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Urbanisation and the political demography of African cities[☆]Nick Dorward^{a,*}, Kristian Hoelscher^b^a School of Geography and Environmental Science, University of Southampton, UK^b Peace Research Institute Oslo, Norway

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ABSTRACT

Africa is undergoing a rapid process of urban demographic change. Increasingly youthful population structures are defining the continent's towns and cities. Scholarship suggests this will be associated with greater protest incidence and lower levels of voting and electoral participation. However, these findings often rely on national-level data, despite there being considerable subnational variation in population structures between African cities. Building on existing theory, we argue that local demographic contexts matter for political behaviour. Specifically, we hypothesise that youthful urban demographic structures will be associated with lower levels of formal political participation (voting) and greater levels of informal contentious mobilisation (protest) for all individuals, and that the magnitude of this effect will be greater for younger people. We test these expectations using novel geospatial data on the spatial extent of unique urban settlements, urban-level age and sex structures, and geolocated individual-level survey data from 399 cities in 36 countries across Africa. Using multilevel regression, we find that individuals are more likely to vote in more youthful urban contexts, with young people no more or less likely to vote than their older counterparts. Conversely, we find no significant relationship between individual protest participation and city youth shares overall. However, young people in more youthful cities are significantly more likely to protest than older people. In light of these findings, we discuss how the demographic composition of individual cities in Africa nuances our understanding of political behaviour and contentious mobilisation.

1. Introduction

Africa is undergoing a rapid demographic transformation characterised by urbanisation, an increase in the share of people living in urban vs rural areas. Yet in contrast to the experiences of Europe and North America, African urbanisation has been largely driven by the natural increase (higher annual rates of births than deaths) of urban populations (Bocquier et al., 2023; Menashe-Oren & Bocquier, 2021). Rather than growing because of rural–urban migration, Africa's cities and emerging urban regions are growing from the in-situ growth and densification of populations (Fox & Goodfellow, 2022; Randolph & Storper, 2023; Dorward et al., 2023), leading to urban and national population structures that are increasingly defined by large youth cohorts, both in absolute and relative terms.

This broad demographic shift that encompasses larger numbers of people residing in urban areas across the continent is having important yet often poorly understood political effects on the continent (Cheeseman, 2022; Gutheil, 2022; Hoelscher et al., 2023). It has also

been of keen interest to political demographers for decades, with particular emphasis placed on the developmental and security implications of large urban youth cohorts (Goldstone, 2002; Urdal, 2006). Yet youthful urban populations also have implications for broader political participation, particularly in relation to how young people will be appealed to by political actors, or how they will engage with formal modes of politics (Gutheil, 2022; Hoelscher et al., 2023).

An extensive literature has since examined the nexus between youth and political participation and mobilisation in Africa, offering a relatively consistent theorisation of why age and demographic structure may matter for formal and informal modes of political behaviour. Broadly, large youth cohorts in urban areas are thought to face crowded labour markets and limited economic opportunities, shaping grievances toward formal political institutions that will: (i) decrease the likelihood of formal political participation through voting (Resnick & Casale, 2014), and (ii) increase the likelihood of informal political participation through protest mobilisation (Branch & Mampilly, 2015; Resnick & Casale, 2011). The implication is that as Africa urbanises and population

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* Corresponding author.

E-mail address: n.dorward@soton.ac.uk (N. Dorward).

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structures become younger; protest will increase while voting intention will wane.

Many of these studies examining the relationship between urbanisation, demographic structure, and political mobilisation and participation do so in one of two ways. One strand of literature has examined youthful demographic composition and in/formal political behaviour (e.g. voting/protest) primarily using national level data (Resnick & Casale, 2011, 2014; Thomson et al., 2021; Urdal & Hoelscher, 2012). A second body of literature has used individual level survey data to focus more exclusively upon the behaviours of individual respondents (Reichenberg & Tambe, 2022; Schmiedl & Liroy, 2024; Tambe & Kopacheva, 2023), often using country level data to explore variation in individual behaviour across national contexts. Very few studies have explicitly pursued a *meso*-level approach as we do, examining how the local demographic context of heterogeneous African cities is associated with individual political participation.

However, situating these studies at the city level is important for two key reasons. First, the theoretical mechanisms that link demographic structure and different forms of political participation generally relate to how individuals experience *local* – not national – economic opportunities, service delivery, and political responsiveness. Secondly, as we show in this paper, there is considerable subnational variation in the demographic structure of individual African cities, both within and between countries. As such, treating African cities as homogenous entities, even within the same country, is imprudent. In aiming to bring further nuance to the literature on the political demography of Africa's rapid urbanisation, this paper offers two key empirical and theoretical innovations.

First, we develop what to our knowledge is the first set of urban population structures at the city level across Africa. Spanning 2000–2020 for 399 African cities in 36 countries, we demonstrate that cities in Africa are not demographically homogenous places but are characterised by diverse population structures. Simply, to understand the political implications of Africa's urban transition, we must account for heterogeneous urban population structures and how they may shape individuals' political participation in specific urban places.

Second, in taking a place-based perspective on the politics of African urbanisation (Hoelscher et al., 2023), we examine how specific demographic characteristics in cities across Africa are associated with political participation and mobilisation *in the places in which this change is occurring*. Building upon recent advances in geospatial data analysis (Dorward & Fox, 2022), we combine fine-grained geospatial data on the spatial extent of unique urban settlements with estimates of urban age structures from WorldPop and geocoded individual-level survey data from three rounds of the Afrobarometer survey. This cross-sectional dataset allows us to measure both formal (voting) and informal (protests) political participation of 18,433 individuals in 399 cities in 36 African countries.

Our analysis fits a series of multilevel regression models with cross-level effects nesting individuals within cities within countries to estimate how the demographic contexts of unique urban settlements are associated with indicators of individual political attitudes and behaviour (voting and protest participation) at the city level. We find that individuals are more likely to participate formally in urban contexts characterised by higher youth shares. Furthermore, young people in more youthful cities are no more or less likely to participate formally than their older counterparts. Conversely, we find no significant relationship between individual protest participation and city youth shares. However, young people in more youthful cities are significantly more likely to protest than older people. Together, this offers a nuanced understanding of how demographic factors underpinning Africa's urban transformation are shaping political behaviour.

2. Political demography in African cities

Previous research on the political impacts of large youth cohorts has

largely focused upon the macro, primarily national, level, emphasising the potential for so called 'youth bulges' to fuel violence and insecurity (Goldstone, 2002, 2010; Urdal, 2006, 2008). Here, high youth shares create crowded labour markets and constrain socioeconomic opportunities leading to grievances against, and disaffection with, formal politics. This results in disengagement from associated institutions and processes (i.e., elections) and a predisposition towards contentious and even violent forms of mobilisation (e.g., protests, riots, and, in extremis, armed conflict). This is thought to be especially salient in urban areas characterised by rapid population growth and upward pressures on labour markets and key urban services including healthcare, education, and infrastructure (Buhaug & Urdal, 2013; Menashe-Oren, 2020; Urdal & Hoelscher, 2012).

