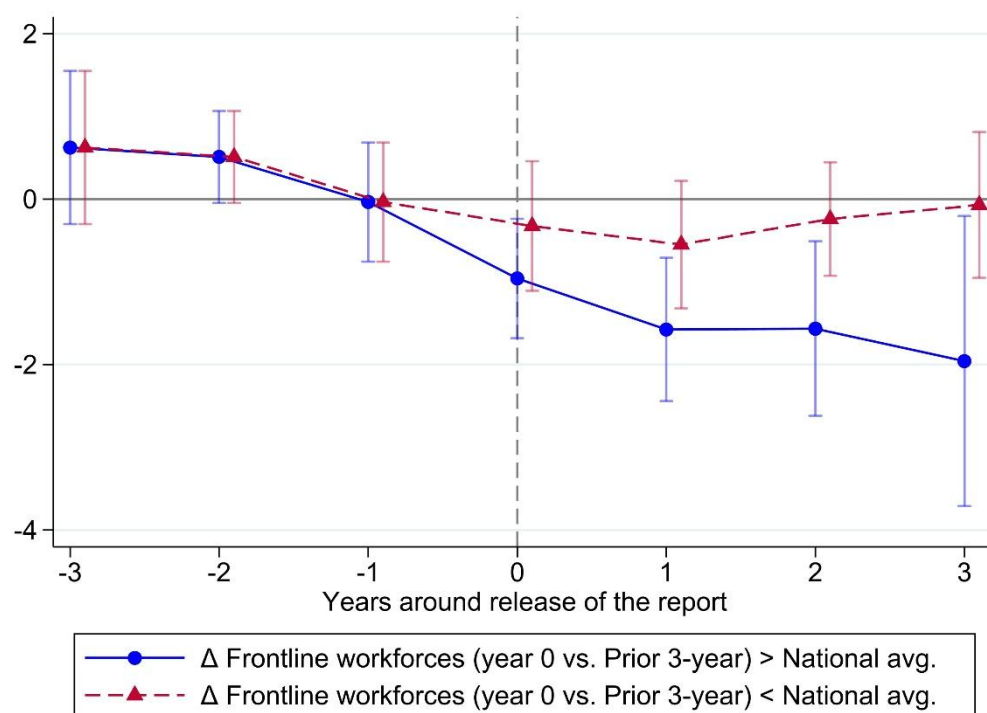
Finally, we examine whether *more rapid* investment in core competencies leads to greater performance improvements (H5). We compare the number of frontline roles in the censure year (year 0) with the average in the three preceding years. This comparison yields the immediate change rate in hiring during the year of censure. We then classify organizations whose change rate exceeds the national average into the “rapid-response” group, and those below this into the “delayed-response” group, and compare their performance in reducing maladministration.<sup>17</sup> The results, in

<sup>17</sup> In the appendix, Table A9 provides the specific coefficient estimates illustrating this effect.



Figure 9, indicate that prompt investments deliver significantly greater performance recovery, consistent with the turnaround adage that “time is of the essence” (H5). In the first year, the rapid-response group saw a more substantial reduction in upheld complaints compared to the delayed-response group. This trend continues over the next three years, with the faster-responding group demonstrating consistently greater decreases in upheld cases. As such, prompt re-allocation of resources appears to accelerate the turnaround process, leading to more significant and sustained performance gains over time.

**Figure 9: Event study - Effects of urgency on maladministration**



**ALT-TEXT: Event study graph showing effect of urgency on maladministration**

## 6. DISCUSSION

### 6.1. Accountability-induced turnarounds

We set out to establish whether accountability can induce turnaround in struggling organizations (or parts thereof), and, further, to see if feedback and standard setting alone can provide sufficient impetus for change. Our results provide consistent evidence that low-stakes accountability, without high-powered incentives, is able to trigger performance recovery in organizational subunits. This is consistent with recent theoretical work (Meier et al., 2015; Woo et al., 2015; Jakobsen et al., 2018; Andersen & Nielsen, 2020), and with legal analysis on the surprising degree of influence obtained by low-powered ombudsmen, including Hertogh's (2001) seminal comparison with administrative courts in the Netherlands. (Our analysis does not, of course, reveal whether high- or low-stakes accountability performs best in a head-to-head comparison. That will be for future research to judge in various scenarios.)<sup>18</sup>

Nevertheless, we were unable to detect any broader performance gains beyond maladministration. This reflects our expectation of some residual tunnel vision, even in low-stakes accountability regimes. Measurement encourages effort prioritization; and narrow feedback from specialist accountability bodies fosters restrictive rather than broad-based learning. Accountability-induced turnarounds thus appear to be *partial* turnarounds, limited by the remit and/or priorities of the account-holder. As Bovens et al. (2008, p. 228) write, "administrators ... get better in meeting the requirements posed by their accountability forums – but not necessarily ... in the real world of ... public service delivery." This suggests that accountability needs to be

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<sup>18</sup> At present, there are few direct comparisons of high- and low-stakes regimes. One example is Hanushek and Raymond's (2005) analysis of different states' enactment of No Child Left Behind legislation, in which low-stakes "report cards" are found to achieve minimal gains compared with more "consequential" systems involving monetary awards or takeover threats.

complemented by other turnaround strategies if it is to improve extra-jurisdictional dimensions.

Besides these main effects, our multidimensional dataset also affords valuable opportunities to test key turnaround behaviors. Our evidence confirms that council leaders significantly increase decision-making in respect of adult social care following the Ombudsman's notifications, consistent with the idea that turnaround needs a "triggering" event. Turnaround also typically involves exercising closer control over problem areas (O'Kane & Cunningham, 2014; Meier et al., 2015), and switching investments from ancillary to core activities (McKiernan, 2006; Boyne & Meier, 2009; Murphy, 2010). Again, we found greater internal audit days allocated toward the responsible subunit following notification, and a large increase in frontline staffing. Moreover, when these investments are implemented rapidly, the long-term improvement in performance is materially greater, consistent with several empirical studies of private sector recoveries (Tangpong et al., 2015; Barbero et al., 2017).