Common to this literature is the assumption that in contexts characterised by high youth shares, young people are more likely to disengage from formal politics and consequently become more predisposed to informal modes of participation. However, such approaches use *national* level demographic data to infer how youth may experience *local* demographic conditions. Essentially, national level studies assume homogeneous demographic structures and labour markets across an entire country, when this is unlikely to be true. Moreover, this assumption risks an ecological fallacy: that large youth cohorts translate to youth disengagement and mobilisation because the demographic context to which they are exposed drives individuals to do so. However, to our knowledge, this assumption has not been tested. While young people are generally more inclined to protest (Branch & Mampilly, 2015; Honwana, 2015; Resnick, 2015; Resnick & Casale, 2014), it remains an open question whether their propensity to disengage from formal politics and engage in protest and violence is further amplified within the demographic context of cities characterised by large youth cohorts.

We argue that local demographic contexts, specifically large urban youth shares, will influence the political behaviour of individuals residing within particular cities. This is because the mechanisms linking demographic structure and political participation are fundamentally subnational, operating at the local job market and service delivery context of individual cities and urban regions. Furthermore, as we demonstrate below, there is significant subnational variation in urban youth shares across African cities, which national-level analyses frequently overlook. Central to our argument, and for understanding of how the demographic structure of a city may influence individual behaviour, is the distinction between compositional and contextual effects in the study of politics (Johnston & Pattie, 1998, 2006).

Composition effects arise when observed geographical differences in political behaviour can be explained by the individual characteristics of the people living within a given area. For example, longstanding evidence demonstrates that young people are significantly less inclined to vote and display greater levels of dissatisfaction with formal political institutions than their older counterparts (Franklin, 2004; Henn & Foard, 2012; Lekalake & Gyimah-Boadi, 2016; Nie et al., 1974; Nkansah, 2022; Nkansah & Bartha, 2023; Nkansah & Papp, 2023; Norris, 2002; Resnick & Casale, 2011, 2014; Tambe & Kopacheva, 2023). We would, therefore, expect a city with a higher proportion of young people to exhibit lower overall voting rates and more informal participation, simply because of its demographic composition. Compositional effects imply that similar people will behave similarly regardless of their geographical setting and that all geographical differences in behaviour are driven by the clustering of particular types of people in specific places (Johnston & Pattie, 2006).

Contextual effects, on the other hand, operate when the characteristics of an individual's local environment itself influences their behaviour, independent of individual attributes. Geographical differences in political behaviour are not solely attributable to aggregate social or demographic characteristics of people within an area, but the characteristics of an area itself (Johnston & Pattie, 1998, 2006). This implies that individuals with similar demographic profiles are likely to behave quite differently depending upon their geographical setting. The key

mechanisms linking local context to political behaviour are social interaction and contact (Johnston & Pattie, 2006; Nathan & Sands, 2023). Local urban environments and the arrangement of individuals within them affect the types of interactions people have, conditioning both the breadth and depth of social contact and fundamentally shaping the political beliefs and behaviours of individuals (Enos, 2017; Nathan, 2019; Nathan & Sands, 2023). Our aim is to understand how living in a city with high youth shares will influence individual political behaviour (context), in addition to how political behaviour is shaped by an individual's own age (composition).

2.1. Individual level effects of city youth shares

The mechanisms linking both *compositional* and *contextual* effects to individual political behaviour are typically framed in terms of the motives, means, and opportunity for distinct types of political action. We take these in turn, discussing the specific mechanisms within each general category and the links to distinct modes of political behaviour – both formal and informal.

2.1.1. Motive – Local economic opportunities, grievances, and dissatisfaction

Large youth cohorts in urban contexts can shape the social and economic environment in which individual political motivations are formed. In cities characterised by high youth shares, individuals, especially youths, are more likely to encounter competitive, overcrowded labour markets leading to higher rates of unemployment, underemployment, or employment in marginal, informal jobs with poor pay and working conditions (Apolte & Gerling, 2018; Bricker & Foley, 2013; Resnick & Casale, 2014). Large youth shares are also likely to result in upward pressures on key urban services including healthcare, education, and other infrastructure (Fetzek & Vivekananda, 2015; Menashe-Oren, 2020; Urdal & Hoelscher, 2009). This environment is likely to amplify dissatisfaction, fostering grievances directed towards formal political institutions or other social groups (Weber, 2019). Such contexts are likely to suppress formal forms of political participation such as voting in national and local elections, while creating an environment conducive to informal modes of expression including participation in protests or violent demonstrations (Akinyetun, 2023; Resnick & Casale, 2014).

Furthermore, the literature suggests that a combination of expanding youth cohorts and increased access to post-primary education can contribute to heightened political violence, even in context of low youth unemployment (Weber, 2019). The resulting gap between educational attainment and available economic opportunities can generate significant frustrations among the youth, serving as a potential motivation for political apathy, disconnection, and disengagement (Nkansah, 2025; Nkansah & Papp, 2023). This is not merely a point about the educational composition of urban youths in African cities but the broader economic climate to which they are exposed. Economic opportunities, or lack thereof are fundamentally local, operating at the city level rather than nationally. The economic productivity of a city and its capacity to provide employment for a growing young population, are, therefore, crucial contextual factors influencing individual grievances.

2.1.2. Means – participation costs and resource constraints

The demographic composition of a city can also influence the resources available to individuals and perceived costs associated with political participation. Large youth cohorts are typically comprised by individuals with comparatively limited social or economic responsibilities, i.e., unmarried, non-primary carers who are for aforementioned reasons more likely to be informally employed if at all (Nkansah, 2022). These individuals have lower opportunity costs vis-à-vis engaging in informal political activities that carry greater risk, such as protests and violent demonstrations (Urdal, 2006). Moreover, younger people without good jobs, familial obligations, or caring responsibilities will also have less stake in, and therefore fewer incentives

to participate in, formal politics. The net result of these reduced costs is to make participation in informal collective action more appealing in cities where there are high concentrations of young people (Hartmann & Biira, 2021; Mueller, 2018; Tambe & Kopacheva, 2023).

Conversely, resource-based theories of political participation present a counterargument. Participation in political processes is costly to individuals in terms of time, financial resources, and the accumulation of political knowledge (Brady et al., 1995; Nkansah, 2022). Young people, therefore, lacking employment opportunities, reliable incomes, and facing significant economic challenges, might prioritise addressing their immediate material needs over political participation, particularly in socioeconomic contexts defined by large urban youth cohorts (Nkansah & Papp, 2023). This could lead to disengagement from political processes altogether, yet less pronounced for informal participation which requires less time, fewer financial resources, and relatively little political knowledge.

2.1.3. The opportunity mechanism – peer effects, social networks, and anonymity

Cities are uniquely conducive to political mobilisation and collective action (Dorward & Fox, 2022; Fox & Bell, 2016). The urban environment, especially when characterised by large youth cohorts, offer unique opportunities for individual engagement and mobilisation. Young people, in general, are more idealistic, are less likely to have significant responsibilities, and typically engage in riskier behaviour than older groups (Goldstone, 2002; Steinberg, 2004; Weber, 2013). Large youth cohorts thus provide a plentiful pool of potential recruits for political actors and entrepreneurs to mobilise in pursuit of their political aims (Urdal, 2006, 2012). However, urban youth can represent a considerable pool of peers, which can significantly amplify peer influence. Young people in communities with a high concentration of their age group tend to interact more frequently, shaping political identities, facilitating information sharing, and mobilising for participation (Nkansah & Papp, 2023; Weber, 2013). On the one hand, these peer effects can motivate political interest and participation. On the other, they may also make individuals more susceptible to anti-establishment or anti-democratic ideologies, giving them a predisposition for informal and contentious modes of political expression (Cincotta & Weber, 2021; Nkansah & Papp, 2023).

Importantly, cities make collective action easier. Higher population densities in urban areas facilitate connectivity between people and foster the formation of dense social networks. These networks are crucial for the exchange of information vital to coordinating protest activities (e.g. transportation, turnout, police presence, and legal access) (Jost et al., 2018). There is also growing evidence that social media platforms, which have higher user rates in urban areas and among youths (Conroy-Krutz et al., 2024; Mast et al., 2025), amplify the coordination and mobilisation capacities of urban social movements (Eltantawy & Wiest, 2007; Hoelscher et al., 2023; Mutsvauro, 2016). Access to social media might also support grievance formation and withdrawal from formal politics through the sharing of emotional, motivational, and sometimes misleading content, such as expressions of anger, social identification, and highlighting concerns about fairness, justice, and deprivation (Boxell & Steinert-Threlkeld, 2022; Garbe et al., 2023; Jost et al., 2018).

We have argued that the demographic context of large youth cohorts might plausibly lead to the fostering grievances, dissatisfaction, and scepticism of formal politics that undermine personal efficacy and reduce formal participation. However, the presence of a large youth cohorts in a city can also foster a heightened sense of collective efficacy, even when there may be scepticism of formal political institutions. Observing large numbers of people with similar demographic profiles willing to engage in coordinated collective action, such as protests, can reinforce the belief that collective action is effective, and increase willingness to participate (McDonnell, 2020). This is particularly relevant for protest mobilisation which, by design, is a more visible form of participation than voting. Finally, the demographic context of large

youth cohorts also provides young people with a greater degree of anonymity. When individuals are part of a large group sharing a similar demographic profile, there is a lower probability of detection associated with protest and informal mobilisation (Dorward & Fox, 2022; Fox & Bell, 2016). This ‘safety in numbers’ effect can lower the psychological costs to participation, making individuals more inclined to join collective action.

2.2. Hypotheses

Our review of the literature on youth and political participation suggests that intensified competition for jobs and services in densely populated, youthful cities generates grievances and dissatisfaction, suppressing formal political participation and predisposing individuals toward informal actions. Furthermore, lower opportunity costs for informal activities among young people in these contexts make informal protest more appealing than formal political engagement. Moreover, urban density and high social media penetration foster dense social networks, amplify peer effects, and offer a degree of anonymity, all of which likely facilitate informal collective action and reinforce collective efficacy. This set of local-level mechanisms help explain why individuals, particularly young people, in cities with higher youth shares may be significantly more likely to engage in informal politics while simultaneously disengaging from formal political activity. This leads us to develop the following set of hypotheses:

- *H1a. Individuals in cities with higher youth shares will be less likely to participate in formal politics*
- *H1b. Individuals in cities with higher youth shares will be more likely to participate in informal politics*

We further anticipate that the association between urban youth shares and political participation will be stronger for young people, such that:

- *H2a. Young people in cities with higher youth shares will be less likely to participate in formal politics than their older counterparts.*
- *H2b. Young people in cities with higher youth shares will be more likely to participate in informal politics than their older counterparts.*

3. Data and methods

We test these expectations with an empirical analysis of individual political participation in 399 distinct African cities. We constructed a dataset that matches individual respondents in the Afrobarometer Survey to individual cities in the Urban Centres Database (UCDB) (Melchiorri et al., 2024) and high-resolution subnational estimates of population age and sex structures from WorldPop (Pezzulo et al., 2017; Tatem et al., 2013). This unique city-level dataset allows us to analyse the effect of a city’s demographic context on individual rates of formal and informal political participation.

3.1. Identifying urban areas

To analyse the relationship between political participation and local urban demographic contexts, we need a consistent method for situating individuals within specific urban areas. While Afrobarometer data provides a rural–urban classifications, the definitions underpinning this classification vary across countries and lack the spatial precision required for subnational analysis. We therefore use the Urban Centres Database (UCDB) (Melchiorri et al., 2024), which applies a globally standardised geospatial definition of urban areas based on the DEGURBA methodology endorsed by the UN (Dijkstra et al., 2021). This defines urban centres as contiguous 1 km² grid cells with at least 1,500 people and a minimum total population of 50,000. We overlay these urban polygons onto georeferenced Afrobarometer respondent data,

linking individuals to specific urban areas and measure contextual characteristics accordingly.

This geospatial approach offers several advantages. First, it ensures cross-national consistency by avoiding reliance on country-specific definitions of ‘urban’, which are often inconsistent or outdated (Dorward et al., 2023). Second, it allows us to align the resolution at which we are measuring demographic context with the scale at which contextual mechanisms operate an impact individual behaviour. Third, it enables inclusion of a broader range of cities, including secondary and smaller urban centres, beyond the handful of major metropolitan areas typically analysed in previous research. This improves the representativeness of the urban sample and, given the rapid population growth and political significance of these emerging towns and cities, captures more variation in demographic conditions across the urban spectrum (Koechlin & Förster, 2018; Kumar & Stenberg, 2023; Zimmer et al., 2020).

There are, however, limitations to this approach. In particular, the UCDB’s relatively high population density thresholds may under-represent urban areas in Africa, where urban growth often occurs at lower densities and in more diffuse settlement patterns (Blei & Angel, 2021; Dorward et al., 2023). As a result, our analysis likely captures more established and densely populated urban centres, potentially excluding newer or *peri*-urban areas where different patterns of political behaviour may emerge. This introduces a conservative bias, meaning our estimates reflect stronger contextual effects in more clearly defined urban environments, but may not generalise to all urbanising areas (see Appendix A in the [supplementary material](#) for detailed discussion of the geospatial data and approach).

3.2. Dependent variables

Our dependent variables are measures of political participation at the individual level. These data are taken from the Afrobarometer geocoded survey Rounds 5 (2011–13), 6 (2015–16) and 7 (2018–20). Afrobarometer is a nationally representative survey conducted in a large sample of African countries through a standardised questionnaire implemented through face-to-face interviews. Individual respondents are selected via full probability sampling of the national population over the age of 18, with a minimum of 1200 respondents per country. We adopt two primary measures capturing an individual’s self-reported levels of formal and informal political participation.

First, formal political participation is primarily measured by whether an individual reports voting in the most recent national or local election. This is operationalised as a dichotomous variable, taking a value of one if an individual voted and zero if they did not. Second, to measure informal political participation, we leverage a question asking respondents to report whether they have attended a demonstration or protest march in the last year. The original response scale for this question ranges from zero (“no, would never do this”) to four (“yes, often”). For our analysis, we have collapsed these responses into a dichotomous variable, taking the value of one if an individual reports having attended a protest in the last year, and zero otherwise.