All that said, the behavioral changes we studied mostly appear to accompany rather than precede the main turnaround outcome. Maladministration is significantly reduced *within* the same year as the public interest notification is received, and the improvement peaks in the follow year. In contrast, measurable changes in leadership attention, centralization, and staff recruitment only emerge a year *after* the notification is given, and do not reach their respective peaks until two years later. This not only indicates that the "Ombudsman effect" is long lasting (and hardly that of an Ombuds*mouse*), but also that the mechanisms of turnaround may be more indirect than previously supposed – something we address in following subsection.

## 6.2 Theoretical implications

Based on these results, we are left, firstly, with the question of why the low-powered Ombudsman is able to induce such a strong reaction. There are several possibilities. The LGSCO has direct access to board-level decision-makers in councils, guaranteed in statute. Many of its investigators are “poachers turned game keepers,” drawn from the ranks of local government and so knowledgeable about the sector and its challenges. The non-coercive and low-stakes environment may encourage dialogue and reflection, rather than blame avoidance and self-preservation (Hertogh, 2001; Gill, 2018). But perhaps most notably, the LGSCO’s feedback to councils is entirely based on the lived experience of individuals who depend upon the organization’s services and have been somehow “let down” by it. As Thomas (2024) writes, “the LGSCO is a casework institution. It sees and approaches local government through individuals’ complaints and how councils’ failings cause people injustice.” Given the dilemma we outlined previously – that the more belief-inconsistent the account-holder’s feedback is, the more *credible* it must be – the rooting of the Ombudsman’s reports in real, personal injustices may offer a significant advantage when it comes to revealing the “gap” between aspirations and achievements necessary to trigger remedial action (Meier et al., 2015; Andersen & Nielsen, 2020). The potential galvanizing effects of this combination of lived experience and moral authority is a promising line of research.

Alternatively, while ombudsmen fail to meet Hibbard et al.’s (2003) four criteria for inflicting external reputational damage on account-givers, *internal* leader reputation among peers may also partly explain our results. Top council officials each oversee subunits that are subject to the Ombudsman’s jurisdiction (housing, highways, etc.); each know what a public interest notification signifies; and, by virtue of their seniority, are each aware every time a report is issued rebuking one of their peer’s

departments. To the extent that this causes embarrassment among colleagues, action may be taken to “save face.” These *internal* reputational effects should be explored in future research.

Finally, while we observe the whole series of changes in performance and behavior predicted by our theory, the temporal ordering is puzzling. Maladministration improves rapidly, whereas organizational changes are slower and more sustained. Our tentative interpretation is that, once councils are notified of failures, staff in responsible subunits proactively seek to recover performance *in anticipation* of the increase in leaders’ attention and control that will now be forthcoming. This is consistent with recent literature on felt accountability (Schillemans et al., 2021), with the Ombudsman’s notification materially increasing the subunit’s *expectation* of being called to account by council leaders. This departs from the conventional understanding in turnaround research that attention and centralization are direct mechanisms by which recovery is achieved, and so requires further quantitative and qualitative investigation.

### 6.3 Limitations

The principal limitation to our study is the extent to which the control group, matched with the treated group on observable characteristics at baseline, provides a convincing counterfactual throughout our study period. There may be unobserved factors that led to the Ombudsman selecting a council for a public interest notification that are not incorporated into our matching strategy. The Ombudsman may want to “ration” its use of these reports, in an effort to bolster their reputational impact. Moreover, some councils that narrowly avoided the notification treatment early in the period may have self-initiated improvements in their adult care services, rendering

themselves less comparable to later-treated councils. To address these challenges, we perform a suite of robustness tests to ascertain the stability of our results. We demonstrate the correlation between our matching covariates in 2012 and their averages from 2013 to 2019. And we focus exclusively on councils that received reports during the study period, thereby isolating the effect of the timing of the event's impact. These robustness checks demonstrate that our baseline empirical strategy, combining CEM and DiD, is effective.

In addition, because the treatment that we analyse above is the public interest notification submitted by the Local Government Ombudsman to a local government in respect of its adult social care provision, we do not take into account the precise nature or content of the complaint that led to the notification. And nor do we analyse different categories of complaint when assessing the effect of notification on council performance. Future research should seek to undertake more granular analysis.

## **7. CONCLUSION**

We theorized that low-stakes accountability would produce turnaround in under-performing local government services, including increased leadership attention to and control over problem areas, and investment in core competencies. Using panel data, coarsened exact matching, and staggered difference-in-differences, we provided strong evidence in support of these hypotheses, whilst also confirming the partial nature of accountability-induced turnarounds. Our study adds to the few existing quantitative evaluations of public service turnaround strategies, including Boyne and Meier's (2009), Rutherford (2014), and Alonso & Andrews's (2021), which is also quasi-experimental. It also confirms the presence of key turnaround behaviors observed in the corporate sector – like centralization and urgent investments – in the

government context, but suggests that some recovery occurs in anticipation of increased leader attention, rather than in consequence of it. And finally, our results challenge the consensus in political science that, for want of strong enforcement mechanisms, the public sector is unable to regulate *itself* (Wilson & Rachal, 1977; Konisky & Teodoro, 2016). Indeed, in spite of – or perhaps because of – its lack of enforcement powers, the ombudsman is remarkably effective at regulation “inside” government (Hood et al., 1999).

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