3.3. Independent variables: Youth and demographic context

We have two main independent variables: the age of an individual respondent and the demographic context of the individual city in which that person lives. Specifically, we examine how the share of young people in a city’s population is associated with individual political behaviour. To measure the age of an individual respondent we created a categorical variable from Afrobarometer’s continuous age attribute. The definition of youth varies substantially depending upon context. The United Nations defines the term ‘youth’ as anyone between the ages of 15–24 years whereas the African Union uses the 15–35 range, acknowledging the extended period of transitioning to adulthood in many African contexts, which can include prolonged education, delayed

employment, and marriage (Honwana, 2014). Following Resnick and Casale (2011) we categorise individuals as: 'young adults' (18–29), 'mid adults' (30–49), and 'older adults' (50 + years). This categorisation is consistent with past literature as well as both the specific conditions of disrupted youth transitions in the African context and wider demographic studies.

To measure local demographic context, we use subnational population estimates from WorldPop's age and sex structure data (Pezzulo et al., 2017; Tatem et al., 2013). These estimates are provided at a 1 km² resolution in raster files disaggregated by year, sex, and age group. We use the unconstrained version of the WorldPop data, which – unlike the constrained version – does not rely on satellite imagery of human settlements and buildings to allocate population. While other gridded population datasets exist, none provide the granular demographic detail required for our analysis. WorldPop remains the best option and has been shown to outperform alternatives like the Global Human Settlement Population Grid and the Gridded Population of the World in estimating subnational populations. Notably, it offers higher accuracy and lower error rates across a variety of countries and income levels (Kuffer et al., 2022; Yin et al., 2021), and performs particularly well in urban areas (Kuffer et al., 2022) and across the city-size distribution (Statham, 2024).

To construct our dataset, we developed a custom program to: (i) automatically download each raster file; (ii) extract pixel-level population values for each year, age, and sex category; (iii) aggregate these values to UCDB city polygons; (iv) merge all files, and (v) calculate zonal statistics to describe city-level demographic characteristics. For both males and females, we calculated total population, total working-age/adult population (ages 15–70), and total youth population (ages 15–29). From these, we computed male and female youth shares as proportions of the total population and of the working-age population. The result is a cross-sectional time-series panel dataset providing total population estimates and detailed demographic pyramids for each city from 2000 to 2020. To our knowledge, this represents the first published use of WorldPop's urban age-sex structures in such a panel format. Given the potential relevance of this data for demographic and development research, we provide a detailed description of the data processing steps in Appendix B.

It is worth noting that descriptive analysis of the dataset reveals a high degree of variation in demographic structures and youth shares both within and between countries (Fig. 1). Importantly, the data support our contention that urban demographic structures in Africa are heterogeneous, showing that national level figures can hide significant subnational variation between cities within regions and countries. This implies markedly different demographic contexts across the spectrum of Africa's urban settlements, which we expect to be associated with local variation in individual political behaviours between cities.

3.4. Controls

At the city level, we control for the total size of the urban population, population density, and population growth. The total population is the sum across all the age-sex bins in the WorldPop data. Population density is this figure divided by the city's total area. Density is likely to be a strong determinant of political participation. However, it is difficult to say in what direction. Elementary economic geography would suggest that density would yield agglomeration effects which should stimulate economic opportunities and improve efficiency in service delivery, promoting formal models of political participation whilst limiting incentives for demonstrations and protest (Glaeser, 2011). However, density may also present distinct challenges including disease burden, crime, congestion pressure on urban resources that represent acute governance issues and stimulate dissatisfaction leading to suppressed formal participation and increased protests (ibid).

The local economic context may also affect both political participation and the demographic structure of a city. More prosperous cities

should see more formal participation and less informal participation because wealthier urban residents, all else equal, should have a greater stake in formal political processes and be less dissatisfied with these processes. Higher economic output also implies higher government revenues, resources that can be diverted into the provision of public goods and services. However, these cities may also have younger population structures given that young people are likely to migrate to prosperous areas in search of jobs.

Local economic output is often measured by Night-time lights data which has been shown to be a reliable proxy in certain applications (Gibson et al., 2020, 2021; Weidmann & Theunissen, 2021). We use a new, high-resolution dataset that estimates local real GDP from night-time emissions (Chen et al., 2022). The dataset calibrates the DMSP/OLS and NPP/VIIRS night-time light series and re-calculated revised measures of local economic output for a global 1 km² grid between 1992–2019. Since our study period extends beyond the end of the local GDP series, we use linear interpolation to fill the missing data from 2020 to 2022. We also log transform this variable to account for the left-skewed distribution.

At the individual level, we also control for a respondent's education, employment status, and experienced poverty taken from Afrobarometer. The education variable accounts for an individual's level of education ranging from 0, 'no formal schooling', 9, indicating that an individual holds a post-graduate qualification. Employment status is a binary variable indicating whether an individual has a job yielding a cash income (0 is no and 1 is yes). Experienced poverty is a composite index measuring the extent to which an individual reported going without one of the following essentials in the month prior to the interview: food, water, medicine, cooking fuel, and a cash income.¹

3.5. Sample and model estimation

Our sample includes all urban residents in cities with populations greater than 50,000 persons, where there is a minimum of three respondents per city. This leaves us with a final sample of 11,831 individuals nested within 354 cities, across 32 countries across Rounds 5, 6, and 7 of Afrobarometer (2011–2020).²

Given that both our dependent variables are binary measures of political participation, we test our hypotheses using a series of Bayesian multilevel logistic regression models. The data have a strict nesting structure – individuals within cities and cities within countries – making multilevel models the most appropriate and statistically robust method, as they inherently account for the dependencies among observations within groups (Gelman & Hill, 2007; Hox et al., 2017; McElreath, 2018).

We specify a random intercept for both cities and countries. This design captures the hierarchical nature of the data, allowing us to estimate the average effect of predictors at both the individual- and city-level while acknowledging city- and country-specific heterogeneity in the baseline level of political participation (Hox et al., 2017). The inclusion of random intercepts best reflects our prior knowledge of the data generating process, while explicitly accounting for unobserved heterogeneity at both the city and country levels, which would

¹ The index was computed using inverse distance weighting and the components showing a high degree of internal consistency with a coefficient alpha of $\alpha = 0.79$.

² All cities with a population less than 50,000 people have been dropped as these are settlements for which measurement errors in the UCDB and WorldPop data are likely to be higher (Blei and Angel, 2021). Furthermore, all cities with fewer than three responses have been dropped from the sample ensuring that there is enough within city information for model convergence. The countries covered are Algeria, Cameroon, Morocco, Nigeria, South Africa, Tanzania, Uganda, Zambia, Ghana, Kenya, Mozambique, Burkina Faso, Burundi, Niger, Togo, Madagascar, Benin, Mali, Tunisia, Guinea, Sudan, Côte d'Ivoire, Sierra Leone, Zimbabwe, Botswana, Lesotho, Namibia, Liberia, and Gabon.

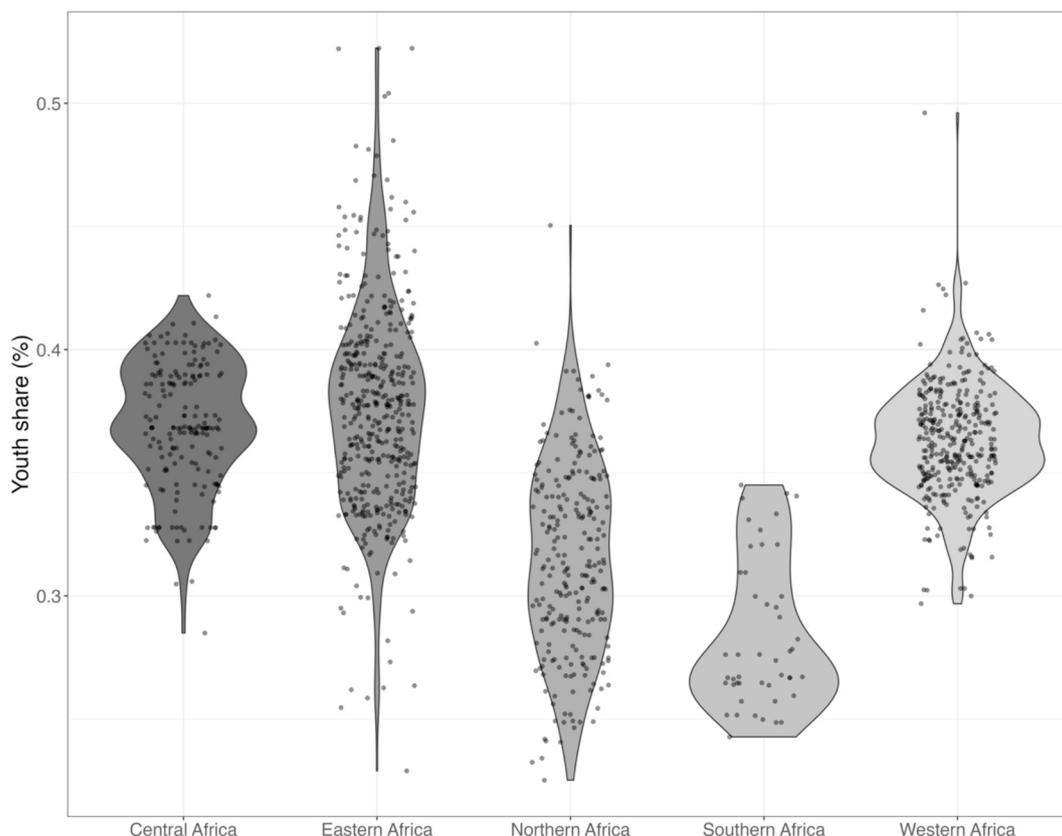


Fig. 1. City youth shares by region. Note: The y-axis displays the share of youth people in a city’s population in percentage terms. Each point represents the average for a unique city between 2000–2020. The shape around the points of each region describes the density of the points at that point on the y-axis.

otherwise lead to biased uncertainty estimates. The choice of random intercepts is also essential for facilitating the estimation of cross-level effects and interactions, which are central to our hypotheses regarding how city youth shares influence the political participation of individuals and of individuals within distinct age groups differently. Multilevel models are specifically designed to analyse the effects of higher-level variables on individual outcomes, inherently addressing within-group (city and country) correlations by modelling variance at each level.

Furthermore, the multilevel structure, particularly the city-level random intercepts, helps to absorb unobserved, time-invariant city characteristics that might be correlated with both youth shares and political participation. This implicitly accounts for some of these confounding factors that remain constant over the relatively short observation period. Similarly, country-level random intercepts account for unobserved national factors such as shared language, history, culture, and institutional conditions, as well as possible cross-national differences in political behaviour during the survey enumeration period (e.g., the impact of ongoing elections in certain countries). This allows the models to implicitly account for this heterogeneity.

Our models thus allow the relationship between each independent variable and the outcome to vary systematically across individuals within cities and cities within countries. This is crucial because the functional relationship between local demographic context and political outcomes may vary depending upon national and subnational contextual factors. The model has the form:

$$\text{logit}(p_{ijk}) = \beta_0 + \beta_1 (\text{Youthshare}_{e_{jk}}) + \beta_2 (\text{Age}_{e_{ijk}}) + \beta_3 (\text{Youthshare}_{e_{jk}} \times \text{Age}_{e_{ijk}}) + \beta X_{ijk} + \beta X_{jk} + \sum_{r=3}^R \gamma(\text{round}_r) + u_{jk} + \epsilon_{ijk}$$

where p_{ijk} is the probability of participation (formal or informal) for an individual, i , in city, j , in country, k . β_0 is the baseline probability of

participation across all individuals when all other variables are zero. $\beta_1 (\text{Youthshare}_{e_{jk}})$ represents the effect of the city's youth share on the outcome and $\beta_2 (\text{Age}_{e_{ijk}})$ gives the effect of an individual's age on the outcome. $\beta_3 (\text{Youthshare}_{e_{jk}} \times \text{Age}_{e_{ijk}})$ is the interaction term, showing whether the effect of city youth share on political participation differs across age groups. βX_{ijk} and βX_{jk} are vectors of individual and country-level control respectively while $\sum_{r=3}^R \gamma(\text{round}_r)$ is the fixed effect accounting for heterogeneity in political behaviour across the distinct survey rounds. u_{jk} is the country random effect, accounting for unobserved heterogeneity at the country level, this is assumed to follow a normal distribution such that $u_{jk} \sim N(0, \sigma_{jk}^2)$. ϵ_{ijk} is the individual error term.³

All models were estimated using 10,000 Markov Chain Monte Carlo (MCMC) iterations, with the first 2,500 iterations discarded as warm-up (burn-in). We employed flat (non-informative) priors for all parameters, allowing the posterior distributions to be predominantly shaped by the data (McElreath, 2020). The Bayesian framework differs from frequentist methods in that it directly estimates full posterior distributions for all model parameters, quantifying uncertainty while accounting for the hierarchical structure of the data. Model results are presented in the usual way with point estimates and uncertainty intervals.

The model coefficients presented are Bayesian point estimates (i.e.,

³ The random effects model implicitly assumes no correlation between the individual level regressors and the random intercept. A violation of this assumption would bias our interpretation of those parameters. In Appendix E we conduct robustness checks explicitly model this assumption using a Correlated Random Effects model. The results of these models are consistent with our main analysis.

Table 1
Age, demographic context, and political participation.

	Dependent variables:					
	Voted			Attended Protest		
	(1)	(2)	(3)	(4)	(5)	(6)
Youth share (%)	1.01 (0.92 – 1.11)	1.01 (0.92 – 1.11)	1.03 (0.93 – 1.14)	0.80 (0.71 – 0.89)	0.78 (0.70 – 0.87)	0.84 (0.75 – 0.95)
Total population (ln)	1.04 (0.90 – 1.20)	1.04 (0.90 – 1.20)	1.03 (0.90 – 1.20)	1.16 (0.98 – 1.39)	1.13 (0.96 – 1.35)	1.13 (0.95 – 1.35)
Population density	0.95 (0.83 – 1.09)	0.95 (0.82 – 1.09)	0.95 (0.83 – 1.09)	0.82 (0.69 – 0.97)	0.85 (0.71 – 1.00)	0.85 (0.72 – 1.00)
Local GDP (ln)	0.83 (0.75 – 0.93)	0.82 (0.73 – 0.92)	0.82 (0.73 – 0.92)	0.92 (0.80 – 1.03)	0.91 (0.80 – 1.03)	0.91 (0.80 – 1.03)
Mid-age adults	1.78 (1.65 – 1.92)	1.78 (1.65 – 1.93)	1.78 (1.65 – 1.93)	0.86 (0.78 – 0.95)	0.89 (0.80 – 0.99)	0.89 (0.81 – 0.99)
Older adults	2.57 (2.30 – 2.90)	2.64 (2.35 – 2.97)	2.64 (2.36 – 2.97)	0.54 (0.46 – 0.62)	0.60 (0.51 – 0.69)	0.59 (0.50 – 0.69)
Gender	0.88 (0.82 – 0.94)	0.90 (0.84 – 0.97)	0.90 (0.84 – 0.97)	0.61 (0.56 – 0.67)	0.64 (0.58 – 0.71)	0.64 (0.58 – 0.70)
Round 6	0.63 (0.57 – 0.70)	0.62 (0.56 – 0.69)	0.62 (0.56 – 0.69)	1.22 (1.07 – 1.39)	1.26 (1.11 – 1.44)	1.26 (1.11 – 1.44)
Round 7	0.59 (0.53 – 0.66)	0.59 (0.53 – 0.66)	0.59 (0.53 – 0.66)	1.39 (1.22 – 1.59)	1.34 (1.17 – 1.53)	1.33 (1.17 – 1.52)
Education		1.02 (1.00 – 1.04)	1.02 (1.00 – 1.04)		1.14 (1.11 – 1.17)	1.14 (1.11 – 1.16)
Experience poverty (ln)		0.87 (0.79 – 0.96)	0.87 (0.80 – 0.95)		1.69 (1.50 – 1.90)	1.70 (1.51 – 1.91)
Employment Status		1.08 (1.00 – 1.17)	1.09 (1.00 – 1.18)		0.91 (0.82 – 1.00)	0.91 (0.82 – 1.00)
Youth share x Mid adults			0.98 (0.90 – 1.06)			0.89 (0.81 – 0.98)
Youth share x Older adults			0.97 (0.87 – 1.08)			0.86 (0.75 – 0.99)
Observations	18,321	18,321	18,321	18,321	18,321	18,321
Marginal R ² / Conditional R ²	0.036 / 0.095	0.038 / 0.096	0.038 / 0.096	0.017 / 0.049	0.027 / 0.060	0.027 / 0.060

the mean of the posterior distribution) of the odds ratios. An odds ratio (OR) more/less than one indicates that an outcome is more/less likely given the level of exposure to an independent variable. 95% credibility intervals representing uncertainty around point estimates are presented in parenthesis. In the context of odds ratios, credibility intervals crossing one indicate a statistically significant association in the traditional, frequentist sense. Unlike frequentist confidence intervals, 95% credibility intervals directly imply a 95% probability that the true parameter value lies within that range.

Full posterior distributions and convergence diagnostics are presented in Appendix D. All variables have been mean-centred and rescaled to enable model convergence and comparability of coefficients. Individual-level variables have been group-mean-centred (Bell et al., 2018). All analysis was undertaken in R Version 2025.05.0 + 496. Multilevel models were estimated using the lme4 (Version 1.1–37) (Bates et al., 2015) and brms (Version 2.22.0) (Bürkner, 2017) packages. Model diagnostics and trace plots indicate stable, well-mixed Markov chains that converge to the target posterior distribution with no divergent transitions (see Appendix D).

4. Empirical findings

Table 1 presents the results of the Bayesian multilevel logistic regression models. Models 1–3 present results for voting behaviour whilst Models 4–6 present result for protest participation. Model 1 reports the baseline association between individual age and city-level youth shares with respect to individual self-reported voting behaviour.

The results in Model 1 indicate that individuals living in cities with high youth shares are no more or less likely to vote than individuals in less youthful urban contexts. Conversely, the coefficients for mid-age and older adults suggest that individuals both categories are significantly more likely to vote than their younger counterparts. The odds of voting for mid-aged adults in the sample are 78% higher than for young

adults while the odds for older adults are 157% higher. Fig. 2 visualises the marginal effects of age group of voting showing that older adults are more likely again to vote than mid-aged adults.

Model 2 includes education, experienced poverty, and employment status as further controls. These were excluded from Model 1 as they may be determined, at least in part, by both the age of an individual and the age structure of the city in which they live and are therefore post-treatment variables. However, there are strong theoretical reasons for including them as controls since they represent fundamental individual-level resources and grievances that could directly influence political participation, irrespective of an individual's specific age or the city's overall youth share. Excluding them would lead to omitted variable bias, as they are well-established determinants of political behaviour. That said, results remain consistent across models with and without these controls, suggesting that the observed associations are not biased by their inclusion.

Models 4 and 5 report results for individual protest participation. Both models indicate a statistically significant, negative link between city youth shares and protest participation, with each percentage point increase in the city youth share associated with a 78–80% decrease in the odds of protest participation for all individuals (see Fig. 3). Similarly, the coefficient for mid-age and older adults also reveals strong negative associations between age and protest behaviour (Fig. 2). The odds of mid-age adults reporting having attended a protest are 11–14% lower than for young adults whilst the odds for older adults protesting are 40–46% lower.

With respect to our hypotheses, these models do not provide support for H1a. The observed odds ratios for youth share with respect to voting are not statistically significant and are close in magnitude to one, indicating no discernible effect of youthful urban demographic structure on individual voting behaviour. Furthermore, we do not find support for H1b. The hypothesis posits that individuals in cities with higher youth shares will be more likely to participate in informal politics (protesting).

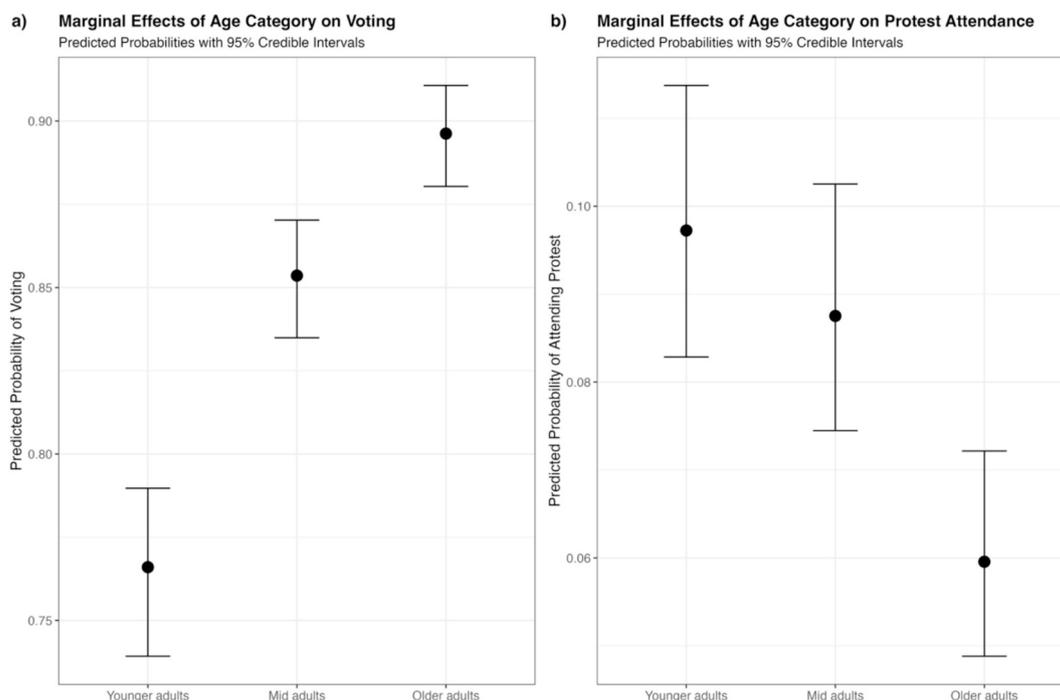


Fig. 2. Marginal effect of age on voting and protest participation. Note: Marginal effects of individual age with respect to the predicted probability of (a) voting and (b) protest participation. Estimates are taken from Models 3 and 6. 95% credibility intervals are reported around Bayesian point estimates.

In contrast, the results consistently show a statistically significant *negative* relationship, indicating that higher youth shares are associated with a *lower* likelihood of protest attendance. This is notable as it is in opposition to what much of the literature would suggest and directly contrasts our own H1b.

Columns 3 and 6 present the results of models featuring cross-level interaction terms between city youth shares and individual age. The interaction terms Youth share x Mid-age adults and Youth share x Older adults in both models are designed to capture whether the effect of city youth shares on individual voting and protest participation are different for younger adults than for mid-age and older adults, as their political behaviour is theoretically more likely to be impacted by such demographic contexts. With respect to voting, both coefficients have odds ratios close to 1 with credibility intervals crossing 1, indicating no significant interaction effect. For protest attendance, on the other hand, both coefficients have odds ratios less than 1 and their 95% credibility intervals do not cross 1. This indicates statistically significant negative interaction effects. The odds ratio for the Youth share x Mid-age adults coefficient is 0.89 (95% CI: 0.81–0.98) and 0.86 (95% CI: 0.75–0.99) for Youth share x Older adults. This indicates that the gap in protest likelihood between young people and older people gets wider in cities with more youthful age structures. While all age groups are less likely to protest in more youthful cities, this negative effect is much stronger for mid-age and older adults. While younger people are still the most likely to protest, the difference between them and older people becomes most pronounced in cities with high youth share.

Fig. 4 plots this interaction effect graphically. The negative slopes for all age categories across the range of youth share support the main results from models 4–5 indicate a negative association between city youth shares and protest participation. With younger adults being more likely to protest than mid adults who are, in turn, more likely to protest than older adults. However, the plot shows that the gap in protest probability between ‘Younger adults’ relative to Mid-age and older adults appears to widen with these groups having a steeper rate of decrease compared to ‘Younger adults’, especially at higher youth share values.

With respect to hypotheses H2a and H2b, these interaction models do not provide support for H2a. The data does not indicate that any

difference in the odds of voting participation between young people and their older counterparts is significantly modified by the youth share of a city. In other words, the relative likelihood of young people voting compared to older adults does not significantly decrease as youth share increases.

The significant interaction effect for in Model 6 does not support H2b in the expected way. H2b implies that in more youthful cities, young people's propensity to protest increases relative to their older counterparts. However, the negative Age group coefficients and interaction term suggest that, while younger adults always show a higher probability of protesting, the relative difference in protest likelihood between younger adults and older adults actually becomes more pronounced as youth share increases. This is because the probabilities for mid-age and older adults decline more steeply or fall to lower absolute values faster than for younger adults.

5. Discussion and conclusion

Our results present a nuanced picture of the political demography of African cities. Overall, after controlling for other urban and individual characteristics, we find that the demographic structure of a city does not make individuals more or less likely to vote. Conversely, individuals are less likely to protest in more youthful cities, a statistically significant negative association that runs counter to expectations.

We also find that political behaviour is shaped by age. All else being equal, and as expected based on the existing literature, younger people are more likely to protest and less likely to vote. This effect is linear with age, with older people being more likely to vote and less likely to protest than mid-age people, who in turn are more likely to vote and less likely to protest than younger people.⁴

Finally, while living in a youthful urban area does not seem to

⁴ For age and voting, our findings contrast slightly with evidence from Europe and North America suggesting a curvilinear relationship between age and voting, with highest rates in the mid-age groups (Franklin, 2004; Henn & Foard, 2012; Wattenberg, 2003).

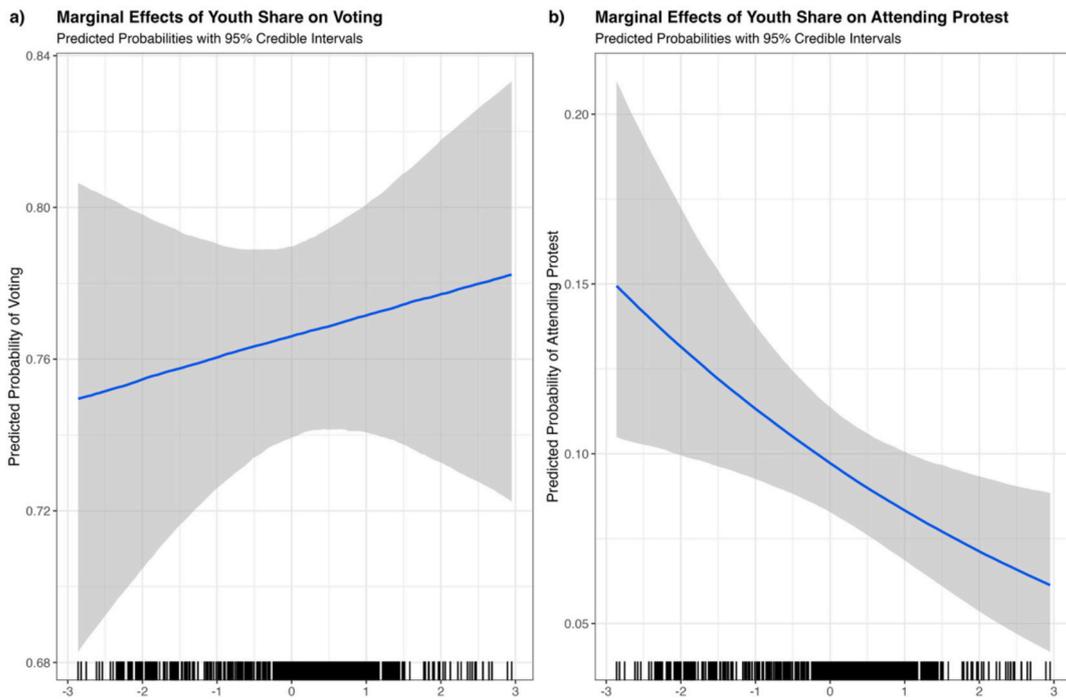


Fig. 3. Marginal effect of city youth shares on voting and protest participation. Note: Marginal effects of city-level youth shares with respect to the predicted probability of an individual reporting having (a) voted and (b) attending a protest. Estimates are taken from Models 3 and 6. 95% credibility intervals are reported around Bayesian point estimates.

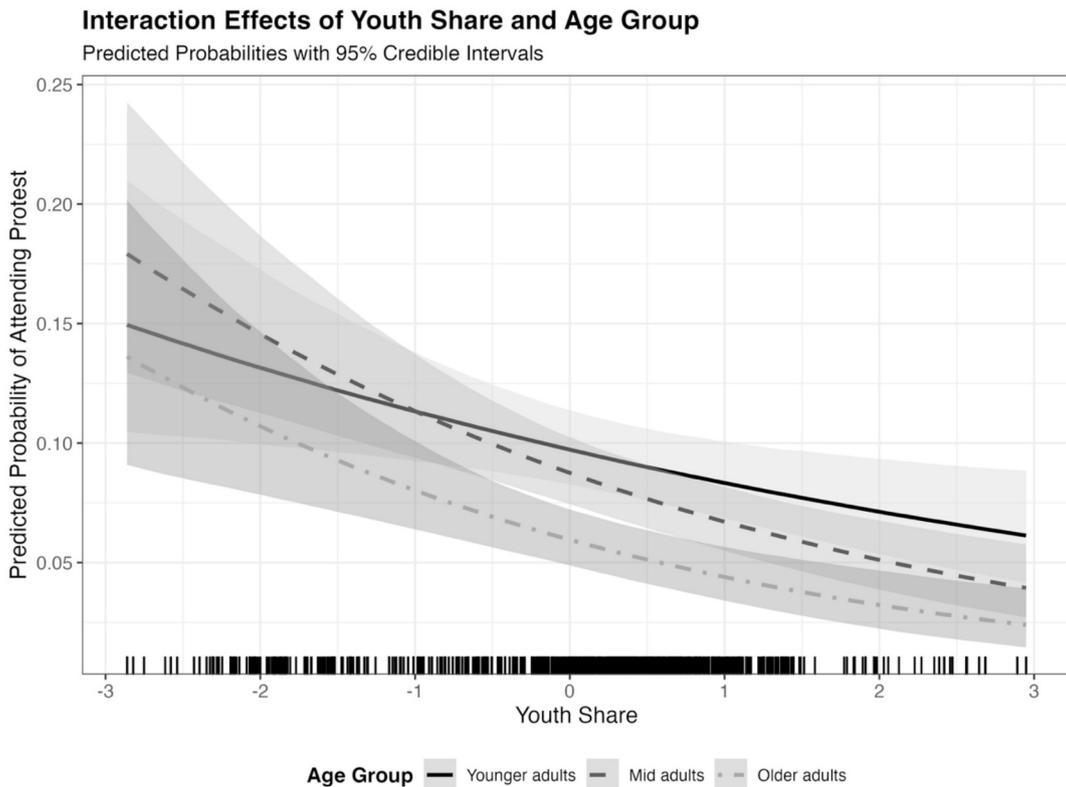


Fig. 4. Cross-level interaction of city youth shares and age group on protest participation. Note: Cross-level interaction effect of city-level youth share and individual age with respect to the predicted probability of protest participation. Estimates are taken from Model 6. 95% credibility intervals are reported around Bayesian point estimates.

mediate the voting behaviour of specific age groups (young, mid-age, and older residents), it does have a differential effect on protest behaviour. The relationship between age and protest remains negative for all age groups, but this negative effect is weaker for younger people. In other words, while all age groups protest less in more youthful cities, the difference between young people and older people is at its greatest in these contexts. Importantly, young people do not become more likely to protest relative to their older counterparts; instead, their older counterparts become even less likely, widening the age-cohort gap.

How might we understand these somewhat counter-intuitive findings, and what do they mean for future research? We identify two theoretically situated propositions that may underlie our unexpected finding that youthful cities have no effect on individual voting and a negative effect on individual protest.

The first relates to the nature of urban growth and economic conditions in African cities as distinct geographies. While large urban youth shares are often assumed to lead to an oversupply of labour and fewer economic opportunities (ref), some urban areas may be places of opportunity, attracting and/or retaining large youth shares precisely because they offer wide labour markets, improved livelihoods, and social mobility (McKenzie, 2017; Sladoje et al., 2020). This is likely to have differential effects on dynamics of political action (Elfverson et al., 2023). If more youthful cities are indeed becoming places of broader economic opportunity, citizens should be more likely to be invested in the maintenance of these local political-economic systems, leading to higher formal engagement and lower protest activity.

A second potential explanation may reflect the changing incentives for political actors to appeal to younger voters. As Africa urbanises, young urban dwellers are becoming increasingly important constituencies to mobilize in order to win elections (Resnick & Casale, 2011). Given this, political entrepreneurs are incentivised to increasingly target appeals to younger people and/or direct public goods and services in ways that benefit youth cohorts (Cheeseman, 2018; Macdonald et al., 2023; Resnick, 2013, 2015, 2017). Our results – that both residents overall and younger people protest less and that voting is not negatively affected – are consistent with political operators targeting their appeals to younger people where large youth shares represent a larger voting bloc. If more inclusive or ‘youth-focused’ political appeals are indeed present in cities with large youth cohorts, then individuals may be more (or not less) likely to engage in formal political behaviours and less likely to engage in informal protest behaviour.

While our findings provide important nuance to the political demography of African cities, they also point to several limitations of this study and avenues for future research. One key limitation is that we do not directly test the causal mechanisms linking city-level youth shares to political behaviour. For example, changes in local economic conditions or political appeals that prime the motive, means, and opportunities for distinct types of political engagement. Instead, we situate these mechanisms at the appropriate spatial scale, demonstrating that local context matters. Building off of our study, future research could conduct survey work in demographically different cities to directly probe these causal mechanisms.

Furthermore, future research could build upon these results by gathering data on the type and nature of urban growth (e.g., migration, natural increase, or reclassification) to determine how this relates to demographic structure and political behaviours. Similarly, studies on the political demography of African urbanisation would benefit from being attuned to the different types of political appeals directed at young people and how this might shape participation and mobilisation in various institutional contexts. A more nuanced picture of these dynamics would be a significant scholarly advance and have important implications for devising sustainable urban policy.

In closing, this article has made three key contributions to the literature. First, we develop and operationalise new subnational estimates of the demographic composition of individual urban settlements, the first article to do so to our knowledge. Second, we use these to

empirically demonstrate how the local demographic characteristics of particular cities may shape the political behaviours of individuals living in these places. The results nuance and challenge common theoretical assumptions and empirical findings about demographic structure and political behaviour and point to future research directions. Third, by demonstrating the benefits of a place-based analysis of urban demographic structure and political behaviour, we contribute to research agendas (Hoelscher et al., 2023) taking a subnational perspective when studying the political implications of urbanisation in Africa.

CRediT authorship contribution statement

Nick Dorward: Writing – review & editing, Writing – original draft, Visualization, Validation, Software, Methodology, Investigation, Funding acquisition, Formal analysis, Data curation, Conceptualization.
Kristian Hoelscher: Writing – review & editing, Writing – original draft, Investigation, Funding acquisition, Conceptualization.

Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: Nick Dorward reports financial support was provided by UK Research and Innovation Economic and Social Research Council. Kristian Hoelscher reports financial support was provided by Norges forskningsråd. If there are other authors, they declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.worlddev.2026.107344>.

Data availability

Data will be made available on request.

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